

Time series applied to finance

Marouane Arab

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Abstract

This document is generated by an R script and a brew file (latex file which embeds R code). All sources are available at https://github.com/arabm/multifractal_model. There are questions about time series:

- (i) Is the data relevant ? coherent ?
- (ii) Which model is adapted ? How validate or reject it ?
- (iii) What is the error committed by using it ?
- (iv) How forecast is usefull ?
- (v) How much day forecast is relevant ?
- (vi) How to validate backtesting ?

I'm counting on your feedback. Do concepts are well introduced, developed and explained ? Do you want more details ? On what ? Are they missing concepts ? Is a concept wrongly explained ? And of course, are there english mistakes ?

Correct english mistakes !

Todo list

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1 Quick start

1.1 Usage

Source The repository github contains two files :

- *script.r* : an R script (official R website is <https://cran.r-project.org/>)
- *template.brew* : a latex template which embeds R commands. Brew is an R package (see <https://cran.r-project.org/web/packages/brew/index.html>).

Pre-required Install R, latex

Complete
pre-required

First command Once R, latex and every pre-required are done, you can launch the script with *from the clone directory*:

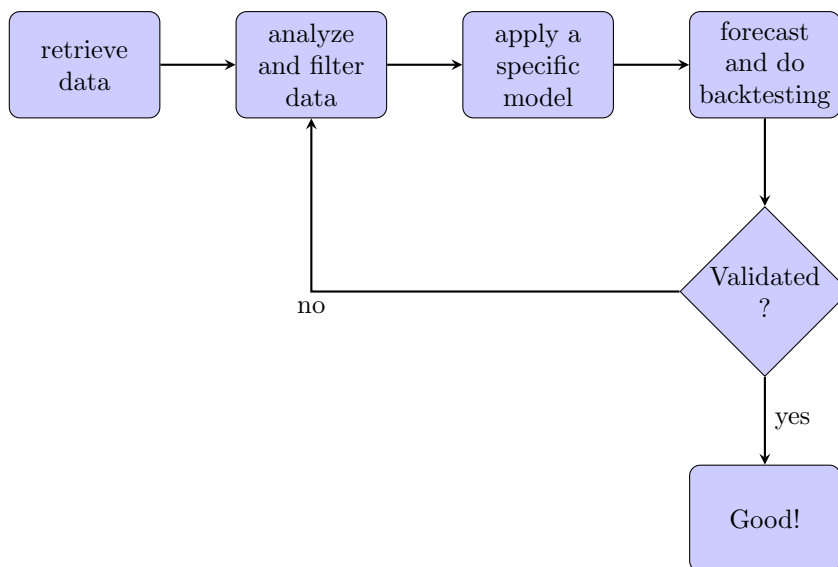
```
R -f script.r
```

It generates *generated_report.pdf* (similar to the one you're reading right now). It also creates a lot of repositories :

```
clone repo
├── output
│   ├── report
│   ├── example
│   └── FCHI
```

Currently there isn't an *example* repository, but rather various example repositories : *example_uniform*, *example_normal*, etc. Same apply for *FCHI*, it creates a repository by financial underlying.

1.2 Modelisation process



Workflow The workflow shows linear tasks, while indeed they're strongly linked. Analyzing, filtering and applying a model can be done at the same time. We look at the data. Then apply a model. Readjust the model or the filter applied on data.

Validation The validation is done by analysing residuals. Residuals are errors or noises if we consider that the data follows our model. Backtesting is also important. It's how the model from an extract of the history forecast correctly what happened.

Parameters Model can have various parameters. The model which fits the best the data is the date itself. Of course, that's not the aim of modelisation. We rather prefer to minimize the number of parameters. A good model, is one which has a good ratio between number of paramaters (lower the better) and residuals (lower the better too).

Complete
quick start

2 Statistical tests

2.1 White noises

Definition Be ϵ_t a white noise. It's independent through time t and got the same law.

white noise
definition

Example with $\mathcal{N}(0,1)$ law We consider a white noise which follow $\mathcal{N}(0,1)$, and generate it for 500 dates.

```
>white_noise<-rnorm(500,0,1)
>plot(white_noise,type="l")
>plot(density(white_noise))
>plot(qqnorm(white_noise))
>qqline(white_noise)
```

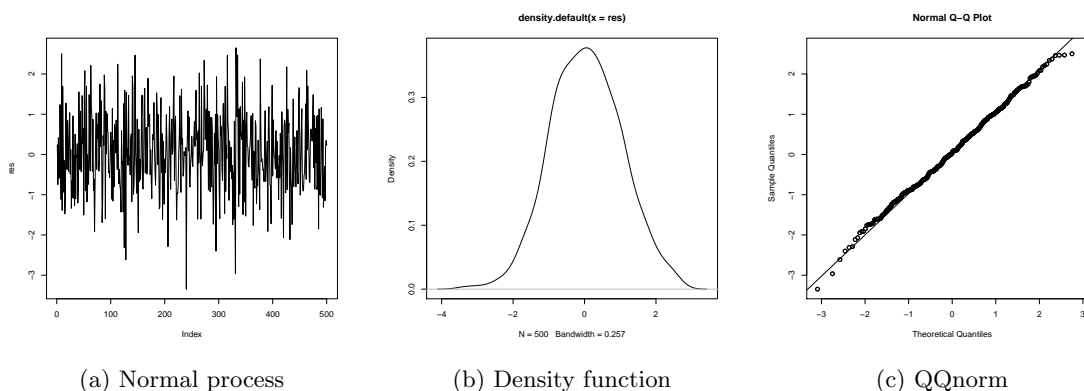


Figure 1: Example of $\mathcal{N}(0,1)$

Test of Box-Ljung for independence :

```
>Box.test(white_noise,type='Ljung')
```

Test of Shapiro for Normal law :

```
>shapiro.test(white_noise)
```

Shapiro-Wilk normality test

```
data:  res
W = 0.99746, p-value = 0.6467
```

link with
Box-Ljung,
shapiro,qqnorm

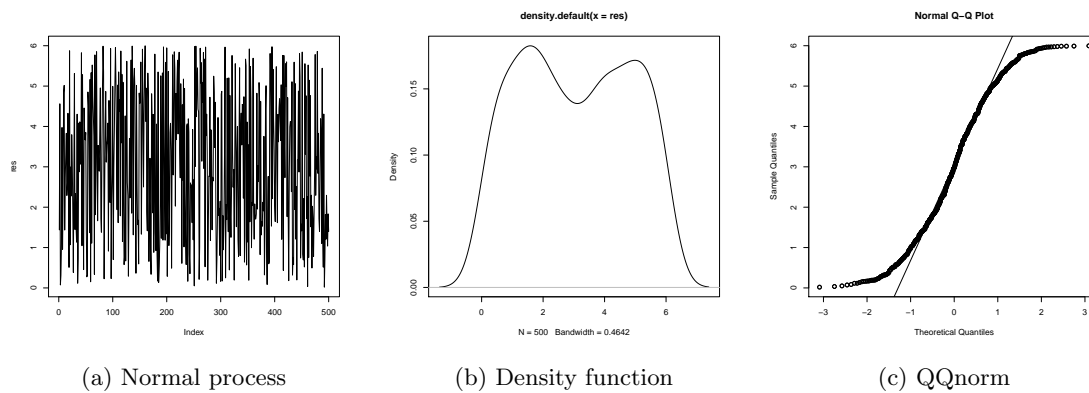


Figure 2: Example of $\mathcal{N}(0,6)$

Test of Box-Ljung for independance : Test of Shapiro for Normal law :

Shapiro-Wilk normality test

data: res
W = 0.94555, p-value = 1.376e-12

2.2 BIC

2.3 AIC

2.4 Likelihood

Add
example
for normal,
AR/-
MA/ARCH/-
Garch,
normal
square

do BIC

do AIC

do Likeli-
hood

3 Model

introduction
Model

3.1 Auto Regressive Moving Average (ARMA)

3.1.1 AR(1)

More ver-
bose AR(1)

Definition Be $c, \beta \in \mathbb{R}^2$, X_t is an AR(1) process if

$$\begin{aligned} X_t &= c + \beta X_{t-1} + \epsilon_t \\ &= c + \sum_{i=1}^{\infty} \beta^i \epsilon_{t-i} + \epsilon_t \end{aligned}$$

Finite variance X_t has a finite variance if $|\beta| < 1$:

$$\begin{aligned} \text{Var}(X_t) &= \text{Var}\left(\sum_{i=0}^{\infty} \beta^i \epsilon_{t-i}\right) \\ &= \sum_{i=0}^{\infty} \beta^{2i} \sigma^2 \\ &= \frac{1 - \beta^\infty}{1 - \beta} \sigma^2 \end{aligned}$$

When $|\beta| > 1$, $\text{Var}(X_t) = \infty$.

When $|\beta| < 1$, $\text{Var}(X_t) = \frac{\sigma^2}{1-\beta}$

Expectation doesn't depend of t It implies $\mathbb{E}\{X_t\} = \mathbb{E}\{X_{t-1}\} = \mu$

$$\begin{aligned} \mathbb{E}\{X_t\} &= c + \beta \mathbb{E}\{X_{t-1}\} \\ \Leftrightarrow \mu &= \frac{c}{1 - \beta} \end{aligned}$$

We could rewrite $\mathbb{E}\{X_t\}$ as :

$$\begin{aligned} \mathbb{E}\{X_t\} &= c + \mathbb{E}\{\beta X_{t-1} + \epsilon_t\} \\ &= c + \mathbb{E}\left\{\sum_{i=1}^{\infty} \beta^i (c + \epsilon_{t-i})\right\} \\ &= \frac{c}{1 - \beta} \end{aligned}$$

Covariance (ACF) There is a linear relation between lag correlation γ_l .

$$\begin{aligned}
 \gamma_l &= \text{cov}(X_t, X_{t-l}) \\
 &= \mathbb{E}\{(X_t - \mu)(X_{t-l} - \mu)\} \\
 &= \mathbb{E}\left\{\left(\epsilon_t + \sum_{i=1}^{\infty} \beta^i \epsilon_{t-i}\right)\left(\sum_{i=1}^{\infty} \beta^i \epsilon_{t-l-i}\right)\right\} \\
 &= \mathbb{E}\left\{\sum_{i,j=1}^{\infty} \beta^i \epsilon_{t-i} \beta^j \epsilon_{t-l-j}\right\} \\
 &= \mathbb{E}\left\{\sum_{i=1}^{\infty} \beta^{i+l} \epsilon_{t-l-i} \beta^i \epsilon_{t-l-i}\right\} \\
 &= \sigma^2 \sum_{i=1}^{\infty} \beta^{i+l} \beta^i \\
 &= \sigma^2 \frac{\beta^l}{1 - \beta^2} \\
 &= \beta \gamma_{l-1}
 \end{aligned}$$

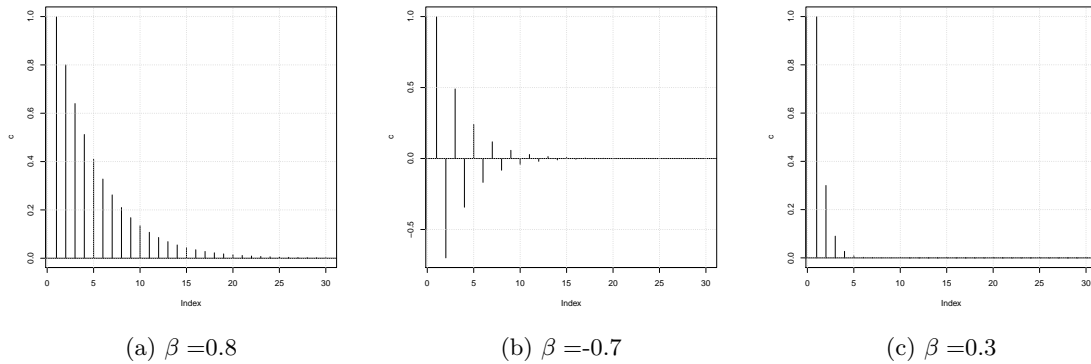


Figure 3: Expected ACF for AR(1) process

In practice, there are noises which deform the theoretical linearity relation.

Partial covariance (PACF) TODO

PACF
AR(1)

Examples In example below, check that the ACF got the linear relation expected, and that PACF got the correct value for lag 1 and approximately 0 for others.

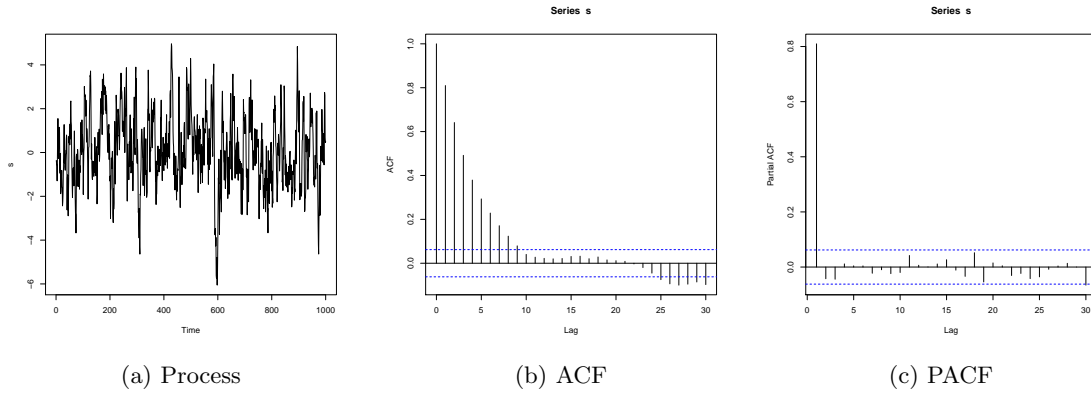


Figure 4: Example of AR(1) process, coefficient AR(0.8) (n=1000)

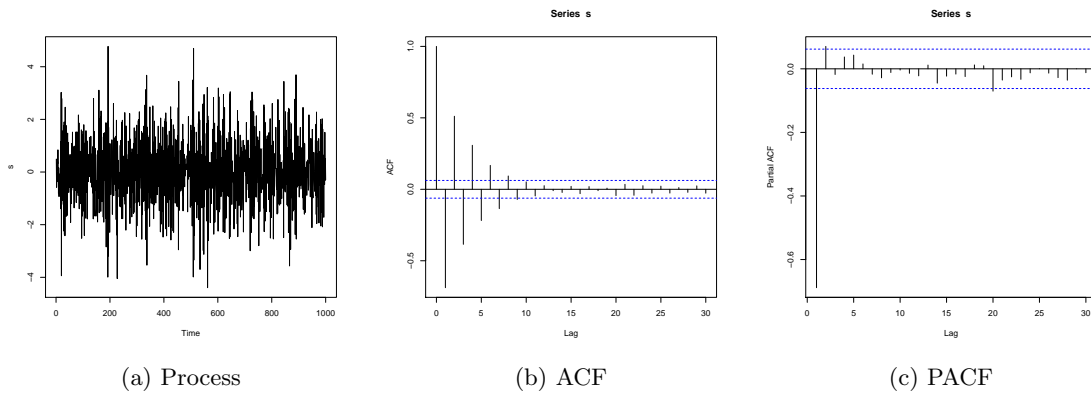


Figure 5: Example of AR(1) process, coefficient AR(-0.7) (n=1000)

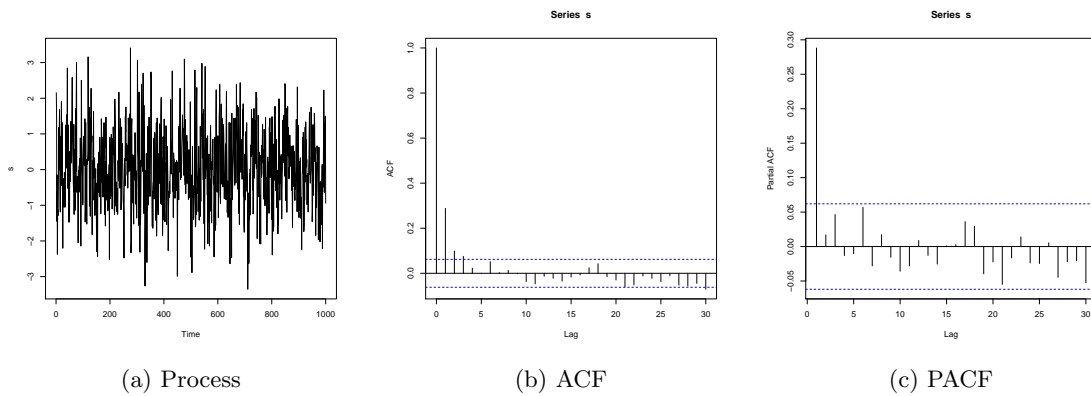


Figure 6: Example of AR(1) process, coefficient AR(0.3) (n=1000)

3.1.2 AR(p)

More ver-
bose AR(p)

Definition Be $c, \beta \in \mathbb{R}^2$, X_t is an AR(p) process if

$$\begin{aligned} X_t &= c + \sum_{i=1}^p \beta_i X_{t-i} + \epsilon_t \\ &= c + \sum_{i=1}^p \beta_i \end{aligned}$$

Stationary condition To be stationary, roots of the polynom $z^p - \sum_{i=1}^p \beta_i z^{p-i}$ must be within the unit circle, $|z_i| < 1$

Examples See below.

