

NT4

Реш

$$p(x, \theta) = \theta \cdot \{(-1, 1)/0\} + f(\theta) \cdot \{0, 2\}$$

$$\int_{-1}^1 \theta dx + f(\theta) \cdot 2 = 1$$

$$2f(\theta) = \frac{1}{2} - \theta$$

$$p(x, \theta) = \theta \cdot \{(-1, 1)/0\} + \left(\frac{1}{2} - \theta\right) \{0, 2\}$$

\vec{X}_n - выборка

$$\mathcal{L}_1 = M[S] = \int_{-1}^1 \theta x dx + 2\left(\frac{1}{2} - \theta\right) = \frac{1}{2} - 2\theta$$

$$\mathcal{L}_2 = M[S^2] = \int_{-1}^1 \theta x^2 dx + 4\left(\frac{1}{2} - \theta\right) = \theta \frac{1}{3} \cdot 2 + 2 - 4\theta = \frac{-10}{3}\theta + 2$$

$$D[S] = \mathcal{L}_2 - \mathcal{L}_1^2 = \frac{-10}{3}\theta + 2 - \left(\frac{1}{2} - 2\theta\right)^2 =$$

$$= \frac{10}{3}\theta + 1 - 4\theta^2$$

невозможность.

$$\alpha_1 = \frac{1}{n} \sum_{i=1}^n x_i = \bar{x}$$

$$\hat{\alpha} - 2\theta = \bar{x} \Rightarrow$$

$$\Rightarrow \hat{\theta}_1 = \frac{\hat{\alpha} - \bar{x}}{2}$$

вероятность:

~~MMN~~ MMN:

m-число "2"
k-число "0" } u

$$L(\theta) = \theta^{n-u} \cdot \left(\frac{1}{2} - \theta\right)^u$$

$$\ln L = (\ln \theta) \cdot (n-u) + (\ln(\frac{1}{2} - \theta)) \cdot u \rightarrow \max$$

$$0 = ()' = (n-u) \frac{1}{\theta} - \frac{u}{\frac{1}{2} - \theta}$$

$$0 = (n-u) \left(\frac{1}{2} - \theta\right) - u\theta$$

$$0 = \frac{n}{2} - n\theta - \frac{u}{2} + u\theta - u\theta$$

$$0 \Rightarrow n\theta = \frac{n}{2} - \frac{u}{2}$$

$$\hat{\theta}_2 = \frac{1}{2} - \frac{u}{2n} = \frac{1}{2} - \frac{1}{2}$$

b) Herd:

$$M[\tilde{\theta}_1] = M\left[\frac{1}{2} - \frac{1}{2} \bar{x}\right] = \frac{1}{2} - \frac{1}{2} M[\bar{x}] =$$

$$= \frac{1}{2} - \frac{1}{2} M[s] = \frac{1}{2} - \frac{1}{2} (1 - 2\theta) = \theta //$$

$$M[\tilde{\theta}_2] = M\left[\frac{1}{2} - \frac{1}{2} v\right] = \frac{1}{2} - \frac{1}{2} M[v] =$$

$$= \frac{1}{2} - \frac{1}{2} \left(\frac{1}{2} - \theta\right) \cdot 2 = \theta$$

Coem:

$$D[\tilde{\theta}_1] = D\left[\frac{1}{2} - \frac{1}{2} \bar{x}\right] = \frac{1}{4} D[\bar{x}] = \frac{1}{4n} D[s] =$$

$$= \frac{1}{4n} \left(\frac{1}{3} \theta - 1 - 4\theta^2 \right) =$$

$$= \frac{1}{4n} \left(\frac{2}{3} \theta + 1 - 4\theta^2 \right) \xrightarrow{n \rightarrow \infty} 0 \Rightarrow \text{Coem}$$

$$D[\tilde{\theta}_2] = D\left[\frac{1}{2} - \frac{1}{2} v\right] = \frac{1}{4} D[v] =$$

$$\frac{1}{4} \cdot \frac{(1-2\theta)2\theta}{n} \xrightarrow{n \rightarrow \infty} 0 \Rightarrow \text{Coem.}$$

c) $D[\tilde{\theta}_1] \cup D[\tilde{\theta}_2]$ вып. на $\forall [a, b] \in (0, \frac{1}{2})$

$\Rightarrow \tilde{\theta}_1$ и $\tilde{\theta}_2$ — вып. по гом. ун.

