

## Econometrics 1, Class Week 7

Back-up Video: Class week 7 video ([click here](#))

### Learning Outcomes

- (a) Forecasting in Autoregressive Integrated Moving Average (ARIMA) models.
- (b) Intertemporal Dependence.
- (c) First Moments of the Stationary Distribution.
- (d) Second Moments of the Stationary Distribution.
- (e) Order of Integration

### Prerequisites

1. Conditional and unconditional moments.
2. Law of Iterated Expectations (not strictly necessary, but may be helpful).

## Univariate Time-Series Models

### Models

$$y_t = \alpha + \rho y_{t-1} + \epsilon_t \quad (1)$$

$$y_t = \alpha + y_{t-1} + \epsilon_t \quad (2)$$

$$y_t = \alpha + \epsilon_t + \mu \epsilon_{t-1} \quad (3)$$

$$\begin{aligned} \Delta y_t &= y_t - y_{t-1} \\ &= \alpha + \rho \Delta y_{t-1} + \epsilon_t + \mu \epsilon_{t-1}, \end{aligned} \quad (4)$$

where  $\epsilon_t$  is white noise (WN), i.e.  $\epsilon_t \stackrel{i.i.d.}{\sim} \mathbb{E}[\epsilon_t] = 0, \mathbb{E}[\epsilon_t^2] = \sigma^2$ .

Model interpretation:

(1) : Autoregressive process of order 1, AR(1),

(2) : Random walk with drift  $\alpha$ .

(3) : Moving Average of order 1, MA(1).

(4) : ARIMA(1,1,1).

(a) Forecasting  $y_{T+1}$  and  $y_{T+2}$ .

Estimate model parameters and residuals on the basis of a sample  $\{y_t, t = 0, \dots, T\}$ ; use estimates to forecast  $y_{T+1}$  and  $y_{T+2}$ .

Denote one-step forecast by  $\hat{y}_{T+1|T}$ , and two-step forecast by  $\hat{y}_{T+2|T}$ , given information up to (and including) time  $T$ .

General approach:

- Obtain expression for  $y_{T+s}$ ;
- replace parameters by their respective estimates and any variable dated  $s > T$  by its conditional expectation, given information up to  $T$ .

AR(1)

$$y_{T+1} = \alpha + \rho y_T + \epsilon_{T+1} \quad (5)$$

$$\begin{aligned} y_{T+2} &= \alpha + \rho y_{T+1} + \epsilon_{T+2} \\ &= \alpha + \rho(\alpha + \rho y_T + \epsilon_{T+1}) + \epsilon_{T+2} \\ &= \alpha(1 + \rho) + \rho^2 y_T + \epsilon_{T+2} + \rho \epsilon_{T+1}. \end{aligned} \quad (6)$$

Hence, by properties of WN,

$$\hat{y}_{T+1|T} = \hat{\alpha} + \hat{\rho} y_T \quad (7)$$

$$\begin{aligned} \hat{y}_{T+2|T} &= \hat{\alpha}(1 + \hat{\rho}) + \hat{\rho}^2 y_T \\ &= \hat{\alpha} + \hat{\rho} \hat{y}_{T+1|T}. \end{aligned} \quad (8)$$

RW with drift: special case of AR(1), with  $\rho = 1$ , so

$$\hat{y}_{T+1|T} = \hat{\alpha} + y_T \quad (9)$$

$$\begin{aligned} \hat{y}_{T+2|T} &= 2\hat{\alpha} + y_T \\ &= \hat{\alpha} + \hat{y}_{T+1|T}. \end{aligned} \quad (10)$$

MA(1)

$$y_{T+1} = \alpha + \epsilon_{T+1} + \mu\epsilon_T \quad (11)$$

$$y_{T+2} = \alpha + \epsilon_{T+2} + \mu\epsilon_{T+1}. \quad (12)$$

Hence,

$$\hat{y}_{T+1|T} = \hat{\alpha} + \hat{\mu}\hat{\epsilon}_T \quad (13)$$

$$\hat{y}_{T+2|T} = \hat{\alpha}. \quad (14)$$

ARIMA(1,1,1)

$$\begin{aligned}
 y_{T+1} &= y_T + \Delta y_{T+1} \\
 &= y_T + \alpha + \rho \Delta y_T + \epsilon_{T+1} + \mu \epsilon_T
 \end{aligned} \tag{15}$$

$$\begin{aligned}
 y_{T+2} &= y_T + \Delta y_{T+1} + \Delta y_{T+2} \\
 &= y_{T+1} + \Delta y_{T+2} \\
 &= y_{T+1} + \alpha + \rho \Delta y_{T+1} + \epsilon_{T+2} + \mu \epsilon_{T+1} \\
 &= y_T + \Delta y_{T+1} + \alpha + \rho \Delta y_{T+1} + \epsilon_{T+2} + \mu \epsilon_{T+1} \\
 &= y_T + \alpha + (1 + \rho) \Delta y_{T+1} + \epsilon_{T+2} + \mu \epsilon_{T+1}
 \end{aligned} \tag{16}$$

Hence,

$$\hat{y}_{T+1|T} = y_T + \hat{\alpha} + \hat{\rho} \Delta y_T + \hat{\mu} \hat{\epsilon}_T \tag{17}$$

$$\begin{aligned}
 \hat{y}_{T+2|T} &= \hat{y}_{T+1|T} + \hat{\alpha} + \hat{\rho} (\hat{y}_{T+1|T} - y_T) \\
 &= y_T + \hat{\alpha} + (1 + \hat{\rho}) (\hat{y}_{T+1|T} - y_T).
 \end{aligned} \tag{18}$$

