

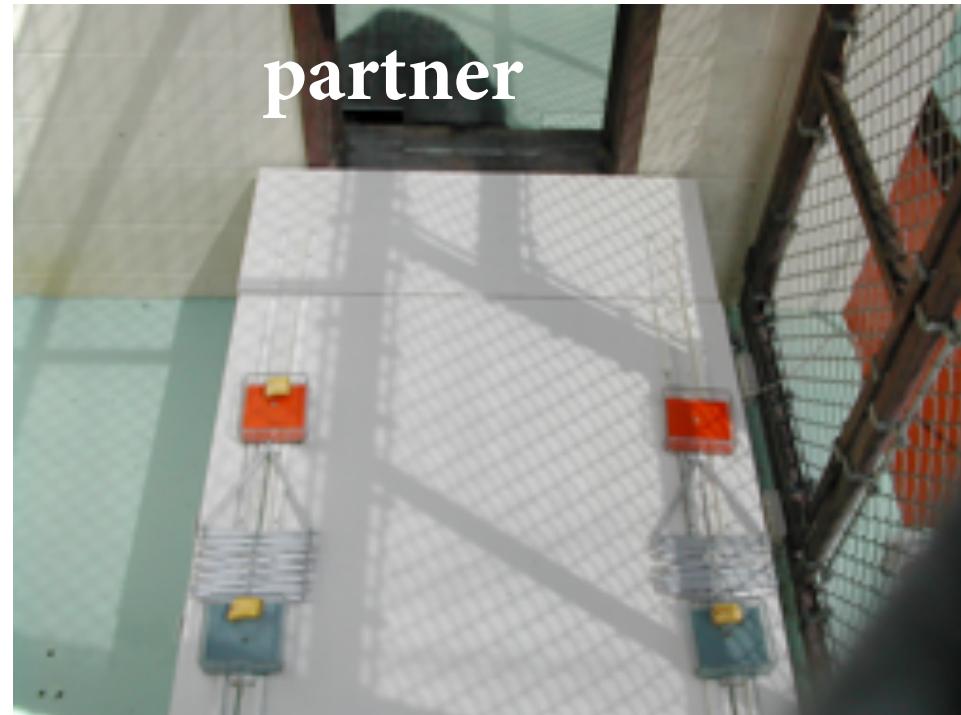
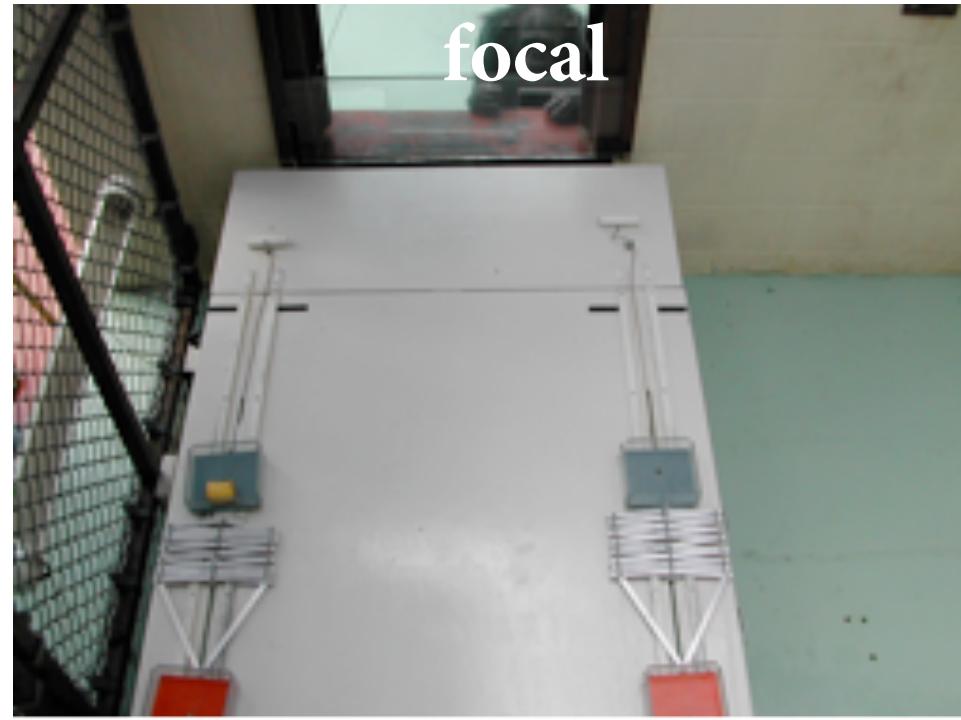
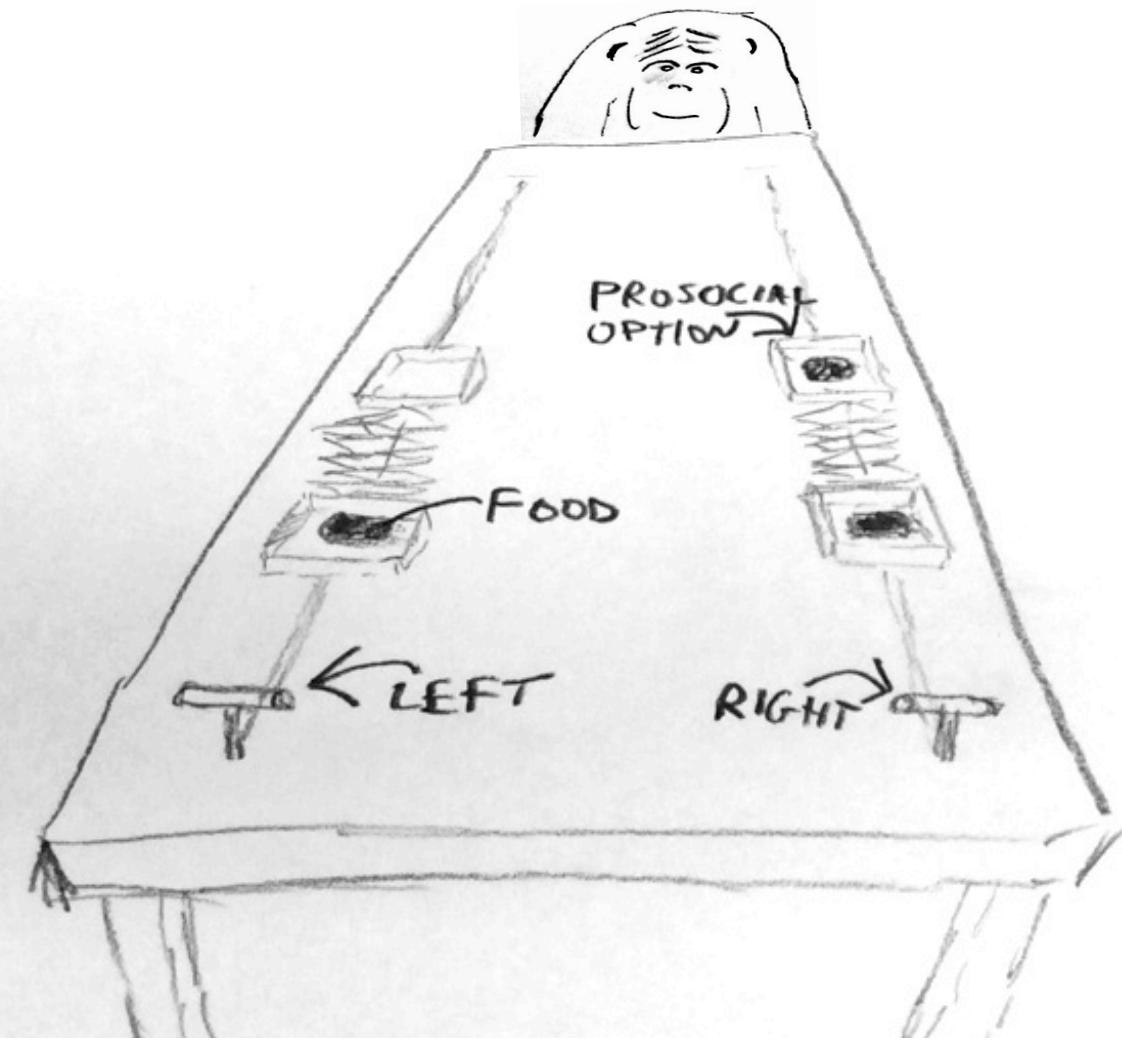
Statistical Rethinking

