Multivariate Subgaussian Stable Distributions in R

by Bruce J. Swihart, John P. Nolan

Abstract We introduce and showcase mvpd (an acronym for multivariate product distributions), a package that uses a product distribution approach to calculating multivariate subgaussian stable distribution functions. The family of multivariate subgaussian stable distributions are eliptically contoured multivariate stable distributions that contain the multivariate Cauchy and the multivariate normal distribution. These distributions can be useful in modeling data and phenomena that have heavier tails than the normal distribution (more frequent occurrence of extreme values). Application areas include log returns for stocks, signal processing for radar and sonar data, astronomy, and hunting patterns of sharks.

Introduction

Multivariate subgaussian stable distributions are the elliptically contoured subclass of general multivariate stable distributions. To begin a brief introduction to multivariate subgaussian stable distributions we start with univariate stable distributions which may be more readily familiar and accessible. Univariate stable distributions are a flexible family and have four parameters, α , β , γ , and δ , and at least eleven parameterizations (!) which has led to much confusion (Nolan, 2020). Here we focus on the 1-parameterization of the Nolan style. Location is controlled by δ , scale by $\gamma \in (0, \infty)$, while $\alpha \in (0, 2]$ and $\beta \in [-1, 1]$ can be considered shape parameters. Being a location-scale family, a "standard" stable distribution will be when $\gamma = 1$ and $\delta = 0$. A solid introduction to univariate stable distributions can be found in the recent textbook *Univariate Stable Distributions* (Nolan, 2020) and its freely available Chapter 1 online (https://edspace.american.edu/jpnolan/stable/).

Univariate symmetric stable distributions are achieved by setting the skew parameter $\beta=0$, which gives symmetric distributions that are bell-shaped like the normal distribution. A way to remember that these are called subgaussian is to see that as $\alpha\in(0,2]$ increases from 0 it looks more and more normal until it is normal for $\alpha=2$ (Figure 1). The sub in subgaussian refers to the tail behavior in that the rate of decrease in the tails is less than that of a gaussian – note how the tails are on top of the gaussian for $\alpha<2$ in Figure 1. Equivalently, as α decreases, the tails get heavier. A notable value of α for subgaussian distributions is $\alpha=1$ which is the Cauchy distribution. The Cauchy and Gaussian distribution are most well-known perhaps because they have closed-form densities, which all other univariate symmetric stable distributions lack.

Therefore, numerically computing the densities is especially important for application. For univariate stable distributions, there is open-source software to compute modes, densities, distributions, quantiles and random variates, including a number of R packages (stabledist, stable, libstableR, for example – see CRAN Task View: Probability *Distributions* for more).

As generalizations of the univariate stable, multivariate stable distributions are a very broad family encompassing many complicated distributions. A subclass of this family is the multivariate subgaussian stable distributions. Multivariate subgaussian stable distributions are symmetric and elliptically contoured. Similar to the aforementioned univariate symmetric stable distributions, the value $\alpha=2$ is the multivariate gaussian and $\alpha=1$ is the multivariate Cauchy. Being that they are elliptically contoured and symmetric makes them applicable to finance where joint returns have an (approximately) elliptical joint distribution (Nolan, 2020). Signal processing, such as with radar and sonar data, tasks itself with filtering impulsive noise from a signal of interest and linear filters in the presence of extreme values tend to underperform, whereas using multivariate stable distributions have been fruitful Tsakalides and Nikias (1998); Nolan (2013). The (multivariate) Holtsmark distribution is a multivariate subgaussian stable distribution ($\alpha=1.5$) that has applications in astronomy, astrophysics, and plasma physics. Lévy flights, which are random walks with steps having a specific type of multivariate subgaussian stable distribution, are used to model interstellar turbulance as well as hunting patterns of sharks (Boldyrev and Gwinn, 2003; Sims et al., 2008).

For multivariate subgaussian stable distributions, the parameter α is a scalar as in the univariate family, while δ (location) becomes a d-dimensional vector and the analog for the scale parameter is a $d \times d$ shape matrix Q. The shape matrix Q needs to be semi-positive definite and is neither a covariance matrix nor covariation matrix. An introduction to multivariate subgaussian stable distributions can be found in Nolan (2013).

