TD 2: Research Design ECO 567A

Geoffrey Barrows¹

¹CREST, CNRS, Ecole polytechnique, Université Paris-Saclay

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- Unbiasedness: Expectation of the sampling distribution is the true value
- ▶ Effiency: OLS has the smallest variance among unbiased linear estimators

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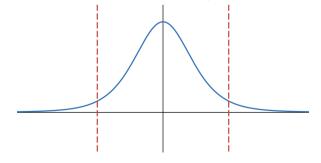
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► This allows hypothesis testing:

$$H0: \theta = \theta_0 \text{ vs. } H1: \left\{ egin{array}{ll} \theta > \theta_0 & (i) \\ \theta < \theta_0 & (ii) \end{array}
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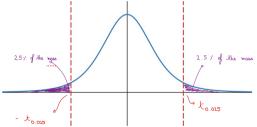
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- Choose a confidence level, say 95%, and then:

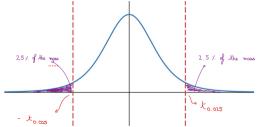


► E.g., if $\frac{\hat{\theta}_{OLS} - \theta_0}{\sec(\hat{\theta}_{OLS})} > t_{0.025}$, accept H1(i) and reject H0

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- ► E.g., if $\frac{\hat{\theta}_{OLS} \theta_0}{\text{se}(\hat{\theta}_{OLS})} > t_{0.025}$, accept H1 (i) and reject H0
- ▶ 95% confidence interval: $(\hat{\theta}_{OLS} \pm t_{.025} * se(\hat{\theta}_{ols}))$

But...

► Can we assume by default?

$$\varepsilon | X \sim \mathcal{N}\left(0, \sigma^2 I\right)$$

▶ Or more generally, can we assume?

$$E(\varepsilon|X)=0$$

► Today: (i) why we need a **research design** and (ii) how it could look like

Outline

```
When Do We Have E\left[\varepsilon|X\right] \neq 0?
Omitted Variable Bias
Reverse Causality / Simultaneity
In the Potential Outcome Framework
```

What Happens When $E[\varepsilon|X] \neq 0$?

What To Do When $E[\varepsilon|X] \neq 0$? Instrumental Variables Randomized Control Trials (RCTs) Fixed Effects Difference-in-Differences

Review Question

When Do We Have $E[\varepsilon|X] \neq 0$?

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Omitted Variable Bias

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▶ And, by doing so, we "forget" about parents' education z:

$$\begin{cases} u_i &= \gamma z_i + \epsilon_i \\ x_i &= \lambda + \mu z_i + \eta_i \end{cases}$$

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$$\begin{cases} u_i &= \gamma z_i + \epsilon_i \\ x_i &= \lambda + \mu z_i + \eta_i \end{cases}$$

- Parents' education directly affects income: Better professional network, etc.
- Parents' education also affects child's education: Role model, information, etc.

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$$= \mu \gamma Var(z_i, z_i) \neq 0$$

▶ When we estimate our wrong model, we have:

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- In the above, we see that there is no issue if:
 - $\mu = 0$: Omitted variable is unrelated to education
 - $ightharpoonup \gamma = 0$: Omitted variable does not affect income directly
 - $ightharpoonup \operatorname{Var}(z_i,z_i)=0$: Omitted variable is a constant (not super interesting)

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- Places with more pollution also have more labor demand
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- And hence better health outcomes
- ▶ So, "income" is an omitted variable related to both pollution and health

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- ▶ So, "expected growth" could be a problematic omitted variable

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▶ But we "forget" that criminality also affects people's movements:

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- Crowd density has a causal effect on crime: Deterrence?
- Crime has a causal effect on density: Lower attractivity

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 - $\delta = 0$, which essentially denies reverse causality

When Do We Have $E[\varepsilon|X] \neq 0$?