Winter 2019

Lecture 16 / Week 8

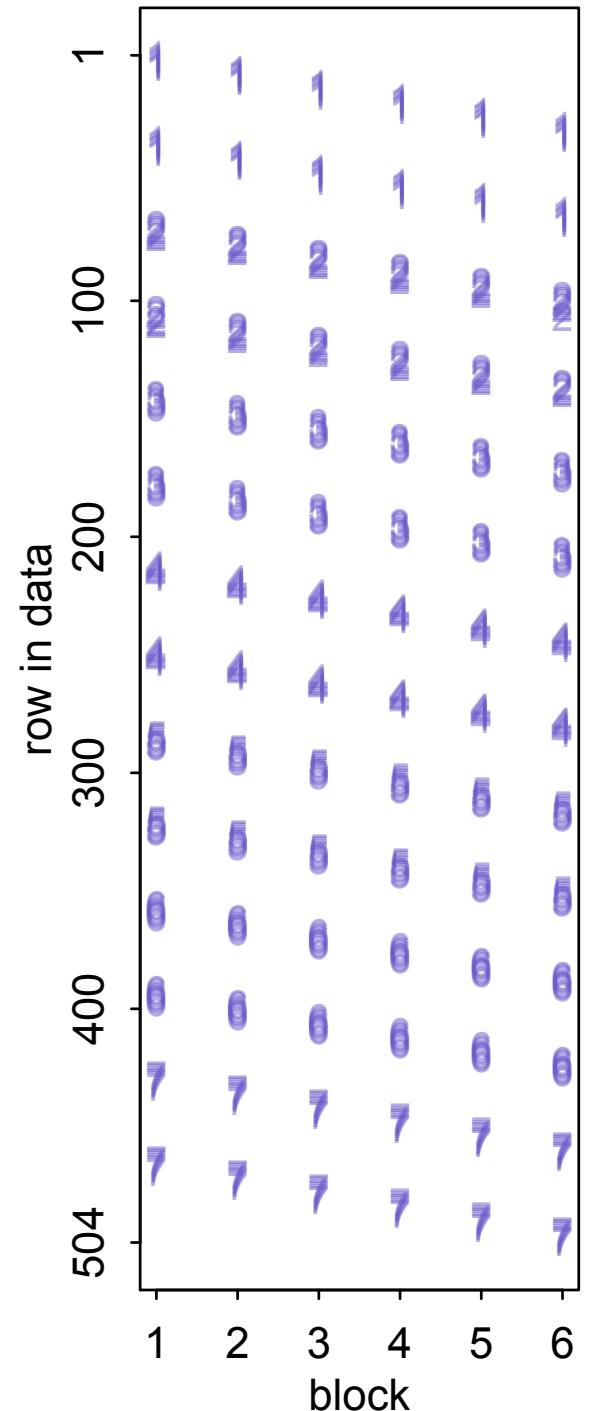
Multilevel Models

Prosocial chimpanzees



Cross-classification

- Can use more than one cluster type
- Chimpanzee experiment data
 - Pulls in chimpanzees
 - Pulls in blocks
 - Each chimp in each block
 - Not nested, but *cross-classified*



Multilevel chimpanzees

$$L_i \sim \text{Binomial}(1, p_i)$$

$$\text{logit}(p_i) = \alpha_{\text{ACTOR}[i]} + \gamma_{\text{BLOCK}[i]} + \beta_{\text{TREATMENT}[i]}$$

$$\beta_j \sim \text{Normal}(0, 0.5) \quad , \text{ for } j = 1..4$$

varying intercepts on actor —— $\alpha_j \sim \text{Normal}(\bar{\alpha}, \sigma_\alpha) \quad , \text{ for } j = 1..7$

$$\gamma_j \sim \text{Normal}(0, \sigma_\gamma) \quad , \text{ for } j = 1..6$$

$$\bar{\alpha} \sim \text{Normal}(0, 1.5)$$

$$\sigma_\alpha \sim \text{Exponential}(1)$$

$$\sigma_\gamma \sim \text{Exponential}(1)$$

Multilevel chimpanzees

$$L_i \sim \text{Binomial}(1, p_i)$$

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$$\beta_j \sim \text{Normal}(0, 0.5) \quad , \text{ for } j = 1..4$$

varying intercepts on actor $\longrightarrow \alpha_j \sim \text{Normal}(\bar{\alpha}, \sigma_\alpha) \quad , \text{ for } j = 1..7$

varying intercepts on block $\longrightarrow \gamma_j \sim \text{Normal}(0, \sigma_\gamma) \quad , \text{ for } j = 1..6$

$$\bar{\alpha} \sim \text{Normal}(0, 1.5)$$

$$\sigma_\alpha \sim \text{Exponential}(1)$$

$$\sigma_\gamma \sim \text{Exponential}(1)$$

Multilevel chimpanzees

```
m13.4 <- ulam(  
  alist(  
    pulled_left ~ dbinom( 1 , p ) ,  
    logit(p) <- a[actor] + g[block_id] + b[treatment] ,  
    b[treatment] ~ dnorm( 0 , 0.5 ) ,  
  
    # adaptive priors  
    a[actor] ~ dnorm( a_bar , sigma_a ) ,  
    g[block_id] ~ dnorm( 0 , sigma_g ) ,  
  
    # hyper-priors  
    a_bar ~ dnorm( 0 , 1.5 ) ,  
    sigma_a ~ dexp(1) ,  
    sigma_g ~ dexp(1)  
  ) , data=dat_list , chains=4 , cores=4 , log_lik=TRUE )
```

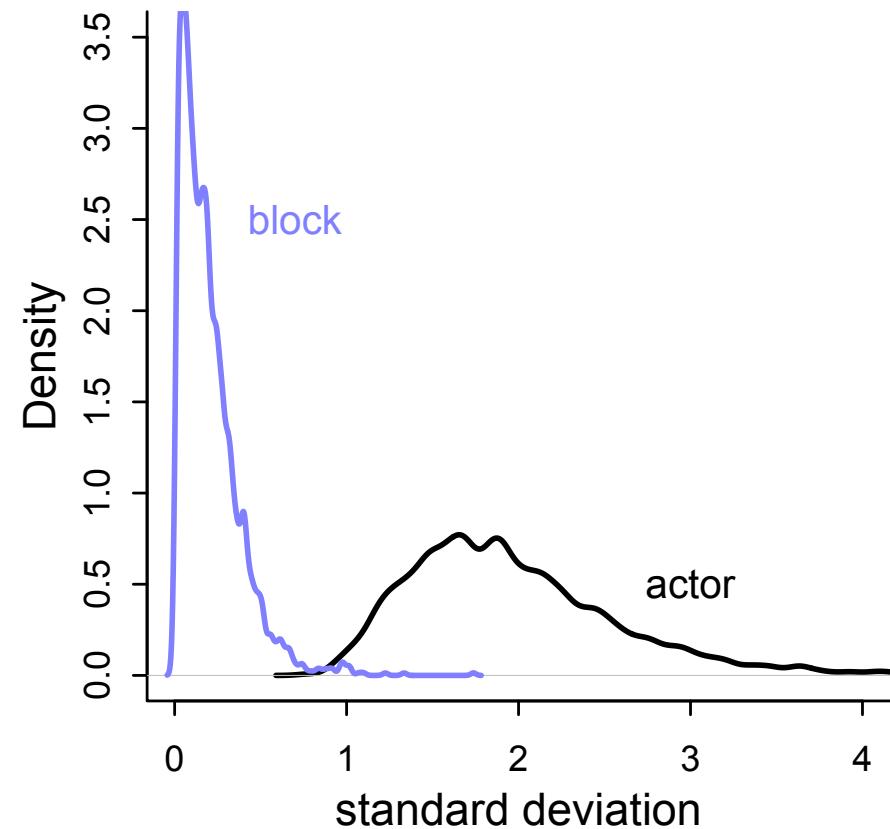
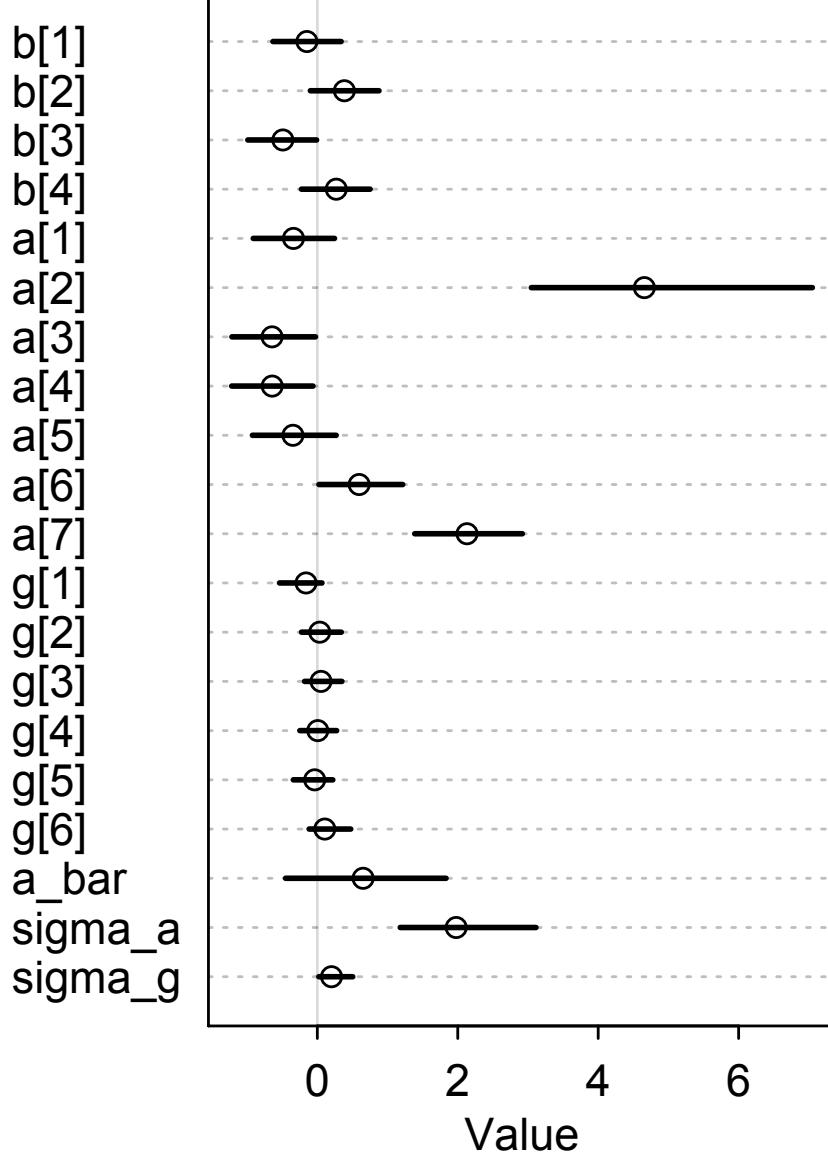
Multilevel chimpanzees

