

Lecture 5: Generalized Linear Models, MLP, and Back-Propagation

Tao LIN

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Reading materials

- Chapter 3, Stanford CS 229 Lecture Notes,
https://cs229.stanford.edu/notes2022fall/main_notes.pdf
- Lecture 4, Stanford CS 231n, <http://cs231n.stanford.edu/schedule.html>

Reference

- EPFL, CS-433 Machine Learning, https://github.com/epfml/ML_course

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- Generalization Gap and Model Selection
- Bias-Variance Decomposition
- Logistic Regression

2 Exponential Families and Generalized Linear Models

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 - The Basic Structure of MLP
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Generalization gap: How far is the test from the true error?

- **True Error:**

$$L_{\mathcal{D}}(f) = \mathbb{E}_{(y, \mathbf{x}) \sim \mathcal{D}} [\ell(y, f(\mathbf{x}))]. \quad (1)$$

- **Test/Empirical Error:**

$$L_{S_{\text{test}}}(f) = \frac{1}{|S_{\text{test}}|} \sum_{(\mathbf{x}_n, y_n) \in S_{\text{test}}} \ell(y_n, f(\mathbf{x}_n)). \quad (2)$$

- **Generalization Error:**

$$|L_{\mathcal{D}}(f) - L_{S_{\text{test}}}(f)|. \quad (3)$$

- **In Expectation:**

$$L_{\mathcal{D}}(f) = \mathbb{E}_{S_{\text{test}} \sim \mathcal{D}} [L_{S_{\text{test}}}(f)], \quad (4)$$

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The variation of the $|L_{\mathcal{D}}(f) - L_{S_{\text{test}}}(f)|$ matters.

Theorem 1

Given a model f and a test set $S_{test} \sim \mathcal{D}$ i.i.d. (not used to learn f) and a loss $\ell(\cdot, \cdot) \in [a, b]$:

$$\Pr \left[|L_{\mathcal{D}}(f) - L_{S_{test}}(f)| \geq \sqrt{\frac{(b-a)^2 \ln(2/\delta)}{2|S_{test}|}} \right] \leq \delta \quad (5)$$

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Given a predictor f and a dataset S , we can control the expected risk:

$$\Pr \left[\underbrace{L_{\mathcal{D}}(f)}_{\text{not computable}} \geq \underbrace{L_{S_{\text{test}}}(f)}_{\text{computable}} + \underbrace{\sqrt{\frac{(b-a)^2 \ln(2/\delta)}{2 |S_{\text{test}}|}}}_{\text{deviation}} \right] \leq \delta. \quad (6)$$

How far is each of the K test errors $L_{S_{\text{test}}}(f_k)$ from the true $L_{\mathcal{D}}(f_k)$?

Theorem 2

We can bound the maximum deviation for all K candidates, by

$$\Pr \left[\max_k |L_{\mathcal{D}}(f_k) - L_{S_{\text{test}}}(f_k)| \geq \sqrt{\frac{(b-a)^2 \ln(2K/\delta)}{2|S_{\text{test}}|}} \right] \leq \delta \quad (7)$$

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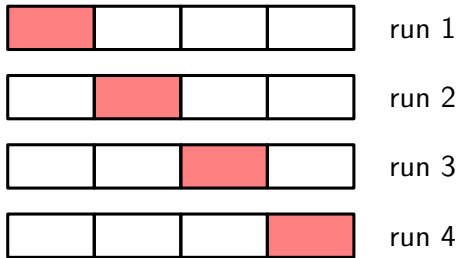
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- The error decreases as $\mathcal{O}(1/\sqrt{|S_{\text{test}}|})$ with the number test points.
 - When testing K hyper-parameters, the error only goes up by $\sqrt{\ln(K)}$.
- ⇒ So we can test many different models without incurring a large penalty.

Issues: Splitting the data once into two parts (one for training and one for testing)
is not the most efficient way to use the data!

