

# Testing Epistemic Democracy's Majority Rule Result

## Abstract

While epistemic democrats have claimed that majority rule comes to recruit the wisdom of the crowds to identify correct answers to political problems, it has yet to be demonstrated that elections actually produce outcomes in this way. This paper illustrates how elections can plausibly be thought to leverage the epistemic capacity of the electorate to enhance the instrumental value of elections via majority rule. To do so we offer a set of sufficient conditions that effect such 'majority rule results,' even when the decision in question is multidimensional. We then look to the case of sociotropic economic voting in U.S. presidential elections to provide empirical support for the tractability of these conditions. We find that absent such an epistemic capacity, seven presidential elections since 1956 might well have been decided differently. By generating clear conditions for the plausibility of claims made by epistemic democrats, and demonstrating their correspondence to empirical data, this paper strengthens the broader instrumental grounds which recommend democracy.

**Keywords.** Democratic theory, epistemic democracy, majority rule, economic voting, plausibility analysis, Condorcet Jury Theorem

## 1 Introduction

Democratic theory’s epistemic turn ought to be exciting for political scientists and political philosophers alike. Beyond the standard endorsements of democracy on procedural, outcome independent grounds, epistemic democracy offers insight into how and why democracy succeeds at getting things “right.” We’ll refer to these sorts of mechanisms as *veritistic* since they get at the truth. In this way, democratic decision making can be recommended not only for fostering autonomy and fairness, but also on these instrumental grounds. Epistemic democrats have pointed toward what we will refer to broadly as ‘majority rule results’ such as the Condorcet Jury Theorem (CJT) and the Miracle of Aggregation (MA) in order to argue that democracy has a tendency to select the correct candidate or proposal in an election (Landemore (2013); Surowiecki (2004); List and Goodin (2001)). Since free and fair elections have become democracy’s *sine qua non*, these majority rule results provide good grounds for recommending democracy more broadly.

Indeed, Aristotle already noted in book three of the *Politics* that “the many, who are not as individuals excellent men, nevertheless can, when they have come together, be better than the few best people, not individually but collectively, just as feasts to which many contribute are better than feasts provided at one person’s expense” (Reeve (1998): 83). This contention, that the fruits of the many are superior to the fruits of the few, is foundational to politics. The collective domain of politics is efficacious in virtue of its propensity to generate collective outcomes superior to those that an individual alone could achieve.

But it is not obvious when Aristotle’s contention is correct, or even what constitutes correctness. A nagging worry remains that beyond aggregating beliefs, votes also tally people’s divergent values. The examples epistemic democrats frequently use are cases where the criteria of correctness are uncontested, such as a guilty verdict or an ox’s weight. Political disagreements, however, are commonly disputes of value rather than of attendant facts or consequences. Disagreements such as gun control, abortion, and civil rights all lack a clear right answer. For this reason many have argued that the machinery of epistemic democracy cannot plausibly endorse electoral outcomes (Anderson (2006); Estlund (1997); Estlund (2008); Ingham (2012); Urbinati (2014)).

In this paper we push back against such skepticism and argue that though

some political disagreements may be disputes of value, other disagreements pertain to facts; there is broad consensus on ends, but not on which means are most likely to effect those ends. Regarding these latter disputes, the result of a democratic election can be considered veritistic. Were voters epistemically biased against the candidates that advanced their commitments, the institution would systematically produce ‘unwanted effects,’ where a constituency would be less likely to obtain its preferred outcome by voting for it. We provide a set of conditions where, to the extent that voters substantively agree on some issue that affects an election’s outcome, their votes can be thought to nudge the election in favor of the ‘right’ candidate—the candidate that they would endorse on some particular grounds given full information—thereby inhibiting unwanted effects.<sup>1</sup>

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Since the claim is instrumental in kind—democracy is good in virtue of the outcomes it produces—it is still necessary to demonstrate that the processes in question arrive at the stipulated ends, as a matter of fact. Previous treatments of majority rule results don’t provide sufficient evidence to support their empirical relevance (Schwartzberg (2015)). And though there are matters on which we imagine consensus exists, such as public health or the competence of political officials (e.g. Page (2007): 256-7; Landemore (2013): 145), there has been no move to identify a case in which the conditions of majority rule results are met. By going beyond the formal results to provide an empirical case that tracks the conjectures of epistemic democrats, this paper both makes good on admonishments to demonstrate the plausibility of normative theory (e.g. Rehfeld (2010), Wiens (2015)), while also pointing to the purchase of arguments made by epistemic democrats to recommend democracy against alternative regimes.

Our plan for this paper is to first explain how majority rule yields veritistic outcomes on a single dimension of a decision function, independent of other contested components. In general, no one issue will itself determine the results of an election, regardless of the number of voters, their competence, or the extent of their agreement. However, when voters agree on a the desirability of a particular outcome—though they may disagree on how best to achieve it—a candidate’s superiority on that issue may increase her chances of winning the election. We specify three jointly sufficient conditions under which this property will hold. The formal proposition we present shares the core probabilistic assumptions as CJT and MA, albeit in a weaker form, but allows for multidimensional considerations. (Section 7.2 addresses the similarities and differences

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<sup>1</sup>Indeed, assumptions such as single-peakedness can also be employed to make sense of aggregating values. These routes, however, would preclude the kind empirical analysis we provide, since they would necessitate considering both the strength and accuracy of the “yeas” and “nays,” which our data just don’t permit.

between CJT and our approach.)

Next, we turn to the case of sociotropic economic voting in U.S. presidential elections to illustrate the empirical plausibility of the majority rule results. That voters by and large want economic growth provides a ready case where a majority of voters agree on some value, leaving that facts to be contested. Our analysis here builds on a trove of both theoretical and empirical evidence for the retrospective economic voting model (Duch and Stevenson (2008), Lewis et al. (2009)). Beyond their own economic well-being, voters select the U.S. president with an eye to recent economic trends. We recruit these findings and models in the service of our point here: that given the electorate’s preference for economic growth, their epistemic capacities allow them to select the better candidate.

This serves as empirical example of the epistemic quality of the majority rule results. Here, insofar as people substantively agree on a particular value (increased  $\Delta RDI$ ), and are in a position to judge that fact of economic growth (whether  $\Delta RDI$  has increased), the election can be thought to be biased in favor of the right candidate—“better” candidate—the one voters themselves would have chosen given full information. Though this effect needn’t be decisive, we estimate that without it the result of seven presidential elections from 1956–2012 would have been different. The formal proof of our result and technical discussion of the models are reserved for the appendix.