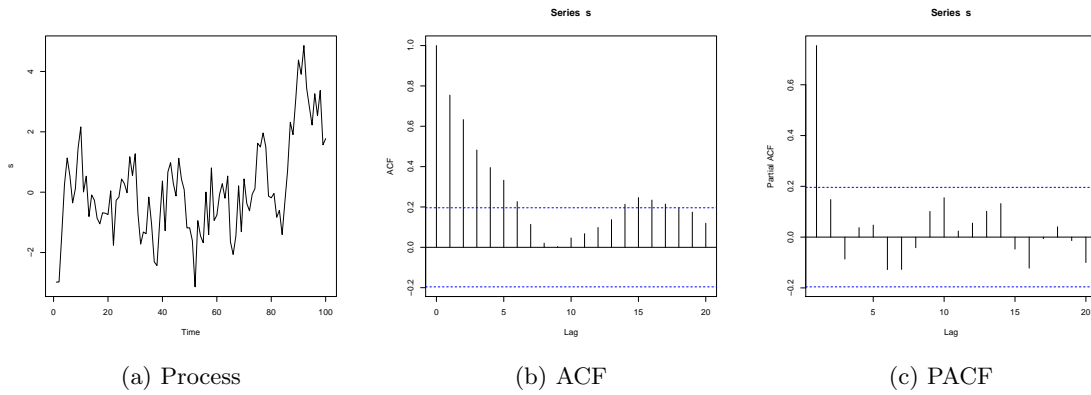


Figure 7: AR(2) process with coefficient AR(0.6,0.3) (n=100)

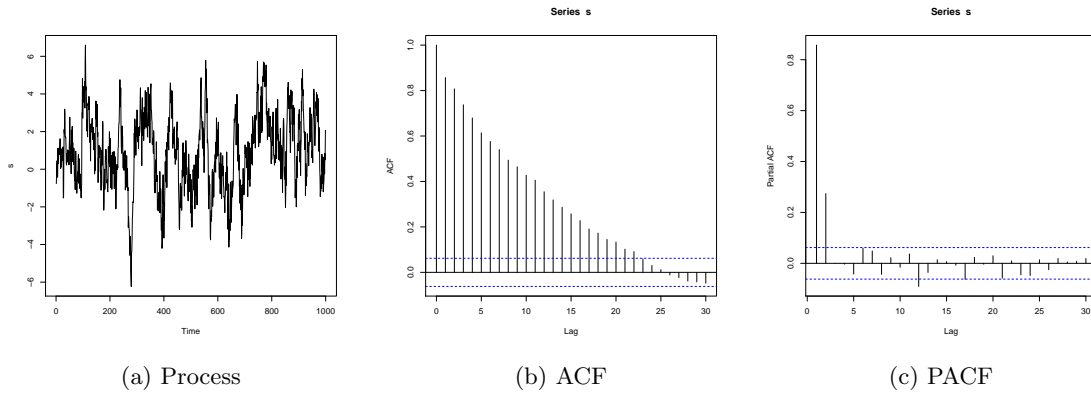


Figure 8: AR(2) process with coefficient AR(0.6,0.3) (n=1000)

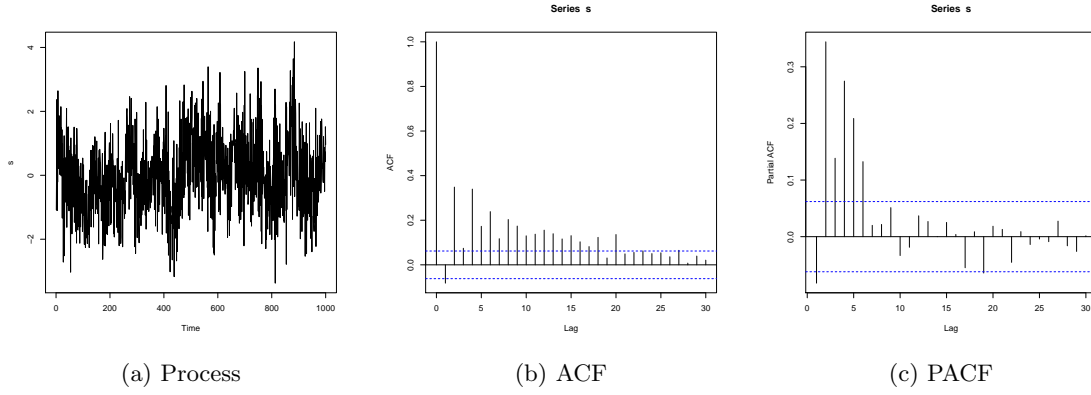


Figure 9: AR(6) process with coefficient $AR(-0.2, 0.2, 0.1, 0.3, 0.2, 0.1)$ ($n=1000$)

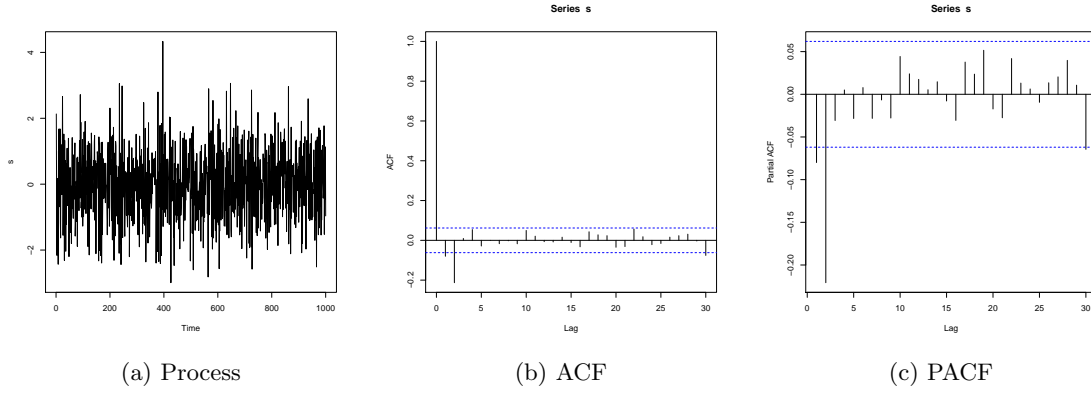


Figure 10: AR(2) process with coefficient $AR(-0.1, -0.2)$ ($n=1000$)

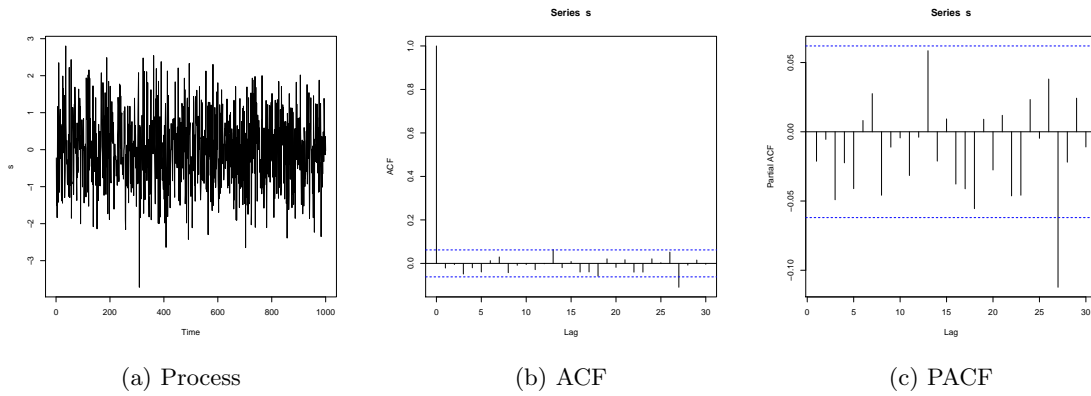


Figure 11: AR(2) process with coefficient $AR(0.01, 0.02)$ ($n=1000$)

3.1.3 MA(1)

Definition Be $\theta \in \mathbb{R}$, X_t is an MA(1) process if

$$\begin{aligned} X_t &= c + \theta \epsilon_t \\ &= c + \beta^t X_0 + \sum_{i=0}^t \beta^i \epsilon_{t-i} \end{aligned}$$

Stationary condition X_t has a finite variance : $\text{Var}(X_t) = \theta^2 \sigma^2$

Examples See below.

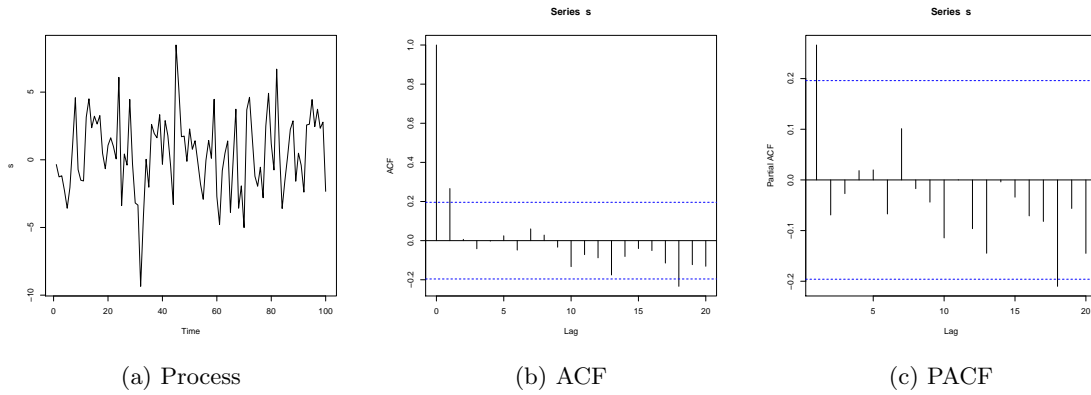


Figure 12: MA(1) process with coefficient MA(3) (n=100)

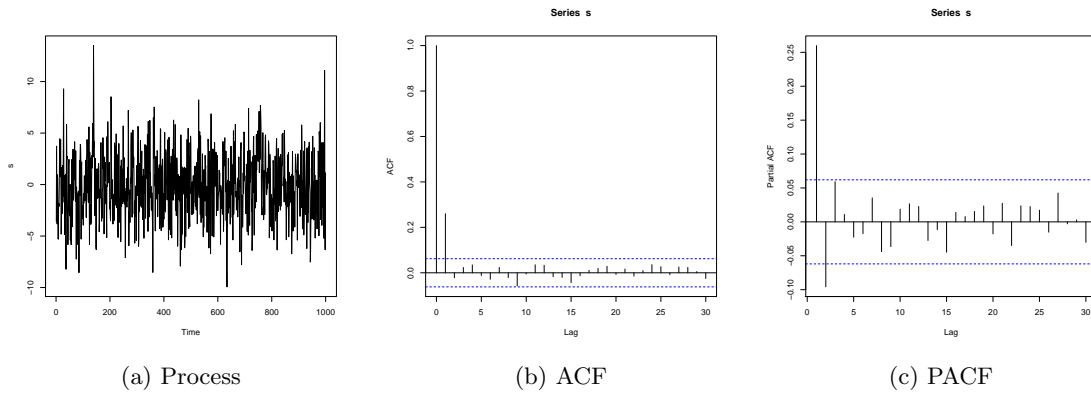


Figure 13: MA(1) process with coefficient MA(3) (n=1000)

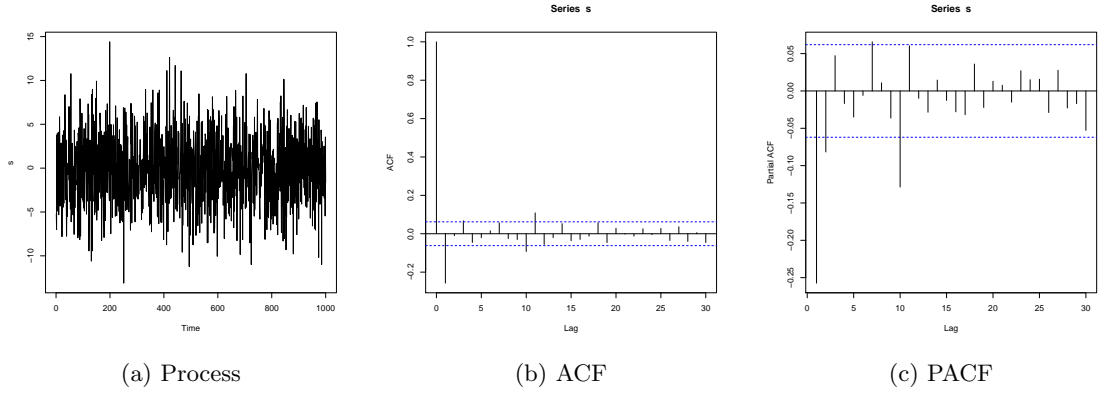


Figure 14: MA(1) process with coefficient MA(-4) (n=1000)

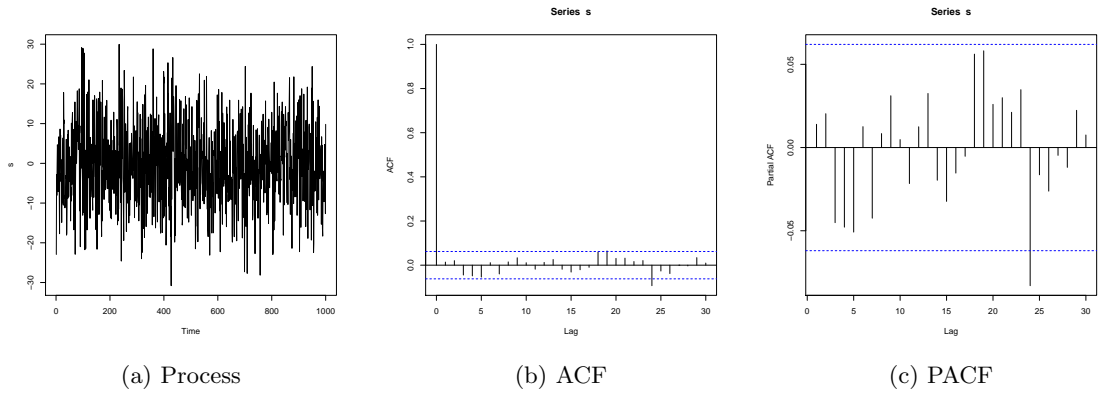


Figure 15: MA(1) process with coefficient MA(10) (n=1000)