Including mvpd, the focus of this paper, if one wanted R functions to interact with multivariate subgaussian stable distributions they have three R package options. These packages are compared in

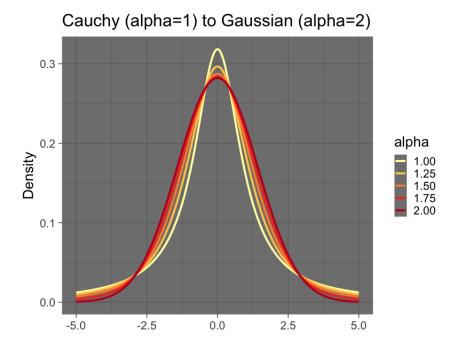


Figure 1: Standard univariate subgaussian distributions for various values of alpha. Cauchy is alpha=1 and normal is alpha=2.

Functionality	alphastable	stable	mvpd
random variates	✓	√	√
parameter estimation	✓	\checkmark	\checkmark
density	✓	\checkmark	\checkmark
cumulative distribution (monte carlo)	x	\checkmark	\checkmark
cumulative distribution (integrated)	x	X	\checkmark
multivariate subgaussian stable	✓	\checkmark	\checkmark
multivariate independent stable	x	\checkmark	X
multivariate isotropic stable	x	\checkmark	X
multivariate discrete-spectral-measure stable	x	\checkmark	X

Table 1: Summary of functionality among R packages. Note: The stable package referenced is not the one on CRAN.

Table 1 and detailed below:

- alphastable provides random univariate and mutlivariate generation, density calculation, and parameter fitting (albeit for modest sample sizes) via an EM alogrithm method (Teimouri et al., 2018, 2019).
- **stable** provides support for all stable univariate distributions and multivariate subgaussian stable distributions. Other cases are handled, to varying degrees, such as isotropic, independent, and spectral measure. For the purposes of this paper, we will note that the **stable** package provides random variate generation, density calculation, parameter fitting, distribution calculations via Monte Carlo methods for multivariate subgaussian stable distributions $2 \le d \le 100$. The **stable** package is developed by Robust Analysis and is available for purchase or through a software grant at http://www.robustanalysis.com/. It is distinct from the univariate **stable** package on CRAN authored by Jim Lindsey.
- mvpd provides random variate generation, density calculation, parameter fitting, distribution
 function calculations via Monte Carlo methods, as well as an integrated method for distribution
 calculations that allows tolerance specification.

While the lack of a tractable density and distribution function impedes directly calculating multivariate subgaussian stable distributions, it is possible to represent them in terms of a product distribution for which each of the two distributions in the product has known numerical methods developed and deployed (in R packages on CRAN). This paper utilizes this approach. The next section covers some product distribution theory.

Product distribution theory

This section reviews the known results of *product distributions* and describes our notation. Allow univariate random variable A with density $f_A(x)$ and d-dimensional random variable G to have density $f_G(x)$ and distribution function $F_G(v, w) = P(v < x \le w)$. Consider the d-dimensional product $H = A^{1/2}G$. From standard product distribution theory we know the density f_H is represented by 1-dimensional integral:

$$f_H(\mathbf{x}) = \int_0^\infty f_B(u) f_G(\mathbf{x}/u) \frac{1}{|u|^d} du \tag{1}$$

where $B = A^{1/2}$ so that $f_B(x) := 2xf_A(x^2)$. Consequently the distribution function F_H (with lower bound v and upper bound w) of the r.v. H is represented by

$$F_{H}(v,w) = \int_{0}^{\infty} f_{B}(u) \int_{v_{1}}^{w_{1}} \cdots \int_{v_{d}}^{w_{d}} f_{G}(t/u) \frac{1}{|u|^{d}} dt_{1} \dots dt_{d} du$$

$$= \int_{0}^{\infty} f_{B}(u) \int_{v_{1}/u}^{w_{1}/u} \cdots \int_{v_{d}/u}^{w_{d}/u} f_{G}(t) dt_{1} \dots dt_{d} du$$

$$= \int_{0}^{\infty} f_{B}(u) F_{G}(v/u, w/u) du$$
(3)

Take note of the representation in (1) and (3). From a practical standpoint, if we have a (numerical) way of calculating f_A , f_G , and F_G we can calculate f_H and F_H . Different choices can be made for f_A , f_G , and F_G in this general setup. The choices required for multivariate subgaussian stable distributions are covered in the following Implementation section.

