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# Causal Inference

a summary

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# 1 General

## Causal Roadmap (Petersen and van der Laan, 2014)

systematic approach linking causality to statistical procedures

**1. Specifying Knowledge.** structural causal model (unifying counterfactual language, structural equations, & causal graphs): a set of possible data-generating processes, expresses background knowledge and its limits

**2. Linking Data.** specifying measured variables and sampling specifics (latter can be incorporated into the model)

**3. Specifying Target.** define hypothetical experiment: decide

1. variables to intervene on: one (point treatment), multiple (longitudinal, censoring/missing, (in)direct effects)
2. intervention scheme: static, dynamic, stochastic
3. counterfactual summary of interest: absolute or relative, marginal structural models, interaction, effect modification
4. population of interest: whole, subset, different population

**4. Assessing Identifiability.** are knowledge and data sufficient to derive estimand and if not, what else is needed?

**5. Select Estimand.** current best answer: knowledge-based assumptions + which minimal convenience-based assumptions (transparency) gets as close as possible

**6. Estimate.** choose estimator by statistical properties, nothing causal here

**7. Interpret.** hierarchy: statistical, counterfactual, feasible intervention, randomized trial

**Notation** chapter 1.1

**average causal effect** chapter 1.2 and 1.3 and 1.4 and 1.5

**randomized experiments (target trial)** 2.1 and 2.2; 3.6

**Standardization** plug-in (or parametric) g-formula

$$E[Y|A=a] = \int E[Y|L=l, A=a] dF_L[l]$$

weighted average of stratum-specific risks; unknowns can be estimated non-parametrically or modeled

**no need to estimate  $f_L[l]$**  as empirical distribution can be used: estimate outcome model  $\rightarrow$  predict counterfactuals  $\rightarrow$  average the results ( $\rightarrow$  CI by bootstrapping)

**for discrete  $L$**   $E[Y|A=a] = \sum_l E[Y|L=l, A=a] \Pr[L=l]$

**IP Weighting** adjust for (surrogate) confounders  $L$

$$E[Y|A=a] = E\left[\frac{I(A=a)Y}{f[A|L]}\right]; W^A = \frac{1}{f[A|L]}; SW^A = \frac{f(A)}{f[A|L]}$$

unknowns can be estimated non-parametrically or modeled

**pseudo-population:** everyone is treated & untreated ( $L \nrightarrow A$ )

**FRCISTG** (*fully randomized causally interpreted structured graph*): probability tree for  $L \rightarrow A \rightarrow Y$ , can be used to calculate/visualize simulation of values for  $A$

**for discrete  $A, L$**   $f[a|l] = \Pr[A=a, L=l]$

**estimators:** Horvitz-Thompson; Hajek (modified version)

**stabilized weights  $SW^A$**  should have an average of 1 (check!)  $\rightarrow$  pseudo-population same size  $\rightarrow$  CI width  $\downarrow$

**Standardization and IP Weighting** are equivalent,

**but** if modeled, different “no misspecification” assumptions:

standardization: outcome model

IP weighting: treatment model

**doubly robust estimators:** reduce model misspecification bias,

consistent if either model is correct; **e.g.:**

1. fit outcome regression with variable  $R = \begin{cases} +W^A & \text{if } A=1 \\ -W^A & \text{if } A=0 \end{cases}$
2. calculate counterfactual results

**identifiability conditions** most of 3

positivity: p. 155, p. 162

additional conditions: chapter 13.5

exchangeability: p 172f

**effect modification** chapter 4

**interaction** chapter 5

**causal diagrams** chapter 6, include swigs from 7.5 and that one technical point

**confounding** chapter 7

**selection bias** chapter 8

**measurement bias** chapter 9

**random variability** chapter 10

## 2 Models

**Modeling** data are a sample from the target population

*estimand*: quantity of interest, e. g.  $E[Y|A=a]$   
*estimator*: function to use, e. g.  $\hat{E}[Y|A=a]$   
*estimate*: apply function to data, e. g. 4.1

**model**: a priori restriction of joint distribution/dose-response curve; *assumption*: no model misspecification (usually wrong)

**non-parametric estimator**: no restriction (saturated model) = *Fisher consistent estimator* (entire population data  $\rightarrow$  true value)

**parsimonious model**: few parameters estimate many quantities

**bias-variance trade-off**:

wiggleness  $\uparrow \rightarrow$  misspecification bias  $\downarrow$ , CI width  $\uparrow$

**Marginal Structural Models** association is causation in the IP weighted pseudo-population

associational model  $E[Y|A] =$  causal model  $E[Y^a]$

*step 1*: estimate/model  $f[A|L]$  (and  $f[A]$ )  $\rightarrow$  get  $(SW)^A$

*step 2*: estimate regression parameters for pseudo-population

**effect modification** variables  $V$  can be included (e. g.

$\beta_0 + \beta_1 a + \beta_2 V a + \beta_3 V$ ; technically not marginal anymore),  
 $SW^A(V) = \frac{f[A|V]}{f[A|L]}$  more efficient than  $SW^A$