K-fold cross-validation:

- 1 Randomly partition the data into K groups
- 2 Train K times. Each time leave out exactly one of the K groups for testing and use the remaining $K - 1$ groups for training.
- 3 Average the K results

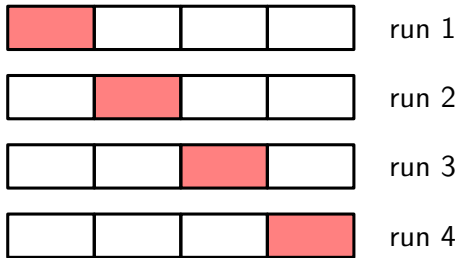


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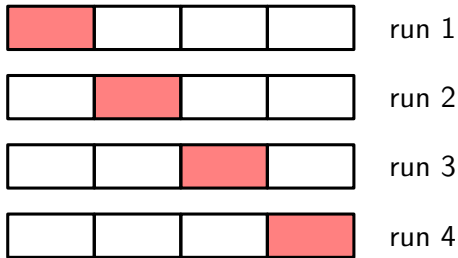
Benefits:

- We have used all data for training, and all data for testing, and used each data point the same number of times.

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- 1 Randomly partition the data into K groups
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Benefits:

- We have used all data for training, and all data for testing, and used each data point the same number of times.
- Cross-validation returns an unbiased estimate of the generalization error and its variance.

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Bias-Variance Decomposition

$$\begin{aligned}\mathbb{E}_{S_{\text{train}} \in \mathcal{D}} [L(f_{S_{\text{train}}})] &= \text{Var}_{\epsilon \sim \mathcal{D}_\epsilon} [\epsilon] && \text{(noise variance)} \\ &+ (f(\mathbf{x}_0) - \mathbb{E}_{S_{\text{train}}} [f_{S_{\text{train}}}(\mathbf{x}_0)])^2 && \text{(Bias)} \\ &+ \mathbb{E}_{S_{\text{train}} \sim \mathcal{D}} \left[(\mathbb{E}_{S_{\text{train}}} [f_{S_{\text{train}}}(\mathbf{x}_0)] - f_{S_{\text{train}}}(\mathbf{x}_0))^2 \right], && \text{(Variance)}\end{aligned}$$

which always lower-bounds the true error.

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⇒ To minimize the true error, we need to select a method that **simultaneously achieves low bias and low variance**.

Double descent curve in Deep Learning

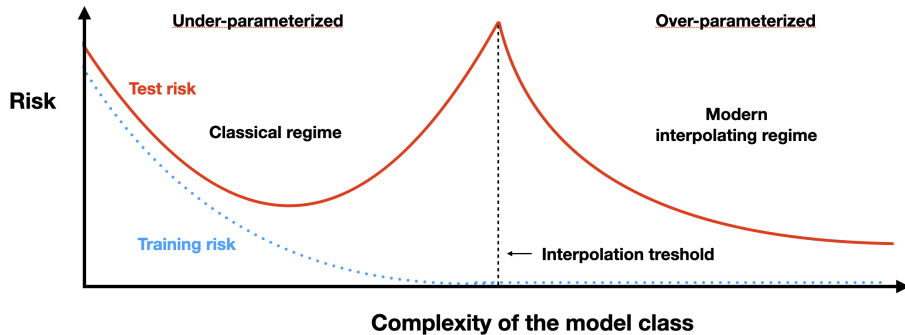


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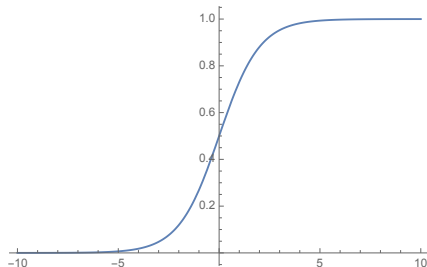
The logistic function

Consider first of all the case of two classes.

The posterior probability for class \mathcal{C}_1 :

$$p(\mathcal{C}_1|\mathbf{x}) = \frac{p(\mathbf{x}|\mathcal{C}_1)p(\mathcal{C}_1)}{p(\mathbf{x}|\mathcal{C}_1)p(\mathcal{C}_1) + p(\mathbf{x}|\mathcal{C}_2)p(\mathcal{C}_2)} \quad (8)$$

$$= \frac{1}{1 + \exp(-\eta)} = \sigma(\eta) \quad (9)$$



Properties of the logistic function:

- $1 - \sigma(\eta) = \sigma(-\eta)$
- $\sigma'(\eta) = \sigma(\eta) (1 - \sigma(\eta))$