Multivariate elliptically contoured stable distributions

From Nolan (2013), H is a d-dimensional subgaussian stable distribution if A is a univariate stable distribution

$$A \sim S\left(\frac{\alpha}{2}, 1, 2\cos\left(\frac{\pi\alpha}{4}\right)^{\left(\frac{2}{\alpha}\right)}, 0; 1\right)$$

and G is a d-dimensional multivariate normal $G \sim MVN(0,Q)$, where the shape matrix Q is positive semi-definite. This corresponds to Example 17 in Hamdan (2000).

Implementation

Using the aforementioned theory of product distributions, we can arrive at functions for a multivariate subgaussian stable density and distribution function thanks to established functions for univariate stable and multivariate normal distributions. A key package in the implementation of multivariate subgaussians in R is **mvtnorm** (Genz et al., 2020; Genz and Bretz, 2009). In the basic product-distribution approach of **mvpd**, f_G and F_G are mvtnorm::dmvnorm and mvtnorm::pmvnorm respectively. Allow the density of A, f_A (to be numerically calculated in R) using stable::dstable or libstableR::stable_pdf (del Val et al., 2017). Presented as pseudo-code:

- $f_A(x,\alpha) := \text{libstableR}:: \text{stable_pdf}(x, \text{pars} = \left(\frac{\alpha}{2}, 1, 2\cos(\frac{\pi\alpha}{4})^{\left(\frac{2}{\alpha}\right)}, 0\right); \text{pm} = 1)$
- $f_B(x,\alpha) := 2xf_A(x^2,\alpha)$
- $f_H(x,\alpha,Q) = \int_0^\infty f_B(u;\alpha) \times \text{mvtnorm}:: \text{dmvnorm}(x/u, \text{sigma} = Q) \frac{1}{|u|^d} du$
- $F_H(v, w, \alpha, Q) = \int_0^\infty f_B(u; \alpha) \times \text{mvtnorm}: \text{pmvnorm}(\text{lower} = v/u, \text{upper} = w/u, \text{sigma} = Q) du$

The (outermost) univariate integral is numerically evaluated with stats::integrate.

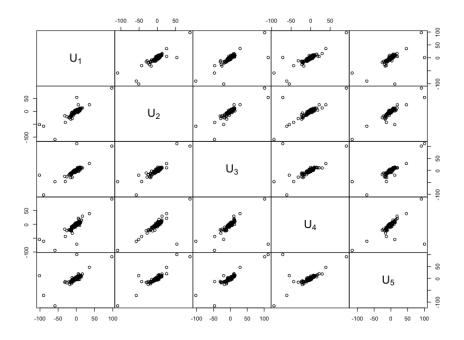


Figure 2: A lattice of scatterplots of the 5-dimensional subgaussian stable distribution.

Quick start

We present an outline of the <code>[dpr]mvss</code> (multivariate subgaussian stable) functions, and walk through the code in the subsequent sections. As an overview, we generate 5000 5-dimensional subgaussian variates with $\alpha=1.7$ and an "exchangeable" shape matrix using rmvss. We then recover the parameters with an illustrative call to <code>fit_mvss</code>. We can calculate the density (dmvss) at the center of the distribution and get a quick estimate of the distribution between -2 and 2 for each member of the 5-dimensional variate using <code>pmvss_mc</code>. We investigate a refinement of that quick distribution estimate using <code>pmvss_mc</code>.

Random variates generation with rmvss

We'll generate 5000 5-dimensional subgaussian random variates with a specified α and shape matrix. They are pictured in Figure 2. In the next section we will fit a distribution to these.

The ability to simulate from a distribution is useful for running simulations to test different scenarios about the phenomena being modeled by the distribution, as well as in this case, to generate a dataset with a known shape matrix and alpha to show our fitting software (next section) can recover these parameters. Our quick start code begins with generating a dataset from a known alpha and shape matrix. However, often a practicioner might start with a dataset from which parameters are

estimated and then random samples can be generated from the distribution specified with those parameters to learn more about the data generating distribution and the behavior of the phenomena.