**Censoring** measuring joint effect of  $A$  and  $C$

$E[Y^{a,c=0}]$  is of interest

**standardization**  $E[Y|A=a] = \int E[Y|L=l, A=a, C=0] dF_L[l]$

**IP weights**  $W^{A,C} = W^A \times W^C$  (uses  $n$ ) or  
 $SW^{A,C} = SW^A \times SW^C$  (uses  $n^{c=0}$ )

**g-estimation** can only adjust for confounding, not selection bias  
 $\rightarrow$  use IP weights

**G-Methods** generalized treatment contrasts: adjust for (surrogate) confounders  $L$

- **standardization**
- **IP weighting**
- **g-estimation**: not needed unless longitudinal

**G-Estimation** (additive) structural nested models

$$\text{logit Pr}[A=1|H(\psi^\dagger), L] = \alpha_0 + \alpha_1 H(\psi^\dagger) + \alpha_2 L$$

$$H(\psi^\dagger) = Y - \psi_1 A$$

find  $\psi^\dagger$  which renders  $\alpha_1 = 0$ ; 95%-CI: all  $\psi^\dagger$  for which  $p > 0.05$   
closed-form solution for linear models

**derivation**:  $H(\psi^\dagger) = Y^{a=0}$

$$\text{logit Pr}[A=1|Y^{a=0}, L] = \alpha_0 + \alpha_1 Y^{a=0} + \alpha_2 L$$

$Y^{a=0}$  unknown, but because of exchangeability  $\alpha_1$  should be zero

$$Y^{a=0} = Y^a - \psi_1 a$$

equivalent to  $Y^{a=0} = Y^{a=1} - \psi_1$ , but using no counterfactuals

**structural nested mean model**

$$\text{additive: } E[Y^a - Y^{a=0}|A=a, L] = \beta_1 a + (\beta_2 a L)$$

$$\text{multiplicative: } \log\left(\frac{E[Y^a|A=a, L]}{E[Y^{a=0}|A=a, L]}\right) = \beta_1 a + (\beta_2 a L)$$

multiplicative is preferred if  $Y$  always positive, but does not

extend to longitudinal case

semi-parametric: agnostic about  $\beta_0$  and effect of  $L \rightarrow$  robust  $\uparrow$

**no time-varying**: no nesting; model equals marginal structural models with missing  $\beta_0, \beta_3$  (unspecified “no treatment”)

**sensitivity analysis**: unmeasured confounding ( $\alpha_1 \neq 0$ ) can be examined: do procedure for different values of  $\alpha_1 \rightarrow$  plot  $\alpha_1$  vs.  $\psi^\dagger \rightarrow$  how sensitive is estimate to unmeasured confounding?

**effect modification**: add  $V$  in both g-estimation equations

**doubly robust estimators** exist

**Outcome regression** chapter 15

**instrumental variable estimation** chapter 16

**causal survival analysis** chapter 17

**Variable Selection** can induce bias if  $L$  includes:

(decendant of) collider: *selection bias under the null*

noncollider effect of  $A$ : *selection bias under the alternative*

mediator: *overadjustment for mediators*

temporal ordering is not enough to conclude anything

**bias amplification**: e.g. by adjusting for an instrument  $Z$  (can also reduce bias)

**Machine Learning**  $L$  is high-dimensional

use lasso or ML for IP weighting/standardization

**but**: ML does not guarantee elimination of confounding and has largely unknown statistical properties

$\rightarrow$  **doubly robust estimator**: consistent if bias  $< \frac{1}{\sqrt{n}}$

*sample splitting*: train estimators on training sample, use

resulting estimators for doubly robust method on estimation sample (CIs on estimation sample are valid, but  $n$  halved)

*cross-fitting*: do again the other way round, average the two estimates, get CI via bootstrapping

**problems**: unclear choice of algorithm, is bias small enough?

# 3 Longitudinal Data

**Time-Varying Treatments** compare 2 treatments

treatment history up to  $k$ :  $\bar{A}_k = (A_0, A_1, \dots, A_k)$

shorthand: always (never) treated  $\bar{A} = \bar{1}$  ( $\bar{0}$ )

**static strategy:**  $g = [g_0(\bar{a}_{-1}), \dots, g_K(\bar{a}_{K-1})]$

**dynamic strategy:**  $g = [g_0(\bar{l}_0), \dots, g_K(\bar{l}_K)]$

**stochastic strategy:** non-deterministic  $g$

optimal strategy is where  $E[Y^g]$  is maximized (if high is good)

**identifiability** middle chapter 19

**treatment-confounder feedback** end chapter 19 and chapter 20

**g-formula** chapter 21.1

**IP weighting** chapter 21.2

**doubly robust estimators** chapter 21.3

**g-estimation** chapter 21.4

**censoring** chapter 21.5

**target trial** chapter 22 (does that even really fit in here, maybe push to 3rd paragraph in without models)

# References

*If no citation is given, the source is (Hernán and Robins, 2023)*

Hernán, M. A. and Robins, J. M. (2023). *Causal inference: what if*. Boca Raton: Chapman & Hall/CRC.

Petersen, M. L. and van der Laan, M. J. (2014). Causal models and learning from data: integrating causal modeling and statistical estimation. *Epidemiology (Cambridge, Mass.)*, 25(3):418–426.