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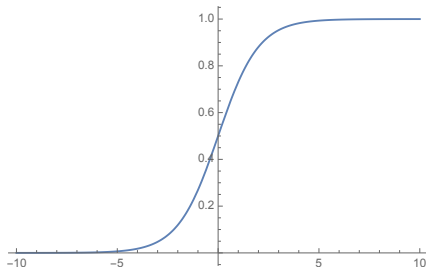
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where we have defined

$$\eta = \ln \frac{p(\mathbf{x}|\mathcal{C}_1)p(\mathcal{C}_1)}{p(\mathbf{x}|\mathcal{C}_2)p(\mathcal{C}_2)} \text{ and } \sigma(\eta) := \frac{e^\eta}{1 + e^\eta} \quad (10)$$



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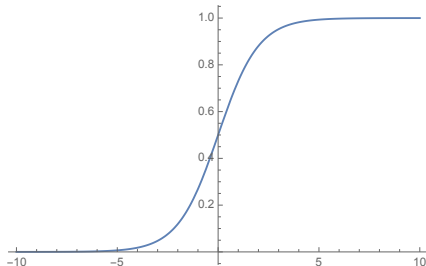
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For the case of $K > 2$ classes, we have

$$p(\mathcal{C}_k|\mathbf{x}) = \frac{p(\mathbf{x}|\mathcal{C}_k)p(\mathcal{C}_k)}{\sum_j p(\mathbf{x}|\mathcal{C}_j)p(\mathcal{C}_j)} = \frac{\exp(\eta_k)}{\sum_j \exp(\eta_j)} \quad (11)$$



Properties of the logistic function:

- $1 - \sigma(\eta) = \sigma(-\eta)$
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Logistic Regression

Given a “new” feature vector \mathbf{x} , we predict the (posterior) probability of the two class labels given \mathbf{x} by means of

$$p(1|\mathbf{x}) := \Pr[Y = 1|\mathbf{X} = \mathbf{x}] = \sigma(\mathbf{x}^\top \mathbf{w} + w_0) \quad (12)$$

$$p(0|\mathbf{x}) := \Pr[Y = 0|\mathbf{X} = \mathbf{x}] = 1 - \sigma(\mathbf{x}^\top \mathbf{w} + w_0) , \quad (13)$$

where we predict a real value (a probability) and not a label.

MLE is a method of estimating the parameters of a statistical model

The MLE finds the parameters \mathbf{w}^* under which $\{\mathbf{y}, \mathbf{X}\}$ are the most likely:

$$\mathbf{w}^* = \arg \max_{\mathbf{w}} \left(\mathcal{L}(\mathbf{w}) := \prod_{n=1}^N p(\{\mathbf{x}_n, y_n\} | \mathbf{w}) \right) = \arg \min_{\mathbf{w}} [-\log \mathcal{L}(\mathbf{w})] . \quad (14)$$

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The likelihood of the data $\{\mathbf{y}, \mathbf{X}\}$ given the parameter \mathbf{w} , i.e., $p(\mathbf{y}, \mathbf{X} | \mathbf{w})$.

$$p(\mathbf{y}, \mathbf{X} | \mathbf{w}) = p(\mathbf{X} | \mathbf{w}) p(\mathbf{y} | \mathbf{X}, \mathbf{w}) = p(\mathbf{X}) p(\mathbf{y} | \mathbf{X}, \mathbf{w}) , \quad (15)$$

where \mathbf{X} does not depend on \mathbf{w} .

MLE for Logistic Regression

For Logistic Regression, we have:

$$p(\mathbf{y}|\mathbf{X}, \mathbf{w}) = \prod_{n=1}^N p(y_n|\mathbf{x}_n) = \prod_{n:y_n=1} p(y_n = 1|\mathbf{x}_n) \prod_{n:y_n=0} p(y_n = 0|\mathbf{x}_n) \quad (16)$$

$$= \prod_{n=1}^N \sigma(\mathbf{x}_n^\top \mathbf{w})^{y_n} [1 - \sigma(\mathbf{x}_n^\top \mathbf{w})]^{1-y_n} \quad (17)$$

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$$= \prod_{n=1}^N \sigma(\mathbf{x}_n^\top \mathbf{w})^{y_n} [1 - \sigma(\mathbf{x}_n^\top \mathbf{w})]^{1-y_n} \quad (17)$$

Minimizing $\mathcal{L}(\mathbf{w})$ through the property of stationary points.

$$\nabla \mathcal{L}(\mathbf{w}) = \frac{1}{N} \sum_{n=1}^N \mathbf{x}_n (\sigma(\mathbf{x}_n^\top \mathbf{w}) - y_n) = \frac{1}{N} \mathbf{X}^\top [\sigma(\mathbf{X}\mathbf{w}) - \mathbf{y}] , \quad (18)$$

where $\mathbf{X} \in \mathbb{R}^{N \times d}$. It has no closed-form solution to $\nabla \mathcal{L}(\mathbf{w}) = 0$.

