machine learning prerequisites workshop probability and statistics

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فهرست مطالب

فضای نمونه (Sample Space) پیشامد (Event)

اصول احتمال (Probability Axioms)

احتمال شرطى (Conditional Probability)

پیشامدهای مستقل (Independent Events)

قضیه بیز (Bayes Theorem)

(Law of total probability) قانون احتمال کل

متغير تصادفي (Random Variable)

تابع جرم احتمال (PMF)

تابع چگالی احتمال (PDF)

تابع توزیع تجمعی (CDF)

امید ریاضی (Expected Value)

Law of the unconscious statistician - LOTUS

واريانس (Variance)

کوواریانس (Covariance)

همبستگی (Correlation)

(Bias) اریب

بایاس-واریانس تریدآف (bias-variance tradeoff)

متغير تصادفي برنولي (Bernoulli)

متغير تصادفي يكنواخت (Uniform Distribution)

متغیر تصادفی دوجملهای (Binomial)

توزیع نرمال (Normal Distribution)

قضیه حد مرکزی (Central Limit Theorem)

قانون اعداد بزرگ (law of large numbers)

Maximum Likelihood Estimation (MLE)

Maximum A Posteriori (MAP)



فضای نمونه (Sample Space)

(شماره شرکت کننده سوم , شماره شرکت کننده دوم , شماره شرکت کننده اول)

Ω

$$(1,2,3) (2,3,1) (3,2,1) (3,1,2) (2,1,3)$$

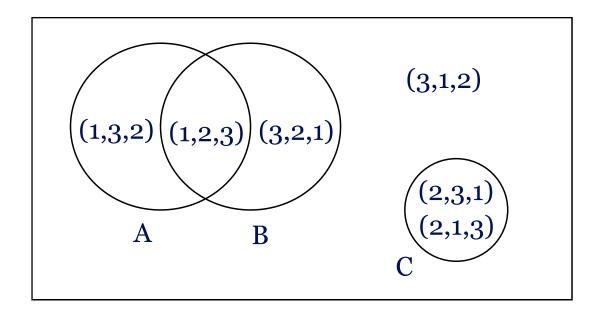


$$\Omega = \{(1,2,3), (1,3,2), (2,1,3), (2,3,1), (3,1,2), (3,2,1)\}$$



پیشامد (Event)

 Ω



$$A = \{(1,2,3), (1,3,2)\}$$

$$B = \{(1,2,3), (3,2,1)\}$$

$$C = \{(2,3,1), (2,1,3)\}$$



(Probability Axioms) اصول احتمال

$$0 \leq P(A) \leq 1$$
 : ابه ازای هر پیشامد داریم (۱

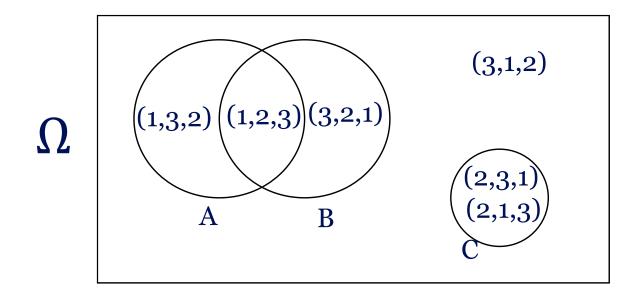
$$P(\Omega) = 1$$
 (Y

$$P(igcup_{i=1}^\infty A_i) = \sum_{i=1}^\infty P(A_i)$$
 برای پیشامدهای ناسازگار داریم: (۳



احتمال شرطى (Conditional Probability)

$$P(A|B) = \frac{P(A \cap B)}{P(B)}$$



$$P(A|B) = \frac{\frac{1}{6}}{\frac{1}{3}} = \frac{1}{2}$$



پیشامدهای مستقل (Independent Events)

$$P(A \cap B) = P(A)P(B)$$



قضیه بیز (Bayes Theorem)

$$P(B|A) = \frac{P(A|B)P(B)}{P(A)}$$

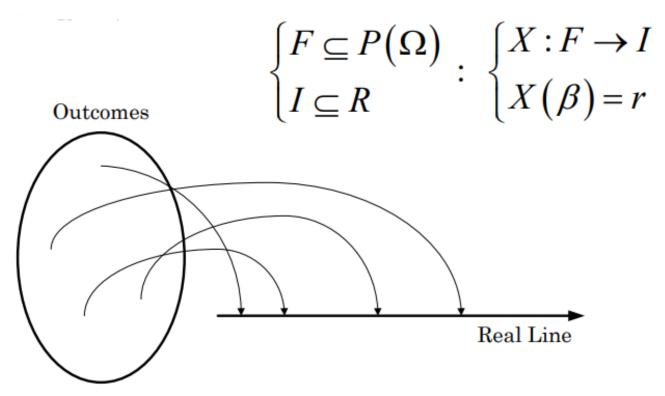


(Law of total probability) قانون احتمال کل

$$P(A) = \sum_{i=1}^{n} P(A \cap B_i)$$



متغير تصادفي (Random Variable)