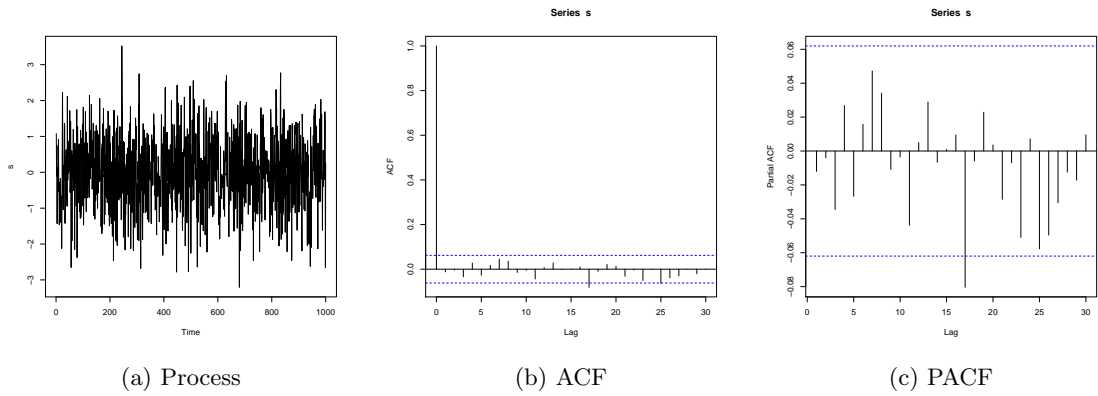


Figure 16: MA(1) process with coefficient MA(0.01) (n=1000)

3.1.4 MA(q)

Definition Be $\theta \in \mathbb{R}$, X_t is an MA(q) process if

$$X_t = c + \sum_{i=0}^q \theta_i \epsilon_{t-i}$$

Stationary condition X_t has a finite variance : $\text{Var}(X_t) = \sum_{i=0}^q \theta_i^2 \sigma^2$

Examples See below.

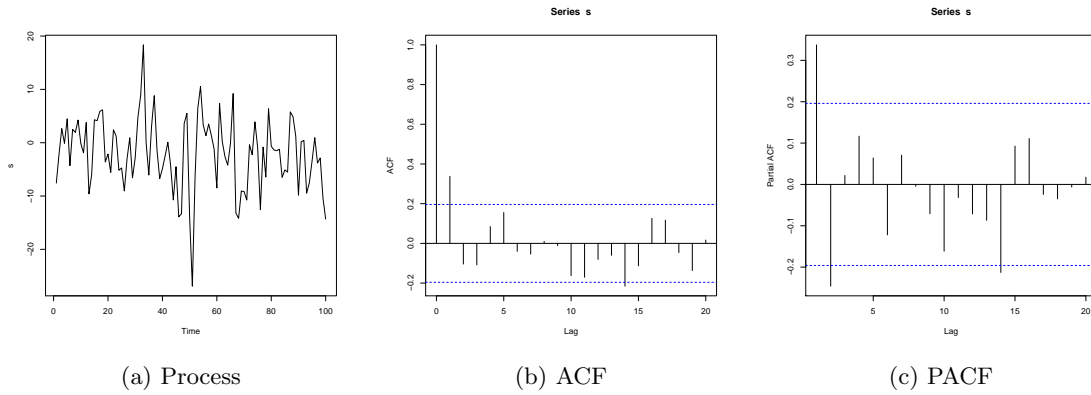


Figure 17: MA(2) process with coefficient MA(3,6) (n=100)

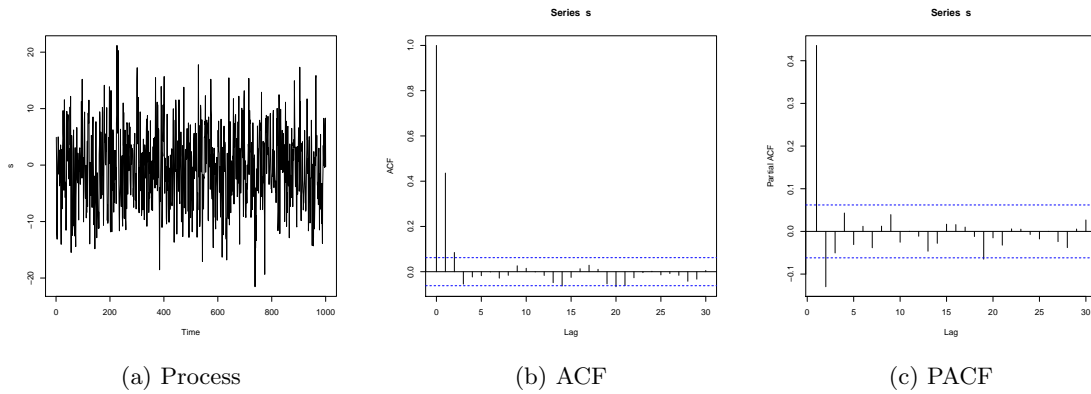


Figure 18: MA(2) process with coefficient MA(3,6) (n=1000)

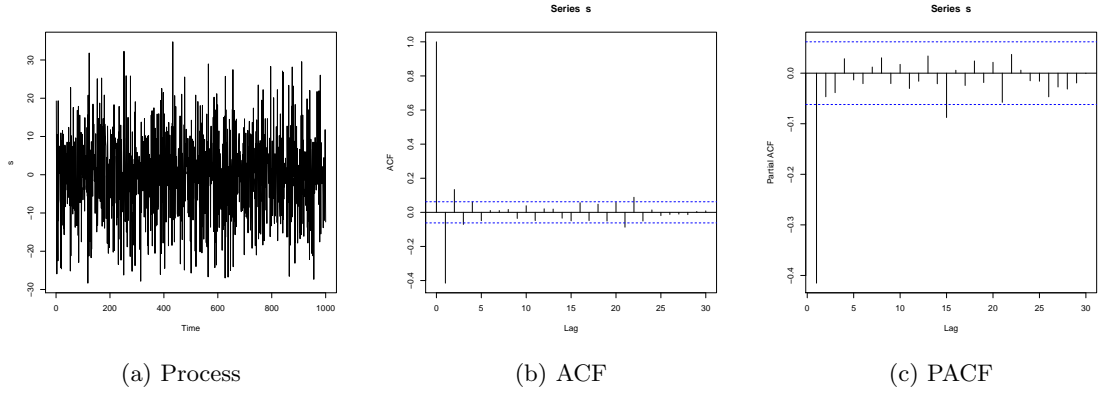


Figure 19: MA(2) process with coefficient MA(-4,10) (n=1000)

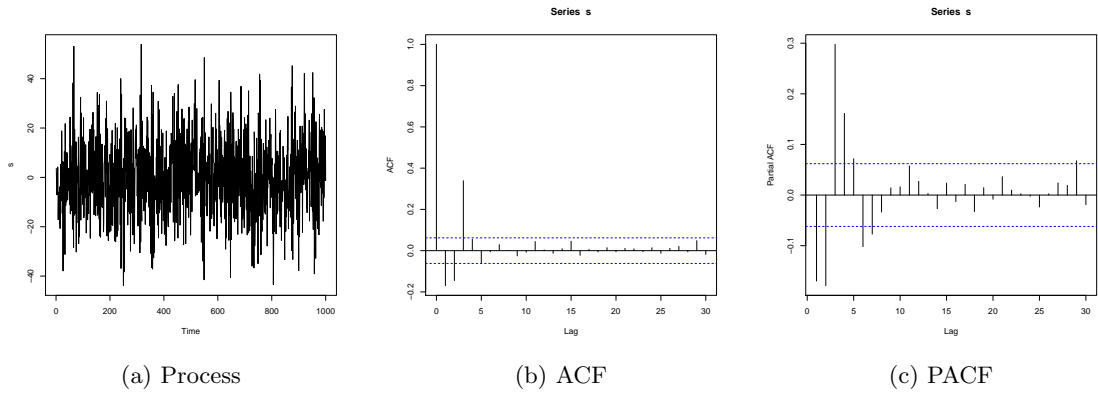


Figure 20: MA(4) process with coefficient MA(10,4,-7,9) (n=1000)

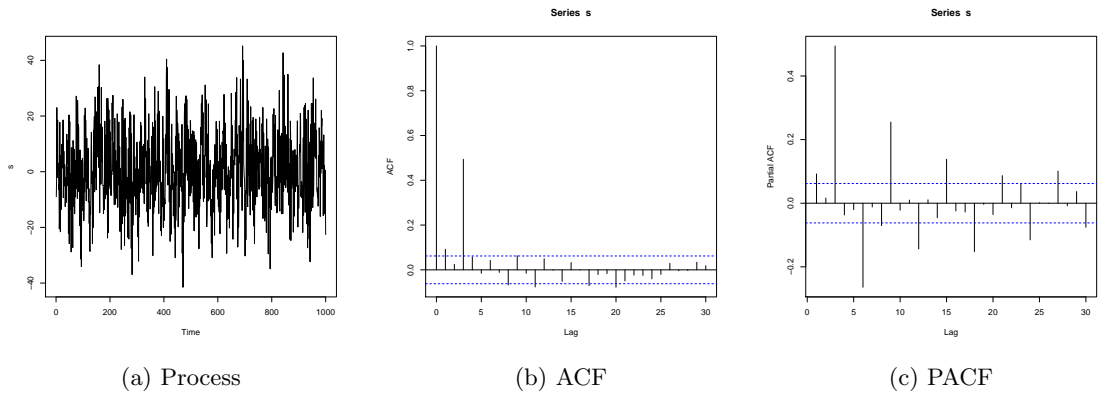


Figure 21: MA(4) process with coefficient MA(10,0,0,9) (n=1000)

3.1.5 ARMA(p,q)

Definition Be $\theta \in \mathbb{R}$, X_t is an MA(q) process if

$$X_t = c + \sum_{i=0}^q \theta_i \epsilon_{t-i} + \sum_{i=1}^p \beta_i X_{t-i}$$

stationary To see if X_t is stationnary we look at its variance.

$$\text{Var}(X_t) = \sum_{i=0}^q \theta_i^2 \sigma^2$$

Examples See below various example of ARMA(p,q)

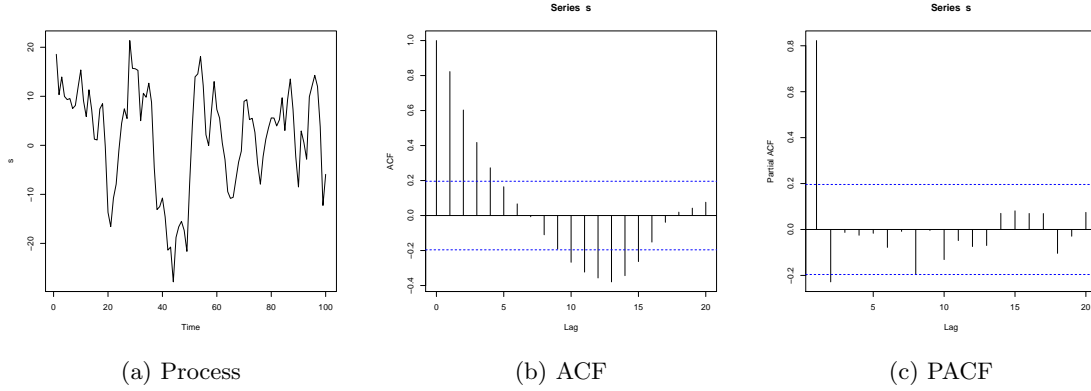


Figure 22: ARMA(1,2) process with coefficient AR(0.8), MA(3,6) (n=100)

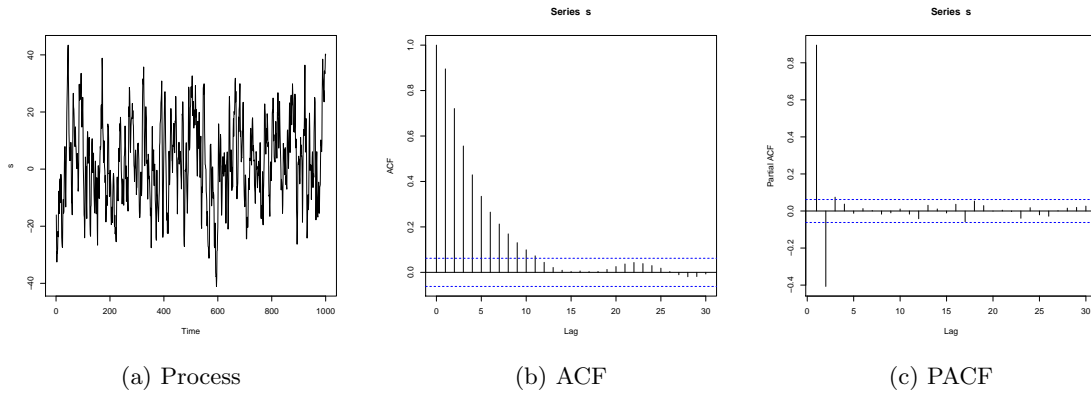


Figure 23: ARMA(1,2) process with coefficient AR(0.8), MA(3,6) (n=1000)

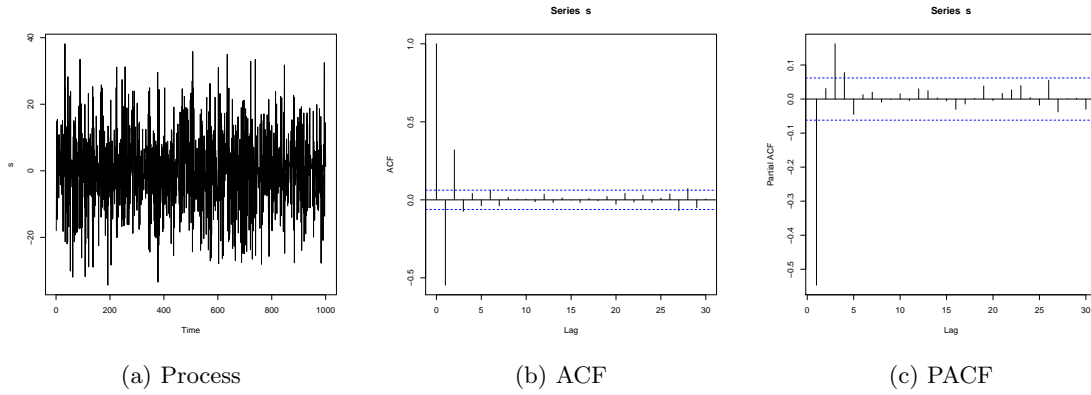


Figure 24: ARMA(3,2) process with coefficient AR(-0.1,0.2,0.1), MA(-4,10) (n=1000)

