1. In a second-order Markov chain, each state depends on the two previous states, i.e.,

$$P\left[X^{(t+1)} = S_k | X^{(t)} = S_j, X^{(t-1)} = S_i, \cdots\right] = P\left[X^{(t+1)} = S_k | X^{(t)} = S_j, X^{(t-1)} = S_i\right].$$

Show that the second-order Markov chain can always be converted to a first-order Markov chain. (Hint: by redesign the states)

A1:If State={A,C,T,G,...}, we can redesign $Z^t=\{X^t,X^{t-1}\}$ To prove the $P(Z^t|Z^{t-1})$ is exist. I show the $P(Z^t|Z^{t-1})$ by simple permutation at here.(just {A,C,T,G})

1	AA AC	AT AG	CA		\dots GG
AA					
CA	A		0	0	0
TA					
GA					
AC					
:	0		C	0	0
:					
:	0		0	T	0
	0		0	0	G
GG					

Eg:
$$P(Z^t = CA|Z^{t-1} = AT) = P(C|A)$$

2. Prove that a Bayesian network must be a Directed Acyclic Graph (DAG).

A2:At Bayesian Network , we can let $P(X_1, X_2, ..., X_N)$ be factorized into $\prod_{i=1}^N P(X_i | Parient(X_i))$. And draw a factorization as a graph. (if the node X_i don't have the Parients. We can use the notation $P(X_i)$) I assume the graph of Bayesian Network is a Directed \underline{cyclic} Graph , and then I got a contradiction because the factorization can't be finished...(every node has a parent but I just have N node.)

So the graph of Bayesian Network is a Directed Acyclic Graph.

- Given random variables A, B, C, and D, answer true or false and justify your answer:
 - (a) $\{A\} \perp \!\!\! \perp \{B\} | \{C\} \text{ implies } \{A\} \perp \!\!\! \perp \{B\};$
 - (b) $\{A\} \perp \!\!\!\perp \{B\} \text{ implies } \{A\} \perp \!\!\!\perp \{B\} | \{C\};$
 - (c) $\{A\} \perp \!\!\!\perp \{B,C\} | \{D\} \text{ implies } \{A\} \perp \!\!\!\perp \{B\} | \{D\}.$

A3(a): false , because we just need to consider the situations

1.tail to tail
$$\rightarrow P(A,B) = \int P(A,B,C)d_C = \int P(A|C)P(B|C)P(C)d_C \neq P(A)P(B)$$

2.head to tail ->
$$P(A,B) = \int P(A,B,C)d_C = \int P(B|C)P(C|A)P(A)d_C \neq P(A)P(B)$$

A3(b): false, because we just need to consider the situation

1. head to head -> P(A,B|C)=
$$\frac{P(A,B,C)}{P(C)} = \frac{\frac{P(A,B,C)}{P(A,B)}*P(A,B)}{P(C)} = \frac{P(C|A,B)*P(A)*P(B)}{P(C)} \neq P(A|C)P(B|C)$$

A3(c): yes___, because we just need to consider the situations

1. tail to tail ->
$$P(A,B|D) = \frac{P(A,B,D)}{P(D)} = \frac{P(A,B|D)P(D)}{P(D)} = \frac{P(A|D)P(B|D)P(D)}{P(D)} = P(A|D)P(B|D)$$

2. head to tail ->P(A,B|D)=
$$\frac{P(A,B,D)}{P(D)} = \frac{P(A|B,D)P(B,D)}{P(D)} = \frac{P(A|D)P(B,D)}{P(D)} = \frac{P(D|A)P(A)P(B|D)}{P(D)} = P(B|D)P(A|D)$$

4. Given a Hidden Markov model with time homogeneous Gaussian emission probability $P[x^{(t)}|z_i^{(t)}, \theta_i] = \frac{1}{(2\pi)^{d/2}det(\Sigma_i)^{1/2}}e^{-\frac{1}{2}(x^{(t)}-\mu_i)^T}\Sigma_i^{-1}(x^{(t)}-\mu_i)dx$, where $\theta_i = (\mu_i, \Sigma_i)$. Consider the problem finding $\Theta = (\pi^{(1)}, A, \{\theta_k\}_{k=1}^K)$ using the EM algorithm. Show that maximizing $\mathcal{Q}(\Theta; \Theta^{old})$ in the M-step gives $\mu_i = \frac{\sum_{t=1}^T \gamma_i^{(t)}x^{(t)}}{\sum_{t=1}^T \gamma_i^{(t)}}$ and $\Sigma_i = \frac{\sum_{t=1}^T \gamma_i^{(t)}(x^{(t)}-\mu_i)(x^{(t)}-\mu_i)^T}{\sum_{t=1}^T \gamma_i^{(t)}}$.

A4: let
$$\gamma_i = \left\{ egin{matrix} 1 & x^t & is & state \ i \\ 0 & Others \end{array} \right.$$
 , and $N_i = \sum_{t=1}^T {\gamma_i}^t$

To Max
$$\log P(X|Z_i, \Theta_i) = \log \left(\prod_{t=1}^T \left(\left(\frac{1}{(2\pi)^{\frac{d}{2}} \cdot \det(\Sigma_i)^{\frac{1}{2}}} \right) e^{-\frac{1}{2}(x^t - u_i)^T \Sigma_i^{-1}(x^t - u_i)} \right)^{\gamma_i^t} \right)$$

$$= -\frac{N_i d}{2} \log(2\pi) - \frac{N_i}{2} \log \left(det(\Sigma_i) \right) - \frac{1}{2} \sum_{t=1}^{T} \gamma_i^{\ t} (x^t - u_i)^T \Sigma_i^{\ -1} (x^t - u_i)$$

$$\frac{d \log P(\mathbf{X}|Z_k,\Theta_k)}{du} = -\sum_{t=1}^{T} \gamma_i^t (\mathbf{x}^t -)^T \Sigma_i^{-1} = 0 \quad \rightarrow \quad \mathbf{u}_i = \frac{\sum_{t=1}^{T} \gamma_i^t \mathbf{x}^t}{\sum_{t=1}^{T} \gamma_i^t}$$

$$\frac{d \log P(X|Z_i, \Theta_i)}{d\Sigma_i^{-1}} = \frac{N_i}{2} \Sigma_i - \frac{1}{2} \sum_{t=1}^{T} \gamma_i^{t} (x^t - u_i)(x^t - u_i)^T = 0$$