Регулярность Могули

1) $p(x, \theta)$ вып. по θ на Π

$$2) \int_K \frac{\partial}{\partial \theta} p(x, \theta) dx = \int_{x=1}^1 dx + (-1) \cdot 2 = 0 //$$

$$\frac{\partial}{\partial \theta} \int_K p(x, \theta) dx = 0 //$$

$$3) I(\theta) = M \left[\frac{\partial \ln p(x, \theta)}{\partial \theta} \right] =$$

$$= 2 \left(\frac{\partial \ln(\frac{1}{2} - \theta)}{\partial \theta} \right)^2 \left(\frac{1}{2} - \theta \right) + \int_{-1}^1 \left(\frac{\partial \ln \theta}{\partial \theta} \right)^2 \theta dx =$$

\uparrow вып.

$$= 2(1 - 2\theta + 2\theta) = 1 - \text{вып. гом. и } \theta > 0$$

$$= 2 \left(\frac{1}{\frac{1}{2} - \theta} \right)^2 \left(\frac{1}{2} - \theta \right) + \frac{1 + 4 - 2\theta}{\theta(\frac{1}{2} - \theta)} - \text{вып. гом.}$$

$$\theta > 0$$

Найдем правдоподоб. кр. б.

$$\frac{(\frac{1}{2} - \theta)\theta}{2(3\theta\theta\theta)n} = P[\tilde{\theta}_2]$$

$\tilde{\theta}_2$ - оценка по средним.

\sqrt{TE}

\vec{X}_n - выбор

$$L_1 = M[S] = \int_0^{\theta} \frac{1}{\theta} x dx = \frac{1}{2} \theta$$

$$L_2 = M[S^2] = \int_0^{\theta} \frac{1}{\theta} x^2 dx = \frac{1}{3} \theta^2$$

$$D[S] = L_2 - L_1^2 = \frac{1}{3} \theta^2 - \left(\frac{1}{2} \theta\right)^2 = \frac{1}{12} \theta^2$$

Нам нужно найти

$$\frac{1}{12} \theta^2 = \bar{X}$$

$$\tilde{\theta}_1 = \sqrt{\frac{12}{1} \bar{X}} = \frac{2}{3} \sqrt{3 \bar{X}}$$

или: n ММН:

$$L(\theta) = \frac{1}{\theta^n}$$

$$\ln L(\theta) = -n \ln \theta$$

$$\frac{d}{d\theta} \ln L(\theta) = -\frac{n}{\theta} = 0 \Rightarrow \theta = \frac{n}{n} = 1$$

b) Herleitung:

$$M[\tilde{\theta}_1] = M\left[\frac{2}{3}\sqrt{\frac{4}{3}}X\right] = \frac{2}{3} M[S] =$$

$$= \frac{2}{3} \cdot \frac{3}{2} \theta = \theta //$$

$$M[\tilde{\theta}_1] = M[X_{\max}^{-x_{\min}}] = \int_{\theta}^{2\theta} X d\left(\frac{X}{\theta}\right)^n =$$

$$= \int_{\theta}^{2\theta} \frac{1}{\theta} X^n \left(\frac{X}{\theta}\right)^{n-1} dX = \frac{n(2^{n+1} - 1) \theta}{n+1}$$

$$\tilde{\theta}_2' = \frac{n+1}{n(2^{n+1} - 1)} X_{\max}$$

Corr:

$$D[\tilde{\theta}_1] = D\left[\frac{2}{3}\sqrt{\frac{4}{3}}X\right] = \frac{9}{4n} \cdot D[S] = \frac{9}{4n} \left(\frac{7}{3} - \frac{3}{4}\right) \theta^2 //$$

$$D[\tilde{\theta}_2'] = D[X_{\max}] =$$

Here we have.

$$M[\tilde{\theta}_2] = M[X_{\min} - X_{\min}] = 0$$

$$P(y) = (f(y))^n (1 - f(\theta - y))^{n-1} =$$

$$= \left(\frac{1}{\theta} \left(\frac{y}{\theta}\right)\right)^n \left(1 - \frac{y}{\theta}\right)^{n-1} = \left(\frac{1}{\theta}\right)^n \left(\frac{y}{\theta}\right)^n \left(\frac{\theta - y}{\theta}\right)^{n-1}$$

$$P'(y) = \frac{1}{\theta} n \left(\frac{1}{\theta}\right)^n y^{n-1} \left(\frac{\theta - y}{\theta}\right)^{n-1}$$

$$M[\tilde{\theta}_2] = \int_0^\theta \frac{1}{\theta} n \left(\frac{1}{\theta}\right)^n y^n \left(\frac{\theta - y}{\theta}\right)^{n-1} dy = \frac{1}{\theta} n \left(\frac{1}{\theta}\right)^n \int_0^\theta y^n \left(\frac{\theta - y}{\theta}\right)^{n-1} dy =$$

$$= \int_0^\theta \frac{1}{\theta} n \left(\frac{1}{\theta}\right)^n x^n dx = \frac{1}{\theta} n \left(\frac{1}{\theta}\right)^n \frac{\theta^{n+1}}{n+1} = \frac{n}{n+1} \theta$$

$$\tilde{\theta}_2 = \frac{n}{n+1} (X_{\max} - X_{\min})$$

Conc.