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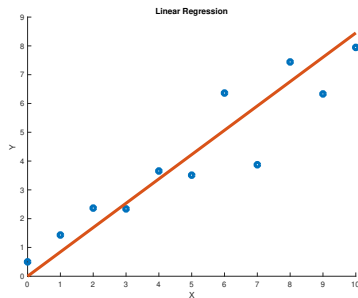
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The Least-Squares can be defined in two different ways

- **Geometric way:**

Minimizing the sum of the squares of the residuals:

$$\hat{\mathbf{w}} = \arg \min_{\mathbf{w}} \frac{1}{2N} \sum_{n=1}^N (y_n - \mathbf{x}_n^\top \mathbf{w})^2 \quad (19)$$



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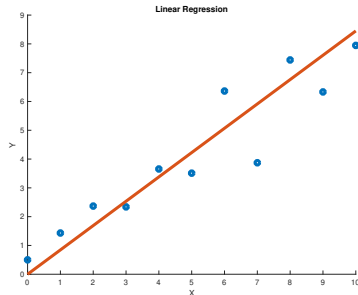
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- **Probabilistic way:**

Assume the data follow a linear Gaussian model:

$$\mathbf{y} = \mathbf{x}^\top \mathbf{w} + \epsilon \text{ where } \epsilon \sim \mathcal{N}(0, \sigma^2) \quad (20)$$



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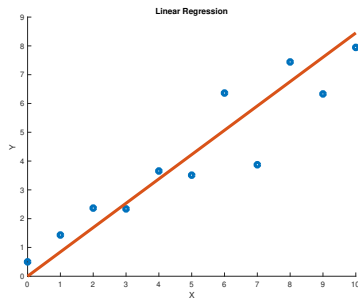
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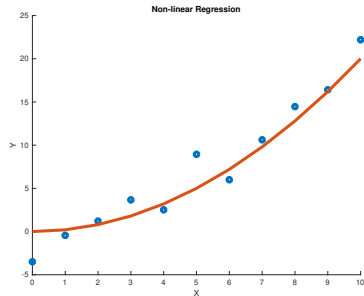
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Doing MLE recovers the LS estimator $\hat{\mathbf{w}}$.



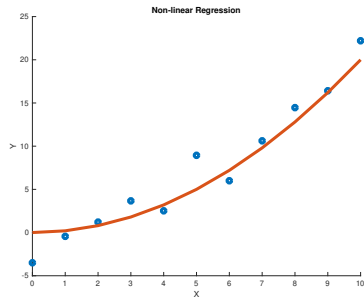
How to get non-linear models?

- **Features augmentations:** add non-linear features
(x, x^2, x^3, \dots)



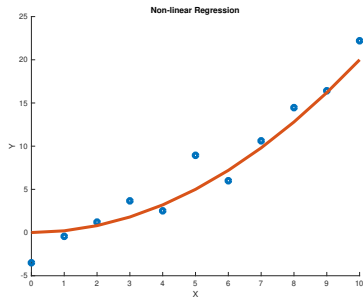
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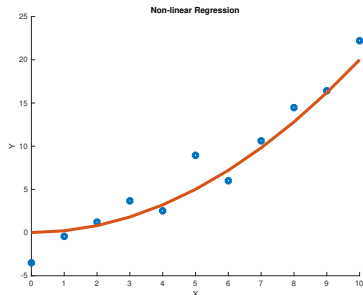
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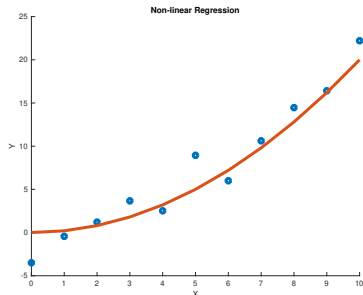
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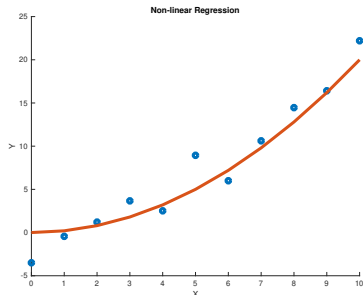
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 \Rightarrow The linear model predicts the mean of a distribution μ
(from which the data are sampled).
 - Logistic Regression: $y \sim \mathcal{B}(\sigma(\mathbf{x}^\top \mathbf{w}))$
 \Rightarrow The linear model predicts another quantity $\eta := \mathbf{x}^\top \mathbf{w}$.