- پیوسته: تابع چگالی احتمال
 گسسته: تابع جرم احتمال



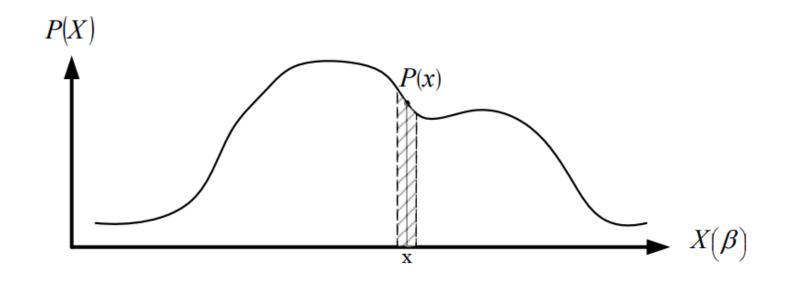
(probability mass function - PMF) تابع جرم احتمال

$$f_X(x) = P(X = x)$$



(probability density function - PDF) تابع چگالی احتمال

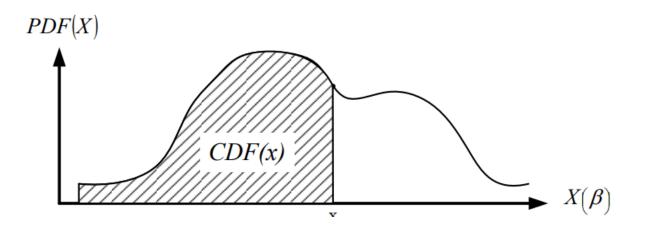
$$P(a \le X \le b) = \int_{a}^{b} f_X(x) dx$$





(cumulative distribution function - CDF) تابع توزیع تجمعی

$$F_X(x) = P(X \le x)$$



$$F_X(x) = \begin{cases} F_X(x) = \sum_{X \le x} f_X(x) \\ F_X(x) = \int_{-\infty}^x f_X(x) \implies \frac{d}{dx} F_X(x) = f_X(x) \end{cases}$$



امید ریاضی (Expected Value)

$$E[X] = \mu = \begin{cases} \sum_{x \in \Omega} x f_X(x) \\ \int_{-\infty}^{\infty} x f_X(x) \end{cases}$$

$$E[aX + bY + c] = aE[X] + bE[Y] + c$$

$$\widehat{\mu} = \frac{\sum_{i=1}^{n} x_i}{n}$$



Law of the unconscious statistician - LOTUS

$$E[g(X)] = \begin{cases} \sum_{x \in \Omega} g(x) f_X(x) \\ \int_{-\infty}^{\infty} g(x) f_X(x) \end{cases}$$



واريانس (Variance)

$$Var(X) = \sigma^2 = E[(X - \mu)^2] = E[X^2] - E[X]^2$$

 $\mu = E[X]$

$$Var(X) = \begin{cases} \sum_{x \in \Omega} (X - \mu)^2 f_X(x) \\ \\ \int_{-\infty}^{\infty} (x - \mu)^2 f_X(x) \end{cases}$$

$$\widehat{\sigma^2} = \frac{\sum_{i=1}^n (x_i - \widehat{\mu})}{n-1}$$

$$Var(aX + b) = a^2 Var(X)$$



(Covariance) کوواریانس

$$Cov(X,Y) = E[(X - \mu_X)(Y - \mu_Y)]$$

$$Cov(X,Y) = \frac{\sum ((x_i - \bar{x})(y_i - \bar{y}))}{n-1}$$

$$Var(X + Y) = Var(X) + Var(Y) + 2Cov(X, Y)$$



همبستگی (correlation)

