EM (cont'd) and Hidden Markov Models

Latent variable models

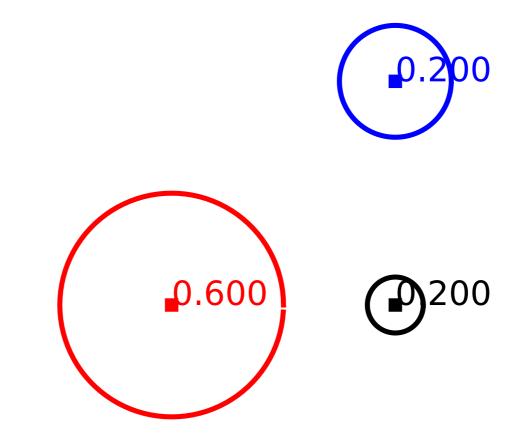
- Latent variable models allow us to hypothesize different types of structures that may underlie the observed data and then recover those structures
- We have understand how to
 - specify them (variables, distributions with parameters)
 - sample from them (as generative models)
 - estimate them from data
- Recall mixture models

$$y y \in \{1, \dots, K\} P(y) = p_y$$

$$\downarrow x \in \mathbb{R}^d P(x|y) = N(x; \mu^{(y)}, \sigma_y^2 I)$$

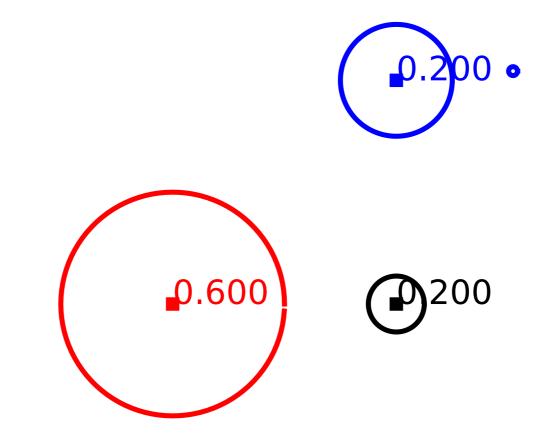


- We can sample points from a mixture model in two steps
 - 1) Sample $y \sim \text{Categ}(p_1, \ldots, p_K)$ which cluster
 - 2) Sample $x \sim N(x; \mu^{(y)}, \sigma_y^2 I)$ which point from the chosen cluster



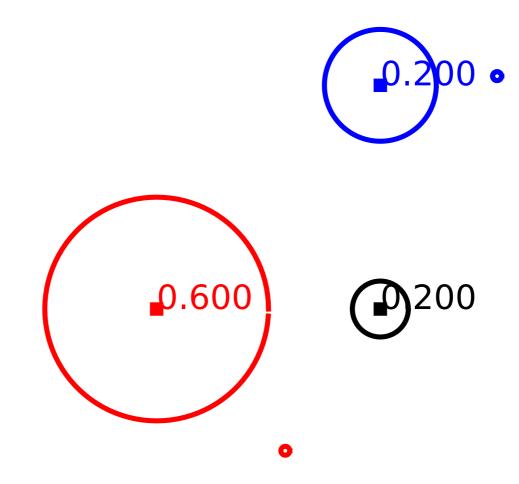


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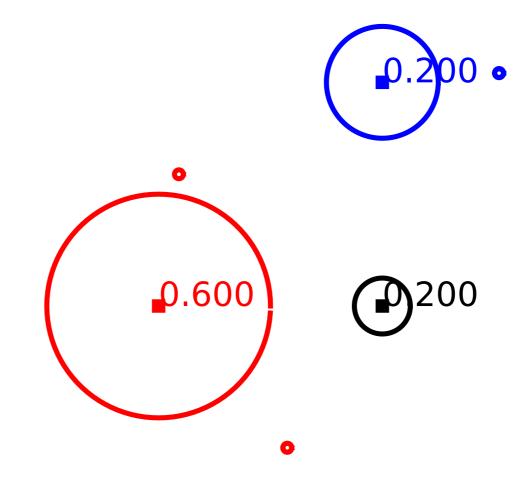