In the Potential Outcome Framework

Introducing the Potential Outcome Framework

▶ Imagine we study a treatment (e.g., ban on polluting cars)

Treatment
$$X_i = \begin{cases} 1 & \text{if treated} \\ 0 & \text{if otherwise} \end{cases}$$

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Potential Outcome =
$$\begin{cases} Y_{1i} & \text{if } X_i = 1 \\ Y_{0i} & \text{if } X_i = 0 \end{cases}$$

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$$\begin{cases} Y_{1i} & \text{if } X_i = 1 \\ Y_{0i} & \text{if } X_i = 0 \end{cases}$$

We only observe the outcome that is actually realized

Observed Outcome
$$Y_i = \begin{cases} Y_{1i} & \text{if } X_i = 1 \\ Y_{0i} & \text{if } X_i = 0 \end{cases} = Y_{0i} + \underbrace{(Y_{1i} - Y_{0i})}_{\text{Treatment effect}} *X_i$$

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Amounts to:
$$\hat{\beta} = \underbrace{\hat{E}(Y_i|X_i=1)}_{\text{Mean of }Y \text{ among }X=1} - \underbrace{\hat{E}(Y_i|X_i=0)}_{\text{Mean of }Y \text{ among }X=0}$$

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- \blacktriangleright Expectation of the sampling distribution: $E(\hat{\beta}) = E(Y_i|X_i=1) E(Y_i|X_i=0)$
 - Where: $E[Y_i|X_i=1] = \alpha + \beta + E[\varepsilon_i|X_i=1]$
 - And: $E[Y_i|X_i=0] = \alpha + E[\varepsilon_i|X_i=0]$

Estimate with OLS:

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 - And: $E[Y_i|X_i=0] = \alpha + E[\varepsilon_i|X_i=0]$
- We can assess unbiasedness:

$$E(\hat{\beta}) = E[Y_i|X_i = 1] - E[Y_i|X_i = 0]$$

$$= \beta + E[\varepsilon_i|X_i = 1] - E[\varepsilon_i|X_i = 0]$$

$$= \beta + \underbrace{E[Y_{0i}|X_i = 1] - E[Y_{0i}|X_i = 0]}_{\text{Risk of selection bias}}$$

What Happens When $E\left[\varepsilon|X\right] \neq 0$?

Monte Carlo Experiment

► Assume a true model

$$y_i = \alpha + \beta * x_i + a_i + \varepsilon_i$$

 $\alpha = 10$
 $\beta = 2$

Monte Carlo Experiment

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$$y_i = \alpha + \beta * x_i + a_i + \varepsilon_i$$

 $\alpha = 10$
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- Examine three cases:
 - Case 1: Exogenous a
 - Variable is omitted but it is unrelated to the regressor of interest
 - Case 2: Endogenous a, and we don't know it
 - ▶ Variable is omitted, and it is related to the regressor of interest
 - Case 3: Endogenous a, but we know it
 - Variable is related to the regressor of interest, but it is not omitted

► True model:

$$y_i = \alpha + \beta x_i + a_i + \varepsilon_i$$

$$E(a + \varepsilon | x) = E(a | x) + E(\varepsilon | x) = E(a | x) = 0$$

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Estimate:

$$y_i = \alpha + \beta x_i + \underbrace{u_i}_{=a_i + \varepsilon_i}$$

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Prediction:

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$$E(a + \varepsilon | x) = E(a | x) + E(\varepsilon | x) = E(a | x) = 0$$

Estimate:

$$y_i = \alpha + \beta x_i + \underbrace{u_i}_{=a_i + \varepsilon_i}$$

Prediction: OLS estimator should be unbiased

Case 1: Exogenous ability — Experiment

► Monte Carlo experiment:

$$x_i \sim \mathcal{U}(0,20)$$

 $a_i \sim \mathcal{N}(0,5)$
 $\varepsilon_i \sim \mathcal{N}(0,5)$
 $y_i = \alpha + \beta x_i + a_i + \varepsilon_i$