Recall the definition of Logistic Regression

Logistic Regression models the probability of the two classes $\{0, 1\}$ by

$$p(1|\eta) = \sigma(\eta) \text{ and } p(0|\eta) = 1 - \sigma(\eta), \quad (21)$$

where $\eta = \mathbf{x}^\top \mathbf{w}$.

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where $\eta = \mathbf{x}^\top \mathbf{w}$. This can be compactly written as

$$p(y|\eta) = \frac{e^{\eta y}}{1 + e^\eta} = \exp(\eta y - \ln(1 + e^\eta)) \quad (22)$$

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where $\eta = \mathbf{x}^\top \mathbf{w}$. This can be compactly written as

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⇒ The *link function*:
the relation between (1) the value η we predict by the linear model and (2) the mean μ .

A unified framework

The distribution used in Logistic Regression can be written in a very specific form:

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- The discussion on a class of distributions, known as *exponential families*.
- Many distributions (but not all) fit into this framework and that distributions in this family have many nice properties.

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Exponential family — Definition

A distribution belongs to the exponential family if it can be written in the form

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¹Assume that we are given independent samples from this distribution. We do know $\boldsymbol{\phi}(y)$ and $h(y)$ but not $\boldsymbol{\eta}$. In order to optimally estimate $\boldsymbol{\eta}$ given these samples, all we need is the empirical average of the $\boldsymbol{\phi}(y)$.

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We will exclude such parameters by only looking at the set of parameters

$$M := \left\{ \boldsymbol{\eta} : \int_y h(y) \exp [\boldsymbol{\eta}^\top \boldsymbol{\phi}(y)] dy < \infty \right\} \quad (27)$$

Why?

Bernoulli distributions belong to the exponential family

Recall: A distribution belongs to the exponential family if it can be written in the form

$$p(y|\boldsymbol{\eta}) = h(y) \exp [\boldsymbol{\eta}^\top \boldsymbol{\phi}(y) - A(\boldsymbol{\eta})] \quad (28)$$

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$$\Pr(Y = 1) = \mu \quad \text{and} \quad \Pr(Y = 0) = 1 - \mu \quad (29)$$

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$$\Rightarrow \eta = g(\mu) = \ln \frac{\mu}{1 - \mu} \iff \mu = g^{-1}(\eta) = \frac{e^\eta}{1 + e^\eta}, \text{ and } g(\mu) \text{ links the mean of } \boldsymbol{\phi}(y) \text{ to } \boldsymbol{\eta}.$$

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Link function ($\eta := g(\mu)$): $\eta_1 = \frac{\mu}{\sigma^2}, \eta_2 = -\frac{1}{2\sigma^2} \iff \mu = -\frac{\eta_1}{2\eta_2}, \sigma^2 = -\frac{1}{2\eta_2}.$

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\Rightarrow The cost function \mathcal{L} is a convex function in η since $A(\eta)$ is convex.

Given the definition

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Closed-form: assume we have determined the link function $\mathbf{g}(\boldsymbol{\mu}) = \boldsymbol{\eta}$

$$\boldsymbol{\eta} = \mathbf{g} \left(\frac{1}{N} \sum_{n=1}^N \boldsymbol{\phi}(y_n) \right) , \quad (44)$$

and justify why we called $\boldsymbol{\phi}(y)$ a sufficient statistics.

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Generalized Linear Models (GLM)

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We can apply a Generalized Linear Model (GLM)!

