Streaming Algorithms

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A Reference Book

• "The Probabilistic Method", Alon and Spencer (2004)

You may find an e-copy of this book on www.lib.nctu.edu.tw

Reschedule

I have to travel abroad for some business in the week of midterm exams, Apr 15 - 20. Here are 2 possible plans. [Voted]

(A) Let the second written assignment be a take-home midterm exam. The number of problem sets will be a little more than an ordinary written assignement, and

the weight of midterm exam $10\% \rightarrow 20\%$ the weight of final project $40\% \rightarrow 30\%$

(B) Find some free time in common to have a 3-hour make-up class.

More on Expectation

Inequalities on Expectation

Let X be a random variable. Then we have:

- (1) $Pr[X \ge E[X]] > 0$
- (2) $Pr[X \le E[X]] > 0$

What's more? We may impose additional restrictions on X.

Inequalities on Expectation

If $X \ge 0$ and E[X] > 0, then we have:

 $Pr[X \ge \lambda E[X]] \le 1/\lambda$ for any $\lambda \ge 1$, (Markov inequality)

Inequalities on Expectation

If $X \ge 0$ and E[X] > 0, then we have:

 $\Pr[X \ge \lambda E[X]] \le 1/\lambda$ for any $\lambda \ge 1$, (Markov inequality)

together with $X \le cE[X]$ for some $c \ge 1$, we get:

 $E[X] \le cE[X]Pr[X > (1-\varepsilon)E[X]] + (1-\varepsilon)E[X]Pr[X \le (1-\varepsilon)E[X]]$

 \Rightarrow Pr[X > (1- ϵ)E[X]] $\geq \epsilon/(c-1+\epsilon)$ for any ϵ in (0, 1)

Sum-free Subsets

<u>Theorem 1</u>. [Erdős 1965] Every set $B = \{b_1, b_2, ..., b_n\}$ of n non-zero integers has a sum-free subset A of size |A| > n/3.

<u>Definition</u>. A set S is sum-free if \forall a, b, c \in S (possibly repeat), $a+b \neq c$.

Sum-free Subsets

Proof Strategy.

If S is sum-free in \mathbb{Z}_p , then S is sum-free in \mathbb{Z} . It suffices to show that B has a sum-free subset A of size |A| > n/3 in \mathbb{Z}_p .

Let p be a prime in the form of 3k+2 so that p > 2 max $|b_i|$. (By Dirichlet Theorem, such p exists.)

Then $C = \{k+1, ..., 2k+1\}$ is sum-free in \mathbb{Z}_p $\Rightarrow xC = \{xc \pmod{p}: c \in C\}$ is sum-free in \mathbb{Z}_p for any non-zero x's.

Find such an x (exists?) that $|B \cap xC| > n/3$. Then output $B \cap xC$ as the sum-free subset.

Sum-free Subsets

X	xc ₁	xc ₂	•••
1	$\pi_1(1)$	$\pi_2(1)$	
2	$\pi_1(2)$	$\pi_2(2)$	
•••		•••	
p-1	$\pi_1(p-1)$	$\pi_2(p-1)$	

For an x picked uniformly at random from $\{1, ..., p-1\}$, $Pr[b \in xC] = |C|/(p-1)$.

Sum-free Subsets

X	xc ₁	xc_2	•••
1	$\pi_1(1)$	$\pi_2(1)$	
2	$\pi_1(2)$	$\pi_2(2)$	
•••		•••	
p-1	$\pi_1(p-1)$	$\pi_2(p-1)$	

Every b in {1, ..., p-1} appears exactly once in each column, and appears at most once in each row.

Sum-free Subsets

X	xc ₁	xc ₂	
1	$\pi_1(1)$	$\pi_2(1)$	
2	$\pi_1(2)$	$\pi_2(2)$	
•••	•••	•••	
p-1	$\pi_1(p-1)$	$\pi_2(p-1)$	

Let $X = \sum_b X_b$. where X_b be the indicator random variable denoting whether $b \in xC$ for a random x.