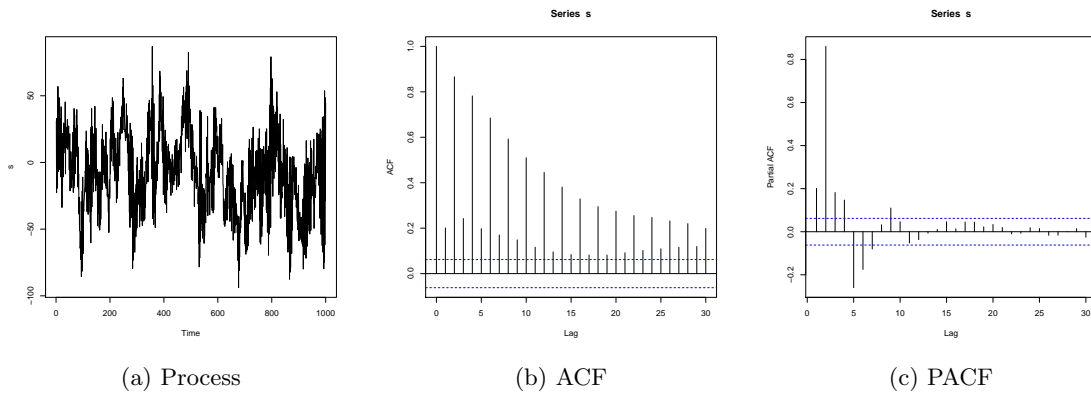


Figure 25: ARMA(2,4) process with coefficient AR(0,0.9), MA(10,4,-7,9) (n=1000)

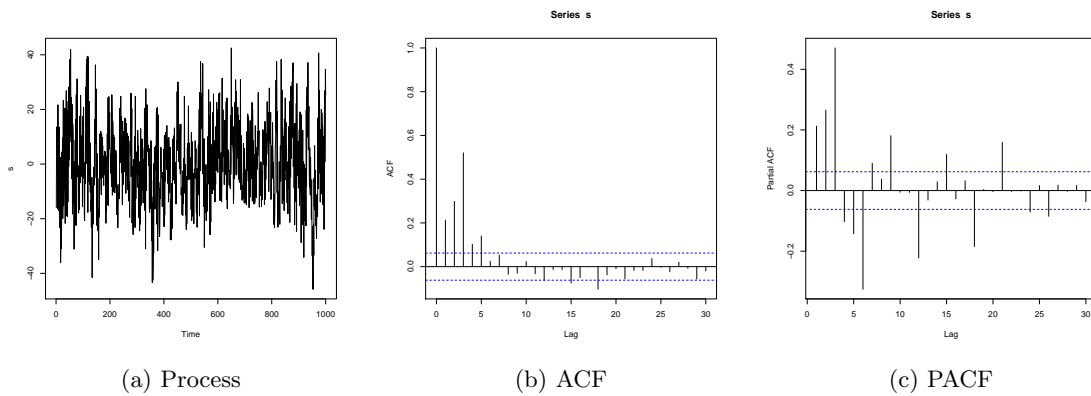


Figure 26: ARMA(2,4) process with coefficient AR(0,0.3), MA(10,0,0,9) (n=1000)

3.2 Autoregressive conditional heteroskedasticity (ARCH)

Definition X_t is an ARCH(q) model if

$$\begin{aligned}X_t &= \epsilon_t \sigma_t \\ \sigma_t^2 &= \alpha_0 + \sum_{i=1}^q \alpha_i \epsilon_{t-i}^2\end{aligned}$$

3.3 Generalized autoregressive conditional heteroskedasticity (GARCH)

GARCH
details

Definition X_t is an GARCH(p,q) model if

$$\begin{aligned}X_t &= \epsilon_t \sigma_t \\ \sigma_t^2 &= \alpha_0 + \sum_{i=1}^q \alpha_i \epsilon_{t-i}^2 + \sum_{i=1}^p \beta_i \sigma_{t-i}^2\end{aligned}$$

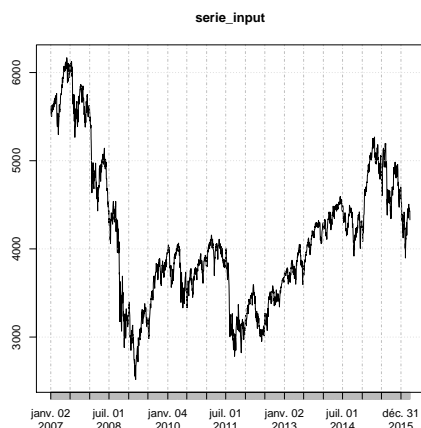
3.4 Markov Switching Model

MSM details

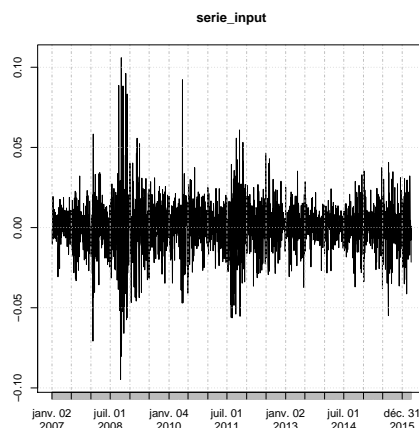
4 Application

4.1 CAC40

This is the historical close quotation for CAC40. Data has been retrieved from yahoo.



(a) Close level of CAC40



(b) log return CAC40

Looking at original

Is it an AR model ? We look for the AR coefficient, and minimize it with the AIC indicator. The AIC is a ratio between a good calibration and the number of parameters. The AR model implies that residual should be white noise (iid normal law). It's tested with the QQnorm. Results:

Call:

```
ar(x = value)
```

Coefficients:

1	2	3	4	5	6
0.9586	0.0052	-0.0014	0.0545	-0.0643	0.0442

Order selected 6 σ^2 estimated as 4526

Explain why we work on log return instead of close level ?

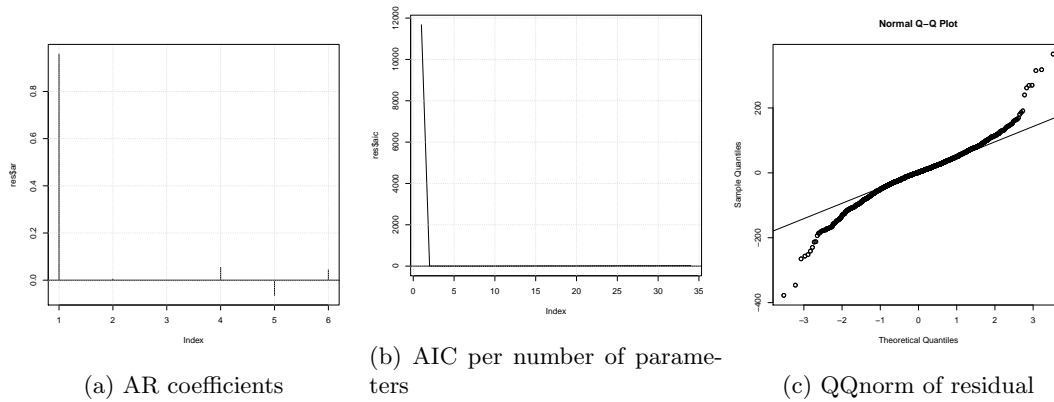
Maybe skip the AR/MA model, and just go to ARMA directly. Move AR/MA into example?

Details in annexe how these tests work

Missing MSM details

Missing Forecast / Backtesting

Add references to figure, few introductions ?



Test residual independences with ljung:

Box-Ljung test

```
data: res
X-squared = 0.34186, df = 1, p-value = 0.5588
```

Test residual normal law with Shapiro:

Shapiro-Wilk normality test

```
data: res
W = 0.9639, p-value < 2.2e-16
```

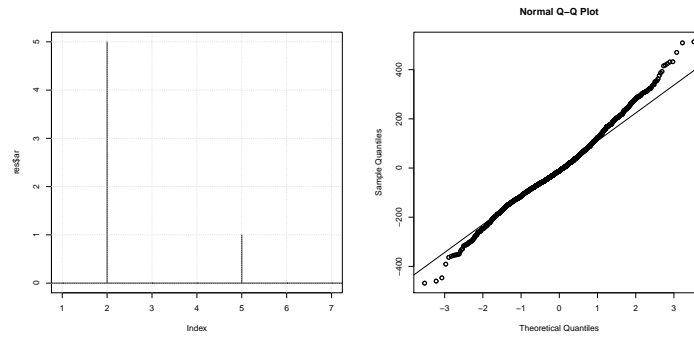
Is it an MA model ? We look for the AR coefficient, and minimize it with the AIC indicator. The AIC is a ratio between a good calibration and the number of parameters. The AR model implies that residual should be white noise (iid normal law). It's tested with the QQnorm.

Results:

```
Series: value
ARIMA(0,0,5) with non-zero mean
```

```
Coefficients:
      ma1      ma2      ma3      ma4      ma5  intercept
      1.7514  2.1009  1.9138  1.3178  0.5338  4139.6223
s.e.    0.0265  0.0439  0.0337  0.0234  0.0180   22.3093
```

```
sigma^2 estimated as 15867: log likelihood=-14782.85
AIC=29579.69 AICc=29579.74 BIC=29620.07
```



(a) MA coefficients

(b) QQnorm of residual

Test residual independences with ljung:

Box-Ljung test

```
data: res
X-squared = 209.85, df = 1, p-value < 2.2e-16
```

Test residual normal law with Shapiro:

Shapiro-Wilk normality test

```
data: res
W = 0.98947, p-value = 3.684e-12
```

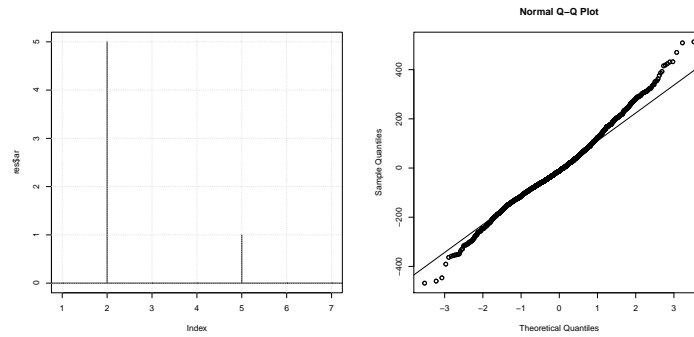
Is it an ARMA model ? We look for the AR coefficient, and minimize it with the AIC indicator. The AIC is a ratio between a good calibration and the number of parameters. The AR model implies that residual should be white noise (iid normal law). It's tested with the QQnorm.

Results:

```
Series: value
ARIMA(0,0,5) with non-zero mean
```

```
Coefficients:
      ma1      ma2      ma3      ma4      ma5  intercept
    1.7514  2.1009  1.9138  1.3178  0.5338  4139.6223
s.e.  0.0265  0.0439  0.0337  0.0234  0.0180   22.3093
```

```
sigma^2 estimated as 15867: log likelihood=-14782.85
AIC=29579.69 AICc=29579.74 BIC=29620.07
```



(a) MA coefficients

(b) QQnorm of residual

Test residual independences with ljung:

Box-Ljung test

```
data: res
X-squared = 209.85, df = 1, p-value < 2.2e-16
```

Test residual normal law with Shapiro:

Shapiro-Wilk normality test

```
data: res
W = 0.98947, p-value = 3.684e-12
```

Markov Switching Model Try with 2 regimes.

Results:

Markov Switching Model

```
Call: msmFit(object = mod, k = 2, sw = c(T, T, T), p = 1, control = list(parallel = F))
```

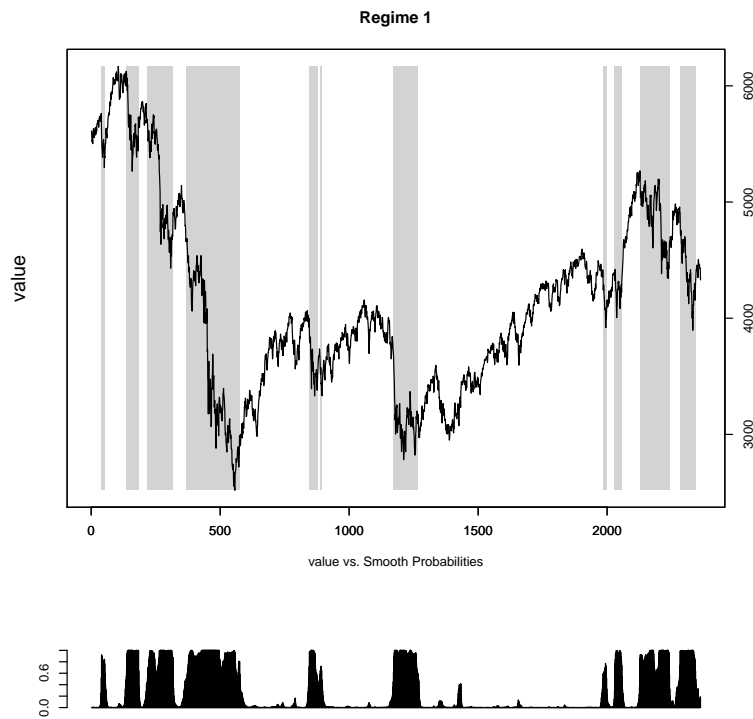
	AIC	BIC	logLik
	25609.13	25663.27	-12800.56

Coefficients:

	(Intercept)(S)	value_1(S)	Std(S)
Model 1	18.121358	0.9938143	86.46705
Model 2	7.304968	0.9989027	42.64456

Transition probabilities:

	Regime 1	Regime 2
Regime 1	0.9788516	0.0093812
Regime 2	0.0211484	0.9906188



(a) Which 2

Looking at logret

Is it an AR model ? We look for the AR coefficient, and minimize it with the AIC indicator. The AIC is a ratio between a good calibration and the number of parameters. The AR model implies that residual should be white noise (iid normal law). It's tested with the QQnorm. Results:

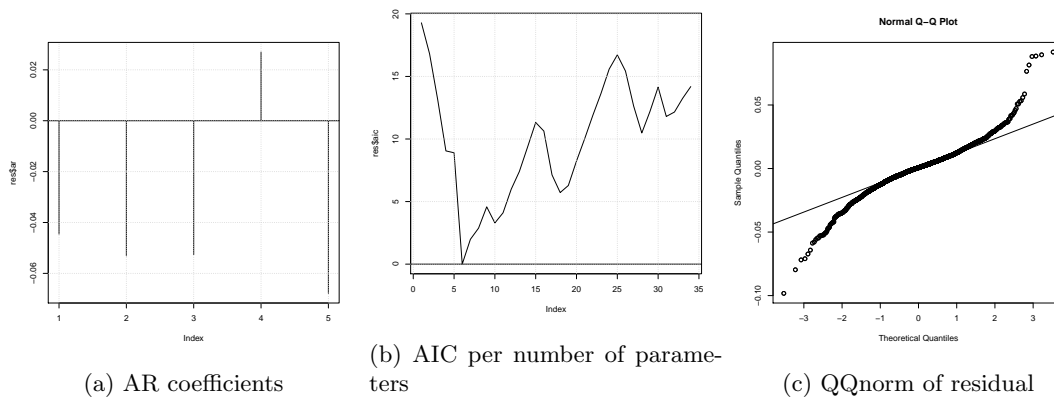
Call:

```
ar(x = value)
```

Coefficients:

1	2	3	4	5
-0.0446	-0.0531	-0.0526	0.0270	-0.0679

Order selected 5 σ^2 estimated as 0.0002385



Test residual independences with ljung:

Box-Ljung test

```
data: res
X-squared = 0.00024827, df = 1, p-value = 0.9874
```

Test residual normal law with Shapiro:

Shapiro-Wilk normality test

```
data: res
W = 0.94774, p-value < 2.2e-16
```

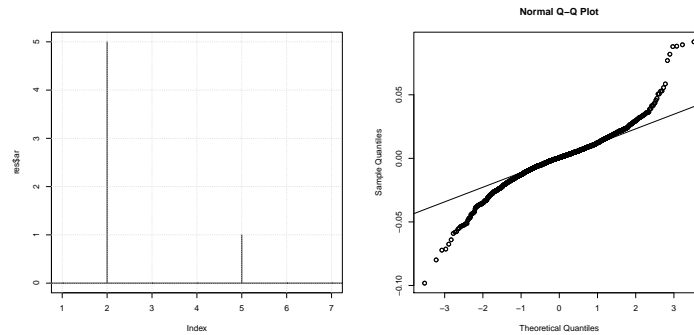
Is it an MA model ? We look for the AR coefficient, and minimize it with the AIC indicator. The AIC is a ratio between a good calibration and the number of parameters. The AR model implies that residual should be white noise (iid normal law). It's tested with the QQnorm.