Fitting a multivariate subgaussian distribution with fit_mvss

If you have data in a $n \times d$ matrix \mathbf{X} and want to fit a d-dimensional multivariate subgaussian distribution to those data, then fit_mvss will return estimates of the parameters using the method outlined in Nolan (2013). The method involves fitting univariate stable distributions for each column and assessing the resulting α , β and δ parameters. The column-wise α estimates should be similar and the column-wise β estimates close to 0. This column-wise univariate fitting is carried out by libstableR::stable_fit_mle2d(W,parametrization = 1L) and the off diagonal elements can be found due to the properties of univariate stable distributions (see Nolan (2013)). For your convenience, the univariate column-wise estimates of α , β , γ and δ are returned in addition to the raw estimate of the shape matrix and the nearest positive definite shape matrix (as computed by Matrix::nearPD applied to the raw estimate).

```
## take X from previous section and estimate
## parameters for the data generating distribution
R> fitmv <- mvpd::fit_mvss(X)</pre>
R> fitmv
Suniv alphas
[1] 1.698617 1.708810 1.701662 1.696447 1.699372
[1] -0.02864287 -0.04217262 -0.08444540 -0.06569907 -0.03228573
$univ_gammas
[1] 1.016724 1.000151 1.008055 1.012017 1.002993
$univ deltas
[1] -0.03150732 -0.06525291 -0.06528644 -0.07730645 -0.04539796
$mult_alpha
[1] 1.700981
$mult_Q_raw
          [,1]
                    [,2]
                              [,3]
                                         [,4]
                                                   Γ.57
[1,] 1.0337276 0.9034599 0.8909654 0.8937814 0.8647089
[2,] 0.9034599 1.0003026 0.9394846 0.9072368 0.8535091
[3,] 0.8909654 0.9394846 1.0161748 0.8929937 0.9037467
[4,] 0.8937814 0.9072368 0.8929937 1.0241777 0.9281714
[5,] 0.8647089 0.8535091 0.9037467 0.9281714 1.0059955
$mult_Q_posdef
          [,1]
                    [,2]
                              [,3]
                                         [,4]
[1,] 1.0337276 0.9034599 0.8909654 0.8937814 0.8647089
[2,] 0.9034599 1.0003026 0.9394846 0.9072368 0.8535091
[3,] 0.8909654 0.9394846 1.0161748 0.8929937 0.9037467
[4,] 0.8937814 0.9072368 0.8929937 1.0241777 0.9281714
[5,] 0.8647089 0.8535091 0.9037467 0.9281714 1.0059955
```

An alternative for fitting this distribution is alphastable::mfitstab.elliptical(X,1.70,shape_matrix,rep(0,5)) and takes 8 minutes (and requires initial values for alpha, the shape matrix, and delta) while fit_mvss(X) took under 2 seconds. For a run of n=1e6,d=20, fit_mvss scales well, taking 60 minutes.

Once the distribution has been fitted, fitmv\$mult_alpha, fitmv\$mult_Q_posdef, and fitmv\$univ_deltas, can be used as the alpha, Q, and delta arguments, respectively, in calls to dmvss to calculate densities and pmvss_mc or pmvss to calculate probabilities. They could also be passed to rmvss to generate random variates for simulations.

Density calculations with dmvss

We can calculate the density at the center of the distribution.

Distribution calculation by Monte Carlo with pmvss_mc

The method of calculating the distribution by Monte Carlo relies on the ability to produce random variates quickly and then calculate what proportion of them them fall within the specified bounds. To generate multivariate subgaussian stable variates, a scalar A is drawn from

libstableR::stable_rnd(
$$n$$
,pars = $\left(\frac{\alpha}{2},1,2\cos\left\{\frac{\pi\alpha}{4}\right\}^{\left(\frac{2}{\alpha}\right)},0\right)$;pm = 1)

and then the square-root of A multiplied by a draw G from

```
mvtnorm: rmvnorm(n, sigma = Q).
```

This allows for quick calculations but to increase precision requires generating larger number of random variates. For instance, if we wanted the distribution between -2 and 2 for each dimension, we could generate 10,000 random variates and then see how many of them fall between the bounds. It looks like 6,820 variates were within the bounds:

We run it again and the answer changes:

With the Monte Carlo method, precision is not specified and no error is calculated. The next section introduces how to use the integrated distribution function F_H from product theory and specify precision.

Distribution function calculation via integration with pmvss

There are three inexact entities involved in the distribution calculation F_H as found in pmvss: the numerically calculated F_G , the numerically calculated F_A , and the outer numerical integration.