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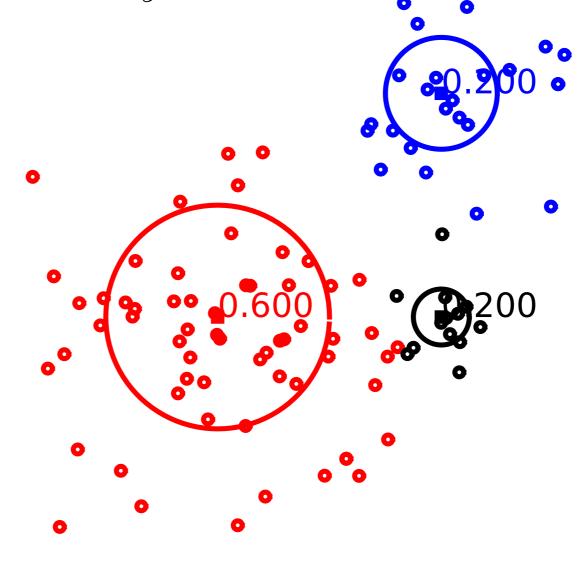


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Mixture model: estimation

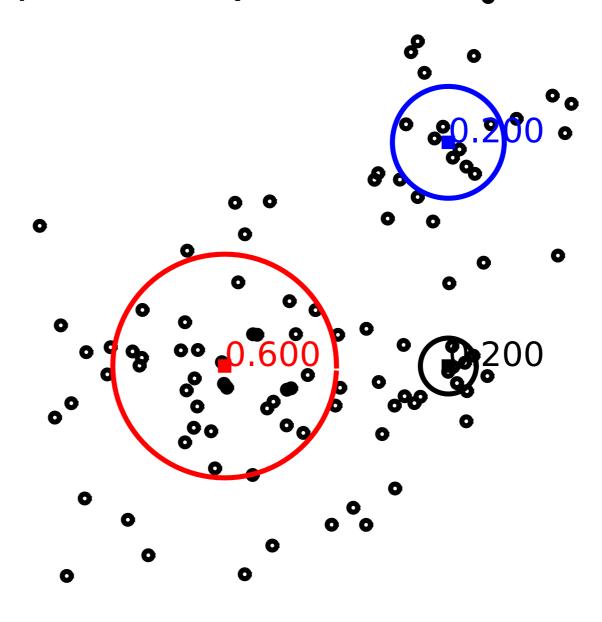
 But we typically don't get the cluster labels in the observed data, only the resulting points x





Mixture model: estimation

 Our goal is to try to estimate the underlying model (mixing proportions, component Gaussians) from the points alone (incomplete data)





Complete data

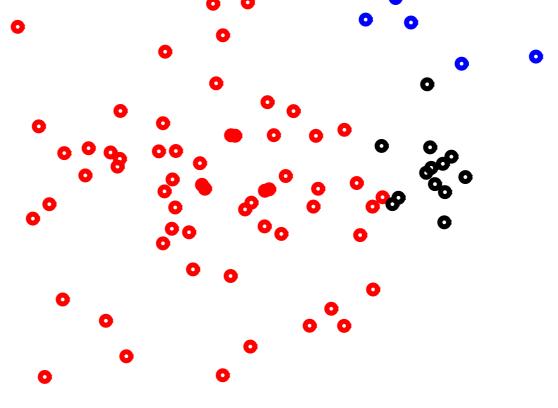
 Estimation would be easy if we had complete data, i.e., pairs of cluster ids y and corresponding points x

complete data =
$$\{x^{(i)}, y^{(i)}\}_{i=1,...,n}$$

 We represent the associated cluster ids in terms of indicators

$$\delta(y|i) = 1 \text{ if } y = y^{(i)}$$

and 0 otherwise



mplete data: mixing proportions

We maximize the log-likelihood of complete data

$$\sum_{y=1}^{K} \sum_{i=1}^{n} \delta(y|i) \log p_y$$

Resulting ML estimate

$$\hat{p}_y = \frac{\sum_{i=1}^n \delta(y|i)}{n}, \quad y = 1, \dots, K$$

Complete data: Gaussians

 We maximize the log-likelihood of complete data (separately for each Gaussian)

