

Chapter 3. part 4.

(1)

§3.7. Extrema and Order Statistics.

This section is concerned with ordering a collection of independent continuous r.v.s. $X_1, X_2, \dots, X_n \stackrel{iid}{\sim} F$, cdf and f. pdf.

$X_{(1)}$ = smallest of X_1, X_2, \dots, X_n ;

$X_{(2)}$ = second smallest of X_1, X_2, \dots, X_n .

.....

$X_{(j)}$ = j-th smallest of X_1, X_2, \dots, X_n

.....

$X_{(n)}$ = largest of X_1, X_2, \dots, X_n .

r-demonstration

$X_{(1)} \leq X_{(2)} \leq \dots \leq X_{(n)}$, order statistics
 $\min(x), \max(x), \text{median}(x)$
 $\text{mean}(x), \text{median}(x) = \text{quantile}(x, 0.5), \text{summary}(x)$, Q3E quantile(x, 0.75)
 $y = \text{sort}(x), Y_{(1)}, Y_{(2)}, \dots$
 $\text{boxplot}(x)$

Expl. X_1, X_2, \dots, X_5 are lifetimes of 5 light bulbs. Assume $X_1, X_2, \dots, X_5 \stackrel{iid}{\sim} \exp(\lambda)$, $X_{(1)}$ is the shortest lifetime of 5. $X_{(5)}$ is the longest one. $X_{(3)}$ is the middle one.

$X_{(1)}, X_{(3)}$ and $X_{(5)}$ are all random variables, which depend on X_1, \dots, X_5 .

e.g. $X_1=10, X_2=30, X_3=5, X_4=25, X_5=12$.

$$\text{then } X_{(1)} = \min\{X_1, X_2, \dots, X_5\} = 5$$

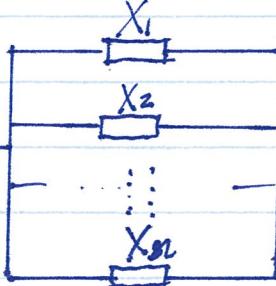
$$X_{(5)} = \max\{X_1, X_2, \dots, X_5\} = 30.$$

$$\text{also } X_{(2)} = 10$$

$$X_{(3)} = \text{third smallest } \{X_1, X_2, \dots, X_5\} = 12.$$

$$X_{(4)} = 25.$$

more examples: A.



X_1, X_2, \dots, X_n lifetime $\sim G(\alpha, \lambda)$.

lifetime of system is

$$X = \max\{X_1, X_2, \dots, X_n\} = X_{(n)}.$$

What is the distribution of $X_{(n)}$? $X_{(n)} \sim f_{X_{(n)}}(t)$?

$$P(X_{(n)} > 300) = ?$$

(2)



X_1, X_2, \dots, X_n n components' lifetimes, $X_i \sim_{\text{iid}} \text{lognormal}(\mu, \sigma^2)$.

What is the lifetime of the above system?

$$X = \min\{X_1, X_2, \dots, X_n\} = X_{(1)} \text{ find } P(X_{(1)} > 300) ?$$

$$X_{(1)} \sim f_{X_{(1)}}(t) ?$$

C. X_1, X_2, \dots, X_n , salaries of n people. a sample for certain group.
what is the median income?

$$\text{if } n=2k, \text{ median} := \frac{1}{2}(X_{(k)} + X_{(k+1)}) =: Y, \text{ find } Y \sim f_Y(y) ?$$

$$n=2k+1, \text{ median} = X_{(k+1)}, \text{ what's distribution?}$$

D. X_1, X_2, \dots, X_n iid SAT scores.

$$\text{range} = \max\{X_1, \dots, X_n\} - \min\{X_1, \dots, X_n\}$$

$$= X_{(n)} - X_{(1)}. \text{ let } R := X_{(n)} - X_{(1)}, \text{ find } R \sim f_R(r) = ?$$

We need 1° $(X_{(1)}, X_{(2)}, \dots, X_{(n)}) \sim f_{X_{(1)}, \dots, X_{(n)}}(x_1, x_2, \dots, x_n) = ?$

2° $X_{(1)} \sim f_{X_{(1)}}(x) ?$ In this book, $V = X_{(1)}$, $V \sim f_V(v)$.

3° $X_{(n)} \sim f_{X_{(n)}}(x) ?$ $U = X_{(n)}$, $U \sim f_U(u)$.

4° any k -th $X_{(k)}$. $X_{(k)} \sim f_{X_{(k)}}(x)$.

5° any $(X_{(j)}, X_{(k)}) \sim f_{(X_{(j)}, X_{(k)})}(x, y) = ?$

Result 1. $X_1, X_2, \dots, X_n \sim F, f$. F is c.d.f, f is p.d.f.

(3)

$U = X_{(n)} = \max\{X_1, X_2, \dots, X_n\}$. find $f_U(u), F_U(u)$.