$$Corr(X,Y) = \frac{Cov(X,Y)}{\sqrt{Var(X)Var(Y)}}$$

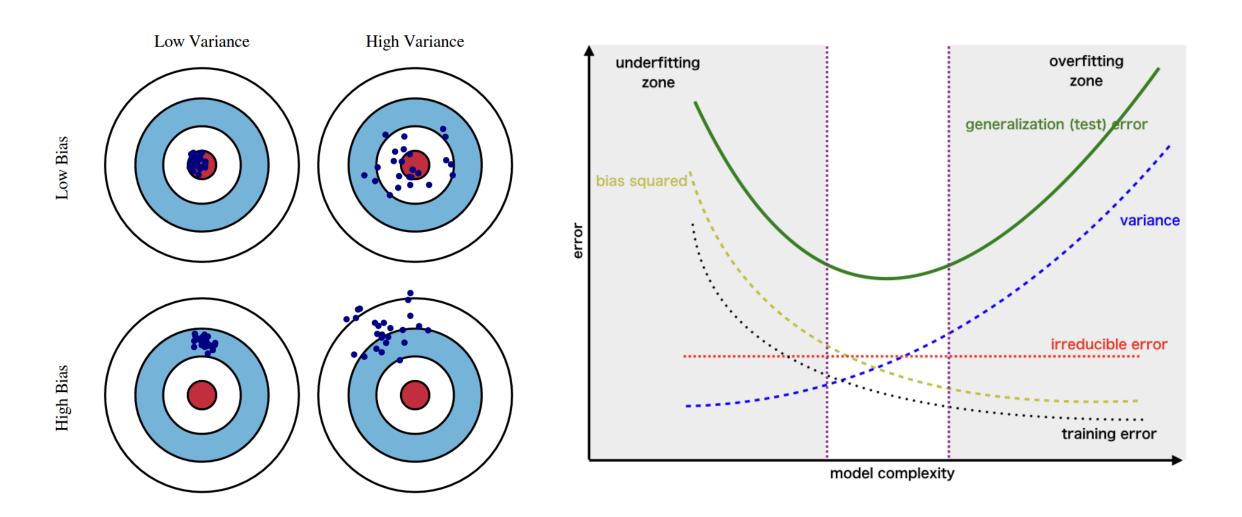


اریب (Bias)

$$Bias(\hat{\theta}) = E[\hat{\theta} - \theta]$$



بایاس – واریانس تریدآف (bias-variance tradeoff)





متغیر تصادفی برنولی (Bernoulli)

$$X \sim Br(p)$$

$$P(X = x) = \begin{cases} p & x = 1 \\ 1 - p & x = 0 \\ 0 & otherwise \end{cases}$$

$$f_X(x) = p^x (1-p)^{1-x}$$

$$E[X] = p$$

$$Var(X) = p(1-p)$$



متغیر تصادفی دوجملهای (Binomial)

$$X \sim Bin(n, p)$$

$$P(X = x) = \begin{cases} \binom{n}{x} p^{x} (1 - p)^{n - x} & 0 \le x \le n \\ 0 & otherwise \end{cases}$$

$$E[X] = np$$

$$Var(X) = np(1-p)$$



متغير تصادفي يكنواخت (Uniform Distribution)

$$X \sim U(a, b)$$

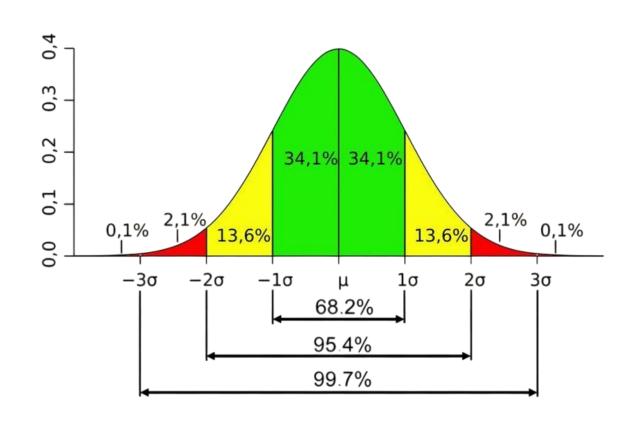
$$f_X(x) = \begin{cases} \frac{1}{b-a} & a \le x \le b \\ 0 & otherwise \end{cases}$$

