

T⁶

$$p(x) = \begin{cases} \frac{\theta-1}{x^\theta}, & x \geq 1 \\ 0, & x < 1 \end{cases} \quad \theta > 1$$

$\bar{x}_n = 6610 \text{ баков}$

a) ОМД

$$L(\theta) = \prod p(x_i, \theta) = \left(\frac{\theta-1}{x_i^\theta} \right)^n \cdot \prod_{i=1}^n x_i^{-\theta} \cdot \mathbb{1}_{x_i \geq 1}$$

$$\ln L = n \ln(\theta-1) - \theta \left(\sum_{i=1}^n \ln x_i \right)$$

$$\frac{d(\ln L)}{d\theta} = \frac{n}{\theta-1} - \sum \ln x_i = 0$$

$$\Rightarrow \theta \approx 1 + \frac{n}{\sum_{i=1}^n \ln x_i}$$

$$(\ln L)'_0 = \frac{-n}{(\theta-1)^2} < 0 \quad \text{б т. экспр}$$

\Rightarrow max.

$$\hat{\theta} \approx 1 + \frac{1}{\ln \bar{x}}$$

b) Проверим условие на однозначность

1) Р - вып. функция

$$2) \int_{-\infty}^{\infty} \left(\frac{2}{\theta} + p \right) dx \rightsquigarrow ?$$

$$p'_{\theta} = \frac{\ln x - \theta \ln x + 1}{x^\theta}$$

$$\int_{-\infty}^{\infty} x^{-\theta} (\ln x / (1-\theta) + 1) dx = \frac{1}{\theta-1} - \int_{-\infty}^{\infty} x^{\theta-1} dx =$$

$$= \frac{1}{\theta-1} - \frac{1}{\theta-1} = 0$$

$$p''_{\theta} = \ln x \left| \frac{\theta \ln x - \ln x - 2}{x^\theta} \right. - \text{hemp. f. f.}$$

$$\int_1^{+\infty} p'' dx = \int_1^{+\infty} \left[(\theta-1)x^{-\theta} \ln^2 x - 2x^{\theta-1} \ln x \right] dx =$$

$$= 0$$

$$3) I(\theta) = \int_1^{+\infty} \left(\frac{1}{(\theta-1)^2} + (-2) \frac{\ln x}{\theta-1} + \ln^2 x \right) \frac{\theta-1}{x^\theta} dx$$

$$= \int_1^{+\infty} \left(\frac{x^{-\theta}}{\theta-1} - 2 \frac{\ln x}{x^\theta} + (\theta-1) \left(\frac{\ln^2 x}{x^\theta} \right) \right) dx =$$

$$= \frac{1}{(\theta-1)^2} - 2 \left(\ln \frac{x^{1-\theta}}{1-\theta} \right) \Big|_1^{+\infty} - \int_1^{+\infty} \frac{x^{-\theta}}{1-\theta} dx +$$

$$+ \int_1^{+\infty} \frac{(\theta-1)\ln^2 x}{x^\theta} dx = \frac{-1}{(\theta-1)^2} - \left(\ln x \cdot x^{1-\theta} \right) \Big|_1^{+\infty}$$

$$= 2 \int_0^\infty \frac{1}{x} x^{1-\theta} e^{-\frac{x}{1-\theta}} dx = \frac{1}{(1-\theta)^2} \times 2 \Gamma\left(\frac{1-\theta}{1-\theta}\right) =$$

$$- \int_0^\infty \frac{x^{-\theta}}{1-\theta} dx = \frac{1}{(1-\theta)^2} - \text{wegen } > 0 \text{ da } (1-\theta)$$

\Rightarrow (nicht unbedingt)

\Rightarrow Okt (wir, aber nicht, zumindest, die Verteilung ist anders)

Mittelwerts: $P(X_{\frac{1}{2}}) = \frac{1}{2}$

$$\int_0^{\frac{1}{2}} \frac{\theta-1}{x^\theta} dx = -\frac{1}{\theta^{1-\theta}} + 1 = \frac{1}{2}$$

$$\text{Mittel } x = \frac{1}{2} \Rightarrow X_{\frac{1}{2}} = 2^{\frac{1}{\theta}-1}$$

$$g(\hat{\theta}) = \theta^{-1/2}$$

$$\frac{1}{\theta} \frac{d\theta}{dx} = 2^{\frac{1}{\theta}-1} \cdot \ln 2 \cdot \frac{-1}{(\theta-1)^2}$$

$$\sqrt{n} \frac{g(\hat{\theta}) - g(\theta)}{\sqrt{\operatorname{Var}[g(\hat{\theta})] I^{-1}(\theta)}} \sim N(0, 1)$$

$$\frac{\sqrt{n}}{2^{\theta-1} \cdot \ln 2} \cdot \frac{(g(\hat{\theta}) - g(\theta)) \cdot (\hat{\theta}^{-1})}{I^{-1}(\hat{\theta})} \sim N(0, 1)$$

$$P\left(\hat{\theta} < \bar{g}(\bar{\theta}) - \frac{1,96 \cdot \ln 2 \cdot 2\bar{\theta}}{\sqrt{n}(\bar{\theta}-1)}\right) < \chi_{\frac{\alpha}{2}}^2 < \bar{g}(\bar{\theta}) + \frac{1,96 \cdot \ln 2 \cdot 2\bar{\theta}}{\sqrt{n}(\bar{\theta}-1)} \quad \text{P}$$

c) Due θ

OMT:

$$\frac{\sqrt{n}(\bar{\theta} - \theta)}{\theta^{-1}} \sim N(0, 1)$$

$$P\left(1 - \frac{1,96 \cdot \sqrt{n}}{\sqrt{n} \cdot \ln x} < \theta < 1 + \frac{1,96 \cdot \sqrt{n}}{\sqrt{n} \cdot \ln x}\right) \approx P$$