Ans:

$$F_U(u) = P(U \leq u) = P(X_{(n)} \leq u) = P(X_1 \leq u, X_2 \leq u, \dots, X_n \leq u)$$

$$\stackrel{iid}{=} P(X_1 \leq u) \cdot P(X_2 \leq u) \cdots P(X_n \leq u) = [F(u)]^n$$

$$f_U(u) = F'_U(u) = n[F(u)]^{n-1} \cdot f(u).$$

e.g. $X_1, \dots, X_n \sim \exp(\lambda)$ with $f(x) = \lambda e^{-\lambda x}$, $F(x) = 1 - e^{-\lambda x}$

$$\text{then } U = X_{(n)} \sim F_U(u) = [1 - e^{-\lambda u}]^n$$

$$U = X_{(n)} \sim f_U(u) = n[1 - e^{-\lambda u}]^{n-1} \cdot \lambda e^{-\lambda u} = n\lambda e^{-\lambda u} [1 - e^{-\lambda u}]^{n-1}, u > 0$$

R-simulation!

Result 2. $V = X_{(1)} = \min\{X_1, X_2, \dots, X_n\}$. find $F_V(v), f_V(v)$.

$$F_V(v) = P(V \leq v) = P(X_{(1)} \leq v) = 1 - P(X_{(1)} > v)$$

$$= 1 - P(X_1 > v, X_2 > v, \dots, X_n > v)$$

$$= 1 - [P(X_1 > v)]^n. \quad \text{note } P(X_1 > v) = 1 - P(X_1 \leq v)$$

$$= 1 - F(v)$$

$$= 1 - [1 - F(v)]^n.$$

$$f_V(v) = F'_V(v) = -\left\{[1 - F(v)]^{n-1}\right\}'_v = -n[1 - F(v)]^{n-1} \cdot (-f(v)) \\ = n \cdot f(v) \cdot [1 - F(v)]^{n-1}$$

again, if $X_i \sim \exp(\lambda)$

with $f(x) = \lambda e^{-\lambda x}, x > 0$, $F(x) = 1 - e^{-\lambda x}, x > 0$

$$\text{One has } F_V(v) = 1 - [1 - (1 - e^{-\lambda v})]^n = 1 - [e^{-\lambda v}]^n = 1 - e^{-\lambda nv}.$$

$$f_V(v) = F'_V(v) = -e^{-\lambda nv}(-\lambda n) = \lambda n e^{-\lambda nv}, \quad v > 0.$$

$V \sim \exp(n\lambda)$.

R demonstration

$$\begin{cases} \lambda=0.1 \\ n=20 \\ \lambda=2 \end{cases}$$

(4)

Result 3. X_1, X_2, \dots, X_n iid $f(x)$, then

$$(X_{(1)}, X_{(2)}, \dots, X_{(n)}) \sim f_{(X_{(1)}, \dots, X_{(n)})}(x_1, x_2, \dots, x_n) = n! \cdot f(x_1) \cdot f(x_2) \cdots f(x_n).$$

$$x_1 \leq x_2 \leq \cdots \leq x_n.$$

Proof. only for $n=2$ case here. show $(X_{(1)}, X_{(2)}) \sim f_{(X_{(1)}, X_{(2)})}(x_1, x_2) = 2f(x_1)f(x_2), x_1 \leq x_2$.

for any $x_1 \leq x_2$,

$$\int_{x_1}^{x_2} \int_{x_1 + \Delta x_1}^{x_2 + \Delta x_2} f(x) dx dy$$

$$P(x_1 \leq X_{(1)} \leq x_1 + \Delta x_1, x_2 \leq X_{(2)} \leq x_2 + \Delta x_2) = \binom{2}{1} P(x_1 \leq X_{(1)} \leq x_1 + \Delta x_1) \cdot \binom{1}{1} P(x_2 \leq X_{(2)} \leq x_2 + \Delta x_2)$$

$$= 2 \int_{x_1}^{x_1 + \Delta x_1} f(x) dx \int_{x_2}^{x_2 + \Delta x_2} f(y) dy$$

