1 Stirling's formula

Start with the definition of the continuous factorial:

$$n! = \int_0^\infty x^n \cdot e^{-x} \cdot dx \tag{1}$$

$$= \int_0^\infty e^{-x+n\log x} \cdot dx \tag{2}$$

$$= n \cdot \int_0^\infty e^{-ny+n\log(ny)} \cdot dy \tag{3}$$

With $f(y) = -y + \log(ny)$ this can be written as

$$= n \cdot \int_0^\infty e^{nf(y)} \cdot dy. \tag{4}$$

The maximum of f(y) is at $y_0 = 1$. Taylor expansion around y_0 :

$$f(y) \approx \log(n) - 1 - \frac{(y-1)^2}{2}.$$
 (5)

Thus

$$n! \approx ne^{n\log(n)-n} \cdot \int_0^\infty e^{-n(y-1)^2/2} \cdot dy$$

$$= ne^{n\log(n)-n} \sqrt{\frac{\pi}{2n}} \left(1 + \operatorname{Erf}\left(\sqrt{\frac{n}{2}}\right) \right)$$

$$\approx \sqrt{2\pi n} e^{n\log(n)-n}, \tag{6}$$

where we have used $\operatorname{Erf}(x) \approx 1$ for large x. Taking the logarithm yields the Stirling formula:

$$\log(n!) = n\log(n) - n + \frac{1}{2}\log(2\pi n). \tag{7}$$



2 Adding two Gaussian distributions

a) Without characteristic function

Let p_x and p_y be Gaussian distributions with parameters μ_x , μ_y and σ .

$$p_x(x) = \frac{1}{\sqrt{2\pi\sigma}} \exp\left(-\frac{(x-\mu_x)^2}{2\sigma^2}\right),\tag{8}$$

$$p_y(y) = \frac{1}{\sqrt{2\pi\sigma}} \exp\left(-\frac{(y-\mu_y)^2}{2\sigma^2}\right). \tag{9}$$

$$p_z(z) = \int dx \int dy \cdot \delta(z - (x + y)) \cdot p(x, y)$$
(10)

$$= \int dx \int dy \cdot \delta(z - (x + y)) \cdot p(x) \cdot p(y)$$
(11)

$$= \frac{1}{2\pi\sigma^2} \int dx \int dy \cdot \delta(z - (x+y)) \cdot \exp\left(-\frac{(y-\mu_y)^2}{2\sigma^2}\right) \cdot \exp\left(-\frac{(x-\mu_x)^2}{2\sigma^2}\right)$$
(12)

$$= \frac{1}{\sqrt{4\pi\sigma^2}} \cdot \exp\left(-\frac{(z - \mu_x - \mu_y)^2}{4\sigma^2}\right). \tag{13}$$

With $\mu_z = \mu_x + \mu_y$ and $\sigma_z = \sqrt{2}\sigma$, this can be rewritten as

$$p_z(z) = \frac{1}{\sqrt{2\pi\sigma_z^2}} \cdot \exp\left(-\frac{(z-\mu_z)^2}{2\sigma_z^2}\right). \tag{14}$$

b) With characteristic function

The Fourier transform of a normal density with mean μ and standard deviation σ is

$$\hat{p}(t) = \int_{-\infty}^{\infty} p(x)e^{-itx}dx = e^{-i\mu t}e^{-\frac{(\sigma t)^2}{2}}.$$
(15)

The product of two distributions with the same σ is

$$\hat{p}_z(t) = \hat{p}_x(t)\hat{p}_y(t) = e^{-i(\mu_x + \mu_y)t}e^{-(\sigma t)^2},$$
(16)

which leads also to the definitions $\mu_z = \mu_x + \mu_y$ and $\sigma_z = \sqrt{2}\sigma$ and a Gaussian distribution.

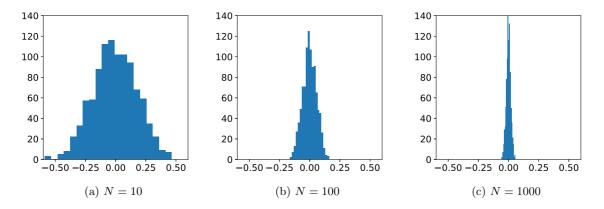
3 Computer exercise on random numbers

Write a little computer program that sums up N random numbers drawn from [-1,1]. Divide by N and compare the obtained distribution to the Gaussian distribution, for N = 10, 100, 1000, by plotting the histogram and the appropriate analytical function together.

To create a distribution we run a loop for 1000 iterations, at each iteration we create a random number by summing up N number between -1 and 1. The computer program can be implemented using the following python snippet:

```
import random
  import matplotlib.pyplot as plt
  import numpy as np
  from scipy.optimize import curve_fit
7~\mathrm{M} = ~1000
  for N in [10, 100, 1000]:
8
       rs = []
       for _{\perp} in range (M):
11
           r = 0
           for i in range (N):
                r += random.uniform(-1, 1)
13
           rs.append(r / N)
14
       plt.figure(figsize=(3, 3))
16
       plt. xlim (-0.6, 0.6)
17
       plt.ylim(0, 130)
18
       plt.hist(rs, bins=20)
19
       plt.savefig(f'../figures/{N}.pdf')
```

This leads to these plots:



As is to be expected, the expectation value of the distrition does not change with increasing N, but stays constant at ≈ 0 . The width of the curve decreases for a large number of summands. This is due to the fact that we divide the sum by N.