Results:

```
Series: value
ARIMA(0,0,5) with zero mean
```

```
Coefficients:
          ma1          ma2          ma3          ma4          ma5
      -0.0446  -0.0485  -0.0482   0.0307  -0.0627
s.e.    0.0205   0.0207   0.0205   0.0205   0.0211
```

```
sigma^2 estimated as 0.000238:  log likelihood=6501.84
AIC=-12991.69  AICc=-12991.65  BIC=-12957.08
```

(a) MA coefficients

(b) QQnorm of residual

Test residual independences with ljung:

Box-Ljung test

```
data: res
X-squared = 0.00020489, df = 1, p-value = 0.9886
```

Test residual normal law with Shapiro:

Shapiro-Wilk normality test

```
data: res
W = 0.94707, p-value < 2.2e-16
```

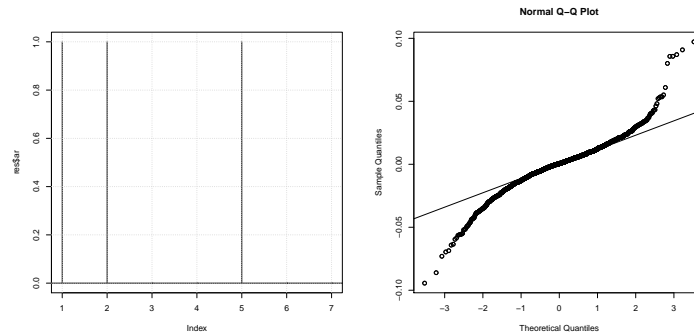
Is it an ARMA model ? We look for the AR coefficient, and minimize it with the AIC indicator. The AIC is a ratio between a good calibration and the number of parameters. The AR model implies that residual should be white noise (iid normal law). It's tested with the QQnorm.

Results:

```
Series: value
ARIMA(0,0,5) with zero mean
```

```
Coefficients:
      ma1      ma2      ma3      ma4      ma5
-0.0446 -0.0485 -0.0482  0.0307 -0.0627
s.e.    0.0205  0.0207  0.0205  0.0205  0.0211
```

```
sigma^2 estimated as 0.000238: log likelihood=6501.84
AIC=-12991.69 AICc=-12991.65 BIC=-12957.08
```



(a) MA coefficients

(b) QQnorm of residual

Test residual independences with ljung:

Box-Ljung test

```
data: res
X-squared = 0.12362, df = 1, p-value = 0.7251
```

Test residual normal law with Shapiro:

Shapiro-Wilk normality test

```
data: res
W = 0.94391, p-value < 2.2e-16
```

Markov Switching Model Try with 2 regimes.

Results:

Markov Switching Model

```
Call: msmFit(object = mod, k = 2, sw = c(T, T, T), p = 1, control = list(parallel = F))
```

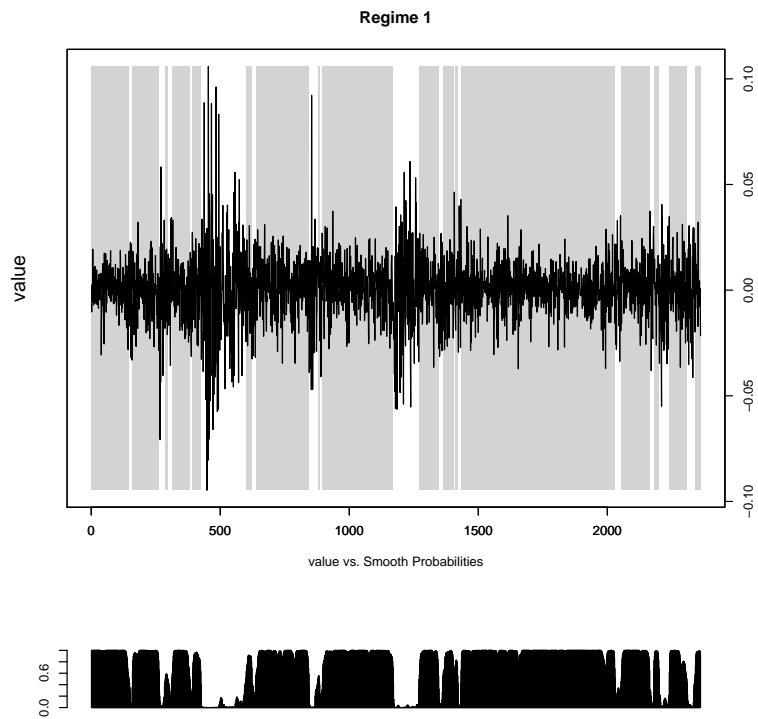
	AIC	BIC	logLik
	-13538.51	-13484.38	6773.255

Coefficients:

	(Intercept)(S)	value_1(S)	Std(S)
Model 1	0.0005173361	-0.03039309	0.01084260
Model 2	-0.0021582307	-0.06141297	0.02489934

Transition probabilities:

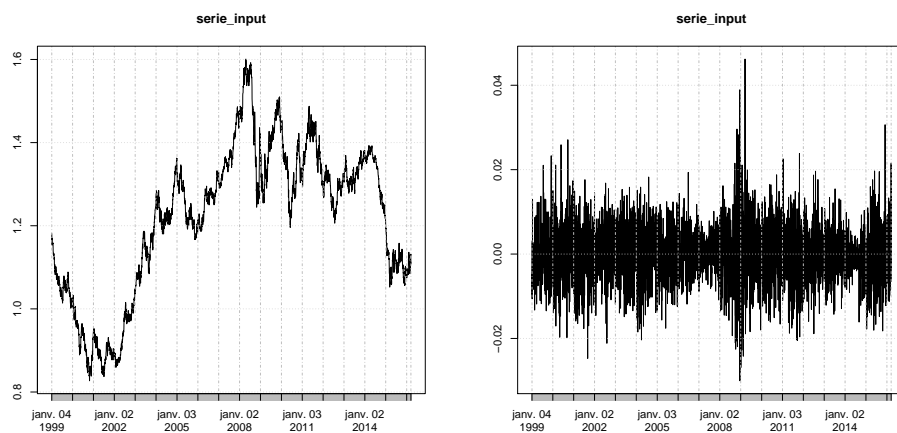
	Regime 1	Regime 2
Regime 1	0.98836266	0.03654694
Regime 2	0.01163734	0.96345306



(a) Which 2

4.2 EURUSD

This is the historical close quotation for EURUSD. Data has been retrieved from FRED.



(a) Close level of EURUSD

(b) log return EURUSD

Looking at original

Is it an AR model ? We look for the AR coefficient, and minimize it with the AIC indicator. The AIC is a ratio between a good calibration and the number of parameters. The AR model implies that residual should be white noise (iid normal law). It's tested with the QQnorm. Results:

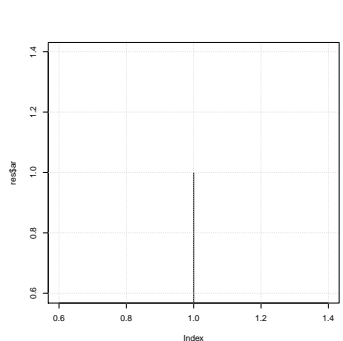
Call:

```
ar(x = value)
```

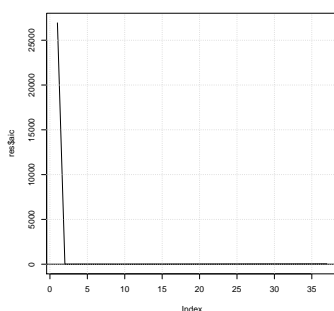
Coefficients:

```
1
0.999
```

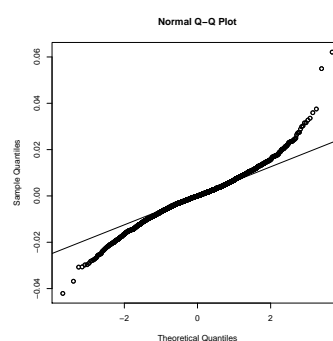
Order selected 1 sigma^2 estimated as 6.313e-05



(a) AR coefficients



(b) AIC per number of parameters



(c) QQnorm of residual

Test residual independences with ljung:

Box-Ljung test

```
data: res
```

```
X-squared = 0.17242, df = 1, p-value = 0.678
```

Test residual normal law with Shapiro:

Shapiro-Wilk normality test

```
data: res
```

```
W = 0.97606, p-value < 2.2e-16
```

Is it an MA model ? We look for the AR coefficient, and minimize it with the AIC indicator. The AIC is a ratio between a good calibration and the number of parameters. The AR model implies that residual should be white noise (iid normal law). It's tested with the QQnorm.

Results:

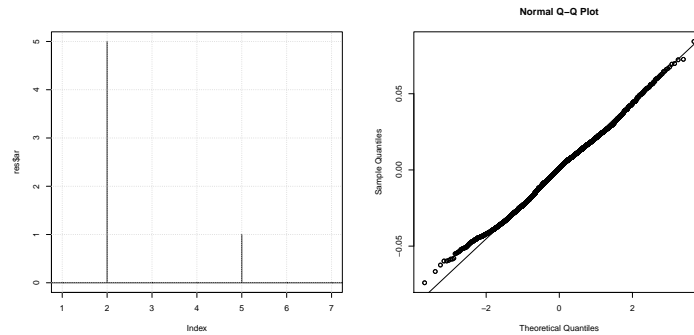
```
Series: value
```

```
ARIMA(0,0,5) with non-zero mean
```

Coefficients:

	ma1	ma2	ma3	ma4	ma5	intercept
	2.1483	2.8344	2.6013	1.6428	0.5836	1.2171
s.e.	0.0182	0.0336	0.0296	0.0182	0.0116	0.0036

sigma² estimated as 0.0004708: log likelihood=10429.22
AIC=-20844.43 AICc=-20844.4 BIC=-20799.82



(a) MA coefficients

(b) QQnorm of residual

Test residual independences with ljung:

Box-Ljung test

data: res
X-squared = 479.06, df = 1, p-value < 2.2e-16

Test residual normal law with Shapiro:

Shapiro-Wilk normality test

data: res
W = 0.99702, p-value = 1.523e-07

Is it an ARMA model ? We look for the AR coefficient, and minimize it with the AIC indicator. The AIC is a ratio between a good calibration and the number of parameters. The AR model implies that residual should be white noise (iid normal law). It's tested with the QQnorm.

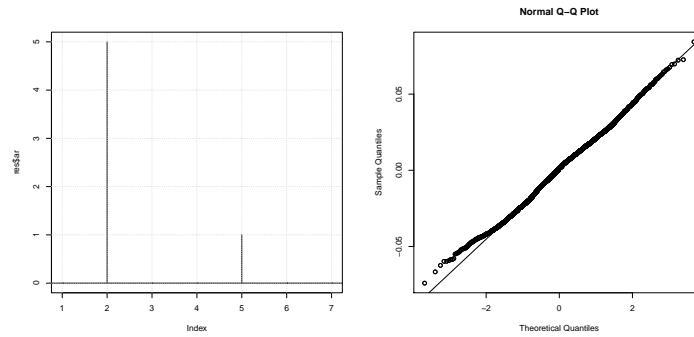
Results:

Series: value
ARIMA(0,0,5) with non-zero mean

Coefficients:

	ma1	ma2	ma3	ma4	ma5	intercept
	2.1483	2.8344	2.6013	1.6428	0.5836	1.2171
s.e.	0.0182	0.0336	0.0296	0.0182	0.0116	0.0036

sigma² estimated as 0.0004708: log likelihood=10429.22
AIC=-20844.43 AICc=-20844.4 BIC=-20799.82



(a) MA coefficients

(b) QQnorm of residual

Test residual independences with ljung:

Box-Ljung test

```
data: res
X-squared = 479.06, df = 1, p-value < 2.2e-16
```

Test residual normal law with Shapiro:

Shapiro-Wilk normality test

```
data: res
W = 0.99702, p-value = 1.523e-07
```

Markov Switching Model Try with 2 regimes.

Results:

Markov Switching Model

```
Call: msmFit(object = mod, k = 2, sw = c(T, T, T), p = 1, control = list(parallel = F))
```

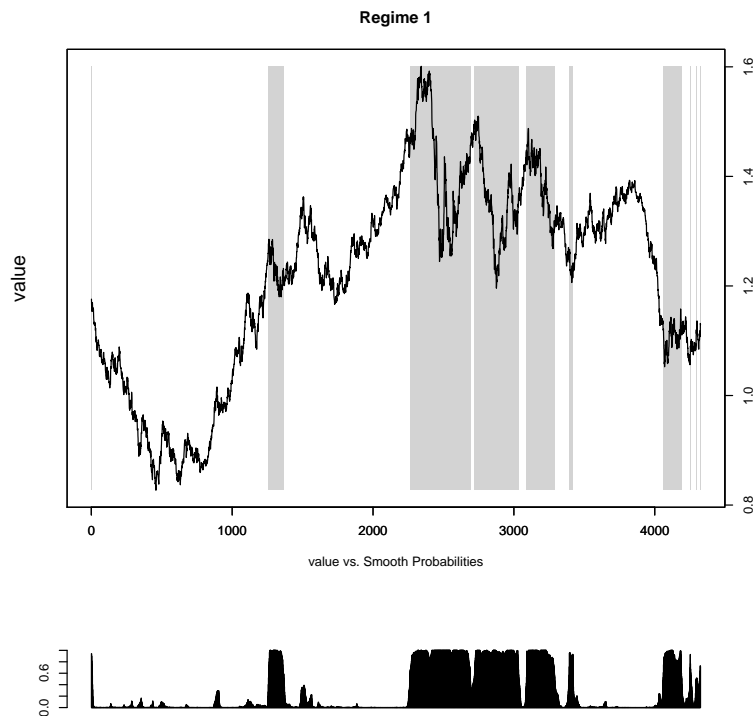
	AIC	BIC	logLik
	-30176.57	-30117.6	15092.29

Coefficients:

	(Intercept)(S)	value_1(S)	Std(S)
Model 1	0.0051391138	0.9959467	0.010649831
Model 2	-0.0001810602	1.0002282	0.006159484

Transition probabilities:

	Regime 1	Regime 2
Regime 1	0.98867243	0.005043856
Regime 2	0.01132757	0.994956144



(a) Which 2

Looking at logret

Is it an AR model ? We look for the AR coefficient, and minimize it with the AIC indicator. The AIC is a ratio between a good calibration and the number of parameters. The AR model implies that residual should be white noise (iid normal law). It's tested with the QQnorm. Results:

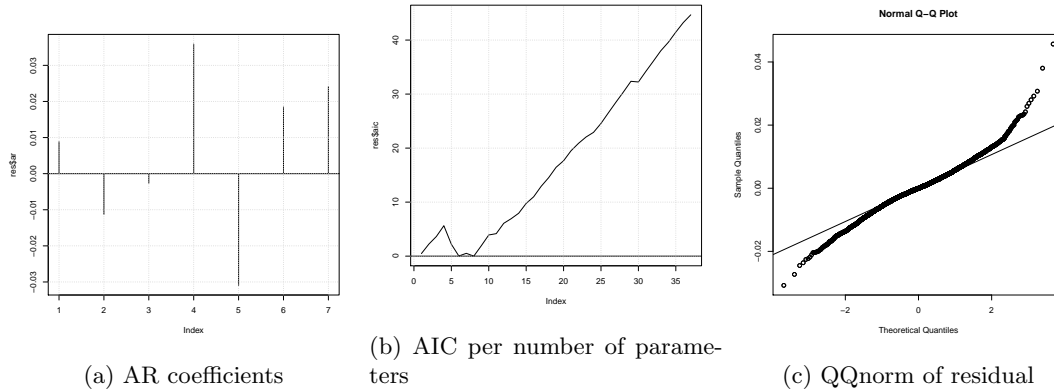
Call:

```
ar(x = value)
```

Coefficients:

1	2	3	4	5	6	7
0.0088	-0.0113	-0.0026	0.0359	-0.0309	0.0185	0.0241

Order selected 7 sigma² estimated as 4.069e-05



Test residual independences with ljung:

Box-Ljung test

```
data: res
X-squared = 1.1456e-05, df = 1, p-value = 0.9973
```

Test residual normal law with Shapiro:

Shapiro-Wilk normality test

```
data: res
W = 0.98309, p-value < 2.2e-16
```

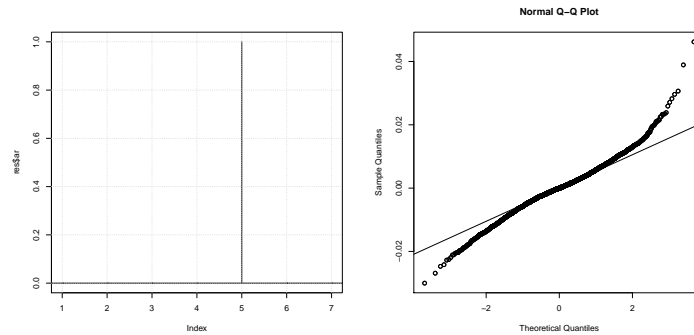
Is it an MA model ? We look for the AR coefficient, and minimize it with the AIC indicator. The AIC is a ratio between a good calibration and the number of parameters. The AR model implies that residual should be white noise (iid normal law). It's tested with the QQnorm.

Results:

```
Series: value
ARIMA(0,0,0) with non-zero mean
```

```
Coefficients:
      intercept
              0e+00
s.e.          1e-04
```

```
sigma^2 estimated as 4.075e-05: log likelihood=15721.53
AIC=-31439.06 AICc=-31439.06 BIC=-31426.32
```

(a) MA coefficients

(b) QQnorm of residual

Test residual independences with ljung:

Box-Ljung test

```
data: res
X-squared = 0.24259, df = 1, p-value = 0.6223
```

Test residual normal law with Shapiro:

Shapiro-Wilk normality test

```
data: res
W = 0.98276, p-value < 2.2e-16
```

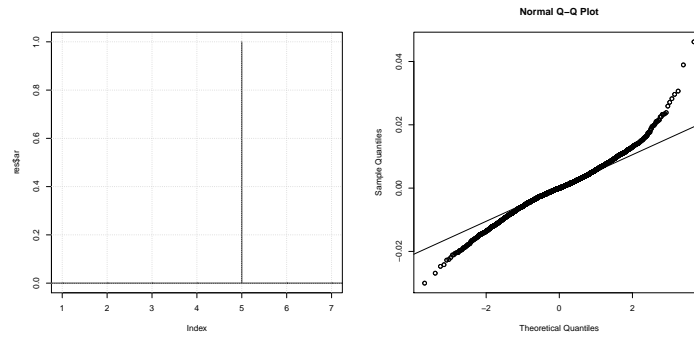
Is it an ARMA model ? We look for the AR coefficient, and minimize it with the AIC indicator. The AIC is a ratio between a good calibration and the number of parameters. The AR model implies that residual should be white noise (iid normal law). It's tested with the QQnorm.

Results:

```
Series: value
ARIMA(0,0,0) with non-zero mean
```

```
Coefficients:
      intercept
              0e+00
s.e.          1e-04
```

```
sigma^2 estimated as 4.075e-05:  log likelihood=15721.53
AIC=-31439.06  AICc=-31439.06  BIC=-31426.32
```



(a) MA coefficients

(b) QQnorm of residual

Test residual independences with ljung:

Box-Ljung test

```
data: res
X-squared = 0.24259, df = 1, p-value = 0.6223
```

Test residual normal law with Shapiro:

Shapiro-Wilk normality test

```
data: res
W = 0.98276, p-value < 2.2e-16
```

Markov Switching Model Try with 2 regimes.

Results:

Markov Switching Model

```
Call: msmFit(object = mod, k = 2, sw = c(T, T, T), p = 1, control = list(parallel = F))
```

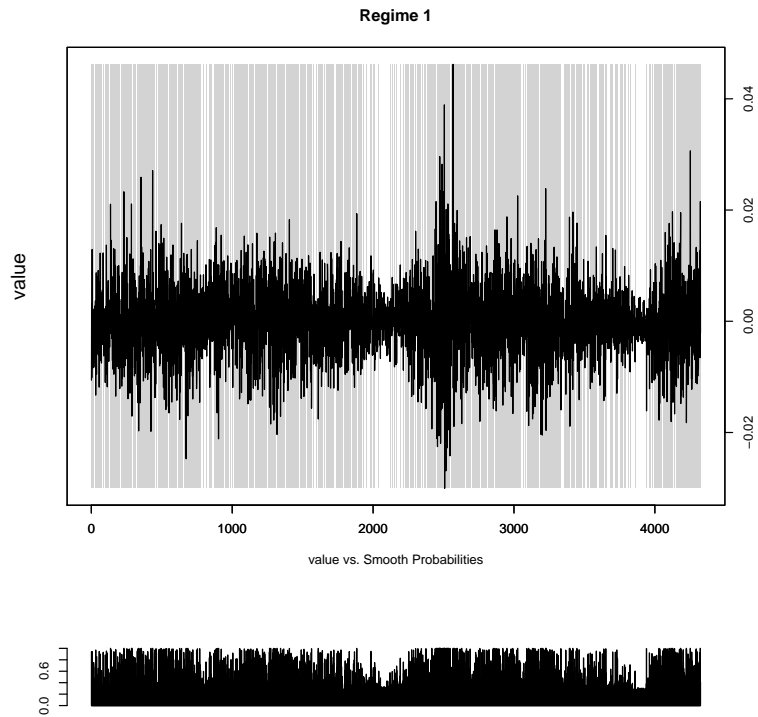
	AIC	BIC	logLik
	-31659.28	-31600.3	15833.64

Coefficients:

	(Intercept)(S)	value_1(S)	Std(S)
Model 1	-0.0002184246	0.06645888	0.008243233
Model 2	0.0001969037	-0.06868736	0.003960878

Transition probabilities:

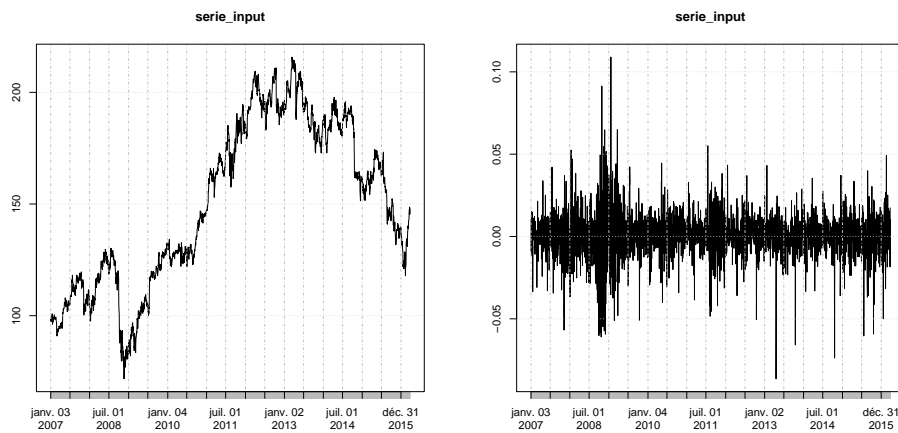
	Regime 1	Regime 2
Regime 1	0.6089057	0.3545759
Regime 2	0.3910943	0.6454241



(a) Which 2

4.3 IBM

This is the historical close quotation for IBM. Data has been retrieved from yahoo.



(a) Close level of IBM

(b) log return IBM

Looking at original

Is it an AR model ? We look for the AR coefficient, and minimize it with the AIC indicator. The AIC is a ratio between a good calibration and the number of parameters. The AR model implies that residual should be white noise (iid normal law). It's tested with the QQnorm. Results:

Call:

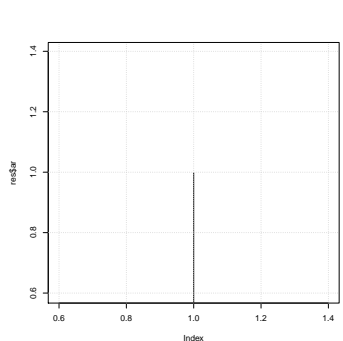
```
ar(x = value)
```

Coefficients:

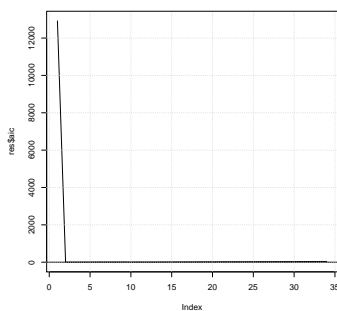
1

0.9981

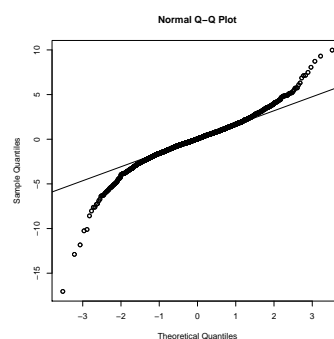
Order selected 1 sigma^2 estimated as 5.151



(a) AR coefficients



(b) AIC per number of parameters



(c) QQnorm of residual

Test residual independences with ljung:

Box-Ljung test

data: res

X-squared = 0.098621, df = 1, p-value = 0.7535

Test residual normal law with Shapiro:

Shapiro-Wilk normality test

data: res

W = 0.9508, p-value < 2.2e-16

Is it an MA model ? We look for the AR coefficient, and minimize it with the AIC indicator. The AIC is a ratio between a good calibration and the number of parameters. The AR model implies that residual should be white noise (iid normal law). It's tested with the QQnorm.

Results:

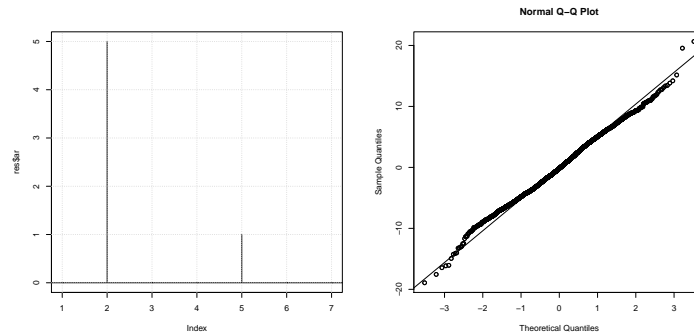
Series: value

ARIMA(0,0,5) with non-zero mean

Coefficients:

	ma1	ma2	ma3	ma4	ma5	intercept
	2.0207	2.6273	2.4116	1.5390	0.568	150.1698
s.e.	0.0216	0.0375	0.0352	0.0242	0.015	1.0180

sigma² estimated as 23.34: log likelihood=-6958.04
AIC=13930.08 AICc=13930.13 BIC=13970.33



(a) MA coefficients

(b) QQnorm of residual

Test residual independences with ljung:

Box-Ljung test

data: res
X-squared = 213.87, df = 1, p-value < 2.2e-16

Test residual normal law with Shapiro:

Shapiro-Wilk normality test

data: res
W = 0.99703, p-value = 0.0001803

Is it an ARMA model ? We look for the AR coefficient, and minimize it with the AIC indicator. The AIC is a ratio between a good calibration and the number of parameters. The AR model implies that residual should be white noise (iid normal law). It's tested with the QQnorm.

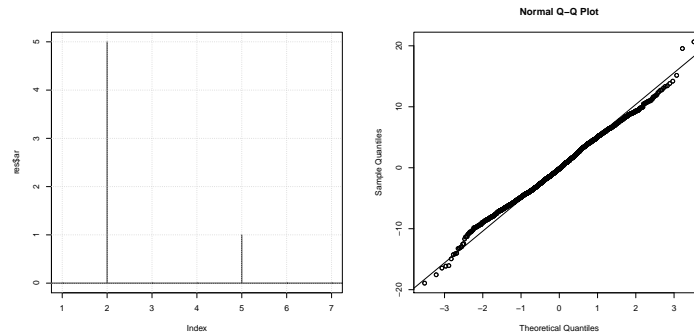
Results:

Series: value
ARIMA(0,0,5) with non-zero mean

Coefficients:

	ma1	ma2	ma3	ma4	ma5	intercept
	2.0207	2.6273	2.4116	1.5390	0.568	150.1698
s.e.	0.0216	0.0375	0.0352	0.0242	0.015	1.0180

sigma² estimated as 23.34: log likelihood=-6958.04
AIC=13930.08 AICc=13930.13 BIC=13970.33



(a) MA coefficients

(b) QQnorm of residual

Test residual independences with ljung:

Box-Ljung test

```
data:  res
X-squared = 213.87, df = 1, p-value < 2.2e-16
```

Test residual normal law with Shapiro:

Shapiro-Wilk normality test

```
data:  res
W = 0.99703, p-value = 0.0001803
```

Markov Switching Model Try with 2 regimes.