$$E[X] = \frac{a+b}{2}$$

$$Var(X) = \frac{(b-a)^2}{12}$$



(Normal Distribution) توزیع نرمال



$$X \sim N(\mu, \sigma^2)$$

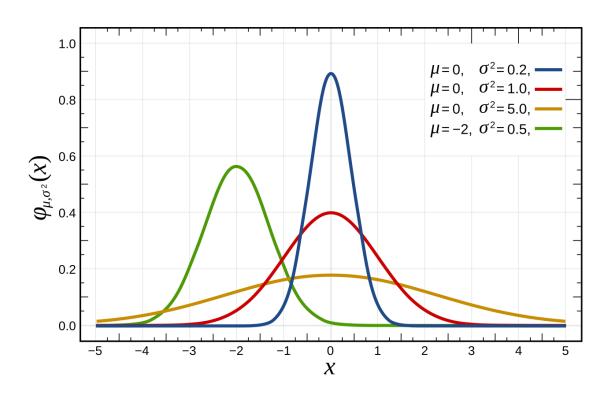
$$f_X(x) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2}$$

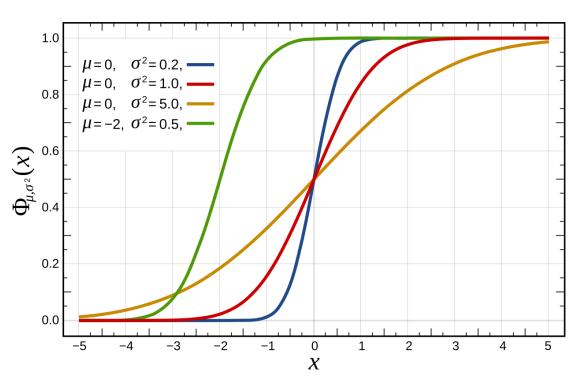
$$E[X] = \mu$$

$$Var(X) = \sigma^2$$



(Normal Distribution) توزیع نرمال







توزيع احتمال توام (Joint probability distribution)

$$F_{X,Y}(x,y) = P(X \le x, Y \le y) = \int_{-\infty}^{x} \int_{-\infty}^{y} f_{X,Y}(x,y) dy dx$$

$$f_{X,Y}(x,y) = \frac{\partial^2}{\partial x \partial y} F_{X,Y}(x,y)$$



قضیه حد مرکزی (Central Limit Theorem)

و $E[X_i]=\mu$ مستقل و دارای توزیع یکسان (idd) باشند به طوری که X_1,X_2,\dots,X_n اگر دارای $Var[X_i]=\sigma^2$

$$\bar{X} \sim N\left(\mu, \frac{\sigma^2}{n}\right)$$

* idd: Independent and Identically Distributed



(law of large numbers) قانون اعداد بزرگ

arepsilon>0 به ازای هر $E[X_i]=\mu$ مستقل و دارای توزیع یکسان (idd) باشند به طوری که $E[X_i]=\mu$ ، به ازای هر داریم :

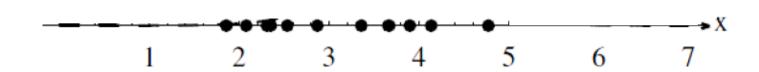
$$P\left\{\left|\frac{X_1+X_2+\cdots+X_n}{n}-\mu\right|<\varepsilon\right\}\to 0 \quad \text{ if } n\to\infty$$

یا به صورت دیگر :