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  ) , data=dat_list , chains=4 , cores=4 , log_lik=TRUE )
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Multilevel chimpanzees

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  alist(  
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    a_bar ~ dnorm( 0 , 1.5 ) ,  
    sigma_a ~ dexp(1) ,  
    sigma_g ~ dexp(1)  
  ) , data=dat_list , chains=4 , cores=4 , log_lik=TRUE )
```

Cross-classified chimpanzees



Cross-classified chimpanzees

- Incorporating block: no benefits; little cost

```
set.seed(14)
m13.5 <- ulam(
  alist(
    pulled_left ~ dbinom( 1 , p ) ,
    logit(p) <- a[actor] + b[treatment] ,
    b[treatment] ~ dnorm( 0 , 0.5 ) ,
    a[actor] ~ dnorm( a_bar , sigma_a ) ,
    a_bar ~ dnorm( 0 , 1.5 ) ,
    sigma_a ~ dexp(1)
  ) , data=dat_list , chains=4 , cores=4 , log_lik=TRUE )
```

R code
13.23

R code
13.24 compare(m13.4 , m13.5)

	WAIC	pWAIC	dWAIC	weight	SE	dSE
m13.5	531.0	8.5	0	0.73	19.23	NA
m13.4	532.9	10.9	2	0.27	19.39	1.64

Everything is random

- “Random effects” have many definitions
- One common & bad definition:
 - Random effects are not fixed by the experiment
 - Does not matter—we want pooling for estimation benefits
 - Consider e.g. treatments

```
set.seed(15)
m13.6 <- ulam(
  alist(
    pulled_left ~ dbinom( 1 , p ) ,
    logit(p) <- a[actor] + g[block_id] + b[treatment] ,
    b[treatment] ~ dnorm( 0 , sigma_b ) ,
    a[actor] ~ dnorm( a_bar , sigma_a ) ,
    g[block_id] ~ dnorm( 0 , sigma_g ) ,
    a_bar ~ dnorm( 0 , 1.5 ) ,
    sigma_a ~ dexp(1) ,
    sigma_g ~ dexp(1) ,
    sigma_b ~ dexp(1)
```

R code
13.25

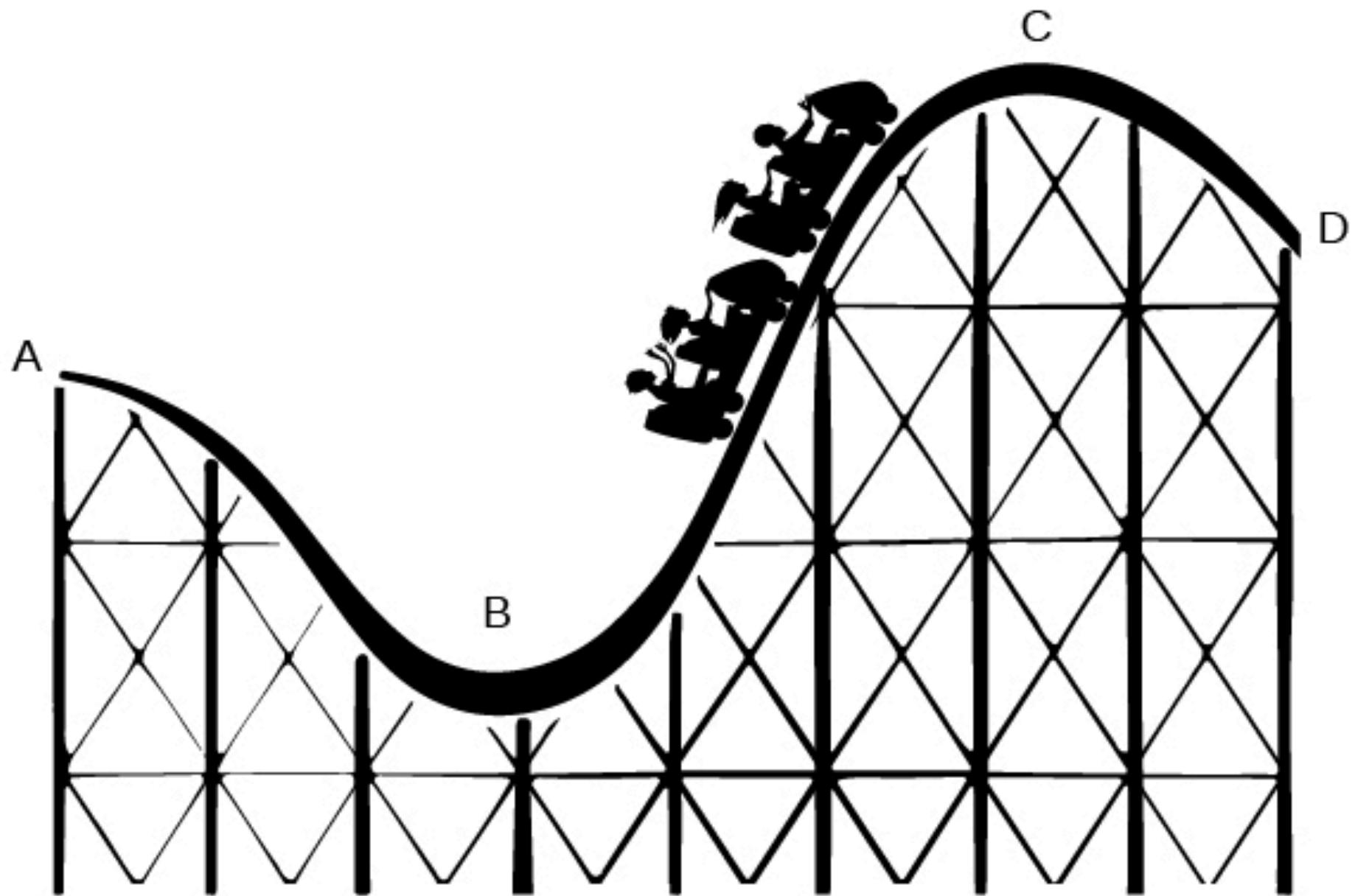
Everything is random