## 2 Evaluating Democracy’s Truth-Tending Properties

While there are many instances in which it is difficult to determine what a community should do or which values it ought to endorse, there might be issues where there is broad normative agreement, leaving the facts of the matter to be contested. Consider a multidimensional decision illustrated here by voters’ choices in the 2012 U.S. Presidential election. Along one dimension Alice and Bernice fundamentally disagree about the access women ought to have to abortions. Since Alice identifies as pro-life, she is more likely to vote for Mitt Romney. Bernice, on the other hand, is pro-choice and feels strongly that President Obama should be reelected. Along another dimension, however, both Alice and Bernice agree that the government should effect the greatest number of Americans to carry health insurance. Given that they both endorse this same value, cashed out using the same metric, their disagreement can be construed as partly factual in kind (Page (2007): 258)—which candidate will in fact lead more Americans to be insured? If Alice and Bernice are each better than random at determining the fact of the matter, we then have reason to think that

the candidate more adept at implementing healthcare policy has his chances of winning boosted. Holding their disagreement about abortion constant, the candidate's fitness on the second dimension enhances his electability.

Modeling votes in a majoritarian election with two candidates as binary random variables, we can formally establish the link between voters' agreement on the value of some issue and electoral results, given some additional assumptions. The assumptions may be regarded as conditional on the facts relevant to the election. The first assumption instrumentalizes voters' agreement on the issue:

**Assumption 1.** *Every voter is at least as—and some more—likely as otherwise to select a candidate if she believes that the candidate is superior to her rival on that issue*

The second assumption characterizes voters' competence: that their beliefs correlate with reality.

**Assumption 2.** *Each voter is more likely to believe a candidate is superior when she actually is than when she is not.*

Were this otherwise, the instrumental value of the contest would be at best null and at worst it would hazard producing unwanted effects, whereby voting for a candidate  $A$  in virtue of some criterion  $c_i$  would diminish the chances that  $c_i$  would be realized (say, in this case, more wide-spread healthcare).

Assumptions 1 and 2 help determine expected vote totals:

**Proposition 1.** *Under assumptions 1 and 2,  $A$  candidate's superiority to her rival on that issue increases her expected vote total.*

Increasing a candidate's expected vote total does not necessarily imply increasing her chances of winning; for that we need an assumption about the statistical relationship between votes. The simplest (if, perhaps, the strongest) such assumption is independence:

**Assumption 3.** *Votes are mutually independent*

Then we have the following proposition:

**Proposition 2.** *Under Assumptions 1, 2, and 3, a candidate's superiority to her rival on that issue increases her probability of being elected.*

Our claim is this: in a binary election (even when the decision function is multidimensional) where independent and minimally competent voters seek to

select the candidate that maximizes a shared value, the majority rule mechanism comes to recommend democracy on veritistic grounds.<sup>2</sup>

Though the result is easiest to see in the case of consensus, where everyone agrees to some value, it's not limited to such cases. The claim here is *insofar as people agree*, the outcome of the election can be thought to be truth-tracking by epistemic means. Instead of thinking of complete consensus of the electorate, consider instead a majority consensus, where a simple majority of Americans want more people to carry health insurance. Here the process can be thought to be veritistic when the majority agrees to the value of some issue and has the capacity to select the candidate that is in fact better on that issue—they have the capacity to obtain the outcomes they so desire. Insofar as this majority wants to elect the candidate that will increase healthcare coverage, the candidate better on this issue will have her chances of victory boosted. Importantly, they need not possess epistemic access to the policy mechanisms—whether single-payer is superior to market-based exchanges—for the election's result to maintain a veritistic quality. It need only be the case that the outcome be the intended one.

This doesn't have to be a trivial result. Sure, just because a candidate is better on *some* issue that a majority agrees to enhances her chances of victory, doesn't mean this bump is close to decisive. That being said, we show here that the effect of sociotropic economic voting in U.S. presidential elections has had consistent and measurable effects on (popular) electoral outcomes.

### 3 The Empirical Plausibility of Democracy's Truth-Tending Properties

The proposition above outlines epistemic considerations relevant to majoritarian voting, stipulating conditions under which a candidate's fitness on some issue increases the probability of her being elected. The practical relevance, however, hinges on whether the conditions actually obtain. To establish this we need to identify a case that tracks how people evaluate some particular and salient dimension of their voting decision. As such, we now turn to the case of sociotropic economic voting in U.S. presidential elections.

Though economic voting is not a part of every election (Stein (1990); Nadeau and Lewis-Beck (2001):171), U.S. presidential elections do appear to turn on economic performance (Miller and Wattenberg (1985), Fiorina (1978); Lockerbie (1992); Lanoue (1994); Lewis-Beck and Stegmaier (2000); Nadeau and Lewis-Beck (2001); Markus (1988)). Voters appear to hold the president responsible for

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<sup>2</sup>The technical restatement of this proposition and its proof for the appendix.

not only their own financial well-being, but also for the health of the domestic economy more broadly—referred to as “sociotropic” economic voting (Kinder and Kiewiet (1981)). Moreover, it is likely that the president has some control over macroeconomic outcomes (Blinder and Watson (2014)). This kind of sociotropic voting reveals that while people collectively seek to endorse the candidate that best serves the American economy, they may disagree about which party is more apt for the task. We look to economic voting, in part, because of the theory’s robustness. Both Duch and Stevenson (2008) and Lewis et al. (2009) review the sizable literature on the matter and argue forcefully that there exists a strong causal relationship between past sociotropic economic performance and people’s assessment of the president. We use this example in part because of the broad theoretical and empirical support for economic voting.<sup>3</sup>

For our purposes is not enough to argue that people share the same value in principle, they must also render it in the same way, that is using the same metric. While any number of criteria plausibly affect people’s judgement about the economy (e.g. inflation, unemployment, trade deficits), we find that the annual change in real disposable income ( $\Delta RDI$ —average income, less taxes) highlights a broad agreement that exists and informs voters’ decisions of whom to vote for.  $\Delta RDI$  is an attractive metric not only because it corresponds to an objective measure of economic performance, but also as people plausibly have epistemic access to it. People can look around and assess the after-tax money that they and those whom they know have earned in recent years. Voters need not diligently read the Wall Street Journal or watch CNBC to know whether the economy around them appears to be growing or shrinking.<sup>4</sup> As evidence for this claim, we show that RDI growth is an important predictor of individual votes, even after controlling for several other canonical variables.<sup>5</sup>

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<sup>3</sup>It is worth noting that while voters respond to continued economic growth, the substance of voters’ judgements isn’t entirely clear. It does not appear, for instance, that the judgement is terribly forward looking—which party *will* handle the economy better (e.g. Nadeau and Lewis-Beck (2001); Woon (2012)). The substance of the judgement—why retrospective economic progress is important—is not, however, of immediate interest here. Because our concern is veritistic (i.e. are people right about the facts which inform their judgements) it is sufficient to note that people belief’s regarding economic growth are accurate and facilitate their decisions.