$$\lim_{n \to \infty} \frac{X_1 + X_2 + \dots + X_n}{n} = \mu$$

* idd: Independent and Identically Distributed





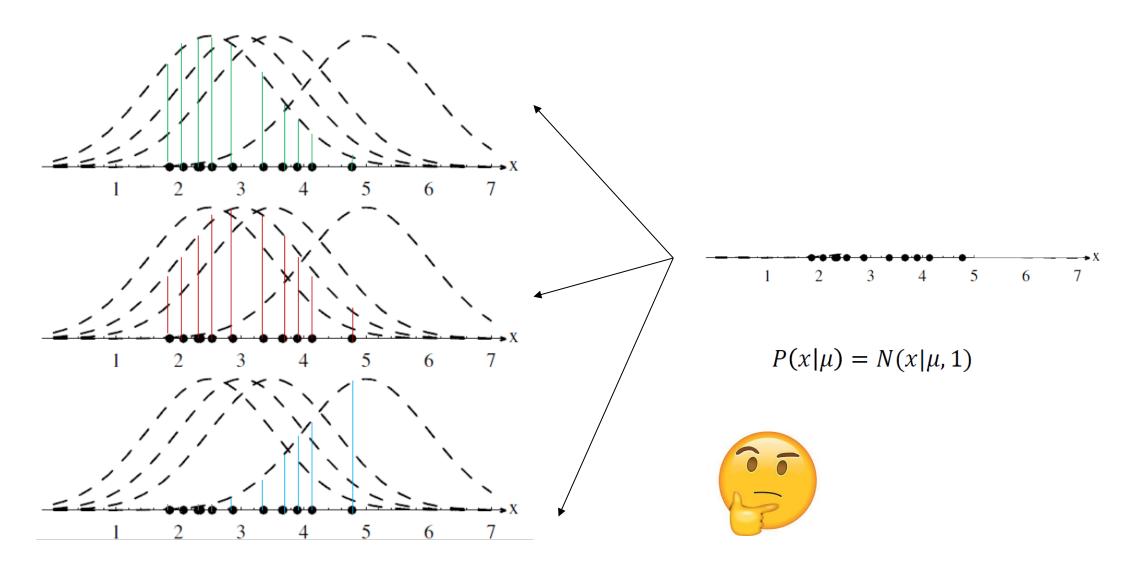
$$P(x|\mu) = N(x|\mu, 1)$$



Likelihood, Posterior and Prior

$$P(heta|Y) = rac{P(Y| heta) P(heta)}{P(Y)}$$
 $\propto P(Y| heta) P(heta)$
 $\Rightarrow P(Y| heta) P(heta)$



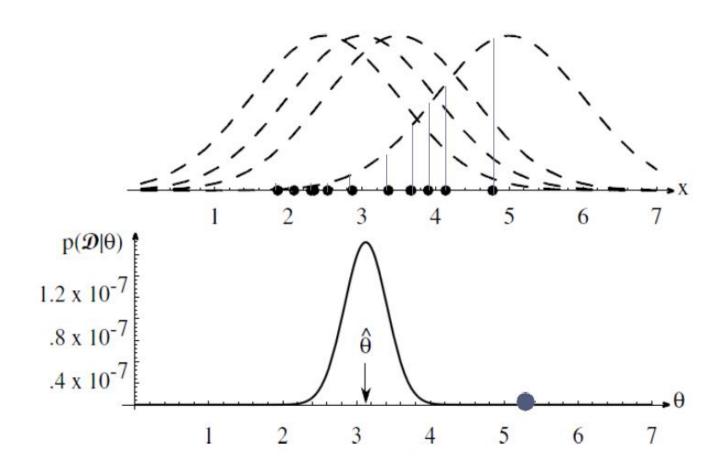




$$P(X|\theta) = \sum_{i=1}^{n} P(x^{(i)}|\theta)$$

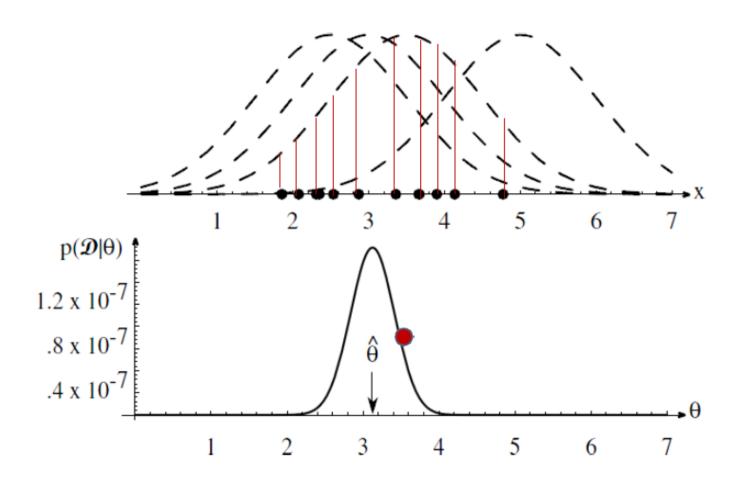
$$\theta_{ML} = argmax P(X|\theta)$$





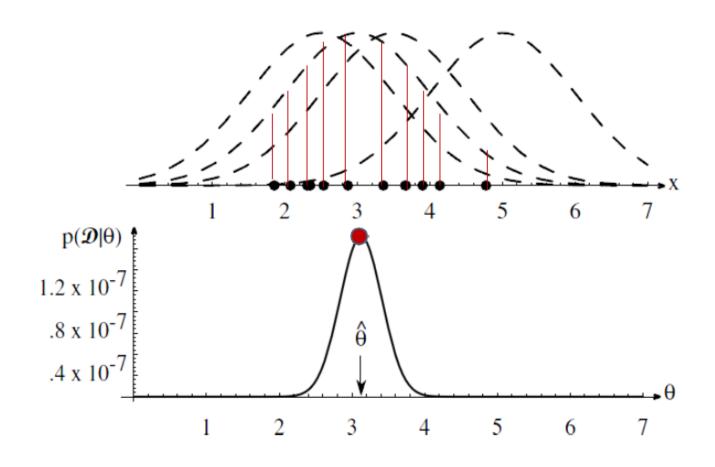


Maximum A Posteriori (MAP)