$$\sum_{i=1}^{n} \delta(y|i) \log N(x^{(i)}; \mu^{(y)}, \sigma_y^2 I)$$

The resulting ML estimates

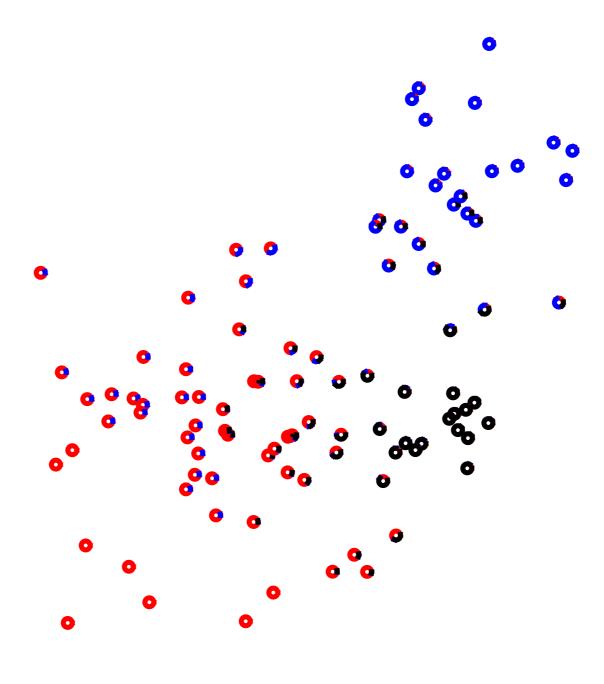
$$\hat{\mu}^{(y)} = \frac{\sum_{i=1}^{n} \delta(y|i)x^{(i)}}{\sum_{i=1}^{n} \delta(y|i)}$$

$$\hat{\sigma}_{y}^{2} = \frac{\sum_{i=1}^{n} \delta(y|i)||x^{(i)} - \hat{\mu}^{(y)}||^{2}}{d\sum_{i=1}^{n} \delta(y|i)}$$



Weighted data

 The estimation would be just as easy if instead of true cluster labels y, we had soft assignments





E-step: weighted data

 The estimation would be just as easy if instead of true cluster labels y, we had soft assignments

$$p(y|i) = \frac{p_y N(x^{(i)}; \mu^{(y)}, \sigma_y^2 I)}{\sum_{j=1}^k p_j N(x^{(i)}; \mu^{(j)}, \sigma_j^2 I)}$$

$$p(y|i) \text{ is the posterior probability}$$
 that the unknown $y^{(i)}$ is y (these assignments will change as the model is updated)

M-step: mixing proportions

We maximize the expected (weighted) log-likelihood

$$\sum_{y=1}^{K} \sum_{i=1}^{n} p(y|i) \log p_y$$

The resulting estimate (cf. before)

$$\hat{p}_y = \frac{\sum_{i=1}^n p(y|i)}{n}, \quad y = 1, \dots, K$$

M-step: Gaussians

 We maximize the expected (weighted) log-likelihood (separately for each Gaussian)

$$\sum_{i=1}^{n} p(y|i) \log N(x^{(i)}; \mu^{(y)}, \sigma_y^2 I)$$

The resulting estimates (cf. before)

$$\hat{\mu}^{(y)} = \frac{\sum_{i=1}^{n} p(y|i)x^{(i)}}{\sum_{i=1}^{n} p(y|i)}$$

$$\hat{\sigma}_{y}^{2} = \frac{\sum_{i=1}^{n} p(y|i) \|x^{(i)} - \hat{\mu}^{(y)}\|^{2}}{d\sum_{i=1}^{n} p(y|i)}$$



What did we learn?

Modeling

- generative models are specified by variables and how the variables are related
- "latent variables" allow us to specify different underlying structures that we can (try to) uncover

Estimation (general)