<sup>4</sup>While wages have largely stagnated over the last four decades, the CBO shows slight growth in even the bottom quartile. Even so, since our inquiry is primarily interested in sociotropic rather than pocketbook voting, we believe that this phenomenon does not undermine the strength of our findings (Congress (2011)).

<sup>5</sup>While we use  $\Delta RDI$  in order to follow the lead of Nadeau and Lewis-Beck (2001) (among other considerations), the results are consistent for GDP growth too.

### 3.1 The Empirical Results

To get a handle on the influence of macroeconomic performance on voters' beliefs and choices, we take a cue from Nadeau and Lewis-Beck (2001) and fit a logit model to American National Election Survey (ANES) data.<sup>6</sup> Our models (Table 1) provide evidence that  $\Delta RDI$  predicts voters' choices as mediated by their beliefs, and corroborates earlier studies which have shown that  $\Delta RDI$  is a strong predictor of the votes for U.S. president.<sup>7</sup> But when the model also accounts for respondents' subjective retrospective economic beliefs (models (2) and (3)), the magnitude of  $\Delta RDI$ 's coefficient diminishes to near zero and becomes statistically insignificant. Consistent with Nadeau and Lewis-Beck's account,  $\Delta RDI$ 's impact on votes is mediated by a person's subjective beliefs of the economy. The level of  $\Delta RDI$  affects subjects' perceptions of the economy, which in turn affect their voting decisions. The role of beliefs here is evidence for the epistemic arguments advanced by democratic theorists.

As in Nadeau and Lewis-Beck's model, subjective measures have a large, positive association with  $\Delta RDI$ . People's assessment of the national change in income levels appears to make up a sizable part of what their retrospective variable measures. We find that retrospective assessments of the economy are correlated with  $\Delta RDI$  at a level of 0.30, consistent with Nadeau and Lewis-Beck's finding that when *retrospective* is regressed on  $\Delta RDI$  the  $R^2$  value is 0.77, indicating that people's economic beliefs correspond to empirical fact (Nadeau and Lewis-Beck (2001): 161, 174). This is consistent with the arguments of epistemic democrats who argue that people's correct beliefs come to be aggregated in democratic contests to bias the outcome in favor of the better candidate. Indeed we estimate that without this effect (that is, were people only no better than random at assessing  $\Delta RDI$ ) elections in 1960, 1968, 1976, 1980, 2000, 2004, and 2012 would have had different outcomes.<sup>8</sup>

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<sup>6</sup>While our models rely on the substantive theory behind those in Nadeau and Lewis-Beck (2001), the fresh analysis allows us to fix some methodological concerns as well as direct attention to the specific ways in which beliefs mediate between voters' observations and decisions (votes). The technical discussion is provided in the appendix.

<sup>7</sup>An increase of one in unit  $\Delta RDI$  increases the odds of a respondent voting for the incumbent by a factor of about 5–36%, adjusting for voters' perceptions of their individual financial situations, their race, their state, and their party identification, as well as the election year. This is an approximate 95% confidence interval: the coefficient estimate plus or minus two standard errors, exponentiated to transform log-odds ratios to odds ratios.

<sup>8</sup>That is, with regard to the popular vote.



	<i>Dependent variable:</i>		
	vote		
	(1)	(2)	(3)
$\Delta$ RDI	0.200** (0.074)	-0.003 (0.085)	0. 0.010 (0.089)
finances	0.303** (0.025)	0.162** (0.032)	0.154** (0.033)
incumbentParty	-0.521** (0.126)	-0.522** (0.139)	-0.497** (0.146)
retrospective		1.019** (0.058)	0.950** (0.062)
prospective			0.259** (0.042)
incumbentParty:race	0.748** (0.054)	0.635** (0.061)	0.574** (0.064)
incumbentParty:partyID	0.824** (0.011)	0.819** (0.014)	0.800** (0.014)
Constant	-0.451* (0.218)	0.029 (0.220)	-0.024 (0.231)
Observations	20,168	13,988	12,746
Log Likelihood	-7,664.552	-4,890.587	-4,508.239
Akaike Inf. Crit.	15,345.100	9,799.174	9,036.478

*Note:* \*p<0.05; \*\*p<0.01

Table 1: Results from three multilevel logistic regressions described in equations (3) and (4) in the appendix.  $\Delta RDI$  is the percent change in national real disposable income per-capita from the previous year; incumbentParty is equal to 1 when Democrats are incumbent and -1 when Republicans are; finances is equal to 1 when respondents answer that their family's financial situation is better than a year ago, -1 when worse and 0 when the same; prospective is 1 when respondents answer that they expect the economy to improve in the following year, -1 when they expect it to get worse and 0 if they expect it to stay the same; retrospective is 1 when respondents answer that they believe the economy improved in the previous year, -1 when they believe it got worse and 0 if they believe it stayed the same; race is equal to 1 if the respondent is non-white and 0 if the respondent is white; partyID is a five-point scale for party identification: positive for Democrats, negative for Republicans; 3 for strong, 1 for weak or leaning, and 0 for apolitical. Models based on ANES data from presidential elections from 1956–2012 (1) or 1980–2012 (2) and (3).

### 3.2 Epistemic Heterogeneity in the Model

The conditions of the majority rule mechanism regard individual level properties (beliefs, values) while models (1)–(3) estimate aggregate association between people’s beliefs of economic progress and actual economic growth. Results in Table 1 suggest that voters *as a whole* are both more likely to vote for the incumbent party the higher  $\Delta RDI$  (model (1)) and more likely to vote for the incumbent party when they believe the economy is growing (models (2) and (3)), from which we conclude that voters as a whole select the candidate they believe is better on economic issues. It’s conceivable, though, that a substantial portion of voters are systematically mistaken regarding economic growth, believing the economy is growing when it is shrinking and vice versa. Such voters violate assumption 2 and, if their number is sufficient, undermine the majority rule mechanism. their individual abilities to judge the facts.