Results:

Markov Switching Model

```
Call: msmFit(object = mod, k = 2, sw = c(T, T, T), p = 1, control = list(parallel = F))
```

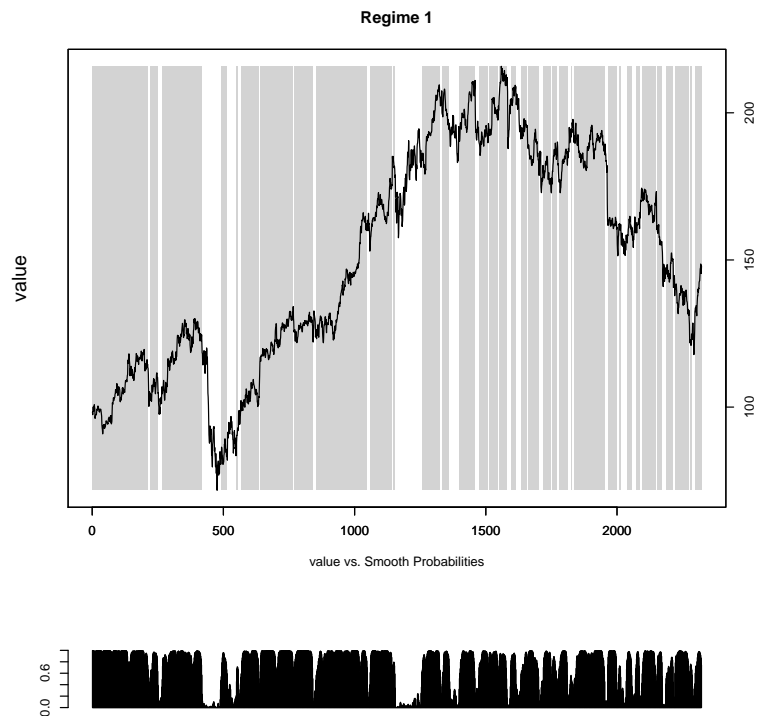
AIC	BIC	logLik
9388.168	9442.169	-4690.084

Coefficients:

	(Intercept)(S)	value_1(S)	Std(S)
Model 1	0.2648856	0.9988069	1.368896
Model 2	0.3165829	0.9969839	3.033226

Transition probabilities:

	Regime 1	Regime 2
Regime 1	0.96502377	0.09019822
Regime 2	0.03497623	0.90980178



(a) Which 2

Looking at logret

Is it an AR model ? We look for the AR coefficient, and minimize it with the AIC indicator. The AIC is a ratio between a good calibration and the number of parameters. The AR model implies that residual should be white noise (iid normal law). It's tested with the QQnorm. Results:

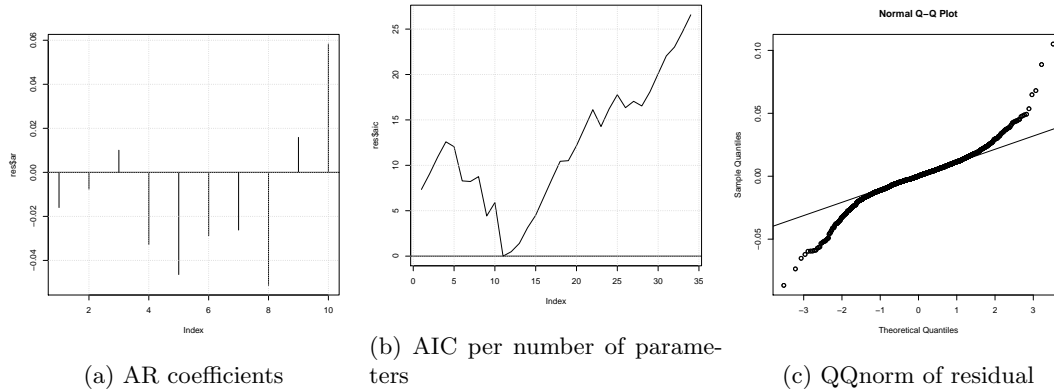
Call:

```
ar(x = value)
```

Coefficients:

1	2	3	4	5	6	7	8
-0.0160	-0.0076	0.0100	-0.0327	-0.0464	-0.0288	-0.0262	-0.0513
9	10						
0.0159	0.0583						

Order selected 10 σ^2 estimated as 0.0002069



Test residual independences with ljung:

Box-Ljung test

```
data: res
X-squared = 0.0038622, df = 1, p-value = 0.9504
```

Test residual normal law with Shapiro:

Shapiro-Wilk normality test

```
data: res
W = 0.94224, p-value < 2.2e-16
```

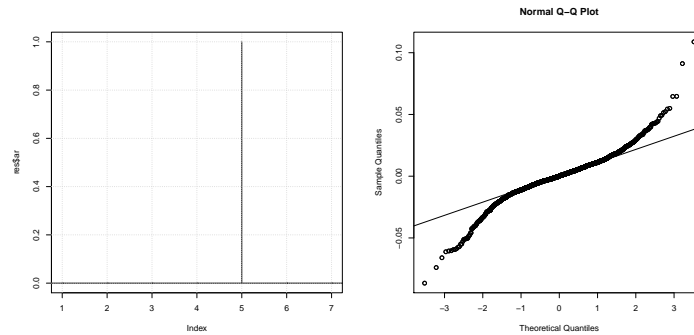
Is it an MA model ? We look for the AR coefficient, and minimize it with the AIC indicator. The AIC is a ratio between a good calibration and the number of parameters. The AR model implies that residual should be white noise (iid normal law). It's tested with the QQnorm.

Results:

```
Series: value
ARIMA(0,0,0) with non-zero mean
```

```
Coefficients:
      intercept
              2e-04
s.e.          3e-04
```

```
sigma^2 estimated as 0.0002084: log likelihood=6545.89
AIC=-13087.77 AICc=-13087.77 BIC=-13076.27
```

(a) MA coefficients

(b) QQnorm of residual

Test residual independences with ljung:

Box-Ljung test

```
data: res
X-squared = 0.28676, df = 1, p-value = 0.5923
```

Test residual normal law with Shapiro:

Shapiro-Wilk normality test

```
data: res
W = 0.94027, p-value < 2.2e-16
```

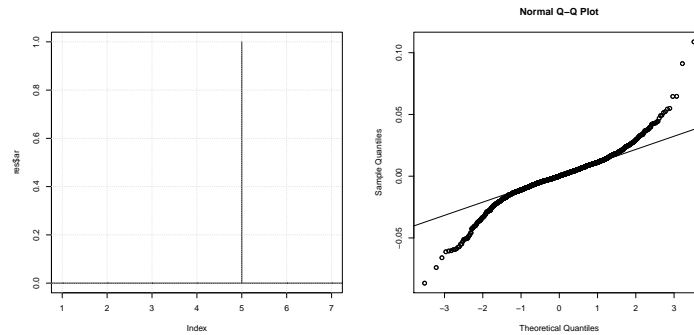
Is it an ARMA model ? We look for the AR coefficient, and minimize it with the AIC indicator. The AIC is a ratio between a good calibration and the number of parameters. The AR model implies that residual should be white noise (iid normal law). It's tested with the QQnorm.

Results:

```
Series: value
ARIMA(0,0,0) with non-zero mean
```

```
Coefficients:
      intercept
              2e-04
s.e.          3e-04
```

```
sigma^2 estimated as 0.0002084: log likelihood=6545.89
AIC=-13087.77 AICc=-13087.77 BIC=-13076.27
```



(a) MA coefficients

(b) QQnorm of residual

Test residual independences with ljung:

Box-Ljung test

```
data: res
X-squared = 0.28676, df = 1, p-value = 0.5923
```

Test residual normal law with Shapiro:

Shapiro-Wilk normality test

```
data: res
W = 0.94027, p-value < 2.2e-16
```

Markov Switching Model Try with 2 regimes.

Results:

Markov Switching Model

```
Call: msmFit(object = mod, k = 2, sw = c(T, T, T), p = 1, control = list(parallel = F))
```

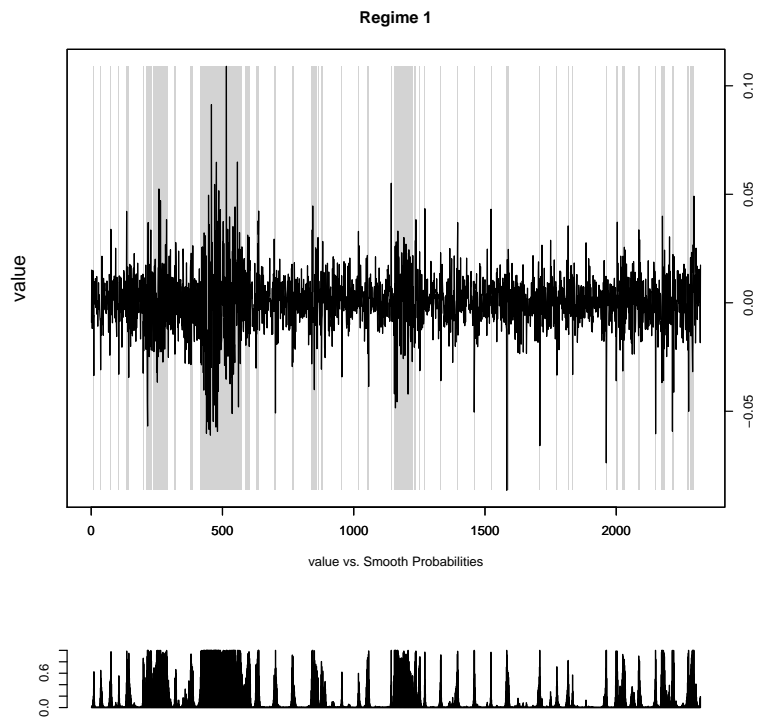
	AIC	BIC	logLik
	-13687.27	-13633.27	6847.635

Coefficients:

	(Intercept)(S)	value_1(S)	Std(S)
Model 1	-0.0007904372	-0.00119046	0.023807260
Model 2	0.0005225907	-0.03309368	0.009267872

Transition probabilities:

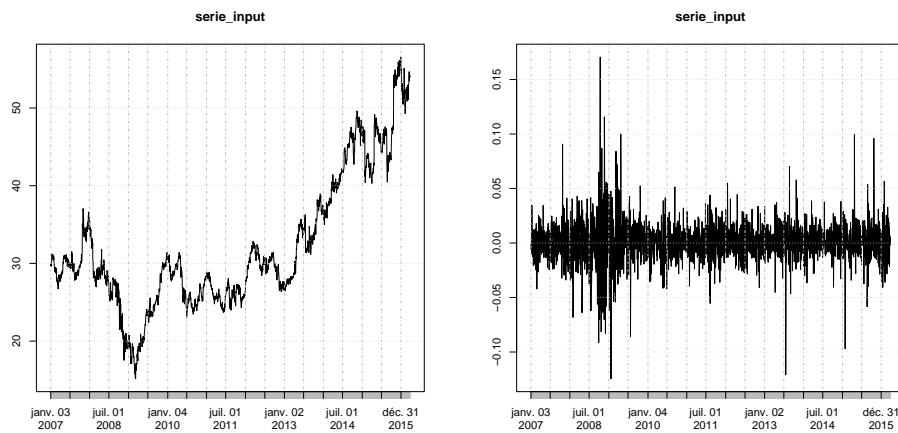
	Regime 1	Regime 2
Regime 1	0.90554306	0.03205306
Regime 2	0.09445694	0.96794694



(a) Which 2

4.4 MSFT

This is the historical close quotation for MSFT. Data has been retrieved from yahoo.



(a) Close level of MSFT

(b) log return MSFT

Looking at original

Is it an AR model ? We look for the AR coefficient, and minimize it with the AIC indicator. The AIC is a ratio between a good calibration and the number of parameters. The AR model implies that residual should be white noise (iid normal law). It's tested with the QQnorm. Results:

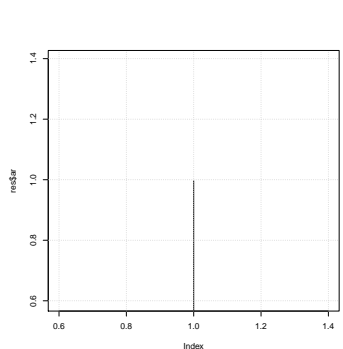
Call:

```
ar(x = value)
```

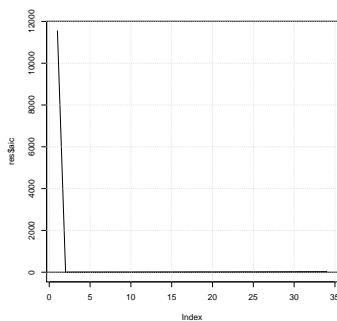
Coefficients:

```
1
0.9965
```

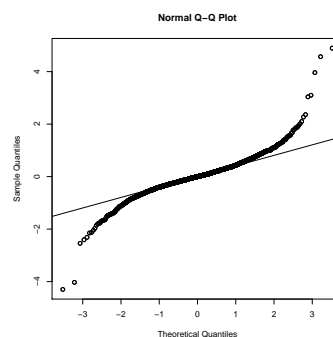
Order selected 1 sigma^2 estimated as 0.5186



(a) AR coefficients



(b) AIC per number of parameters



(c) QQnorm of residual

Test residual independences with ljung:

Box-Ljung test

```
data: res
```

```
X-squared = 0.23417, df = 1, p-value = 0.6284
```

Test residual normal law with Shapiro:

Shapiro-Wilk normality test

```
data: res
```

```
W = 0.91461, p-value < 2.2e-16
```

Is it an MA model ? We look for the AR coefficient, and minimize it with the AIC indicator. The AIC is a ratio between a good calibration and the number of parameters. The AR model implies that residual should be white noise (iid normal law). It's tested with the QQnorm.

Results:

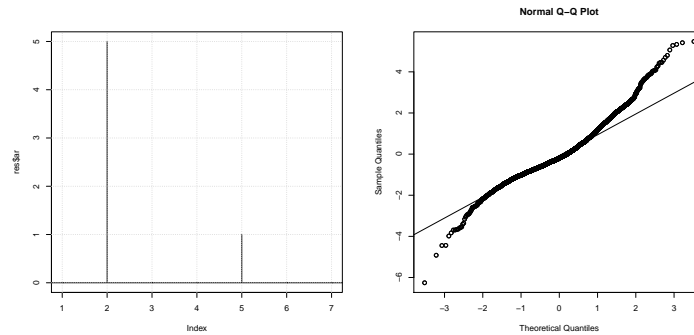
```
Series: value
```

```
ARIMA(0,0,5) with non-zero mean
```

Coefficients:

	ma1	ma2	ma3	ma4	ma5	intercept
	1.8856	2.3315	2.1797	1.4642	0.5639	32.1564
s.e.	0.0188	0.0280	0.0285	0.0251	0.0143	0.2431

sigma² estimated as 1.549: log likelihood=-3807.32
AIC=7628.63 AICc=7628.68 BIC=7668.89



(a) MA coefficients

(b) QQnorm of residual

Test residual independences with ljung:

Box-Ljung test

data: res
X-squared = 210.06, df = 1, p-value < 2.2e-16

Test residual normal law with Shapiro:

Shapiro-Wilk normality test

data: res
W = 0.96066, p-value < 2.2e-16

Is it an ARMA model ? We look for the AR coefficient, and minimize it with the AIC indicator. The AIC is a ratio between a good calibration and the number of parameters. The AR model implies that residual should be white noise (iid normal law). It's tested with the QQnorm.

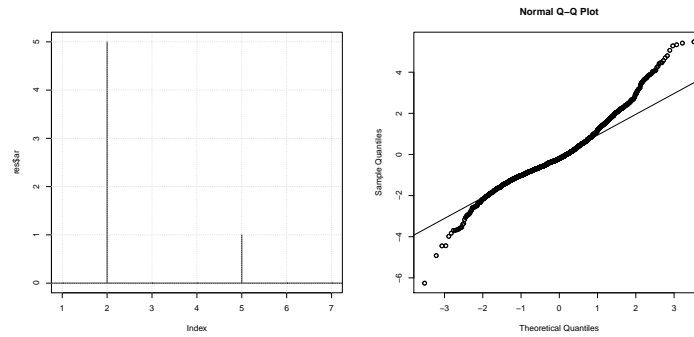
Results:

Series: value
ARIMA(0,0,5) with non-zero mean

Coefficients:

	ma1	ma2	ma3	ma4	ma5	intercept
	1.8856	2.3315	2.1797	1.4642	0.5639	32.1564
s.e.	0.0188	0.0280	0.0285	0.0251	0.0143	0.2431

sigma² estimated as 1.549: log likelihood=-3807.32
AIC=7628.63 AICc=7628.68 BIC=7668.89



(a) MA coefficients

(b) QQnorm of residual

Test residual independences with ljung:

Box-Ljung test

```
data: res
X-squared = 210.06, df = 1, p-value < 2.2e-16
```

Test residual normal law with Shapiro:

Shapiro-Wilk normality test

```
data: res
W = 0.96066, p-value < 2.2e-16
```

Markov Switching Model Try with 2 regimes.