The outer integral by integrate assumes the integrand is calculated without error, but this is not the case. See the supplementary materials section "Thoughts on error propagation in pmvss" for justification and guidance for specifying the values of abs.tol.si, abseps.pmvnorm, and maxpts.pmvnorm. The first of these three arguments is passed to the abs.tol argument of stats::integrate and controls the absolute tolerance of the numerically evaluated outer 1-dimensional integral. The remaining two are passed to maxpts and abseps of mvtnorm::GenzBretz and control the accuracy of mvtnorm::pmvnorm.

Briefly, our experience suggests that to be able to treat abs.tol.si as the error of the result, abseps.pmvnorm should be 1e-2 times smaller than the specified abs.tol.si which may require a

multiple of the default 25000 default of maxpts.pmvnorm — which will lead to more computational intensity and longer computation times as demonstrated below (as conducted on Macbook Intel Core i7 chip with 2.2 GHz):

##	abs.tol.si	abseps.pmnvorm	maxpts	Tin	ne
##	1e-01	1e-03	25000		
##	1e-02	1e-04	25000*10	3	sec
##	1e-03	1e-05	25000*100	22	sec
##	1e-04	1e-06	25000*1000	4	\min
##	1e-05	1e-07	25000*10000	26	min
##	1e-06	1e-08	25000*85000	258	min

With this in mind, the output from the Quick Start code is:

Both pmvss and pmvss_mc take infinite limits. Since pmvss_mc calculates how many random variates H_i , $i \in \{1, ..., n\}$ are within the bounds, pmvss might be preferred to pmvss_mc when calculating the tails of the distribution, unless n is made massively large.

Speed and Accuracy Trials

We provide a sense of accuracy and computational time tradeoffs with a modest simulation experiment (Figure 3, see supplementary materials for code). Estimating these distributions is inherently difficult – difficult in the sense that expecting accuracy farther out than the 5th decimal place for distribution functions is unreasonable. Therefore, we will define our "gold standard" targets for accuracy evaluation as the numerical density produced by Robust Analysis' dstable integrated by cubature::hcubature() with tolerance tol=1e-5.

We will time three functions using bench::mark() in different scenarios. The first function is Robust Analysis' pmvstable.MC() (abbreviated as RAMC, below) and the other two are mvpd::pmvss_mc() (abbreviated as PDMC, below) and mvpd::pmvss() (abbreviated as PD, below). Fixing $\alpha=1.7$ and dimension d=4, the different test scenarios will involve a low level of pairwise association vs. a high level in a shape matrix of the form:

•
$$Q_{exch} = \begin{bmatrix} 1 & \rho & \rho & \rho \\ \rho & 1 & \rho & \rho \\ \rho & \rho & 1 & \rho \\ \rho & \rho & \rho & 1 \end{bmatrix}$$
 for $\rho = 0.1$ and $\rho = 0.9$.

We calculate the distributions in the hypercube bounded by (-2,2) in all four dimensions. The gold standard for the $\rho=0.1$ case was 0.5148227 and 0.7075104 for the $\rho=0.9$ case. The numerical integration of the former took 3 minutes whereas the latter took 1 hour – which portends that higher associations involve more computational difficulty. We back-calculated the number of samples needed to give the methods involving Monte Carlo (RAMC and PDMC) a 95% CI width that would fall within 0.001 and 0.0001 of the gold standard, and display the scatter plots of estimate and computational time in (Figure 3).

From Figure 3, some high-level conclusions can be drawn: higher pairwise associations require more computational resources and time, increasing the precision requires more computational resources and time, and sometimes the Monte Carlo methods are faster than PD, sometimes not. PD seems to be quite precise and possibly underestimating the gold standard.

Of course, we cannot test every possible instance of alpha and shape matrices for all *d* dimensions, integration limits, and specified precision. In our experience, the computational intensity is an interplay between alpha, the integration limits, the shape matrix structure, delta, and the requested precision. We provide the code that we used for our simulation study and encourage the readers who need to explore these issues for their particular integral to edit the code accordingly.