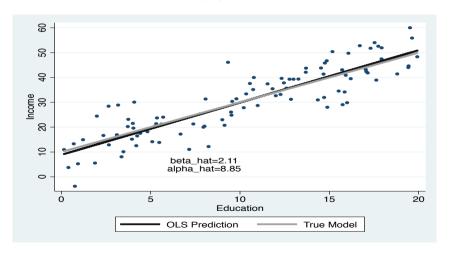
Data:

Table: Summary statistics

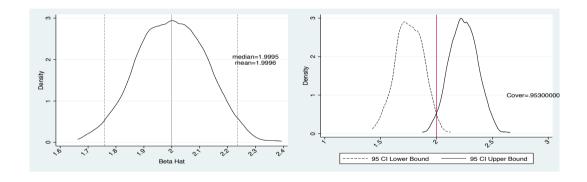
Variable	Mean	Std. Dev.	Min.	Max.	N
у	30.268	13.734	-3.795	60.018	100
×	10.169	5.758	0.156	19.923	100
а	0.012	4.899	-13.5	12.748	100
eps	-0.083	5.251	-11.975	11.012	100

Case 1: Exogenous ability — Results

Estimation equation:
$$y_i = \alpha + \beta x_i + \underbrace{u_i}_{=a_i+\epsilon_i}$$



Case 1: Exogenous ability — Results



► True model:

$$y_i = \alpha_1 + \beta x_i + a_i + \varepsilon_i$$

$$x_i = \alpha_2 + \delta a_i + \eta_i$$

$$\delta = 0.5$$

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Estimate:

$$y_i = \alpha_1 + \beta x_i + \underbrace{u_i}_{=a_i + \varepsilon_i}$$

Prediction: OLS estimator is likely biased

Case 2: Endogenous a, and we don't know it — Experiment

► Monte Carlo experiment:

$$\eta \sim N(0,1)$$
; $a_i \sim N(0,5)$; $\varepsilon_i \sim N(0,5)$
 $x_i = \alpha_2 + \delta a_i + \eta_i$
 $y_i = \alpha_1 + \beta x_i + a_i + \varepsilon_i$

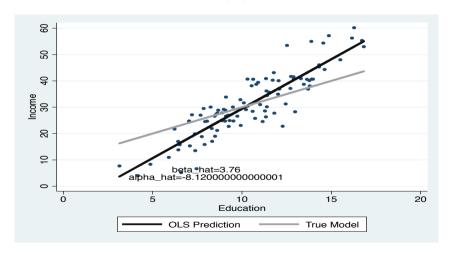
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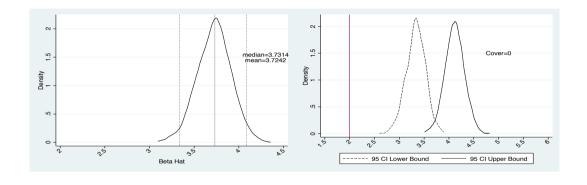
Variable	Mean	Std. Dev.	Min.	Max.	N
у	31.212	12.035	4.037	60.175	100
Χ	10.472	2.852	3.142	16.819	100
a	0.940	5.418	-11.25	14.065	100
eps	-0.672	4.947	-13.5	10.32	100
eta	0.002	1.02	-2.068	2.252	100

Case 2: Endogenous a, and we don't know it — Results

Estimation Equation:
$$y_i = \alpha_1 + \beta * x_i + \underbrace{u_i}_{=a_i+\varepsilon_i}$$



Case 2: Endogenous a, and we don't know it — Results



► True Model

$$y_i = \alpha_1 + \beta x_i + a_i + \varepsilon_i$$

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► True Model

$$y_i = \alpha_1 + \beta x_i + a_i + \varepsilon_i$$

$$x_i = \alpha_2 + \delta a_i + \eta_i$$

$$\delta = 0.5$$

Estimate:

$$y_i = \alpha_1 + \beta * x_i + a_i + \varepsilon_i$$

► True Model

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Estimate:

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Prediction: OLS estimator should be unbiased

Case 3: Endogenous a, but we know it — Experiment

► Monte Carlo experiment:

$$\eta \sim N(0,1)$$
; $a_i \sim N(0,5)$; $\varepsilon_i \sim N(0,5)$
 $x_i = \alpha_2 + \delta a_i + \eta_i$
 $y_i = \alpha_1 + \beta x_i + a_i + \varepsilon_i$

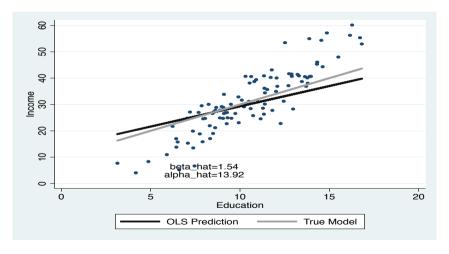
Data:

Table: Summary statistics

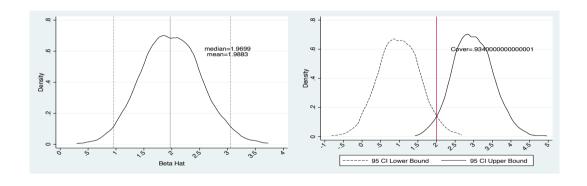
Variable	Mean	Std. Dev.	Min.	Max.	N
у	31.212	12.035	4.037	60.175	100
Χ	10.472	2.852	3.142	16.819	100
a	0.940	5.418	-11.25	14.065	100
eps	-0.672	4.947	-13.5	10.32	100
eta	0.002	1.02	-2.068	2.252	100

Case 3: Endogenous a, but we know it — Results

Estimation equation: $y_i = \alpha_1 + \beta x_i + a_i + \varepsilon_i$



Case 3: Endogenous a, but we know it — Results



What To Do When $E[\varepsilon|X] \neq 0$?

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Instrumental Variables

Context

▶ Suppose $E[\varepsilon|X] \neq 0$. E.g., x correlated with a, which also determines y:

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Then, we fall into omitted variable bias, and OLS is biased

- Instrumental variable approach:
 - \triangleright Find another z variable which correlates with x...
 - ...but not with a...
 - \triangleright ...and use it to recover $\beta!$

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- ► E.g., subway station maintenance operations
 - Strong first stage: Clear negative correlation with crowd density
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- Strong first stage can be tested, not the exclusion restriction in general

In Practice: Two-stage Least Squares (2SLS)

1. First, use z to predict x:

$$x_{i} = \pi_{1} + \pi_{2}z_{i} + \underbrace{\gamma_{i}}_{=\delta a_{i} + \eta_{i}}$$

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- ► Then: $E(u_i|\hat{x}_i) = E(a_i + \varepsilon_i|\hat{\pi}_1 + \hat{\pi}_2 * z_i) = E(a_i|\hat{\pi}_1 + \hat{\pi}_2 * z_i) + E(\varepsilon_i|\hat{\pi}_1 + \hat{\pi}_2 * z_i) = 0$
- \hat{x} uncorrelated with a: Extracted the random component in x, and **OLS** is unbiased

► True model:

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Estimate:

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Prediction: 2SLS estimator should be unbiased

Monte Carlo Experiment — Specification

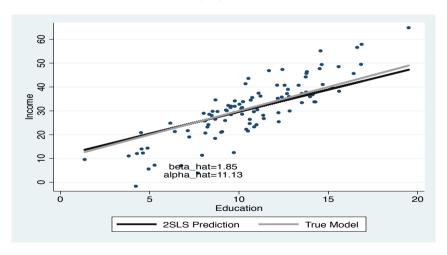
$$\eta \sim N(0,1)$$
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 $a_i \sim N(0,5)$
 $z_i \sim N(0,2)$
 $x_i = \alpha_2 + \delta * a_i + z_i + \eta_i$
 $y_i = \alpha_1 + \beta * x_i + a_i + \varepsilon_i$