Results:

Markov Switching Model

```
Call: msmFit(object = mod, k = 2, sw = c(T, T, T), p = 1, control = list(parallel = F))
```

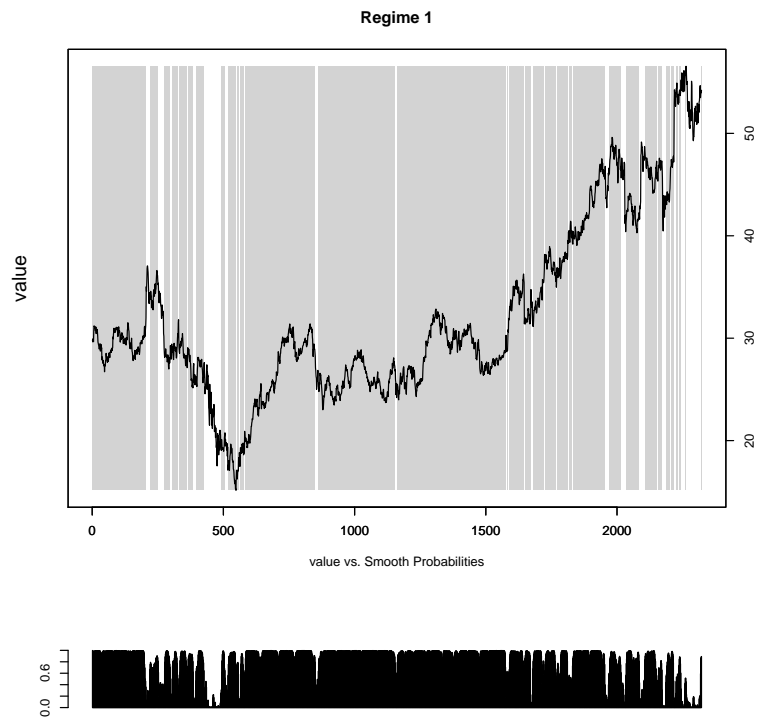
	AIC	BIC	logLik
	3252.634	3306.636	-1622.317

Coefficients:

	(Intercept)(S)	value_1(S)	Std(S)
Model 1	0.03245657	0.9993361	0.3726641
Model 2	0.02607107	0.9994247	0.9869924

Transition probabilities:

	Regime 1	Regime 2
Regime 1	0.96872469	0.1242708
Regime 2	0.03127531	0.8757292



(a) Which 2

Looking at logret

Is it an AR model ? We look for the AR coefficient, and minimize it with the AIC indicator. The AIC is a ratio between a good calibration and the number of parameters. The AR model implies that residual should be white noise (iid normal law). It's tested with the QQnorm. Results:

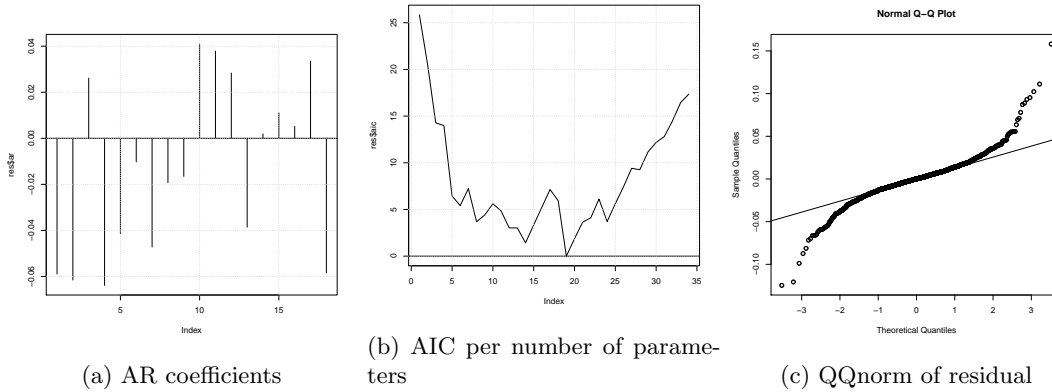
Call:

```
ar(x = value)
```

Coefficients:

1	2	3	4	5	6	7	8
-0.0589	-0.0615	0.0261	-0.0638	-0.0414	-0.0102	-0.0472	-0.0193
9	10	11	12	13	14	15	16
-0.0166	0.0409	0.0378	0.0283	-0.0386	0.0019	0.0110	0.0052
17	18						
0.0335	-0.0584						

Order selected 18 σ^2 estimated as 0.000324



Test residual independences with ljung:

Box-Ljung test

```
data: res
X-squared = 0.00018124, df = 1, p-value = 0.9893
```

Test residual normal law with Shapiro:

Shapiro-Wilk normality test

```
data: res
W = 0.91986, p-value < 2.2e-16
```

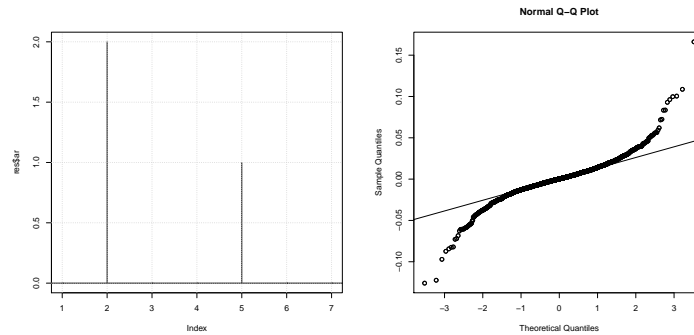
Is it an MA model ? We look for the AR coefficient, and minimize it with the AIC indicator. The AIC is a ratio between a good calibration and the number of parameters. The AR model implies that residual should be white noise (iid normal law). It's tested with the QQnorm.

Results:

```
Series: value
ARIMA(0,0,2) with zero mean
```

```
Coefficients:
          ma1          ma2
      -0.0560  -0.0613
s.e.    0.0209   0.0224
```

```
sigma^2 estimated as 0.0003278: log likelihood=6019.9
AIC=-12033.8  AICc=-12033.79  BIC=-12016.55
```

(a) MA coefficients

(b) QQnorm of residual

Test residual independences with ljung:

Box-Ljung test

```
data: res
X-squared = 0.0083701, df = 1, p-value = 0.9271
```

Test residual normal law with Shapiro:

Shapiro-Wilk normality test

```
data: res
W = 0.91432, p-value < 2.2e-16
```

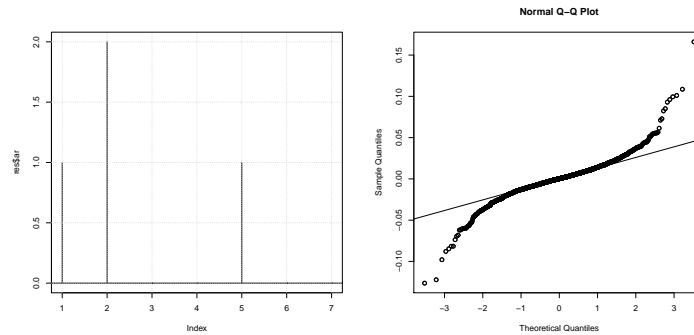
Is it an ARMA model ? We look for the AR coefficient, and minimize it with the AIC indicator. The AIC is a ratio between a good calibration and the number of parameters. The AR model implies that residual should be white noise (iid normal law). It's tested with the QQnorm.

Results:

```
Series: value
ARIMA(0,0,2) with zero mean
```

```
Coefficients:
          ma1      ma2
      -0.0560  -0.0613
s.e.    0.0209   0.0224
```

```
sigma^2 estimated as 0.0003278: log likelihood=6019.9
AIC=-12033.8  AICc=-12033.79  BIC=-12016.55
```



(a) MA coefficients

(b) QQnorm of residual

Test residual independences with ljung:

Box-Ljung test

```
data: res
X-squared = 0.0017872, df = 1, p-value = 0.9663
```

Test residual normal law with Shapiro:

Shapiro-Wilk normality test

```
data: res
W = 0.91427, p-value < 2.2e-16
```

Markov Switching Model Try with 2 regimes.

Results:

Markov Switching Model

```
Call: msmFit(object = mod, k = 2, sw = c(T, T, T), p = 1, control = list(parallel = F))
```

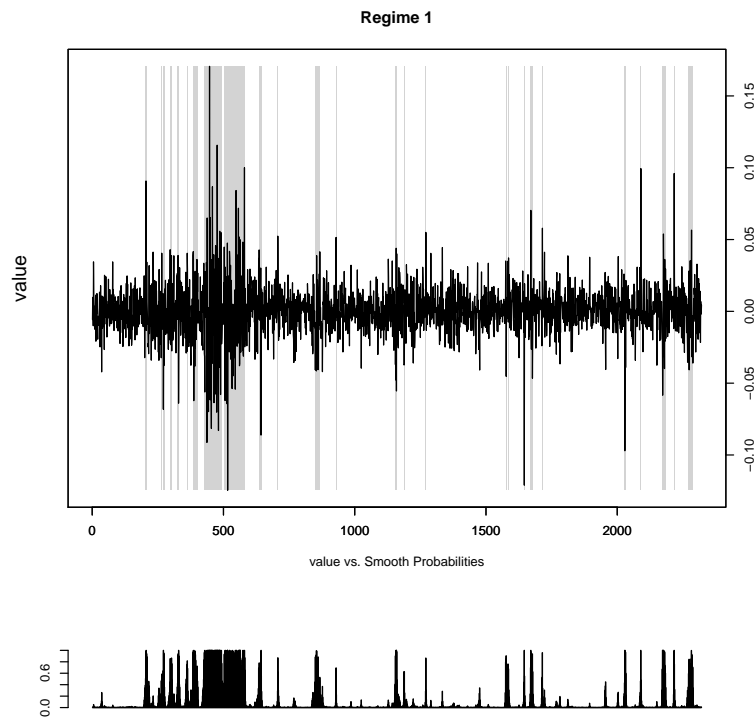
	AIC	BIC	logLik
	-12733.08	-12679.08	6370.541

Coefficients:

	(Intercept)(S)	value_1(S)	Std(S)
Model 1	-0.001015905	-0.09460918	0.03640935
Model 2	0.000462657	-0.01414423	0.01256221

Transition probabilities:

	Regime 1	Regime 2
Regime 1	0.8903983	0.01873014
Regime 2	0.1096017	0.98126986



(a) Which 2

A Notions

Strongly Stationary Be X_t a time series. X_t is strictly stationary process or strongly stationary process if :

$$\forall h, \forall n, X_{t_1} \dots X_{t_n} \text{ has the same law than } X_{t_1+h} \dots X_{t_n+h}$$

Weakly stationary X_t is a weak or wide-sense stationary process if :

- (i) $\forall t, \mathbb{E}\{X_t^2\} < \infty$, finite variance
- (ii) $\forall t, \mathbb{E}\{X_t\} = m$, expectation doesn't depend of time t
- (iii) $\forall t, \text{cov}(X_t, X_{t+h}) = \gamma(h)$, covariance doesn't depend of time t and only on the lag h .

In practice strongly stationary process are hard to demonstrate. In models below we'll demonstrate weak stationarity.

Covariance Be X, Y two random variables. We note $\text{cov}(X, Y)$ the covariance between X and Y

$$\text{cov}(X, Y) = \mathbb{E}\{(X - \mathbb{E}\{X\})(Y - \mathbb{E}\{Y\})\}$$

Correlation Be X, Y two random variables. We note $\rho(X, Y)$ the correlation between X and Y

$$\rho_{X,Y} = \frac{\text{cov}(X, Y)}{\sqrt{\text{Var}(X) \text{Var}(Y)}}$$

Autocorrelation Be X_t a time serie. X_1, X_2, \dots are random variables. We note γ_l the covariance between X_t and X_{t-l}

$$\gamma_l = \frac{\text{cov}(X_t, X_{t-l})}{\sqrt{\text{Var}(X_t) \text{Var}(X_{t-l})}}$$

Note that $\gamma_0 = 1$

Partial autocorrelation (PACF) We note $\pi(k)$ the partial autocorrelation :

$$\pi(k) = \text{corr}(X_t - \mathbb{E}\{X_t | X_{t-1} \dots X_{t-k+1}\}, X_{t-k} - \mathbb{E}\{X_{t-k} | X_{t-1} \dots X_{t-k+1}\}) =$$

Akaike Information Criterion (AIC)

$$AIC = \frac{-2}{T} \ln(\text{likelihood}) + \frac{2}{T} (\text{number of parameters})$$

It misses a real part on AIC, OLS,...

Normal law $\mathcal{N}(\mu, \sigma)$ X follow a normal law $\mathcal{N}(\mu, \sigma)$. The density function is :

$$f(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

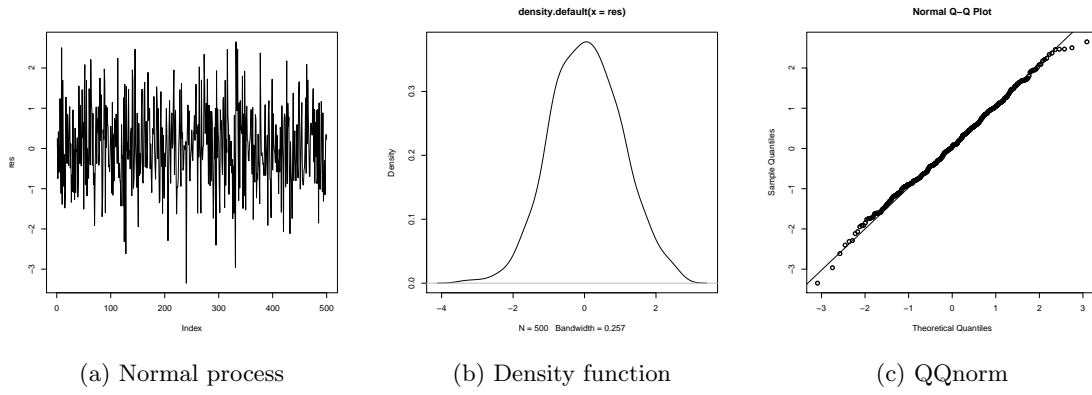


Figure 63: Example of $\mathcal{N}(0,1)$

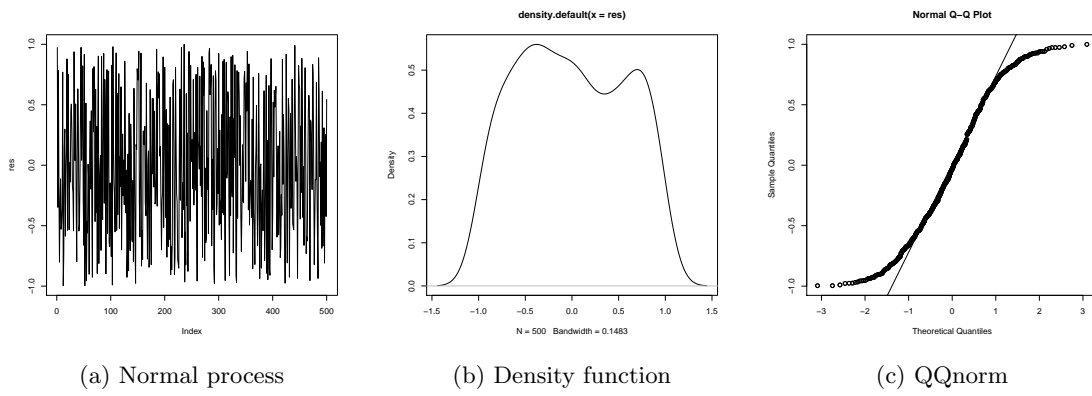


Figure 64: Example of $\mathcal{U}(-1,1)$