```
set.seed(15)
m13.6 <- ulam(
  alist(
    pulled_left ~ dbinom( 1 , p ) ,
    logit(p) <- a[actor] + g[block_id] + b[treatment] ,
    b[treatment] ~ dnorm( 0 , sigma_b ) ,
    a[actor] ~ dnorm( a_bar , sigma_a ) ,
    g[block_id] ~ dnorm( 0 , sigma_g ) ,
    a_bar ~ dnorm( 0 , 1.5 ) ,
    sigma_a ~ dexp(1) ,
    sigma_g ~ dexp(1) ,
    sigma_b ~ dexp(1)
  ) , data=dat_list , chains=4 , cores=4 , log_lik=TRUE )
coef(tab(m13.4,m13.6)
```

R code
13.25

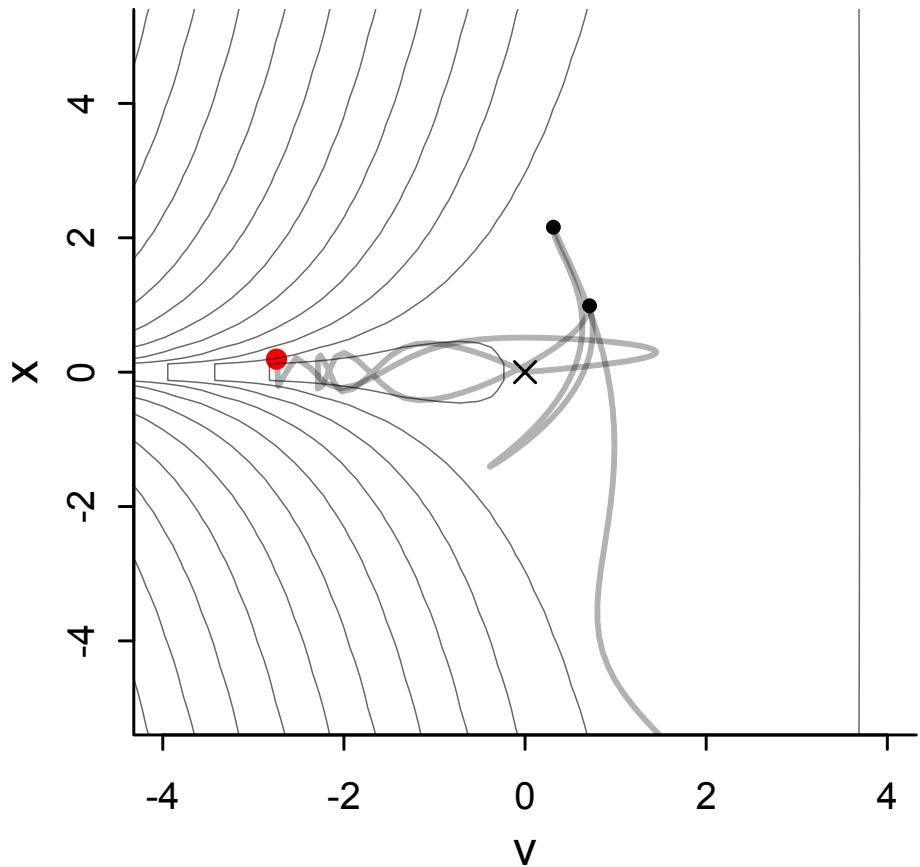
	m13.4	m13.6
b[1]	-0.11	-0.15
b[2]	0.41	0.34
b[3]	-0.46	-0.47
b[4]	0.31	0.25