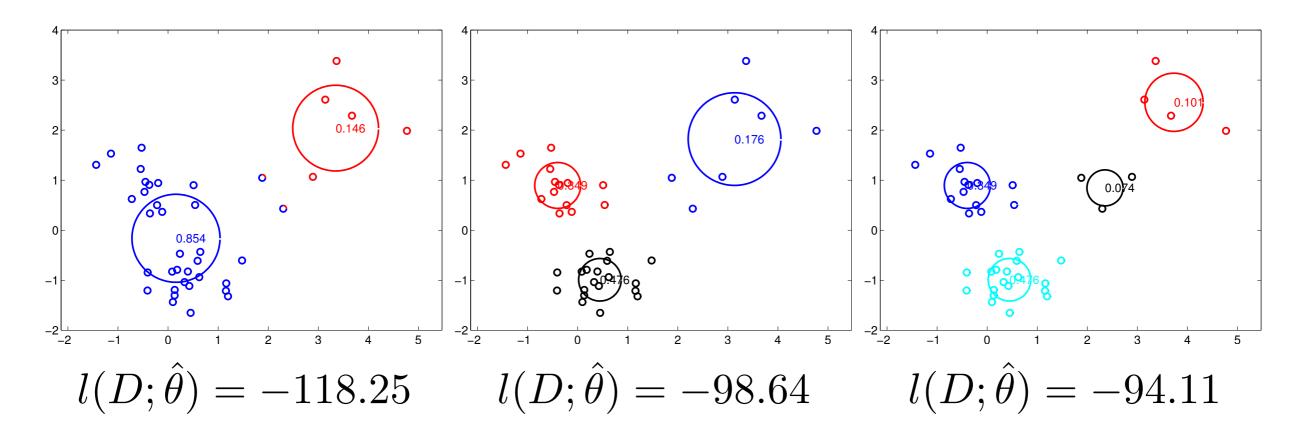
- models can be estimated "in pieces" if we have complete data, i.e., observations with a value assignment for each variable
- easy estimation extends to weighted complete data

Estimation (EM algorithm)

- EM algorithm iteratively creates weighted training sets (based on posterior assignments) and updates the model parameters, separately for each piece, based on the weighted data
- the EM algorithm applies the same to other generative models where y specifies different latent choices (more later)

Model selection

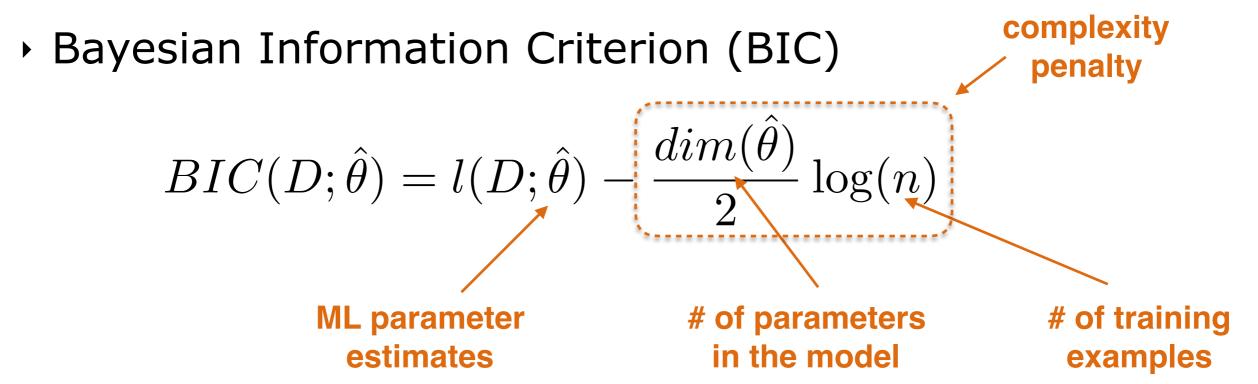
 We can run the EM-algorithm with different numbers of components. Need to specify a criterion for selecting among the different models.



 Basing the selecting on the value of log-likelihood would invariably lead to the largest number of components