While the data don’t allow us to track individual competencies, we can look at subgroups within the population to assess whether any are systematically confused. The model reported substantial variation in the correlation between  $\Delta RDI$  and *retrospective* as a function of a variety of respondents’ demographic features (finances, party identification, race, age, class, education, gender, marital status and urbanism), no combination thereof reported a negative correlation. Though this does not imply that every voter is a good judge of the economy, it suggests that very few voters are particularly poor judges.

Another worry is that it’s *possible* that the result is produced by a preferences of a minority of electors rather than a majority.<sup>9</sup> We try to get a handle on this by estimating the number of those who want the economy to contract is to ask what proportion of survey respondents answered that “the economy was doing much better” given a value of  $\Delta RDI$  in the bottom quartile. This puts the number at 1.2% (with a 95% confidence interval ranging form 0.9-1.5%), but that is likely a large overestimate. The fact that surveys don’t even ask respondents whether they wish the economy would contract is possibly the best evidence of the ubiquity of the assumption. Given this, it is just not possible for the observed effect to be due to a minority of voters who want the economy

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<sup>9</sup>More precisely, the worry is as follows: The proposition claims that insofar as people agree and both know something about economic progress and vote with regard to economic progress, the majority rule results can be thought to obtain. We then specify that the argument is not that everyone need agree that positive  $\Delta RDI$  is a good, but insofar as the majority believe it to be a good, our evidence demonstrates that they have the capacity to select the better candidate in this regard. What we want to show here is that the outcomes of our empirical model are the result of the majority getting it right (they want the economy to grow, they believe it is growing, and it is in fact growing), rather than a minority getting it wrong (they want the economic to shrink, they believe it is shrinking, and it is in fact growing).

to shrink.

## 4 Critiques

There are, however, standing critiques of the economic voting literature. Christopher Anderson, for instance, doubts the link between economic performance and electoral outcomes (Anderson (2007): 272), particularly given the myriad of other metrics that might motivate voters’ decisions (ibid.: 274). Furthermore, he along with Bryan Caplan call into question whether the information voters do have allows them to make such inferences about the economy (Anderson (2007): 279-281; Caplan (2006): chapter 3).

We needn’t suppose that most voters *directly* observe  $\Delta RDI$  in order for them to care about its performance, however.  $\Delta RDI$  might well serve as a proxy voters observations of local and national economic trends. For instance, percent change in  $\Delta RDI$  is correlated with percent change in GDP ( $\rho = 0.71$ ) and inversely correlated with the change in unemployment ( $\rho = -0.43$ ).<sup>10</sup> Yet  $\Delta RDI$  is also something that voters can plausibly look around and observe—are those I see taking home more or less pay than in past? have their consumer habits changed? are they tightening their belts? This contention is consistent with our finding that  $\Delta SDI$  (state-level real disposable income) also has marginally significant effect (at a 10% level) on presidential vote share, loosely suggesting that people make political inferences using both national and regional economic signals (though the former is clearly more pronounced). But we also have good theoretical reasons to think that voters have access to rich empirical knowledge. Arthur Lupia and others have written that voters can key-in to sophisticated information using proxies and shortcuts which allow them to make refined decisions (Kinder and Kiewiet (1981): 130-1; Lupia and McCubbins (2000); Lupia (2006)).

Meanwhile Caplan’s strongest counterexamples indicate that voters are woefully ignorant about economic mechanisms such as whether “technology is displacing workers” (Caplan (2006): 65). As mentioned above, it is not necessary for voters to have access to the causal mechanisms that produces some outcome for the election to be thought veritistic. The question of whether RDI is growing at a normal clip is such a question that regards performance rather than policy, and which economists and lay-voters both do a fairly good job of assessing (Caplan (2006): 78). Whether take-home income has risen or fallen is something

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<sup>10</sup>Data Source: FRED, Federal Reserve Economic Data, Federal Reserve Bank of St. Louis: <http://research.stlouisfed.org/fred2/graph/?g=DQL> accessed 06/20/2014. Estimates are based on quarterly data from 1960–2014.

that people realistically have access to. The 0.68 correlation between RDI in successive years indicates that there is a rather low epistemic burden to predicting whether the economy will grow or shrink in the coming year. Just knowing the previous year’s RDI provides a good deal of predictive power absent any sophisticated training. Indeed, even Achen and Bartels (2016), who are not kind to epistemic interpretations of democracy, carve out an exception for economic voting (ibid.: 97-8).

## 5 Plausibility of the Theory in Light of the Evidence

To review, epistemic democrats claim that if voters are marginally better than random at assessing some set of facts, then they are more likely to select the “right” candidate than not. We stipulate three conditions that are sufficient for the theory to hold, including that voters are independent and minimally competent. To examine plausibility of these claims we offer a linear model of voters’ decisions in U.S. presidential elections. Here voters observe national economic conditions, which inform their beliefs about the economy’s performance. Those beliefs, in turn, affect voters’ choices. Voters who believe that the economy has improved or will improve are more likely to vote for the incumbent party.

These results, together with the rest of the vast economic voting literature, give us good reason to think that the majority rule results do indeed obtain, thereby indicating the epistemic virtues of democracy—that people’s votes can tend to select the candidates they want. The coefficient on  $\Delta RDI$  is substantial and significant only when economic beliefs are excluded from the model. Once included, the subjective measures soak up its explanatory role, illustrating the role of beliefs in affecting political outcomes.

The theory is premised on people substantively agreeing on the value of some issue. While we do not show that everyone wants the economy to grow, we do provide evidence that the vast majority do, indicating that a correspondence between beliefs and economic trends on the one hand, and outcomes (vote share) on the other, result from the choices of a majority of voters. Condition one, that voters are as good or better than random at selecting the superior candidate, is tricky to substantiate with direct observational evidence. Our subgroup analysis, however, goes some distance to allay worries, as we fail to find any combination of subgroups for which a positive  $\Delta RDI$  is anti-correlated with the variable *retrospective*. Our empirical model supports this contention, too. As such we argue that the majority rules results, touted by epistemic democrats, inhibit unwanted effects from voting.