```
mcelreath — R — 80x24  
Chain 2:          2.69352 seconds (Total)  
Chain 2:  
Chain 1: Iteration: 900 / 1000 [ 90%] (Sampling)  
Chain 1: Iteration: 1000 / 1000 [100%] (Sampling)  
Chain 1:  
Chain 1: Elapsed Time: 1.93964 seconds (Warm-up)  
Chain 1:          0.954394 seconds (Sampling)  
Chain 1:          2.89403 seconds (Total)  
Chain 1:  
Chain 4: Iteration: 800 / 1000 [ 80%] (Sampling)  
Chain 4: Iteration: 900 / 1000 [ 90%] (Sampling)  
Chain 4: Iteration: 1000 / 1000 [100%] (Sampling)  
Chain 4:  
Chain 4: Elapsed Time: 1.65333 seconds (Warm-up)  
Chain 4:          1.99443 seconds (Sampling)  
Chain 4:          3.64777 seconds (Total)  
Chain 4:  
Warning messages:  
1: There were 3 divergent transitions after warmup. Increasing adapt_delta above  
  0.95 may help. See  
http://mc-stan.org/misc/warnings.html#divergent-transitions-after-warmup  
2: Examine the pairs() plot to diagnose sampling problems  
>
```



Divergent transitions

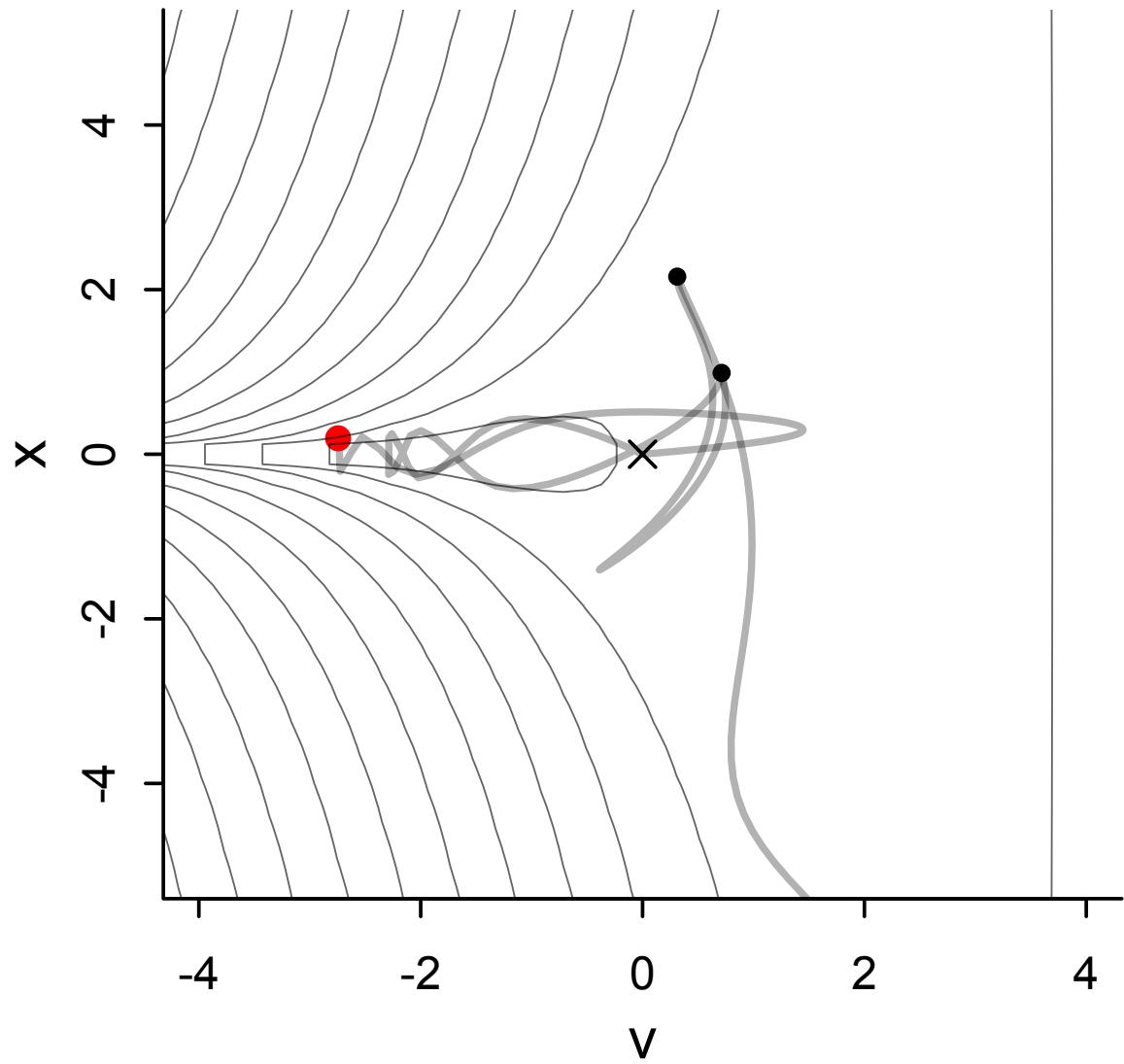
- HMC runs a physics simulation
- Each *transition* is a sample path
- In real physics, energy is conserved
- If energy at end of transition is not equal to energy at start, transition is *divergent*
- Indicates inaccurate approximation
- Tends to happen in regions of strong curvature of log-posterior
- **Other sampling strategies also bad in these cases, but produce no warnings!**



Divergent transitions

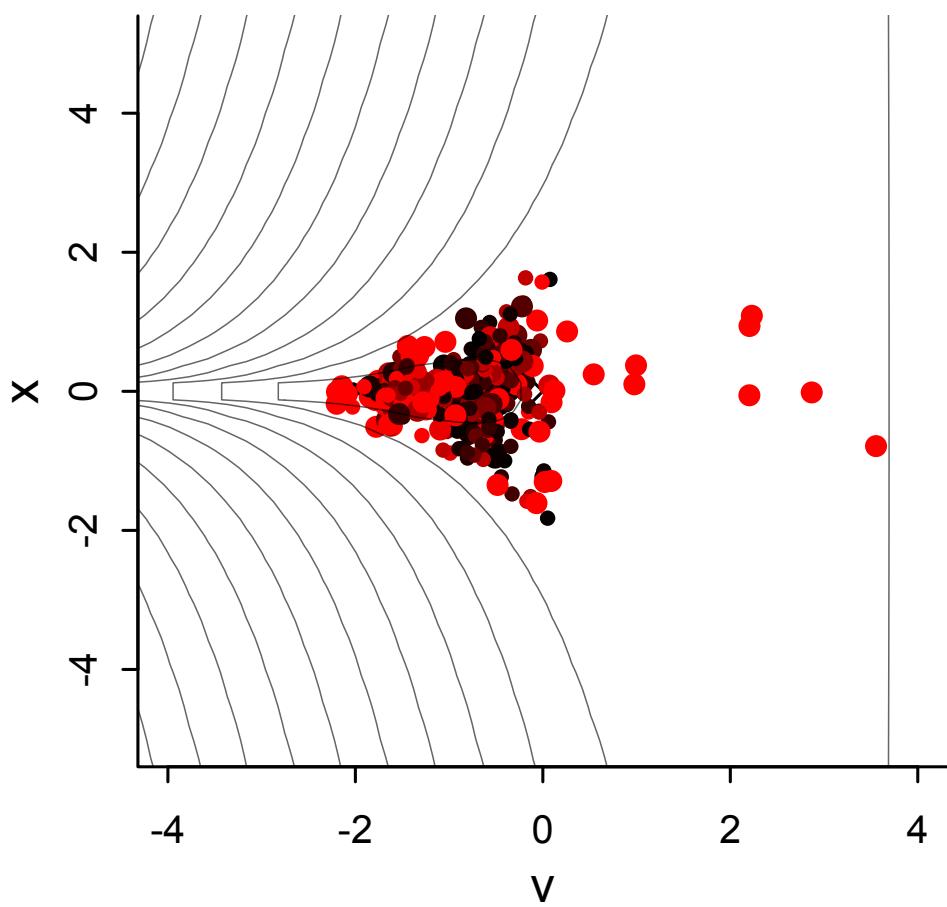
$v \sim \text{Normal}(0, 3)$

$x \sim \text{Normal}(0, \exp(v))$

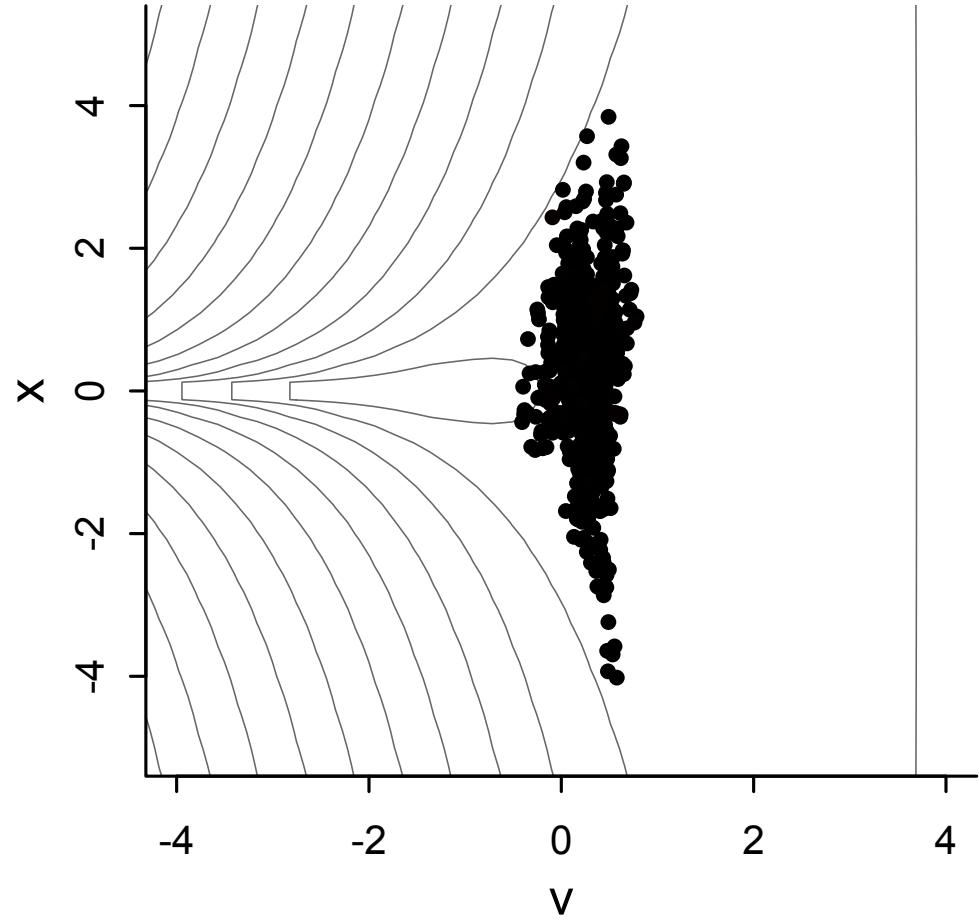


Divergent transitions

large step size

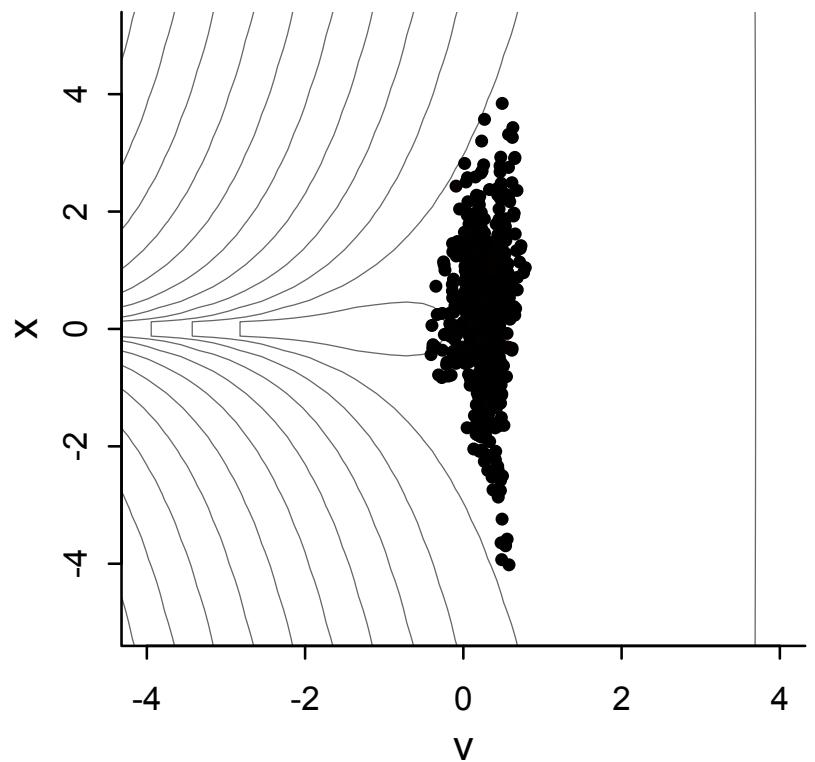


small step size



Divergent transitions

- Two basic strategies:
- (1) Increase Stan's adapt_delta control parameter => better step size adaptation + slower exploration
- (2) Re-parameterize!



Re-parameterize!

- Most any statistical model can be expressed in several mathematically identical ways

$$\alpha \sim \text{Normal}(\mu, \sigma)$$

Re-parameterize!

- Most any statistical model can be expressed in several mathematically identical ways