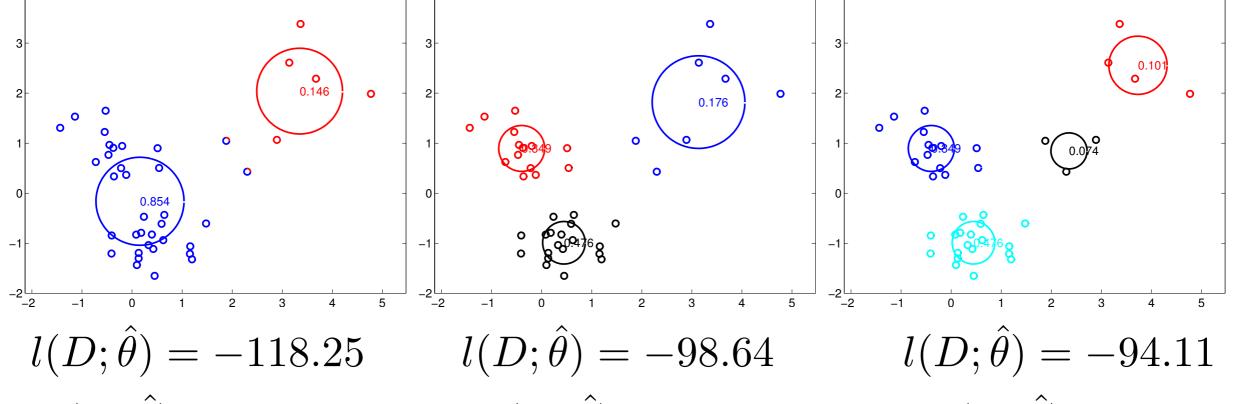
Model selection: BIC



- Many closely related ways to think about "complexity" penalties in model selection criteria
 - how excess parameters help fit the data even in the absence of any signal
 - the cost of communicating the model (likelihood tells us the cost of transmitting the data given the model)
 - fraction of the parameter space that explains the data



Mixture models and BIC



$$BIC(D; \hat{\theta}) = -131.16 \ BIC(D; \hat{\theta}) = -118.93 \ BIC(D; \hat{\theta}) = -121.78$$



Modeling sequences

- Lots of interesting data come in the form of sequences
 - temporal data (financial, monitoring, speech)
 - languages
 - user behavior
 - bio-sequences
 - etc.
- Our goal is to model such sequences, i.e., specify and learn probability distributions over sequences
 - specification (how to define, parameterize)
 - sampling (understand as a generative model)
 - estimation (learn from data)

Modeling sequences

Consider a sequence of (e.g., binary) variables

$$y_1$$
 y_2 y_3 y_4 \dots $y_i \in V$

 We wish to specify a probability distribution over their values. By the chain rule, we can always write

$$P(y_1, \dots, y_n) = P(y_1)P(y_2|y_1)P(y_3|y_2, y_1) \cdots P(y_n|y_{n-1}, \dots, y_1)$$

but, without any assumptions, we would need very large probability tables |V|^n

The 1st order Markov assumption (bigram model):

$$P(y_1, \dots, y_n) \stackrel{def}{=} P(y_1)P(y_2|y_1)P(y_3|y_2)\cdots P(y_n|y_{n-1})$$



Markov Model

 A 1st order (homogenous) Markov Model requires us to specify two sets of distributions

$$y_1 \longrightarrow y_2 \longrightarrow \cdots \longrightarrow y_{n-1} \longrightarrow y_n$$

Initial state distribution:

$$P(y_1 = y) = \pi_y, \ \sum_{y=1}^K \pi_y = 1$$

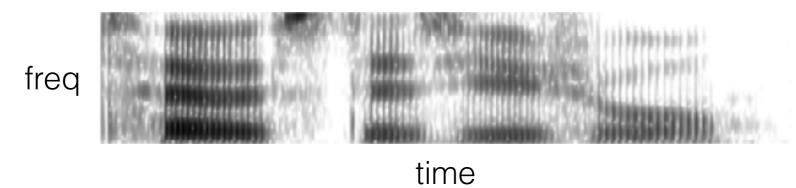
State transition probabilities:

$$P(y_i = y'|y_j = y) = a_{y,y'}, \sum_{y'=1}^{K} a_{y,y'} = 1 \ \forall y$$



Modeling hidden sequences

Speech recognition



Handwriting recognition

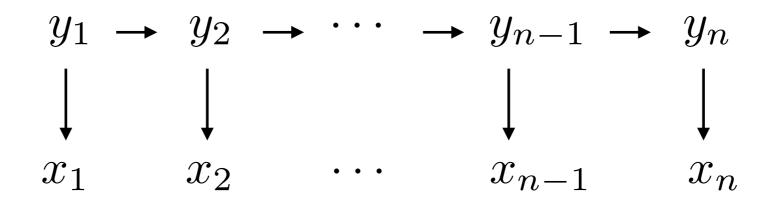
Machine learning is fun

- Information extraction, part of speech tagging, etc.
- Bio-sequence annotation
- Etc.