Our models' results also speak to assumption two, that voters' are not less likely to select a candidate they believe to be better on the issue. Given our finding that voters' subjective impressions of economic performance in models (2)–(3) diminishes  $\Delta RDI$ 's predictive power, we can infer that it is their beliefs that affects their vote, substantiating the epistemic valence of democracy. If we account for the fact that voters' perceptions of the economy are colored by their perceptions of the political parties and candidates (e.g. Evans and Andersen (2006)), then subjective sociotropic variables measure something akin to voters' opinions of which party (and hence candidate) is most fit to manage the macroeconomy. The objective variable  $\Delta RDI$ , short term economic growth, does not reflect voters' subjective beliefs, but, rather, objective conditions. Therefore, since including subjective sociotropic measures in the regression leads to a near-0 coefficient on  $\Delta RDI$ , we conclude that, given voters' *beliefs* about a candidate, the candidate's actual ability does not affect voters' choices.

Finally condition three is a simplifying assumption intended primarily to streamline the proof. Indeed we know that votes are not cast independently (e.g. Sinclair (2012)), though it is hard to know the size of the effect of dependence. We put forward what can be considered the least complicated assumption and leave it to others to identify circumstances where dependent voting would undermine our results. There is no reason to think that our results disappear given small amounts of dependence. And as Dietrich and List point out (2004), independence most plausible when we have reason to think that voters are encountering heterogeneous information, as would be the case in a nation-wide presidential election with comparably high turnout.

The proposition requires assumptions 1–3 after holding all else constant. This caveat is necessary because a candidates' superiority on one issue may correlate with an inferiority (in some voters' minds) on another. This would be a problem if, say, candidates who are better able to increase  $\Delta RDI$  tend to be Republican, so many voters who identify as Democrats would tend to prefer the economically inferior candidate. However, conditional on party identification, Democrats and Republicans alike would prefer economically superior candidates. Seen this way, conditioning variables such as party identification weaken the assumptions. Models (1)–(3) condition on respondents' finances, race, and party identification, and the incumbent party; the model results support the assumptions conditional on those variables.

Our models speak directly to conditions one and two. Condition three must remain assumption, though the diversity of the electorate in U.S. presidential elections makes it an eminently reasonable one. What we do here is outline com-

mon assumptions to gain traction on how democracy can leverage the wisdom of the crowds and demonstrate the theory’s correspondence to data.

## 6 Conclusion

When political theorists argue that majority rule results, such as the Condorcet Jury Theorem or Miracle of Aggregation, pertain to politics they do so assuming there exist cases in which these results actually obtain. We point out that such instrumental claims are weak absent a compelling example, which is non-trivial given that most decisions entail a bundle of propositions, many of which are normatively contested. We show that, given certain conditions, insofar as electors hold some value in common, democracy can be thought to be recommended veritistically, even if voters fundamentally disagree regarding many other aspects of the vote.

The example of sociotropic economic voting in U.S. presidential elections offers an instance where a widely agreed upon value significantly effects an election’s outcome, thereby plausibly meeting these conditions. In this binary contest, the data indicate that voters select candidates with an eye to the annual change in real disposable income, among other considerations. The broad agreement here comes to recommend the election’s result by nudging the outcome in favor of the candidate that voters themselves would have chosen given full information. What’s more, it provides evidence that people’s perceptions affect their votes through epistemic means—that the effect of economic growth on electoral outcomes is mediated by people’s beliefs of how the economy is fairing. Indeed we estimate that absent this effect outcomes of the popular vote in 1960, 1968, 1976, 1980, 2000, 2004, and 2012 would have been different.

Beyond the support of claims made by epistemic democrats, this paper shows the feasibility and usefulness of testing theoretical conjectures against empirical data. Whenever we make instrumental claims in politics, we do so conditional on a certain state of the world being the case. If we are to take these arguments seriously more needs to be done to determine that the empirical conditions are indeed satisfied. We hope to have offered one such effort here.

## 7 Appendix

### 7.1 Proof of Proposition

In order to demonstrate our claim that *in a binary election where votes are independently cast and, in part or whole, seek to select the candidate that maxi-*

mizes a shared value, majority rule results recommends democracy on veritistic grounds we will prove a more general conjecture which admits the proposition in Section 2 as a special case, which we then state as a corollary.

We are interested in the electability of a particular candidate in a binary election. Let  $E = 1$  if she is elected, and 0 otherwise. Say there is a fixed subset of the electorate,  $\mathcal{S}$ , such that  $E = 1$  if and only if the proportion of votes for a candidate from  $\mathcal{S}$  exceeds a pre-determined (if unknown) threshold  $\alpha$ . In other words,

$$E = 1 \text{ iff } \frac{1}{n} \sum_{i \in \mathcal{S}} V_i > \alpha \quad (1)$$

where  $n = |\mathcal{S}|$ , the number of voters in  $\mathcal{S}$ . For voters  $i \in \mathcal{S}$  let  $V_i = 1$  if voter  $i$  votes for the candidate in question, and 0 otherwise. Let  $a_x = 1$  denote the truth of a fact  $x$  that sways voters to vote for the candidate (for instance, she is better than her rival at increasing RDI). For clarity, we will drop the  $x$  subscript, so  $a = 1$  denotes  $a_x = 1$ . Let  $A_i = 1$  if voter  $i$  believes  $a$ .

The consider the following assumptions:

For all  $i \in \mathcal{S}$ :

1.  $Pr(V_i = 1|A_i = 1) \geq Pr(V_i = 1|A_i = 0)$ , where a strict inequality holds for some subset of the population; stated otherwise,  $A$  affects all voters' choices in the same direction.
2.  $Pr(A_i = 1|a = 1) \geq Pr(A_i = 1|a = 0)$ , where a strict inequality holds for some subset of the population; stated otherwise, that voters are better than random at judging  $a$ .
3. For all  $j \neq i$ ,  $V_i \perp V_j|a$
4.  $Pr(V_i = 1|A_i, a) = Pr(V_i = 1|A_i)$

It may be the case that some voters' attitudes on issue  $x$  depend on a number of other facts as well, in such a way that assumption 2 only holds for some voters in  $\mathcal{S}$  under certain additional conditions. For instance, a candidate's superiority on RDI might only appeal to environmentalist voters if that superiority is not due to environmentally harmful policies. To allow for such situations, define a set of "conditioning facts"  $\tilde{a}$  consisting of those facts which may qualitatively moderate the relationship between  $a$  and  $V_i$ , for some  $i$ . The composition of  $\tilde{a}$  depends on the issue in question, and may in some circumstances be empty. In any case, we may restate assumptions 1–4 as conditional on  $\tilde{a}$ .