Table: Summary statistics

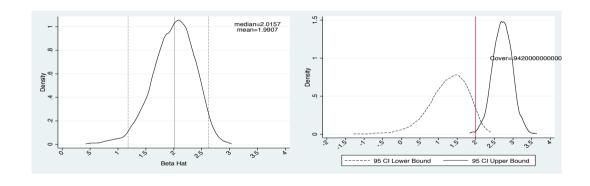
Variable	Mean	Std. Dev.	Min.	Max.	N
у	30.581	12.304	-1.62	64.849	100
Χ	10.502	3.345	1.364	19.494	100
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eps	-0.672	4.947	-13.5	10.32	100
eta	0.002	1.02	-2.068	2.252	100
Z	0.376	2.167	-4.5	5.626	100

Monte Carlo Experiment — Results

Estimation equation:
$$y_i = \alpha + \beta \hat{x}_i + \underbrace{u_i}_{=a_i+\epsilon_i}$$



Monte Carlo Experiment — Results



Monte Carlo Experiment — Results

	First Stage	Second Stage		
	X	у	у	у
Z	0.88***			
	(0.13)			
×		3.16***	1.97***	
		(0.19)	(0.21)	
a			1.11***	
			(0.14)	
â				1.85***
				(0.62)
Constant	10.17***	-2.61	9.59***	11.13*
	(0.28)	(2.09)	(2.26)	(6.62)
# Observations	100	100	100	100
R squared	0.329	0.738	0.840	0.083
Mean Dep. Var	10.502	30.581	30.581	30.581

What To Do When $E[\varepsilon|X] \neq 0$?

Randomized Control Trials (RCTs)

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 - Many possible confounders: Type of job, education, household composition, etc.
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- ► E.g., effect of disposable income on consumption expenditures¹
 - Many possible confounders: Type of job, education, household composition, etc.
 - Solution? Randomize (part of) disposable income thanks to direct transfers
- Sounds crazy?
 - Boehm, Fize, and Jaravel (2025) spent around €300,000
 - ➤ Some customers of a bank receive €300, some don't; allocation is random
 - ▶ Use the transfer as an instrument for income and recover the desired parameter

In Practice

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- 2019 Nobel Prize to Esther Duflo, Abhijit Banerjee, and Michael Kremer
- In terms of identification, the best we can do
- Potential limitations:
 - Feasibility / Cost
 - ► Ethical concerns: € transfers are OK, but could we randomize education choices?
 - External validity

What To Do When $E[\varepsilon|X] \neq 0$?

Fixed Effects

- ► We now move from cross-sectional to panel data
 - ▶ Instead of solely comparing units *i* (individuals, countries, cities, etc.) once in time
 - We observe different units *i* over time *t*

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Some unobserved, time-invariant heterogeneity a_i correlates with x_{it}

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Some unobserved, time-invariant heterogeneity a_i correlates with x_{it}

OLS is biased when estimated on:

$$y_{it} = \alpha + \beta * x_{it} + u_{it}$$

Within estimator:

Let:
$$\bar{y}_{it} = \frac{1}{T} \sum_{t=1}^{T} \left(\alpha + \beta x_{it} + \underbrace{a_i + \varepsilon_{it}}_{u_{it}} \right) = \alpha + \beta \bar{x}_{it} + a_i + \bar{\varepsilon}_{it}$$

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First-difference estimator:

Consider instead:

$$\Delta y_{it} = y_{it} - y_{it-1} = \beta \Delta x_{it} + \Delta \varepsilon_{it}$$

Within estimator:

- Let: $\bar{y}_{it} = \frac{1}{T} \sum_{t=1}^{T} \left(\alpha + \beta x_{it} + \underbrace{a_i + \varepsilon_{it}}_{u_{it}} \right) = \alpha + \beta \bar{x}_{it} + a_i + \bar{\varepsilon}_{it}$ Then: $v_{it} \bar{v}_{it} = \beta * (x_{it} \bar{x}_{it}) + (\varepsilon_{it} \bar{\varepsilon}_{it})$
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• Again, exogeneity condition is satisfied: $E(\Delta \varepsilon_{it} | \Delta x_{it}) = 0$