$$\alpha \sim \text{Normal}(\mu, \sigma)$$

$$\alpha = \mu + \beta$$

$$\beta \sim \text{Normal}(0, \sigma)$$

Re-parameterize!

- Most any statistical model can be expressed in several mathematically identical ways

$$\alpha \sim \text{Normal}(\mu, \sigma) \qquad \qquad \qquad \textit{Centered}$$

$$\alpha = \mu + \beta$$

$$\beta \sim \text{Normal}(0, \sigma)$$

$$\alpha = \mu + z\sigma$$

$$z \sim \text{Normal}(0, 1)$$

Non-centered

Re-parameterize!

- Why would this madness help with sampling?
- HMC sees a different geometry!

Centered

$$v \sim \text{Normal}(0, 3)$$

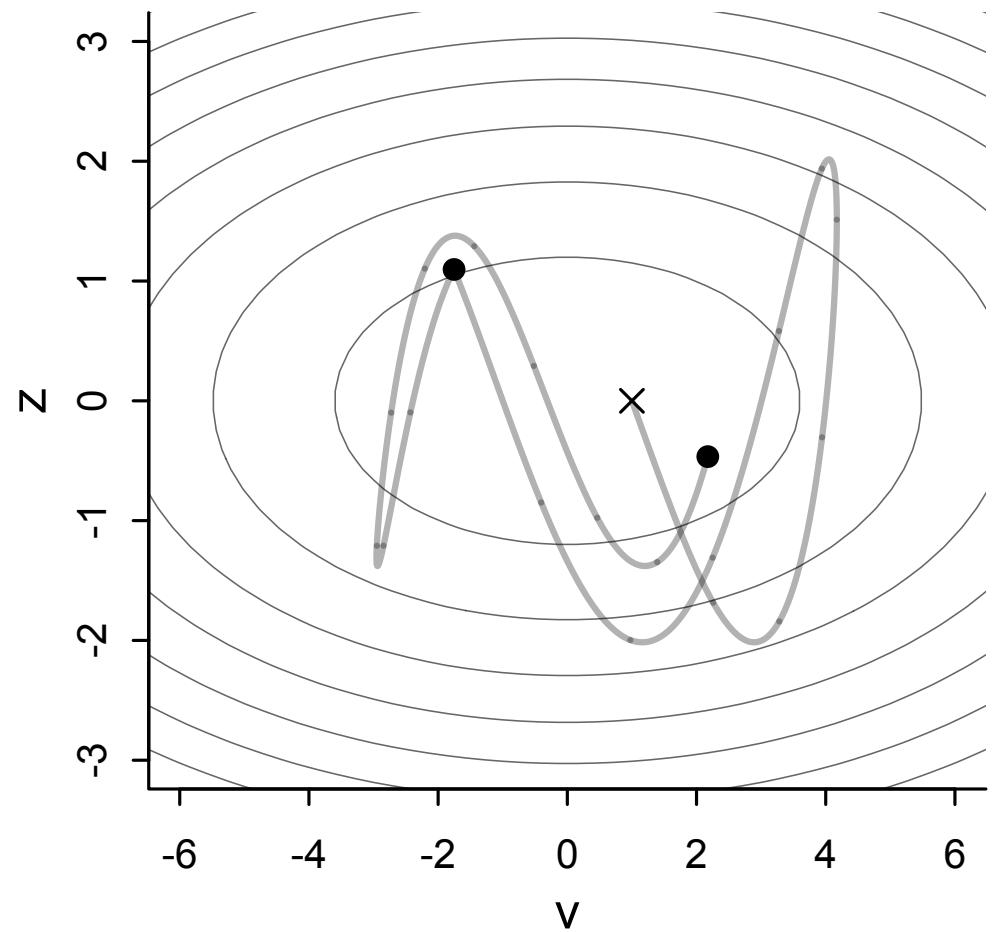
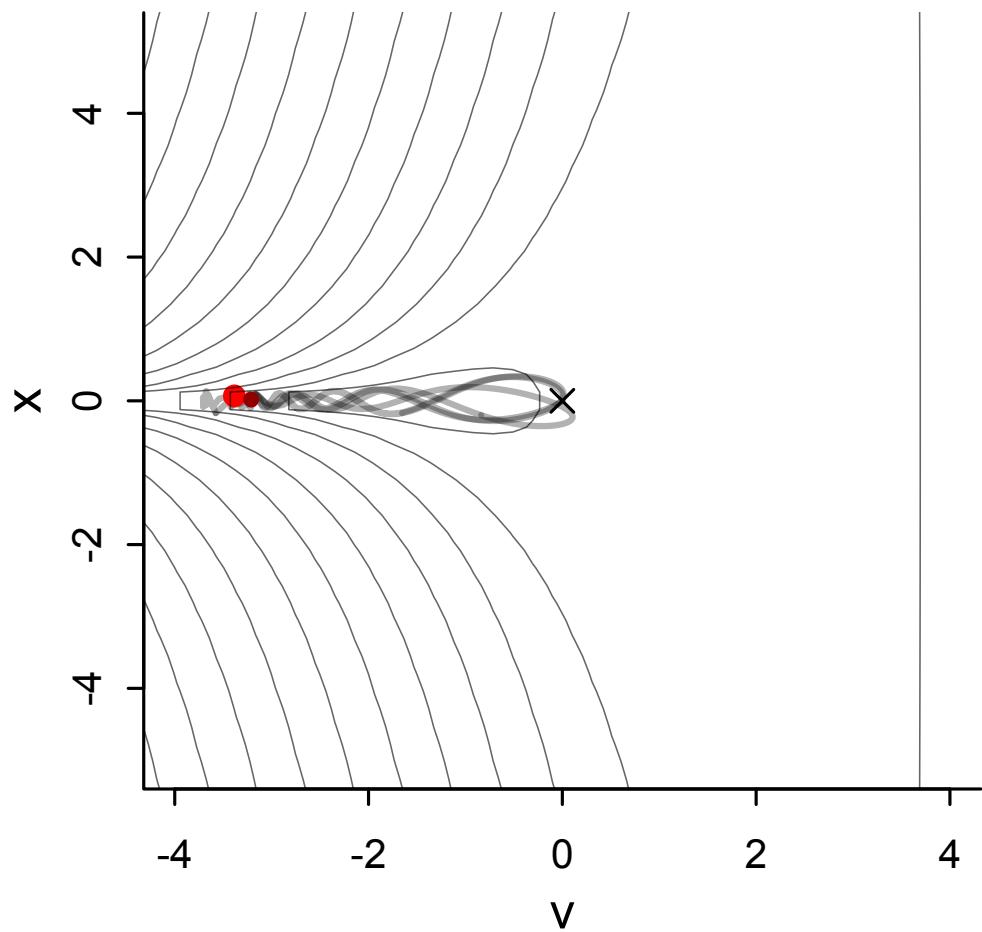
$$x \sim \text{Normal}(0, \exp(v))$$

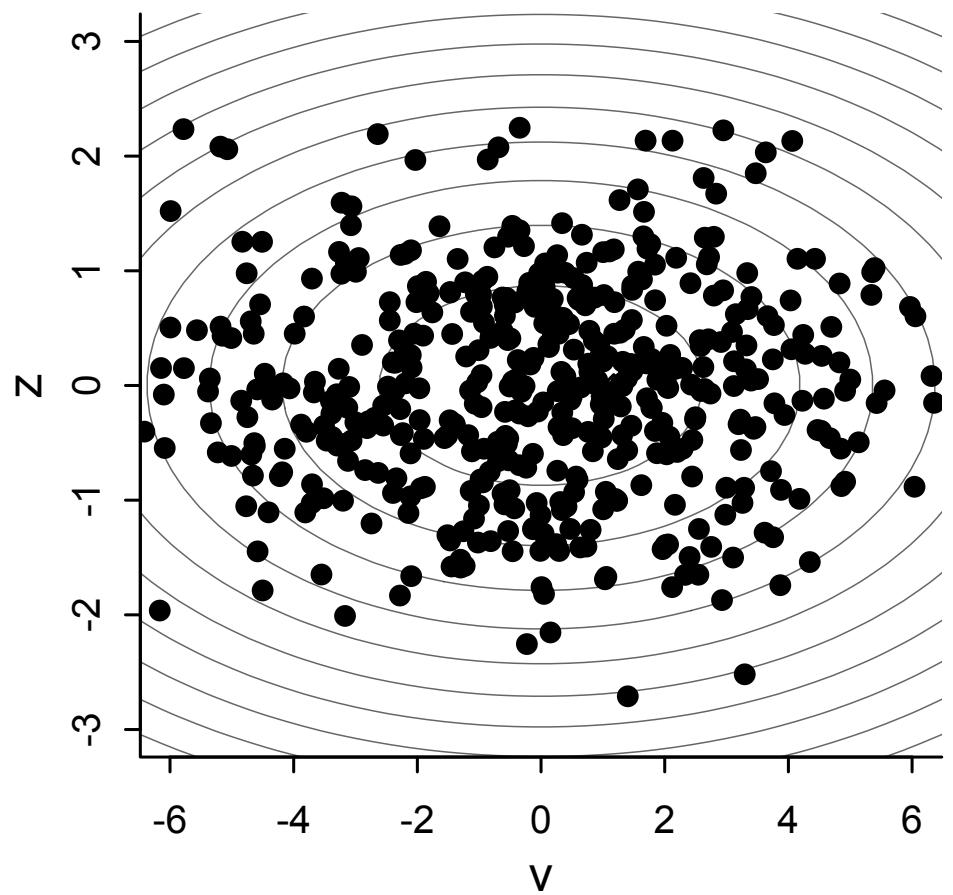
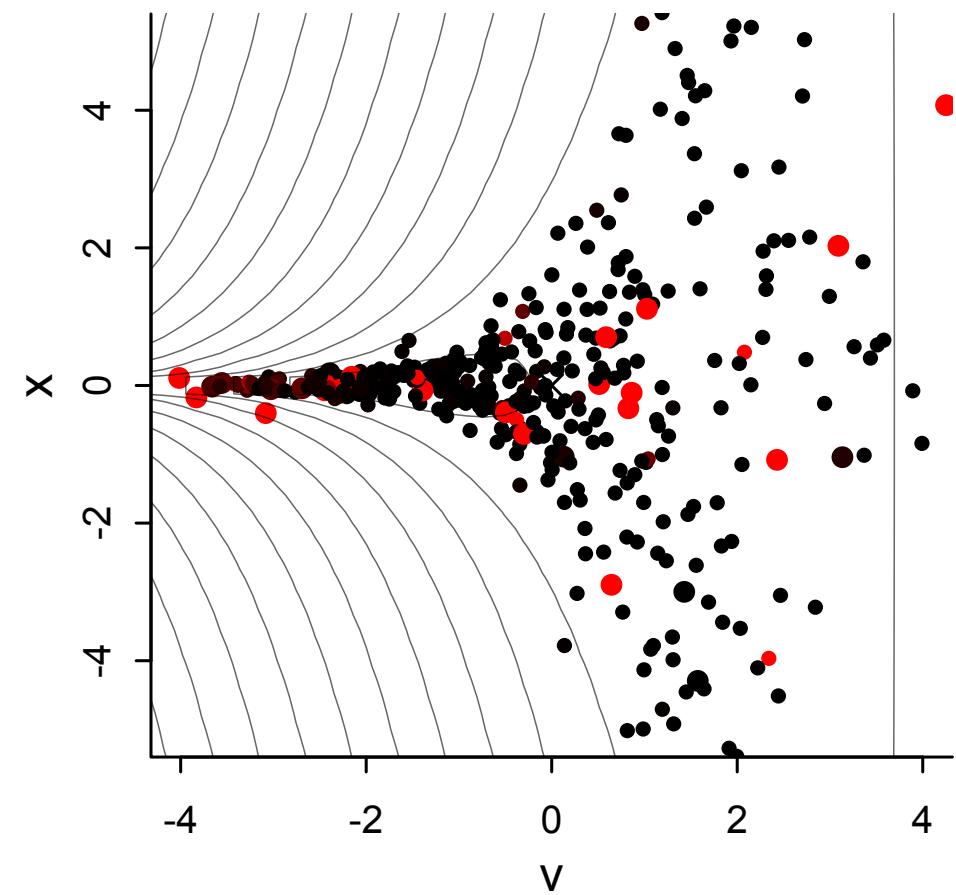
Non-centered

$$v \sim \text{Normal}(0, 3)$$

$$x = z \exp(v)$$

$$z \sim \text{Normal}(0, 1)$$

$v \sim \text{Normal}(0, 3)$ $x \sim \text{Normal}(0, \exp(v))$ $v \sim \text{Normal}(0, 3)$ $x = z \exp(v)$ $z \sim \text{Normal}(0, 1)$ 

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Re-parameterize!

$$L_i \sim \text{Binomial}(1, p_i)$$

$$\text{logit}(p_i) = \alpha_{\text{ACTOR}[i]} + \gamma_{\text{BLOCK}[i]} + \beta_{\text{TREATMENT}[i]}$$

$$\beta_j \sim \text{Normal}(0, 0.5) \quad , \text{ for } j = 1..4$$

$$\alpha_j \sim \text{Normal}(\bar{\alpha}, \sigma_\alpha) \quad , \text{ for } j = 1..7$$

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$$\bar{\alpha} \sim \text{Normal}(0, 1.5)$$

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$$\text{logit}(p_i) = \alpha_{\texttt{ACTOR}[i]} + \color{blue}{\gamma_{\texttt{BLOCK}[i]}} + \beta_{\texttt{TREATMENT}[i]}$$

$$L_i \sim \text{Binomial}(1,p_i)$$

$$\text{logit}(p_i) = \bar{\alpha} + z_{\texttt{ACTOR}[i]} \sigma_\alpha + x_{\texttt{BLOCK}[i]} \sigma_\gamma + \beta_{\texttt{TREATMENT}[i]}$$

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$$L_i \sim \text{Binomial}(1, p_i)$$

$$\text{logit}(p_i) = \bar{\alpha} + z_{\text{ACTOR}[i]} \sigma_\alpha + x_{\text{BLOCK}[i]} \sigma_\gamma + \beta_{\text{TREATMENT}[i]}$$

R code
13.28