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The condition probability is thus modeled as:

$$p(y|\mathbf{x}; \mathbf{w}) = h(y_n) \exp(\eta\phi(y) - A(\eta)) \quad \text{for } \eta = g \circ f(\mathbf{x}^\top \mathbf{w}) \quad (45)$$

Negative log-likelihood estimation:

$$\mathcal{L}(\mathbf{w}) = -\frac{1}{N} \sum_{n=1}^N \ln p(y_n | \mathbf{x}_n^\top \mathbf{w}) \quad (46)$$

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In the case of Logistic Regression:

$$\nabla_{\mathbf{w}} \mathcal{L}(\mathbf{w}) = \frac{1}{N} \mathbf{X}^\top [\sigma(\mathbf{X}\mathbf{w}) - \mathbf{y}] \quad (49)$$

Some examples

- Gaussian distribution



Least Squares

Some examples

- Gaussian distribution



Least Squares

- Bernoulli distribution



Logistic Regression

Some examples

• Gaussian distribution	\Rightarrow	Least Squares
• Bernoulli distribution	\Rightarrow	Logistic Regression
• Multi-nomial distribution	\Rightarrow	Softmax Regression

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- 3 Multi-Layer Perceptron (MLP) and Back-Propagation (BP)**
 - The Basic Structure of MLP
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MLP from a single neuron view

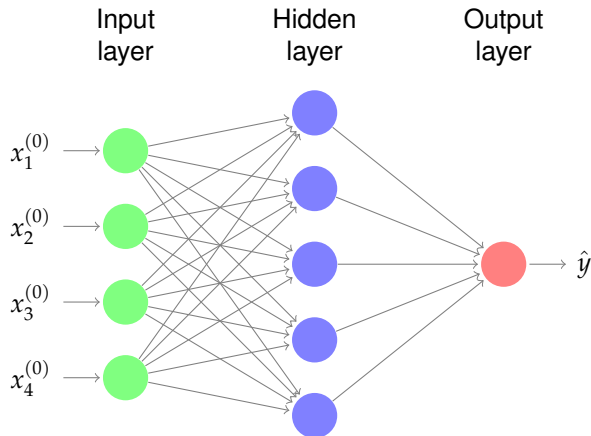


Figure: A simple MLP.

The output at the node j in layer l is denoted by $x_j^{(l)}$ and it is given by

$$x_j^{(l)} = \phi\left(\sum w_{i,j}^{(l)} x_i^{(l-1)} + b_j^{(l)}\right). \quad (50)$$

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- The last layer $\mathbb{R}^K \rightarrow \mathbb{R}$: It performs the desired ML task, either linear regression or classification.

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Training loss for a regression problem with $S_{\text{train}} = \{(\mathbf{x}_n, y_n)\}_{n=1}^N$:

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- f is the function represented by a NN.
- The overall function $y = f(\mathbf{x}^{(0)})$ can then be written as the composition:

$$f(\mathbf{x}^{(0)}) = f^{(L+1)} \circ \dots \circ f^{(2)} \circ f^{(1)}(\mathbf{x}^{(0)}).$$

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- The function that is implemented by each layer in the form

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where $w_{i,j}^{(l)}$ is the edge weight that connects node i on layer $l - 1$ to node j on layer l .

The back-propagation algorithm

Cost function:

$$\mathcal{L}_n = \left(y_n - f^{(L+1)} \circ \dots \circ f^{(2)} \circ f^{(1)}(\mathbf{x}_n^{(0)}) \right)^2,$$

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Recall that we aim to compute:

$$\frac{\partial \mathcal{L}_n}{\partial w_{i,j}^{(l)}}, \quad l = 1, \dots, L+1,$$
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- Quantity computed in the **forward pass**:

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In vector form, we can write this as

$$\delta^{(l)} = (\mathbf{W}^{(l+1)} \delta^{(l+1)}) \odot \phi'(\mathbf{z}^{(l)}), \quad (58)$$

where \odot denotes the Hadamard product (the point-wise multiplication of vectors).

Now that we have both $\mathbf{z}^{(l)}$ and $\delta^{(l)}$ let us get back to our initial goal.

$$\frac{\partial \mathcal{L}_n}{\partial w_{i,j}^{(l)}} = \sum_k \frac{\partial \mathcal{L}_n}{\partial z_k^{(l)}} \frac{\partial z_k^{(l)}}{\partial w_{i,j}^{(l)}} = \underbrace{\frac{\partial \mathcal{L}_n}{\partial z_j^{(l)}}}_{\delta_j^{(l)}} \underbrace{\frac{\partial z_j^{(l)}}{\partial w_{i,j}^{(l)}}}_{\mathbf{x}_i^{(l-1)}} = \delta_j^{(l)} \mathbf{x}_i^{(l-1)}$$

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Final computation: For all parameters compute

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Next lecture ?