**Proposition.** *If assumptions 1,2, and 4 hold conditional on  $\tilde{a}$ , then*

$$Pr(E = 1|a = 1, \tilde{a}) > Pr(E = 1|a = 0, \tilde{a}) \quad (2)$$

*Proof.* First note that under assumptions 1-3, we have (letting all probabilities be conditional on  $\tilde{a}$ )

$$\begin{aligned} & Pr(V_i = 1|a = 1) \\ &= \mathbf{E} Pr(V_i = 1|a = 1, A_i) \\ &= Pr(V_i = 1|A_i = 1)Pr(A_i = 1|a = 1) + Pr(V_i = 1|A_i = 0)Pr(A_i = 0|a = 1) \\ &= Pr(A_i = 1|a = 1)\{Pr(V_i = 1|A_i = 1) - Pr(V_i = 1|A_i = 0)\} + Pr(V_i = 1|A_i = 0) \\ &\geq Pr(A_i = 1|a = 0)\{Pr(V_i = 1|A_i = 1) - Pr(V_i = 1|A_i = 0)\} + Pr(V_i = 1|A_i = 0) \\ &= Pr(V_i = 1|A_i = 1)Pr(A_i = 1|a = 0) + Pr(V_i = 1|A_i = 0)Pr(A_i = 0|a = 0) \\ &= Pr(V_i = 1|a = 0) \end{aligned}$$

Where the inequality is a result of assumptions 1 and 2, and is a strict  $>$  for some members of the population. That is,  $V_i|a = 1$  is stochastically greater than  $V_i|a = 0$ . That being the case,  $\sum_{i=1}^n V_i|a = 1$  is stochastically greater than  $\sum_{i=1}^n V_i|a = 0$  (Shaked and Shanthikumar (2007)). Therefore,  $Pr(E = 1|a = 1) = Pr(\sum_{i=1}^n V_i > n\alpha|a = 1) > Pr(\sum_{i=1}^n V_i > n\alpha|a = 0) = Pr(E = 1|a = 0)$   $\square$

**Corollary 1.** *Let  $\alpha \subset \mathcal{S}$ . If the conditions obtain for all  $i \in \alpha$  then the result is the proposition still holds for members of  $\alpha$ .*

## 7.2 A Comparison with CJT

The proposition described informally in Section 2 and formally in this appendix builds on prior theoretical work in epistemic democracy, perhaps most famously the Condorcet Jury Theorem (CJT). In this section we will briefly compare and contrast our proposition with the CJT. We will argue that our proposition may be thought of as a generalization of the CJT to the multidimensional case.

The CJT may be stated as follows (e.g. Boland, 1989):

**Proposition (CJT).** *A decision-making body is comprised of  $n \geq 3$  voters who cast votes  $V_1, \dots, V_n \in \{0, 1\}$ . The group's decision  $E = 1$  if and only if  $\sum_i V_i > n/2$  and  $E = 0$  otherwise. If:*

1. *There is a unique correct decision  $a \in \{0, 1\}$*
2.  *$p_i = Pr(V_i = a|a) = p > 1/2$  for all  $i$  (Competence) and*



3.  $V_i \perp\!\!\!\perp V_j | a, i \neq j$  (*Independence*)

Then  $Pr(E = a|a) > p$  and  $Pr(E = a|a) \rightarrow 1$  as  $n \rightarrow \infty$

Our proposition extends the CJT to political elections. The existence of a “correct” decision between candidates is arguable, and typically hinges on voters’ personal values, as well as matters of fact. Therefore, assumption 1 in our theorem is not that there is a correct overall decision, but that, on one particular issue values are not in dispute among members of the electorate—for this particular issue, the “correct” decision  $a$  is the one that conforms with the universal value. The formal statement  $Pr(V_i = 1|A_i = 1) > Pr(V_i = 1|A_i = 0)$  means that voter  $i$ ’s perception of candidate 1’s superiority on the particular issue in question increases the chances he’ll vote for her; to state it for all voters  $i$  is to say that they all agree on the issue’s choice-worthiness.

Focusing on one dimension among possibly many that influence votes  $V$  required two additional modifications to the CJT. First, we introduced the “individual belief” variable  $A_i$  encoding  $i$ ’s belief about which candidate is superior on the issue in question. In the classic CJT setup, voters are assumed to vote for the option they believe is best (See Austen-Smith and Banks, 1996, however), so  $V_i \equiv A_i$ , necessarily. However, in a political election  $i$  may believe candidate 1 to be superior to 0 on the issue in question, so  $A_i = 1$ , but believe candidate 0 to be superior on other issues, so that  $V_i = 0$ .

Secondly, we have weakened the relationship between truth  $a$  and, respectively, beliefs  $A$ , votes  $V$ , and the outcome  $E$ . In place of the CJT assumption 2 that voters are more likely to pick the correct than the incorrect option, we assume that a fact’s truth makes voters more likely to believe it than were it false. So, for instance, say  $a = 1$ , so candidate 1 is superior on the issue in question. However, perhaps because of voter  $i$ ’s prejudice or confusion, he is unlikely to believe in 1’s superiority under any circumstances. That said, he is slightly more likely to believe 1 is superior if that is the truth than if 0 is in fact better. For instance,  $Pr(A_i = 1|a = 1) = 0.1$  and  $Pr(A_i = 1|a = 0) = 0.01$ . Voter  $i$  would violate assumption 2 of the CJT, since  $p < 1/2$ , but does not violate our assumption 2. Conversely,  $p > 1/2$  as in the CJT implies that  $Pr(A_i = 1|a = 1) > 1/2 > Pr(A_i = 1|a = 0)$ , so our assumption 2 is strictly weaker.

That said, Dietrich (2008) showed that for CJT to hold, we must merely have  $\bar{p} = \sum_i p_i/n > 1/2$ .<sup>11</sup> This assumption is neither strictly weaker nor stronger than ours. On the one hand, if, say, as above  $Pr(A_i = 1|a = 1) = 0.1$  and

<sup>11</sup>Boland (1989) shows that for  $Pr(E = a|a) > \bar{p}$  to hold in finite samples, we must assume  $\bar{p} > 1/2 + 1/2n$ .