Sum-free Subsets

x	xc ₁	xc_2	•••
1	$\pi_1(1)$	$\pi_2(1)$	
2	$\pi_1(2)$	$\pi_2(2)$	
•••	•••	•••	
p-1	$\pi_1(p-1)$	$\pi_2(p-1)$	

$$E[X] = \sum E[X_b] = n|C|/(p-1) = n(k+1)/(3k+1) > n/3.$$

Alteration

Construct a Large Sum-free Subset

Observe that $X \ge 0$ and E[X] > 0.

Because $X \le n$ and E[X] > n/3, $X \le 3E[X]$.

By the second form of Markov inequality, we have:

 $Pr[X > (1-\varepsilon)E[X]] \ge \varepsilon/(2+\varepsilon)$ for any ε in (0, 1).

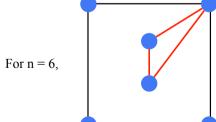
w.h.p. means with probability $1-1/n^{\Omega(1)}$.

By trying O(log n) random x's, we can have a sum-free subset of size $(1-\epsilon)n/3$ for any constant ϵ in (0, 1) w.h.p.

Maximize the Minimum-area Triangle

Let S be a set of n points in a unit square, and let T(S) be the minimum area of traingles whose vertices are three distinct points in S. Let T(n) be the maximum possible T(S) for all S.

Theorem 2. (Thm 3.3.1) There is a set S so that $T(S) \ge 1/(100n^2)$ i.e. $T(n) = \Omega(1/n^2)$.



Maximize the Minimum-area Triangle

<u>Proof Strategy</u>. Randomly sample n+C points, if there are at most C triangles of area $< 1/(100n^2)$, then remove a vertex for each triangle. In this way, all triangles induced by some n points have area $\ge 1/(100n^2)$. This technique is called alteration.

Maximize the Minimum-area Triangle

Let P, Q, R be points sampled independently and uniformly from the unit square. Let μ be the area of triangle P, Q, R.

Claim. $Pr[\mu \le \varepsilon] \le 16\pi\varepsilon$.

Sample 2n points indepently and uniformly from the unit square, then let X be the number of triangles of area $< 1/(100n^2)$.

$$E[X] \le {2n \choose 3} 16\pi/(100n^2) \le n$$

There exists an arrangement of n points that induce no triangle of area $< 1/(100n^2)$.

Maximize the Minimum-area Triangle

Let P, Q, R be points sampled independently and uniformly from the unit square. Let μ be the area of triangle P, Q, R.

Claim. $Pr[\mu \le \epsilon] \le 16\pi\epsilon$.

Sample 2n points indepently and uniformly from the unit square, then let X be the number of triangles of area $< 1/(100n^2)$.

$$E[X] \le \binom{2n}{3} 16\pi/(100n^2) \le n$$

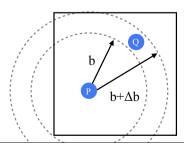
There exists an arrangement of 2n points that induce at most n triangles of area $< 1/(100n^2)$.

Proof of $Pr[\mu \le \epsilon] \le 16\pi\epsilon$

Let ℓ be the distance between P and Q. Then we have

$$\Pr[b \le \ell \le b + \Delta b] = \pi(b + \Delta b)^2 - \pi b^2, \text{ and in the limit}$$

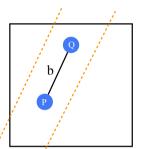
$$\Pr[b \le \ell \le b + db] = 2\pi b \text{ db}.$$



Proof of $Pr[\mu \le \epsilon] \le 16\pi\epsilon$

Given PQ has length b, to make PQR has area $\leq \varepsilon$, then R must fall within the following strip of width = $2*2\varepsilon/b$ and length $\leq \sqrt{2}$.

Thus,
$$\Pr[\mathbf{\mu} \leq \varepsilon] \leq \int_0^{\sqrt{2}} 2\pi b \frac{4\sqrt{2}\varepsilon}{b} db = 16\pi\varepsilon$$



The Second Moment

Exercise 1

What is the running time of finding the smallest-area triangle of n given points?

Suppose that given n points, testing whether there are 3 points colinear requires $\Omega(n^2)$ time. What can we infer?