Modeling hidden sequences

 Similarly to mixture models, we must connect the latent selections to actual observations

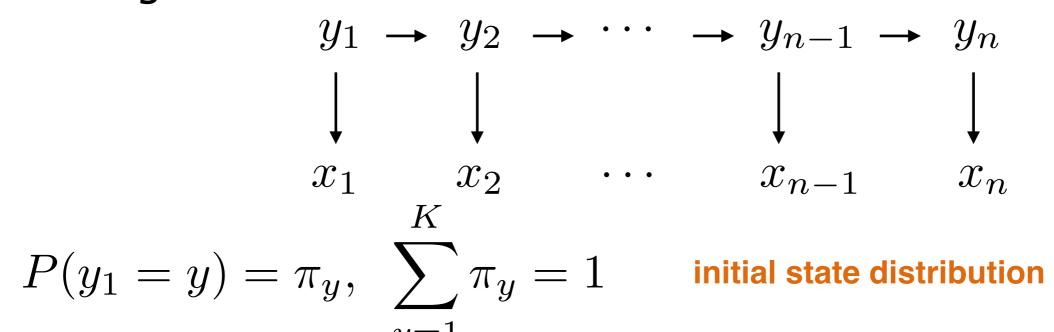


 If the latent sequence is Markov, and the observations are only tied to the state at the corresponding time, we get a Hidden Markov Model (HMM)

$$P(y_1, \dots, y_n, x_1, \dots, x_n) = P(y_1)P(x_1|y_1) \prod_{i=2}^n [P(y_i|y_{i-1})P(x_i|y_i)]$$



We need three sets of probabilities to specify a homogeneous HMM



$$P(y_1=y)=\pi_y, \quad \sum_{y=1}^K \pi_y=1$$
 initial state distribution

$$P(y_i=y'|y_{i-1}=y)=a_{y,y'}, \quad \sum_{y'=1}^{K}a_{y,y'}=1 \quad \forall y \quad \text{state transition}$$

$$P(x_i = x | y_i = y) = \begin{cases} N(x; \mu^{(y)}, \sigma_y^2 I) \\ \text{or discrete} \end{cases} = b_y(x) \quad \text{observation} \quad \text{model}$$



HMM

 We need three sets of probabilities to specify a homogeneous HMM

$$P(y_1=y)=\pi_y, \quad \sum_{y=1}^K \pi_y=1$$
 initial state distribution

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$$P(y_1, \dots, y_n, x_1, \dots, x_n; \theta) = \pi_{y_1} b_{y_1}(x_1) \prod_{i=2} [a_{y_{i-1}, y_i} b_{y_i}(x_i)]$$

HMM problems to solve

Evaluation: calculate the probability of observed values,
 i.e., sum over all the underlying state sequences

$$P(x_1, \dots, x_n) = \sum_{y_1, \dots, y_n} P(y_1) P(x_1|y_1) \prod_{i=2}^n \left[P(y_i|y_{i-1}) P(x_i|y_i) \right]$$

- Estimation: EM algorithm for HMMs
 - E-step now needs a little effort...
 - M-step is easy (separate estimation of pieces)
- ✓ Prediction: find the most likely hidden state sequence responsible for the observations (Viterbi)

$$\hat{y}_1, \dots, \hat{y}_n = \arg\max_{y_1, \dots, y_n} P(y_1, \dots, y_n, x_1, \dots, x_n)$$



Viterbi algorithm

 We can find the most likely hidden sequence using dynamic programming (focusing first on the max value)

$$d_i(y_i) = \max_{y_1, \dots, y_{i-1}} P(y_1, \dots, y_i, x_1, \dots, x_i)$$

Recursion (should be implemented on a log-scale)

$$d_1(y_1) = P(y_1)P(x_1|y_1)$$

$$d_i(y_i) = \max_{y_{i-1}} \left\{ d_{i-1}(y_{i-1})P(y_i|y_{i-1})P(x_i|y_i) \right\}$$

The end result of the recursion is that

$$\max_{y_n} d_n(y_n) = \max_{y_1, ..., y_n} P(y_1, ..., y_n, x_1, ..., x_n)$$



Viterbi, backtracking

 We can reconstitute the maximizing sequence by working backwards, instantiating the argmax's

$$\hat{y}_n = \arg\max_{y_n} d_n(y_n)$$

$$\hat{y}_{i-1} = \arg\max_{y_{i-1}} \left\{ d_{i-1}(y_{i-1}) P(\hat{y}_i | y_{i-1}) P(x_i | y_i) \right\}$$



Key things to know

- Mixture models
 - specification, sampling, estimation
 - the EM algorithm (for mixtures)
- Markov Models
 - specification, sampling, ML estimation
- Hidden Markov Models
 - specification, sampling, (estimation qualitatively)
 - Viterbi algorithm for prediction