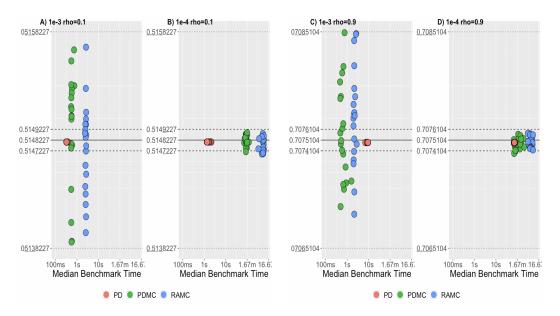


Figure 3: Speed and accuracy of distribution calculations. Consider a multivariate subgaussian stable distribution of d=4, $\alpha=1.7$, with limits of integration being (-2,2) for each dimension. In panels A) and B) we have an exchangeable shape matrix with $\rho=0.1$, and specified precision of 1e-3 and 1e-4 for pmvss (PD), respectively. Analogously, panels C) and D) display results for an exchangeable shape matrix with $\rho=0.9$. Concurrently in each panel, are the results for Robust Analysis' pmvstable.MC (RAMC) and pmvss_mc (PDMC) with enough simulated variates to produce a 95 CI width that matches the precision. Each point is an independent call and the calculated distribution is on the Y-axis vs the median benchmark time on the X-axis. There are 20 calls per function per scenario. The dotted line is the 1e-3 boundary of the gold standard and the dashed line is the 1e-4 boundary.

Bonus: faster distribution calculations via a modified QRVSN algorithm

Insight: multivariate student's t distributions are a product distribution

The derivation of the univariate student's t distribution is commonly motivated with a ratio of two quantitites each involving random variables: a standard normal Z in the numerator and a $V \sim \chi^2(\nu)$ in the denominator:

$$T_{\nu} = \frac{Z}{\sqrt{V/\nu}} = Z\sqrt{\frac{\nu}{V}}$$

but what often is left out of the instruction is that $A_{V} = \frac{\nu}{V} \sim IG\left(\frac{\nu}{2}, \frac{\nu}{2}\right)$ has an inverse-gamma distribution, where $X \sim IG\left(r,s\right)$ with rate r, shape s, and density $f(x;r,s) = \frac{r^s}{\Gamma(s)}x^{(-s-1)}e^{-r/x}$. This implies we equivalently have a product distribution of the type $T_{V} = A_{V}^{1/2}Z$. This notion holds for the multivariate case as well, where for $G \sim MVN(0,Q)$ as before and A_{V} is an inverse-gamma with $r=s=\nu/2$ then $H_{V}=A_{V}^{1/2}G$ is a d-dimensional student's t distribution with ν degrees of freedom and covariance matrix Q. This corresponds to Example 16 in in Hamdan (2000), and is equivalent to the 'chi-normal' (χ - Φ) formulation in Genz and Bretz (2002) (earning the namesake ' χ ' due to the fact that $V \sim \chi^2(\nu) \implies \sqrt{V} \sim \chi(\nu)$). For $d \geq 4$, the QRVSN algorithm is used in mytnorm::pmvt. The reordering and rotational methodology that makes pmvtnorm::pmvt so fast is independent of the part that generates $\sqrt{1/A_{V}}$ random variates. This means that if one replaced $\sqrt{1/A_{V}}$ with $\sqrt{1/A}$ variates, mytnorm::pmvt would produce not multivariate student's t distributions but multivariate subgaussian stable distributions. We implement a modified QRSVN algorithm for multivariate subgaussian stable distributions in a separate package, mvgb in honor of Genz and Bretz.

Implementation of mvgb::pmvss

Generating random variates of *A* requires two independent uniform random variates, and only one of which is Quasi-Random in our implmentation. Regardless, this modified QRSVN approach enables the potential advantage of the rotation of the distribution and the reordering of integration limits. The takeaway is, that for similar precisions, mvgb::pmvss may be much faster than mvpd::pmvss, such as 10 seconds vs 500 seconds for 4 digits of precision in the following example:

```
R> set.seed(321)
R> library(mvgb)
R> tictoc::tic()
## probability calculated by mvgb takes about 10 seconds
R> gb_4digits <-
   mvgb::pmvss(lower = rep(-2,5),
                upper = rep(2,5),
                alpha = fitmv$mult_alpha,
                Q = fitmv$mult_Q_posdef,
                delta = fitmv$univ_deltas,
                abseps = 1e-4,
                maxpts = 25000*350)
R> tictoc::toc()
9.508 sec elapsed
> gb_4digits
[1] 0.6768
## now calculate same probability with similar precision
## in mvpd
R> tictoc::tic()
## probability calculated by mvpd takes about 10 MINUTES
R> pd_4digits <-
    mvpd::pmvss(lower = rep(-2,5),
                upper = rep(2,5),
                alpha = fitmv$mult_alpha,
                Q = fitmv$mult_Q_posdef,
                delta = fitmv$univ_deltas,
                abseps.pmvnorm = 1e-6,
                maxpts.pmvnorm = 25000*1000.
                abs.tol.si = 1e-4)
R> tictoc::toc()
518.84 sec elapsed
R> pd_4digits[1]
[1] 0.6768
```