$$M[\bar{\theta}_2] = \int_{\theta}^{2\theta} 2n \left(\frac{1}{\theta}\right)^n (y-\theta)^{n-1} y^2 dy =$$

$$= \frac{(8n^2 + 8n + 1)\theta^2}{2n^2 + 3n + 1} \cdot \frac{(4n^2 + 8n + 2)\theta^2}{n^2 + 3n + 2}$$

$$D[\bar{\theta}_2] = \left(\frac{4n^2 + 8n + 2}{n^2 + 3n + 2} - \frac{(4n+1)^2}{(4n+1)^2} \right) \theta^2 \xrightarrow{n \rightarrow \infty} 0$$

$n \rightarrow \infty \rightarrow 4 \quad n \rightarrow \infty \rightarrow 4$

for θ_1

$$D[\bar{\theta}_1] \xrightarrow{n \rightarrow \infty} 0 \quad \text{concl.} //$$

c) $D[\bar{\theta}_1] = \frac{18+5}{n} \quad \forall \epsilon > 0 \quad \forall n > 0 \quad D[\bar{\theta}_1] < \epsilon < D[\bar{\theta}_2]$

$D[\bar{\theta}_2] = \dots \Rightarrow \bar{\theta}_2$ - ~~is not~~ better

d) $f(\theta, \bar{x}_n) = \frac{x_{\max}}{\theta}$

$$\frac{x_{\max}}{\theta} \sim P(y) = P(y > \frac{x_{\max}}{\theta}) = P(\theta y > x_{\max})$$

$$= \left(\frac{\theta y - \theta}{\theta} \right)^n = (y-1)^n$$

$$\varphi(y) = \varphi' = n(y-1)^{n-1}$$

$$t_1 = \int_1^{t_1} \varphi(y) dy = \frac{1-\beta}{2} \Rightarrow \text{Korrektur}$$

$$(t_1 - 1)^n = \frac{1-\beta}{2}$$

$$t_1 = 1 + \sqrt[n]{\frac{1-\beta}{2}}$$

analogously

$$t_2 = 1 + \sqrt[n]{\frac{1+\beta}{2}}$$

$$1 + \sqrt[n]{\frac{1-\beta}{2}} < \frac{x_{\max}}{\theta} < 1 + \sqrt[n]{\frac{1+\beta}{2}}$$

$$\frac{x_{\max}}{1 + \sqrt[n]{\frac{1-\beta}{2}}} < \theta < \frac{x_{\max}}{1 + \sqrt[n]{\frac{1+\beta}{2}}}$$

e.) Abschätz. OMM

$$\tilde{\theta} = \frac{2}{3} \bar{x} \Rightarrow g(\bar{x}) = \frac{2}{3} \bar{x}$$

$$g(x) = \frac{2}{3} x$$

$$\sigma(x) = \sqrt{\nabla^T g(x) K \cdot \nabla g(x)} = \sqrt{\frac{2}{3} (x_2 - x_1) \frac{2}{3}}$$

$$\sqrt{n} \frac{g(x) - g(\bar{x})}{\sigma(x)} \rightsquigarrow N(0, 1)$$

$$\sqrt{n} \frac{\frac{2}{3} \tilde{I}_1 - \frac{2}{3} I_1}{\sqrt{\frac{2}{3} (\tilde{I}_2 - \tilde{I}_1) \frac{1}{3}}} = \sqrt{n} \frac{\tilde{I}_1 - I_1}{\sqrt{\tilde{I}_2 - I_1^2}} \rightsquigarrow N(0,1)$$

$$-1,96 < \sqrt{n} \frac{\tilde{I}_1 - \frac{3}{2} \theta}{\sqrt{\tilde{I}_2 - \tilde{I}_1}} < 1,96$$

$$\frac{2}{3} \left(1,96 \sqrt{\frac{\tilde{I}_2 - \tilde{I}_1}{n}} + \tilde{I}_1 \right) < \theta < \frac{2}{3} \left(-1,96 \sqrt{\frac{\tilde{I}_2 - \tilde{I}_1}{n}} + \tilde{I}_1 \right)$$