Maximum A Posteriori (MAP)





$$\mathcal{L}(\boldsymbol{\theta}) = \ln p(\mathcal{D}|\boldsymbol{\theta}) = \ln \prod_{i=1}^{N} p(\boldsymbol{x}^{(i)}|\boldsymbol{\theta}) = \sum_{i=1}^{N} \ln p(\boldsymbol{x}^{(i)}|\boldsymbol{\theta})$$

$$\widehat{\boldsymbol{\theta}}_{ML} = \underset{\boldsymbol{\theta}}{\operatorname{argmax}} \mathcal{L}(\boldsymbol{\theta}) = \underset{\boldsymbol{\theta}}{\operatorname{argmax}} \sum_{i=1}^{N} \ln p(\boldsymbol{x}^{(i)} | \boldsymbol{\theta})$$



MLE Bernoulli

Given:
$$\mathcal{D} = \{x^{(1)}, x^{(2)}, ..., x^{(N)}\}, m \text{ heads (I)}, N - m \text{ tails (0)}$$

$$p(x|\theta) = \theta^x (1-\theta)^{1-x}$$

$$p(\mathcal{D}|\theta) = \prod_{i=1}^{N} p(x^{(i)}|\theta) = \prod_{i=1}^{N} \theta^{x^{(i)}} (1-\theta)^{1-x^{(i)}}$$
$$\ln p(\mathcal{D}|\theta) = \sum_{i=1}^{N} \ln p(x^{(i)}|\theta) = \sum_{i=1}^{N} \{x^{(i)} \ln \theta + (1-x^{(i)}) \ln(1-\theta)\}$$

$$\ln p(\mathcal{D}|\theta) = \sum_{i=1}^{N} \ln p(x^{(i)}|\theta) = \sum_{i=1}^{N} \{x^{(i)} \ln \theta + (1 - x^{(i)}) \ln(1 - \theta)\}$$

$$\frac{\partial \ln p(\mathcal{D}|\theta)}{\partial \theta} = 0 \Rightarrow \theta_{ML} = \frac{\sum_{i=1}^{N} x^{(i)}}{N} = \frac{m}{N}$$



Maximum A Posteriori (MAP)

MAP estimation

$$\widehat{\boldsymbol{\theta}}_{MAP} = \operatorname*{argmax}_{\boldsymbol{\theta}} p(\boldsymbol{\theta}|\mathcal{D})$$

Since
$$p(\boldsymbol{\theta}|\mathcal{D}) \propto p(\mathcal{D}|\boldsymbol{\theta})p(\boldsymbol{\theta})$$

$$\widehat{\boldsymbol{\theta}}_{MAP} = \operatorname*{argmax}_{\boldsymbol{\theta}} p(\mathcal{D}|\boldsymbol{\theta}) p(\boldsymbol{\theta})$$



Gaussian MAP

$$p(x|\mu) \sim N(\mu, \sigma^2)$$
 μ is the only unknown paramete $p(\mu|\mu_0) \sim N(\mu_0, \sigma_0^2)$ μ_0 and σ_0 are known

$$\frac{d}{d\mu} \ln \left(p(\mu) \prod_{i=1}^{N} p(x^{(i)} | \mu) \right) = 0$$

$$\Rightarrow \sum_{i=1}^{N} \frac{1}{\sigma^{2}} (x^{(i)} - \mu) - \frac{1}{\sigma_{0}^{2}} (\mu - \mu_{0}) = 0$$

$$\Rightarrow \hat{\mu}_{MAP} = \frac{\mu_{0} + \frac{\sigma_{0}^{2}}{\sigma^{2}} \sum_{i=1}^{N} x^{(i)}}{1 + \frac{\sigma_{0}^{2}}{\sigma^{2}} N}$$





- Stochastic Processes Hamid R. Rabiee
- machine learning Ali Sharifi-Zarchi
- machine learning Mahdieh Soleymani Baghshah