Although currently on CRAN, we include mvgb::pmvss here as a proof-of-concept and as an area of future work. More research is needed into its computational features and accuracy, and this is only encouraged by promising preliminary results. Additionally, more research may be warranted for other R package methodologies that use a multivariate Gaussian, Cauchy, or Holtsmark distribution to generalize to a multivariate subgaussian stable distribution (a helpful reviewer suggested generalizing the multivariate distributions as used in fHMM (Oelschläger and Adam, 2021) and generalizing the normally mixed probit model in RprobitB). For more about elliptically contoured multivariate distributions in general, consult Fang and Anderson (1990); Fang et al. (2018).

Acknowledgements

This work utilized the computational resources of the NIH HPC Biowulf cluster (http://hpc.nih.gov). We thank Robust Analysis for providing their **stable** R package via a software grant.

Bibliography

- S. Boldyrev and C. R. Gwinn. Lévy model for interstellar scintillations. *Physical review letters*, 91(13): 131101, 2003. [p1]
- J. R. del Val, F. Simmross-Wattenberg, and C. Alberola-López. libstable: Fast, parallel, and high-precision computation of *α*-stable distributions in R, C/C++, and MATLAB. *Journal of Statistical Software*, 78(1):1–25, 2017. doi: 10.18637/jss.v078.i01. [p3]
- K.-T. Fang and T. W. Anderson. *Statistical inference in elliptically contoured and related distributions*. Allerton Press, 1990. [p9]
- K.-T. Fang, S. Kotz, and K. W. Ng. *Symmetric multivariate and related distributions*. Chapman and Hall/CRC, 2018. [p9]

- A. Genz and F. Bretz. Comparison of methods for the computation of multivariate t probabilities. *Journal of Computational and Graphical Statistics*, 11(4):950–971, 2002. [p8]
- A. Genz and F. Bretz. *Computation of Multivariate Normal and t Probabilities*. Lecture Notes in Statistics. Springer-Verlag, Heidelberg, 2009. ISBN 978-3-642-01688-2. [p3]
- A. Genz, F. Bretz, T. Miwa, X. Mi, F. Leisch, F. Scheipl, and T. Hothorn. *mvtnorm: Multivariate Normal and t Distributions*. The organization, 2020. URL https://CRAN.R-project.org/package=mvtnorm. R package version 1.1-0. [p3]
- H. N. Hamdan. *PhD Thesis: Characterizing and approximating scale mixtures*. American University, 2000. [p3, 8]
- J. P. Nolan. Multivariate elliptically contoured stable distributions: theory and estimation. *Computational Statistics*, 28(5):2067–2089, 2013. [p1, 3, 5]
- J. P. Nolan. *Univariate stable distributions*. Springer, 2020. [p1]
- L. Oelschläger and T. Adam. Detecting bearish and bullish markets in financial time series using hierarchical hidden markov models. *Statistical Modelling*, page 1471082X211034048, 2021. [p9]
- D. W. Sims, E. J. Southall, N. E. Humphries, G. C. Hays, C. J. Bradshaw, J. W. Pitchford, A. James, M. Z. Ahmed, A. S. Brierley, M. A. Hindell, et al. Scaling laws of marine predator search behaviour. *Nature*, 451(7182):1098–1102, 2008. [p1]
- M. Teimouri, S. Rezakhah, and A. Mohammadpour. Parameter estimation using the em algorithm for symmetric stable random variables and sub-gaussian random vectors. *Journal of Statistical Theory and Applications*, 17(3):439–461, 2018. [p2]
- M. Teimouri, A. Mohammadpour, and S. Nadarajah. *alphastable: Inference for Stable Distribution*, 2019. URL https://CRAN.R-project.org/package=alphastable. R package version 0.2.1. [p2]
- P. Tsakalides and C. L. Nikias. *Deviation from Normality in Statistical Signal Processing: Parameter Estimation.* Springer Science & Business Media, 1998. [p1]

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