$Pr(A_i = 1|a = 0) = 0.01$ , but this time for all voters  $i$ , our assumption would be satisfied, but Dietrich (2008) would not. On the other hand, it is possible for some subset of the electorate to violate our assumption but be outweighed by the remainder of the electorate, so  $\bar{p} > 1/2$ . In any event, simply substituting  $\bar{p} > 1/2$  into our proposition would not suffice—if the group of voters who are more likely to believe in a candidate’s superiority when the candidate is in fact inferior than when she is actually superior are also more fervent believers in the importance of the issue, they can undermine the result.

Our result of our proposition is weaker as well—indeed, it ought to be. If broad agreement holds for only one of many issues, it would be troubling to state that  $Pr(E = 1|a = 1) \rightarrow 1$  as the electorate grows, implying that all other issues—legitimate, if ambiguous—become irrelevant. Instead, as in the case with belief, we merely show that a candidate’s superiority on the issue increases her probability of being elected, relative to what it would have been had she instead been inferior. That is, we claim that superiority on an agreed-upon issue boosts a candidate’s chances, but provides no guarantees.

Additionally, unlike the CJT, our result holds for finite samples.

Finally, our independence assumption is the same as CJT’s. This suggests that relaxations of independence in the CJT case (Boland et al., 1989, e.g.) may work in our case as well.

### 7.3 Modeling Choices: Logit Model

Our models are based roughly on those presented in Nadeau and Lewis-Beck (2001). In particular, they estimated the following model:

$$\begin{aligned} vote_{ind,elec} = & \alpha + \beta_1 RDI_{elec} + \beta_2 Finances_{ind,elec} + \beta_3 IncumbentParty_{elec} \\ & + \beta_4 PartyID_{ind,elec} + \beta_5 Race_{ind,elec} + \epsilon_{ind,elec} \end{aligned}$$

where  $vote_{ind,elec}$  is an individual’s vote, coded as 1 if she voted for the incumbent party, and 0 otherwise.  $RDI_{elec}$  is the percent change in RDI per capita from the previous year,  $Finances_{ind,elec}$  is an individual’s assessment of his own personal finances, compared to the previous year (1 denotes “better,” 0 “the same,” −1 “worse”).  $IncumbentParty_{elec}$  codes whether the Democratic or Republican party was the incumbent in the election.  $PartyID_{ind,elec}$  is a five-point scale that codes to what extent the voter’s party identification agrees with the incumbent party (3: strongly agrees, 2: weakly agrees, 0: indifferent, −2 weakly disagrees, −3: strongly disagrees) and  $Race_{ind,elec}$  is an indicator for race, also aligned with incumbent party (if the incumbent party is Republican, 1 indicates ‘white’

and 0 'nonwhite,' with the opposite if the incumbent party is Democratic). The parameters  $\alpha$  and  $\beta_k$ ,  $k = 1, \dots, 5$  are estimated with ordinary least squares, and  $\epsilon_{ind,elec}$  is a regression error.<sup>12</sup>

We model each individual's vote as a separate coin-toss. These coins, though, are not "fair," in the usual sense: we assume that each voter has a different probability of voting for the incumbent party. The probability each voter chooses the incumbent party is modeled as a function of several factors, including  $\Delta RDI$  and, in some versions of the model, their subjective beliefs regarding the state of the economy. In particular, the logit of the probability is a linear function of these factors.<sup>13</sup>

Our model has two layers, the first one models an individual's probability of voting for the incumbent party:

$$\text{logit}(Pr(\text{vote}_{ind,elec} = 1)) = \alpha + \nu_{elec} + \eta_{state} + X_{ind,elec}\beta \quad (3)$$

where  $X_{ind,elec}$  is a vector of all of the individual level regressors for voter  $ind$  in election  $elec$  and  $\text{logit}(Pr(\text{vote}_{ind,elec} = 1))$  is the logit of the probability that voter  $ind$  votes for the incumbent party. In some models  $X_{ind,elec}$  includes sociotropic measures: "retrospective" codes whether voters believe the economy has improved over the past year, and "prospective" codes voters' beliefs regarding whether the economy will improve in the coming year. These metrics are important because they give us a handle on the effect beliefs have on votes. The model also allows the states voters reside in,  $\eta_{state}$ , to idiosyncratically influence their voting decisions, with the random intercept  $\eta_{state} \stackrel{iid}{\sim} N(0, \sigma_s)$ .

In the second level of the model,  $\nu_{elec}$  is a random effect for the election,

$$\nu_{elec} = \gamma_1 \Delta RDI_{elec} + \gamma_2 \text{incumbentParty}_{elec} + \zeta_{elec} \quad (4)$$

where  $\zeta_{elec} \stackrel{iid}{\sim} N(0, \sigma_e)$ . Here voters' choices are affected, on an election-to-election basis by  $\Delta RDI$  and the incumbent party, in addition to idiosyncratic factors. The variance parameters  $\sigma_e$  and  $\sigma_s$  are estimated from the data, along with the coefficients  $\alpha$ ,  $\beta$ ,  $\gamma_1$ , and  $\gamma_2$ . Since voters in each election are not independent of each other, we include random effects for each election. Including a random election effect  $\nu_{elec}$  models that dependence and corrects the overly-optimistic standard errors in Nadeau and Lewis-Beck (2001).

Since each election has its own idiosyncrasies—for instance, the Vietnam war

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<sup>12</sup>The individual-level data in the model came from the American National Election Survey (ANES) time series Studies (2010b) and the RDI data is from the National Bureau of Economic Research.

<sup>13</sup>If the probability of voting for the incumbent party is  $p$ , the logit is  $\log(p/(1-p))$ .

and Lyndon Johnson’s surprising decision not to run in 1968, or the fallout from the Lewinsky scandal in the 2000 election—the error term from each election must be independently modeled. We account for this by explicitly modeling the multilevel structure in the data with a mixed-model, measuring some variables at the voter level and others at the election level (e.g. Raudenbush and Bryk, 2002). We also included state random effects, to account for state-level idiosyncrasies that are stable over time.<sup>14</sup>

Another issue with the Nadeau and Lewis-Beck model is its use of OLS to model binary outcomes. Since binary outcomes are restricted to be either one or zero, the regression errors  $\epsilon$  are necessarily heteroskedastic, which can also bias standard error estimates. Some linear probability modeling techniques can overcome this difficulty (see, e.g. Agresti, 2002, sec. 4.6), but another issue remains: the fitted values of a linear probability model might fall outside of the range 0–1, which hinders their interpretation. Indeed, when replicating NLB’s specification, we found that greater than 10% of the fitted values fell outside this range. We avoided these problems by modeling the data to fit a logistic regression.