NT 6

$$a) \quad p(x) = \begin{cases} \frac{\theta-1}{x^\theta} & x \geq 1; \quad \theta > 1 \\ 0 & x < 1; \end{cases}$$

$$L(\theta) = \frac{(\theta-1)^n}{(\prod x_i)^{\theta-1}}$$

$$\ln L(\theta) = n(\theta-1) - \theta \sum x_i$$

$$\frac{\partial}{\partial \theta} (\ln L) = n - \theta \sum x_i = 0$$

$$\frac{n-1}{\theta} = \bar{x}$$

$$\tilde{\theta} = \frac{n-1}{\bar{x}}$$

$$\frac{\partial^2 L}{\partial \theta^2} = \frac{-n}{(\theta-1)^2} < 0 \Rightarrow \max$$

~~1) $p(x|\theta)$~~
~~Bestimmung des Maximums~~

$$b) f(\vec{x}_n, \theta) = ?$$

gew. mit

$$\int_{-\infty}^{\infty} p(x) dx = \frac{1}{2}$$

$$\int_1^{\tilde{x}} \frac{\theta-1}{x^\theta} dx = -\tilde{x} + 1$$

$$\tilde{x} = \left(\frac{1}{2}\right)^{\frac{1}{1-\theta}}$$

$$\text{med}(\theta) = 2^{\frac{1}{\theta-1}}$$

Perfekturform (WMO mit)

1) $p(x, \theta)$ - Kupp. gupp. mit θ an \mathbb{E} von θ

$$2) \int_1^{\infty} \frac{\partial}{\partial \theta} \frac{\theta-1}{x^\theta} dx = \int_1^{\infty} \frac{x^\theta - (\theta-1)\theta x^{\theta-1}}{x^\theta} dx = 0$$

$$\frac{\partial}{\partial \theta} \int_1^{\infty} \frac{\theta-1}{x^\theta} dx = 0$$

$$3) I(\theta) \theta$$

$$\ln(p(x, \theta)) = \ln(\theta - 1) - \theta \ln x$$

$$\frac{\partial \ln(p(x, \theta))}{\partial \theta} = \frac{1}{\theta - 1} - \ln x$$

$$\int_1^{\infty} \left(\frac{1}{\theta - 1} - \ln x \right) \frac{\theta - 1}{x^{\theta}} dx = \frac{1}{(\theta - 1)^2} - \ln 2$$

\Rightarrow непрерывно,

ОМП:

$$\sqrt{n} \frac{f(\hat{\theta}) - f(\theta)}{\sigma(\theta)} \rightsquigarrow N(0, 1)$$

$$\sigma = \sqrt{\nabla^T f(\theta) I^{-1} \nabla f(\theta)}$$

логарифм $f = \ln \theta$

$$\nabla f(\theta) = \frac{1}{\theta - 1} \ln 2 \left(-\frac{1}{(\theta - 1)^2} \right)$$

$$\sigma = \frac{\frac{1}{2(\theta - 1)} \ln 2}{\theta - 1}$$

$$g(\theta) = 2^{\left(\frac{1}{\theta - 1}\right)}$$

$$\sigma = \frac{g(\theta) \ln 2}{\theta - 1}$$

$$\sqrt{n} \frac{g(\tilde{\theta}) - g(\theta)}{\left(\frac{g(\tilde{\theta}) \ln 2}{\tilde{\theta} - 1} \sqrt{n} \right)} \rightsquigarrow N(0, 1)$$

$$g(\tilde{\theta}) - \frac{g(\tilde{\theta}) \cdot \ln 2}{(\tilde{\theta} - 1) \sqrt{n}} < g(\theta) < g(\tilde{\theta}) + \frac{g(\tilde{\theta}) \cdot \ln 2}{(\tilde{\theta} - 1) \sqrt{n}}$$

$$c) \sqrt{n} \frac{\tilde{\theta} - \theta}{\sigma(\theta)} \rightsquigarrow N(0, 1)$$

$$\sigma(\tilde{\theta}) = \theta - 1$$

$$\sqrt{n} \frac{\tilde{\theta} - \theta}{\theta - 1} \rightsquigarrow N(0, 1)$$

$$\frac{-1,96(\theta - 1)}{\sqrt{n}} + \frac{n}{\sum \ln x_i} < \theta < \frac{1,96(\theta - 1)}{\sqrt{n}} + \frac{n}{\sum \ln x_i}$$