Chebyshev Inequality

Recall the definition of the varience,

$$Var(X) = E[(X-\mu)^2] = \sigma^2.$$

By Markov inequality, we get

$$\Pr[(X-\mu)^2 \ge \lambda^2 \sigma^2] \le 1/\lambda^2$$
,

or equivalently

 $\Pr[|X-\mu| \ge \lambda \sigma] \le 1/\lambda^2$. (Chebyshev Inequality)

What's more? We may impose additional restrictions on X.

Chebyshev Inequality

If X is a nonnegative integral-valued random variable and E[X]>0,

$$Pr[X = 0] \le Pr[|X - \mu| \ge \mu] = Pr[|X - \mu| \ge (\mu/\sigma)\sigma] \le Var(X)/E[X]^2$$
.

Can we do something similar for X=k? What is the assumption?

We may like to bound Pr[X = 0] small because we may set X=0 to be an undesirable case.

Distinct Sums

We say a set $x_1, x_2, ..., x_k$ of positive integers has distinct sums if all subset sums

$$\sum_{i \in S} x_i$$
 for all $S \subseteq \{1, 2, ..., k\}$

are distinct. Let f(n) denote the maximum k for the case $x_1, x_2, ..., x_k \le n$.

 $1, 2, ..., 2^{\lfloor \log_2 n \rfloor}$ yields that $f(n) \ge \lceil \log_2 n \rceil + 1$.

Exercise 2

Prove that for nonnegative integral-valued random variable X,

$$Pr[X=0] \leq Var(X)/E[X^2].$$

Distinct Sums

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are distinct. Let f(n) denote the maximum k for the case $x_1, x_2, ..., x_k \le n$.

Since each subset sum is unique and less than nk, $2^{f(n)} < nf(n) \Rightarrow f(n) < \log_2 n + \log_2 \log_2 n + O(1)$

A Tighter Upper Bound

Let $\varepsilon_1, \varepsilon_2, ..., \varepsilon_k$ be independent random variables so that

$$Pr[\varepsilon_i = 1] = 1/2$$
 and $Pr[\varepsilon_i = 0] = 1/2$ for each i in [1, k].

Let $X = \sum_i \varepsilon_i x_i$ i.e. a random subset sum. Thus, $E[X] = (\sum_i x_i)/2$ and $Var(X) = (\sum_i (x_i)^2)/4$ (Why?).

Let $X = \sum X_i$. If all X_i 's are independent, then $Var(X) = \sum Var(X_i)$.

A Tighter Upper Bound

By Chebyshev inequality,

$$\Pr[|X-\boldsymbol{\mu}| \ge \lambda n k^{1/2}/2] \le 1/\lambda^2.$$

or equivalently

$$\Pr[|X - \mu| < \lambda n k^{1/2}/2] \ge 1 - 1/\lambda^2.$$
 (a)

Because each distinct sum occurs with probability either 0 or 2^{-k} ,

$$Pr[|X-\mu| < \lambda nk^{1/2}/2] < (\lambda nk^{1/2}+1)2^{-k}$$
. (b)

Combine (a), (b) and set $\lambda = \sqrt{3}$ yields that

$$f(n) \le \log_2 n + \frac{1}{2} \log_2 \log_2 n + O(1)$$

A Tighter Upper Bound

Let ε_1 , ε_2 , ..., ε_k be independent random variables so that

$$Pr[\varepsilon_i = 1] = 1/2$$
 and $Pr[\varepsilon_i = 0] = 1/2$ for each i in [1, k].

Let $X = \sum_i \varepsilon_i x_i$ i.e. a random subset sum. Thus, $E[X] = (\sum_i x_i)/2$ and $Var(X) = (\sum_i (x_i)^2)/4 \le kn^2/4$.

By Chebyshev inequality,

$$\Pr[|X - \mu| \ge \lambda n k^{1/2/2}] \le 1/\lambda^2.$$

or equivalently

$$\Pr[|X-\mu| < \lambda n k^{1/2}/2] \ge 1-1/\lambda^2.$$
 (a)