We further expanded on NLB’s model in two different ways. First, the model in NLB was fit using years 1956–1996, which we extended through 2012, using, in part, the ANES 2012 time-series file Studies (2010a). We chose to use individual-level data post-1976 because our scientific question fundamentally regards individual voting decisions as a function of their beliefs. As such we gain substantive purchase on our research question by omitting these aggregate level data. By including four more presidential elections than Nadeau and Lewis-Beck (2001), those from 2000–2012, we could also mitigate the consequences of omitting elections from 1956–1976. Second, while their model used an aggregate sociotropic metric for economic voting, we included individual-levels where possible, which conveniently allows us to skirt any worries associated with the ecological fallacy. These sociotropic measures take two forms: The “retrospective” measure is a voter’s answer to the question: “Would you say that over the past year the nation’s economy has gotten better, stayed (all yrs. exc 1984: about) the same or gotten worse?” while the “prospective” measure is subjects’ answer to the question: “Do you expect the economy to get better, get worse, or stay about the same?” (1: better, –1: worse, 0: same). Model (1) draws from data on presidential voting from 1956–2012. Models (2)–(3), which include belief variables, are fit to data from presidential elections from 1980–2012, since

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<sup>14</sup>An alternative approach is to use cluster-robust standard errors (see Freedman, 2006). Doing so, in our model, gave broadly similar results.

the ANES only began measuring individual’s beliefs of sociotropic economic performance following the 1976 election.<sup>15</sup>

For the sake of simplicity, the results in Table 1 came from a “complete case analysis” of the data, which dropped any case with any missing regressors. However, in a supplemental analysis we accounted for item-level non-response with multiple imputation (Rubin, 2004). With the R package *mice* (van Buuren and Groothuis-Oudshoorn, 2011), we created five datasets, with missing values for the regressors randomly imputed. Next, we used the package *lme4* (Bates et al., 2014) to fit mixed-effects models in each imputed dataset, and combined the results to yield estimates very similar to those in Table 1. This relaxes the strong assumption that item-level nonresponse is entirely random, replacing it with an assumption that nonresponse is random after accounting for the observed regressors.

Aside from computational tractability, one of the advantages of the multi-level logistic regression model is that it allows us to independently assess the respective relationships between votes and each variable on the right-hand side. In particular, each coefficient can be interpreted as the relationship between its corresponding variable and votes, after modeling the effects of all of the other variables in the model. For instance, let  $\hat{\gamma}_1$  be the estimated coefficient on  $\Delta RDI$  in Model 1 found in Table 1. One way of interpreting this result would be that, everything else (in the model) being equal, when  $\Delta RDI$  increases by 1 percentage point, one may expect the logit of the probability of a voter choosing the incumbent party to increase by about  $\hat{\gamma}_1$ .

Regarding epistemic heterogeneity, since each wave of the ANES survey selects a different set of voters, the ANES data don’t allow us to identify the individuals who systematically misjudge the economy, and thereby directly rule out such a possibility. However, the ANES data do allow us to identify the correlations between  $\Delta RDI$  and voters’ perceptions of the economy within demographic sub-groups. A demographic whose members’ retrospective judgement of the economy correlate negatively with  $\Delta RDI$  may indicate a problem. To attempt to identify such a sub-group, we fit a multilevel linear model to the data, predicting *retrospective* as a function  $\Delta RDI$ . The model included interactions between  $\Delta RDI$  and a thick set of probative variables including personal finances, incumbent party, race, and party ID (variables incorporated in models

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<sup>15</sup>The lack of data regarding individual beliefs about economic performance led Nadeau and Lewis-Beck (2001) to fit their models to an aggregate sociotropic measure, which they termed the “National Business Index,” computed using data from a separate survey (page 162). We included the retrospective belief measure in alone individually and coupled with a prospective measure too.

(1)–(3)), in addition to social class, education, gender, Hispanic identification, marital status, urbanism, and state. This test seeks to find exceptional cases, subsets of the population that systematically misjudge the health of the economy. Though it is large, we needn’t worry that the model is overfit, since the setup here is meant to serve as a conservative test, identifying subgroups that could challenge our conjecture.

To test the possible heterogeneity of values held among the electorate we estimate the conceivable size of such a minority that might unwittingly vote for the right candidate given mistaken beliefs by asking whether there exists some subset of the sample that prefers a shrinking economy. To do this we fit our multilevel model to the data, this time including interactions between  $\Delta RDI$  and a thick set of probative variables including personal finances, incumbent party, race, and party ID (variables incorporated in models (1)–(3)), in addition to social class, education, gender, Hispanic identification, marital status, urbanism, and state. This test seeks to find exceptional cases, subsets of the population that prefer the economy to shrink.<sup>16</sup> We needn’t worry that we have overfit our model by including so many interaction terms, since the setup here is meant to serve as a conservative test, identifying subgroups that could challenge our conjecture.

Only two groups could be identified at all for which  $\Delta RDI$  was negatively correlated with election outcome (1) Respondents who reported that they do not know if they are Hispanic and (2) respondents who identified as “upper class,” and both of these are likely due to noise. Together, these groups comprised a quarter of a percent of the dataset, neither statistically significance. Even if members of these two groups would prefer that the economy shrink, the minuscule effect size is exceedingly unlikely to practically effect any election outcome.<sup>17</sup>

A final caveat is that, though ANES is a nationally representative survey, our coefficient estimates are design consistent for the subjects in the study who voted or reported their votes. Due to the difficulty of fitting a multilevel model with survey weights (Gelman, 2007), we fit an unweighted model, which limits the external validity of its estimates. That being said, our attempts to fit weighted models in Stata (StataCorp, 2013) with the `gllamm` function (Rabe-Hesketh et al., 2004) yielded broadly similar results.

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<sup>16</sup>This model, in effect, searches for the existence subgroups of the population (defined by these variables) for which the linear relationship between  $\Delta RDI$  and the log-odds of voting for the incumbent party in a presidential election varies considerably from the estimate given in Table 1.

<sup>17</sup>A full report of the model results is available as an online supplement.

We examined several alternative model specifications, as a robustness check. The results of our multiple-imputation model, our mediation model, robustness checks and residual plots, along with the code to produce them, are available at <http://tinyurl.com/EpDemSupplement-pdf>.

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