```
set.seed(13)
m13.4nc <- ulam(
  alist(
    pulled_left ~ dbinom( 1 , p ) ,
    logit(p) <- a_bar + z[actor]*sigma_a + x[block_id]*sigma_g + b[treatment] ,
    b[treatment] ~ dnorm( 0 , 0.5 ) ,
    z[actor] ~ dnorm( 0 , 1 ) ,
    x[block_id] ~ dnorm( 0 , 1 ) ,
    a_bar ~ dnorm( 0 , 1.5 ) ,
    sigma_a ~ dexp(1) ,
    sigma_g ~ dexp(1)
  ) , data=dat_list , chains=4 , cores=4 )
```

Non-centered vs centered

R code
13.29

```
neff_c <- precis( m13.4 , depth=2 )[['n_eff']]
neff_nc <- precis( m13.4nc , depth=2 )[['n_eff']]
par_names <- rownames( precis( m13.4 , depth=2 ) )
neff_table <- cbind( neff_c , neff_nc )
rownames(neff_table) <- par_names
round(t(neff_table))
```

	b[1]	b[2]	b[3]	b[4]	a[1]	a[2]	a[3]	a[4]	a[5]	a[6]	a[7]	g[1]	g[2]	g[3]
neff_c	584	572	587	523	562	759	541	547	582	601	677	418	793	784
neff_nc	1144	1233	1144	1115	592	962	614	614	611	627	772	1811	2137	1748
	g[4]	g[5]	g[6]	a_bar	sigma_a	sigma_g								
neff_c	605	934	603	824		898		243						
neff_nc	1575	2127	1429		573		777		970					

Posterior predictions

- Predictions more subtle: Same clusters or new clusters?
- Same clusters: proceed as usual
- New clusters: should average over distribution of varying effects
- In this case:
 - Same clusters: Predictions for these chimpanzees
 - New clusters: Prediction for a new chimpanzee or rather for population of chimpanzees

Same clusters, new clusters

- Same actors:
 - Really same as before: varying effects are just parameters; you know the model; push samples back through the model
 - `link()` and `sim()` obey this rule
- New actors (counterfactual):
 - which actor (cluster) to use for counterfactual predictions?
 - average actor
 - marginal of actor
 - show sample of actors from posterior

Average actor

- “average actor” means actor with population average intercept, “alpha”
- Strategy:
 - replace varying intercept samples with zeros => all actors have average intercept now
 - compute predictions as usual

```
# replace varying intercept samples with zeros  
# 1000 samples by 7 actors  
a_actor_zeros <- matrix(0,1000,7)
```

R code
12.32

```
# fire up link  
# note use of replace list  
link.m12.4 <- link( m12.4 , n=1000 , data=d.pred ,  
replace=list(a_actor=a_actor_zeros) )
```

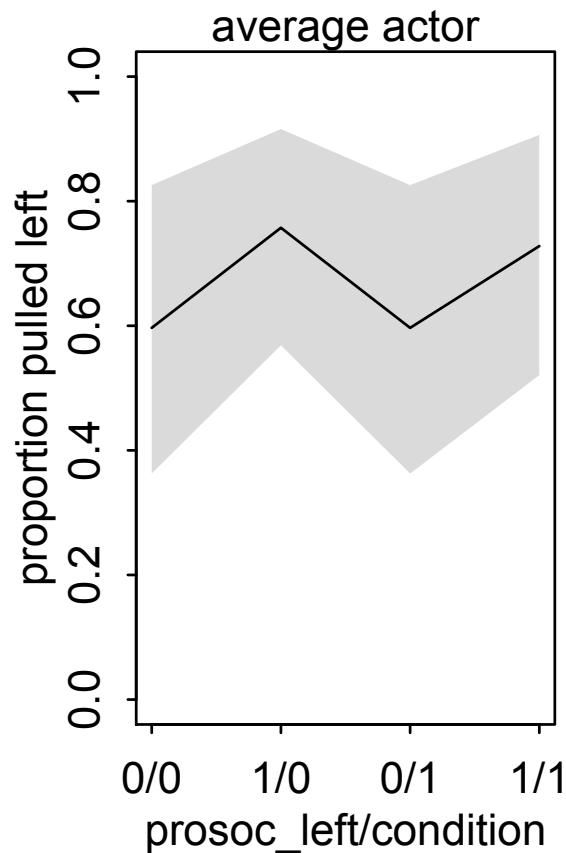
R code
12.33

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# replace varying intercept samples with zeros  
# 1000 samples by 7 actors  
a_actor_zeros <- matrix(0,1000,7)
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R code
12.32