(b) Intertemporal Dependence of AR(1): Recursive Substitution

AR(1), using  $L$  to denote the lag operator,

$$\begin{aligned} y_t &= \alpha + \rho y_{t-1} + \epsilon_t \\ &= \alpha + \rho L y_t + \epsilon_t \end{aligned} \quad (19)$$

$$\begin{aligned} &= \alpha + \rho(\alpha + \rho y_{t-2} + \epsilon_{t-1}) + \epsilon_t \\ &= \alpha + \rho\alpha + \rho^2 y_{t-2} + \epsilon_t + \rho\epsilon_{t-1} \\ &= \dots \\ &= \alpha \left( \sum_{k=0}^{s-1} \rho^k \right) + \rho^s y_{t-s} + \sum_{m=0}^{s-1} \rho^m \epsilon_{t-m}. \end{aligned} \quad (20)$$

and if  $|\rho| < 1$ , i.e. AR(1) is stationary, by taking limits,

$$\begin{aligned} y_t &= \frac{\alpha}{1-\rho} + \sum_{m=0}^{\infty} \rho^m \epsilon_{t-m} \\ &= \frac{\alpha}{1-\rho} + \text{MA}(\infty) \\ &= \frac{1}{1-\rho L} (\alpha + \epsilon_t), \end{aligned} \quad (21)$$

where the last equality can be seen directly from (19).

(20) implies for RW ( $\rho = 1$ ) with drift

$$y_t = s\alpha + y_{t-s} + \sum_{m=0}^{s-1} \epsilon_{t-m}. \quad (22)$$

### (c) Expectations of the Stationary Distribution

Definition: A stochastic process is *covariance-stationary* if the unconditional moments do not depend on calendar time  $t$ .

So can only obtain unconditional moments for covariance-stationary processes, i.e. (1) and (3).

For AR(1), assuming  $\mathbb{E}[y_t]$  exists, taking expectations on LHS and RHS of (1),

$$\begin{aligned}\mathbb{E}[y_t] &= \alpha + \rho\mathbb{E}[y_{t-1}] \\ &= \alpha + \rho\mathbb{E}[y_t] \\ &= \frac{\alpha}{1 - \rho}.\end{aligned}\tag{23}$$

For MA(1),

$$\mathbb{E}[y_t] = \alpha.\tag{24}$$



## (d) Variance, Covariance, Correlations of Stationary Distribution

Recall that additive constants (such as  $\alpha$ ) do not contribute to variances and covariances; and that multiplicative constants (such as  $\rho$  and  $\mu$ ) contribute multiplicatively.

AR(1): Using (20), and assuming second moment exists,

$$\begin{aligned}
 \text{var}(y_t) &= \text{var}(\alpha + \rho y_{t-1} + \epsilon_t) \\
 &= \rho^2 \text{var}(y_{t-1}) + \sigma^2 \quad \text{b/c } \epsilon \text{ is WN} \\
 &= \rho^2 \text{var}(y_t) + \sigma^2 \quad \text{b/c of stationarity} \\
 &= \frac{\sigma^2}{1 - \rho^2}.
 \end{aligned} \tag{25}$$

$$\begin{aligned}
 \text{cov}(y_t, y_{t-s}) &= \text{cov}\left(\alpha \left(\sum_{k=0}^{s-1} \rho^k\right) + \rho^s y_{t-s} + \sum_{m=0}^{s-1} \rho^m \epsilon_{t-m}, y_{t-s}\right) \\
 &= \rho^s \text{var}(y_t) \\
 &= \frac{\rho^s \sigma^2}{1 - \rho^2}
 \end{aligned} \tag{26}$$

$$\text{corr}(y_t, y_{t-s}) = \rho^s. \tag{27}$$

So AR process exhibits exponentially declining, yet infinite dependence (memory).

MA(1)

$$\begin{aligned}\text{var}(y_t) &= \text{var}(\alpha + \epsilon_t + \mu\epsilon_{t-1}) \\ &= \sigma^2 + \mu^2\sigma^2 \text{ b/c } \epsilon \text{ is WN} \\ &= (1 + \mu^2)\sigma^2.\end{aligned}\tag{28}$$

$$\begin{aligned}\text{cov}(y_t, y_{t-1}) &= \text{cov}(\alpha + \epsilon_t + \mu\epsilon_{t-1}, \alpha + \epsilon_{t-1} + \mu\epsilon_{t-2}) \\ &= \mu\text{var}(\epsilon_{t-1}) \\ &= \mu\sigma^2.\end{aligned}\tag{29}$$

$$\text{cov}(y_t, y_{t-s}) = 0 \quad \forall |s| > 1.\tag{30}$$

$$\text{corr}(y_t, y_{t-1}) = \frac{\mu}{1 + \mu^2}\tag{31}$$

$$\text{corr}(y, y_{t-s}) = 0 \quad \forall |s| > 1.\tag{32}$$

So MA(1) exhibits no dependence (memory) beyond first lag.

And MA(s) exhibits no dependence beyond lag s.

(e) Integration of order 1,  $I(1)$

RW with drift is  $I(1)$ :

$$\Delta y_t = (1 - L)y_t = \alpha + \epsilon_t. \quad (33)$$

ARIMA(1,1,1) is  $I(1)$ :

$$\begin{aligned} (1 - L)y_t &= \alpha + \rho(1 - L)y_{t-1} + \epsilon_t + \mu\epsilon_{t-1} \\ (1 - \rho L)(1 - L)y_t &= \alpha + (1 + \mu L)\epsilon_t. \end{aligned} \quad (34)$$