Fixed Effects

► First-difference estimator

$$\Delta y_{it} = \beta * \Delta x_{it} + \Delta \varepsilon_{it}$$

Is functionally equivalent to estimating fixed effect model:

$$y_{it} = \alpha + \beta x_{it} + \sum_{j=1}^{N} \alpha_j \mathbb{1}\{j=i\} + \varepsilon_{it}$$

l.e., controlling for a set of dummy variables for all units $\{\mathbb{1}\{j=i\}\}_{j\in[1;N]}$

▶ **NB**: To lighten notations, we often write that we estimate the model:

$$y_{it} = \alpha + \beta x_{it} + \alpha_i + \varepsilon_{it}$$

What To Do When $E[\varepsilon|X] \neq 0$?

Difference-in-Differences

Consider again panel data; suppose the true model is:

$$y_{it} = \alpha + \beta * x_{it} + \underbrace{a_i + \gamma_t + \varepsilon_{it}}_{u_{it}} \text{ with } E\left[\varepsilon_{it}|x_{it}\right] = 0 \text{ and } E\left[u_{it}|x_{it}\right] \neq 0$$

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- ▶ E.g., think about health (y_{it}) and pollution (x_{it})
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 - ▶ COVID (part of γ_t) affects all regions with an impact on x_{it} and y_{it}
- ▶ Again, OLS is biased when estimated on: $y_{it} = \alpha + \beta * x_{it} + u_{it}$

Spirit

► Take the first difference along the time dimension:

$$\Delta y_{it} = \beta \Delta x_{it} + \Delta \gamma_t + \Delta \varepsilon_{it}$$

► Take another difference between units:

$$\Delta y_{it} - \Delta y_{jt} = \beta \left(\Delta x_{it} - \Delta x_{jt} \right) + \left(\Delta \varepsilon_{it} - \Delta \varepsilon_{jt} \right)$$

▶ Applying OLS to this model yields an unbiased estimator of β if:

$$E\left(\Delta\varepsilon_{it}-\Delta\varepsilon_{jt}\right|\Delta x_{it}-\Delta x_{jt})=0$$

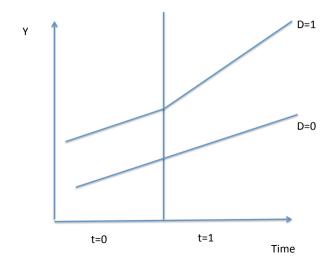
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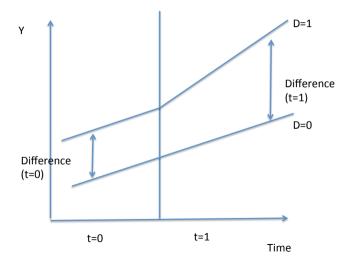
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- Assumption: Absent the treatment, both groups evolve similarly in post-period
 - ► Fundamental "parallel trends" assumption that researchers must defend
- Can we test this assumption?
 - We can check that both groups evolve similarly before the treatment
 - But we cannot test this after the treatment is deployed!

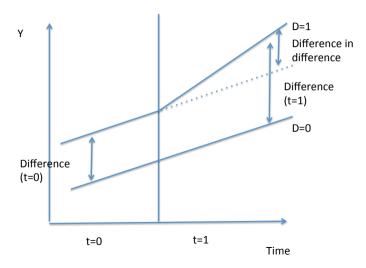
Graphical Illustration



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 - Under what condition is OLS unbiased?
 - Is this condition likely to be met?
 - ▶ What can you do to recover the causal impact of temperature on GDP?

References

▶ Joshua, D. Angrist. MOSTLY HARMLESS ECONOMETRICS. 2009.