```
# fire up link  
# note use of replace list  
link.m12.4 <- link( m12.4 , n=1000 , data=d.pred ,  
    replace=list(a_actor=a_actor_zeros) )
```

R code
12.33



Marginal of actor

- “Marginal of” means “averaging over variation in actors” => shows variation arising from variation across actors

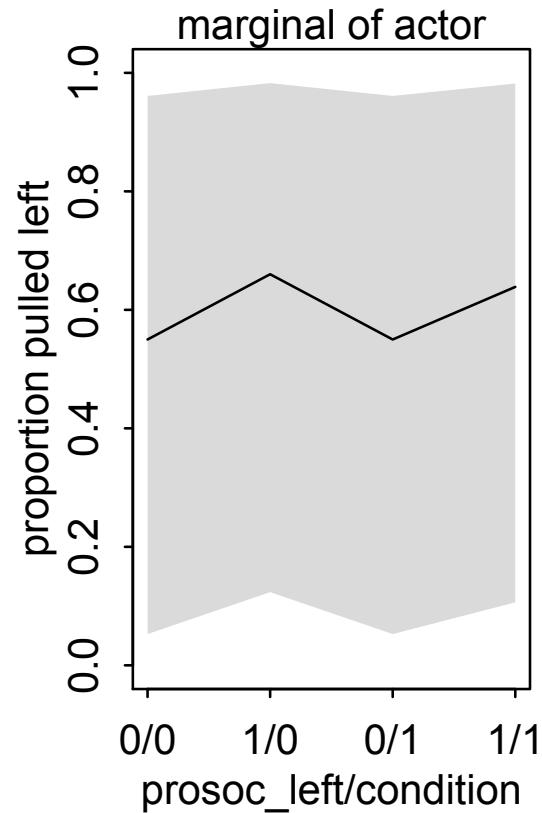
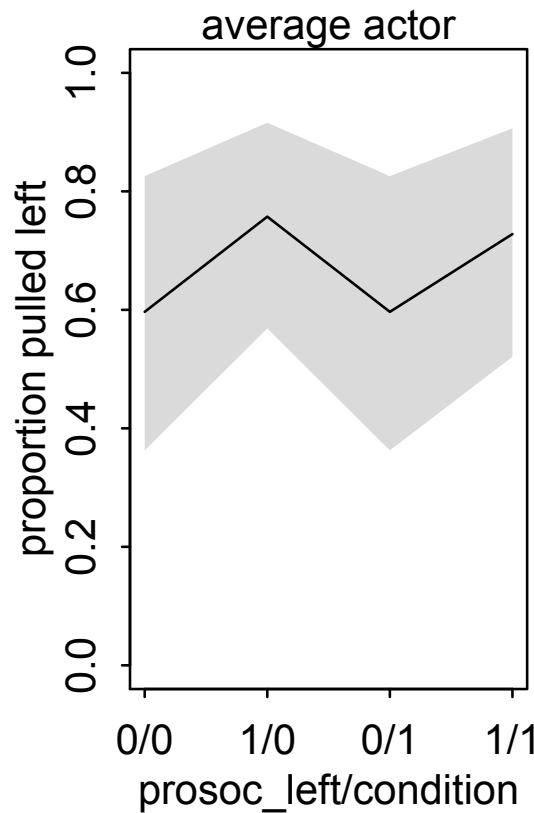
$$\alpha_{\text{ACTOR}} \sim \text{Normal}(0, \sigma_{\text{ACTOR}})$$

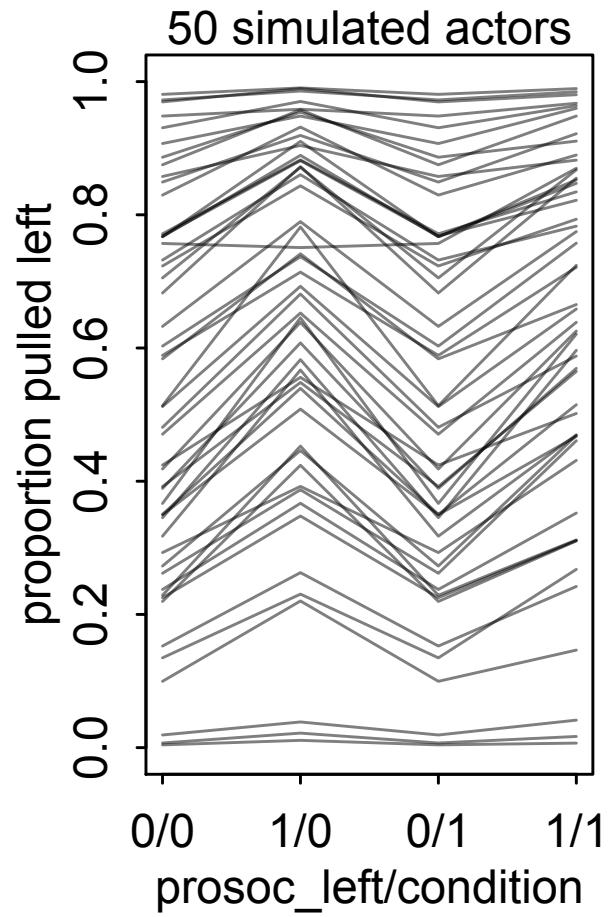
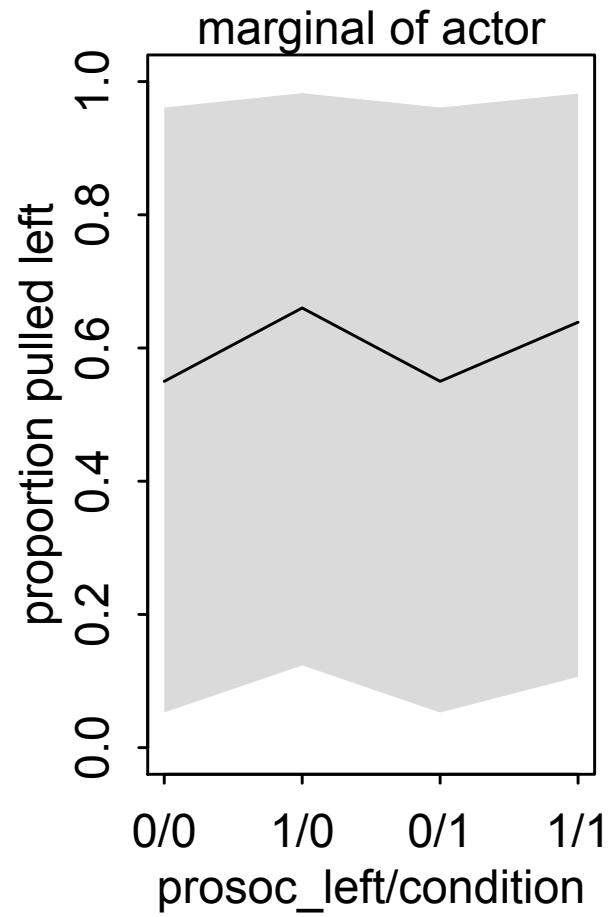
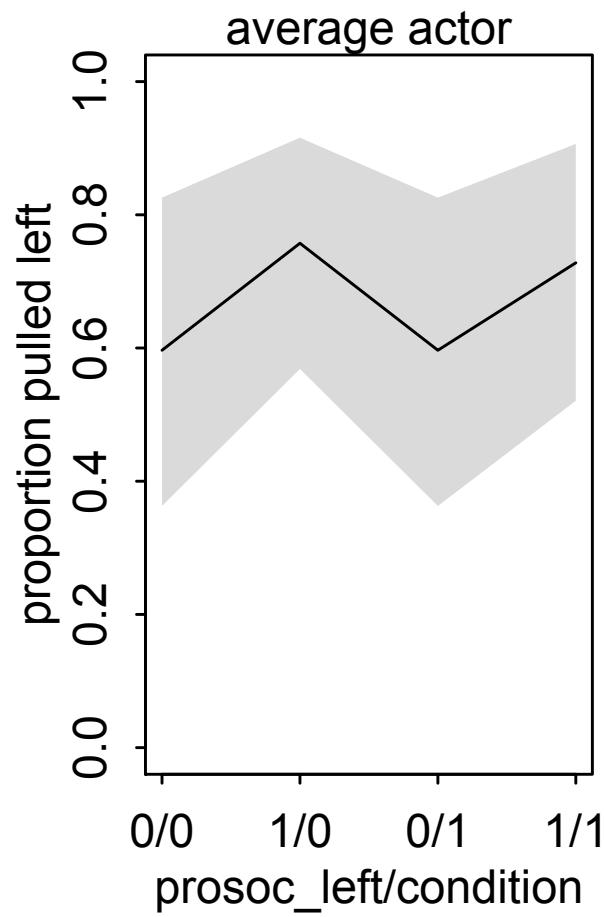
- Strategy:
 - Extract samples for sigma_actor
 - Simulate new varying intercepts
 - Use simulated intercepts to simulate predictions

R code
12.34

```
# replace varying intercept samples with simulations
post <- extract.samples(m12.4)
a_actor_sims <- rnorm(7000,0,post$sigma_actor)
a_actor_sims <- matrix(a_actor_sims,1000,7)

# fire up link
# note use of replace list
link.m12.4 <- link( m12.4 , n=1000 , data=d.pred ,
replace=list(a_actor=a_actor_sims) )
```





Homework

- Frogs & contraception
- Next week: Chapter 14 — varying slopes and other wonders