↑ indicate any one of two r.v's

$$= \int_{x_1}^{x_1 + \Delta x_1} \int_{x_2}^{x_2 + \Delta x_2} 2f(x)f(y) dx dy.$$

however, left side of above equality also

$$= \int_{x_1}^{x_1 + \Delta x_1} \int_{x_2}^{x_2 + \Delta x_2} f(x, y) dx dy$$

let $\Delta x_1 \rightarrow 0, \Delta x_2 \rightarrow 0$

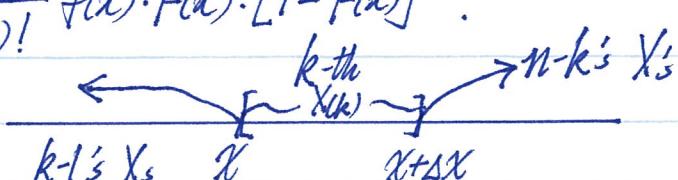
$$\text{thus left side} = f_{(X_{(1)}, X_{(2)})}(x_1, x_2) \cdot \Delta x_1 \Delta x_2. \quad \text{thus, } f_{(X_{(1)}, X_{(2)})}(x_1, x_2) = 2f(x_1)f(x_2).$$

$$\text{right side} = 2f(x_1) \cdot f(x_2) \Delta x_1 \Delta x_2.$$

$$x_1 \leq x_2.$$

Result 4. $X_{(k)}$ = k-th smallest of $\{X_1, X_2, \dots, X_n\}$.

$$\text{then } X_{(k)} \sim f_{X_{(k)}}(x) = \frac{n!}{(k-1)!(n-k)!} f(x) \cdot F(x)^{k-1} \cdot [1 - F(x)]^{n-k}.$$



Proof: $P(X \leq X_{(k)} \leq x + \Delta x)$

$$= \binom{n}{k-1} [P(X \leq x)]^{k-1} \cdot \binom{n-k+1}{1} P(x \leq X \leq x + \Delta x) \cdot \binom{n-k}{n-k} [P(X > x + \Delta x)]^{n-k}$$

$$= \frac{n!}{(k-1)!(n-k+1)!} [F(x)]^{k-1} \cdot (n-k+1) \cdot \int_x^{x+\Delta x} f(y) dy \cdot \frac{(n-k)!}{(n-k)! 0!} [1 - F(x + \Delta x)]^{n-k}$$

$$= \frac{n!}{(k-1)!(n-k)!} [F(x)]^{k-1} \int_x^{x+\Delta x} f(y) dy \cdot [1 - F(x+\Delta x)]^{n-k}$$

(5)

$$\text{note } P(x \leq X_{(k)} \leq x + \Delta x) = \int_x^{x+\Delta x} f(y) dy$$

$$\text{thus } \int_x^{x+\Delta x} f_{X_{(k)}}(y) dy = \frac{n!}{(k-1)!(n-k)!} [F(x)]^{k-1} \int_x^{x+\Delta x} f(y) dy \cdot [1 - F(x+\Delta x)]^{n-k}$$

$$\text{let } \Delta x \rightarrow 0, \quad f_{X_{(k)}}(x) = \frac{n!}{(k-1)!(n-k)!} [F(x)]^{k-1} \cdot f(x) \cdot [1 - F(x)]^{n-k}$$

Expl. $X_1, X_2, X_3 \stackrel{iid}{\sim} U[0, 1]$, find $P(\frac{1}{4} < X_{(2)} < \frac{3}{4})$

Ans. $n=3, k=2$. $f(x)=1, F(x)=x$,

$$f_{X_{(2)}}(x) = \frac{3!}{1!(3-2)!} x^{2-1} \cdot 1 \cdot (1-x)^{3-2} = 6x(1-x), \quad 0 < x < 1.$$

$$\text{i.e. } f_{X_{(2)}}(x) = \begin{cases} 6x(1-x), & 0 < x < 1 \\ 0, & \text{else.} \end{cases}$$

$$P(X_{(2)} < \frac{1}{4}) = \int_0^{\frac{1}{4}} f_{X_{(2)}}(x) dx = \frac{1}{8} = 0.125$$

$$P(\frac{1}{4} < X_{(2)} < \frac{3}{4}) = \int_{1/4}^{3/4} 6x(1-x) dx = \dots = \frac{11}{16} = 0.6875.$$

in R: $1 - \text{sum}(\text{med} < \frac{1}{4}) / k - \text{sum}(\text{med} > \frac{3}{4}) / k$

Result 5. $X_1, X_2, \dots, X_n \stackrel{iid}{\sim} f(x), F(x)$.

$$\text{then } (X_{(j)}, X_{(k)}) \sim f_{(X_{(j)}, X_{(k)})}(x, y) = \frac{n!}{(j-1)!(k-j-1)!(n-k)!} \times [F(x)]^{j-1} \times [F(y) - F(x)]^{k-j-1} \times [1 - F(y)]^{n-k} \cdot f(x) \cdot f(y). \quad x \leq y.$$

Proof:

$$\frac{x^{j-1} - x_{(j)}}{x} \quad \frac{x_{(j)}}{x+\Delta x} \quad \frac{k-1-j}{y} \quad \frac{x_{(k)}}{y+\Delta y} \quad \frac{n-k}{y}$$

$$P(x \leq X_{(j)} \leq x + \Delta x, y \leq X_{(k)} \leq y + \Delta y) =$$

(6)

$$\begin{aligned}
 & \binom{n}{j-1} [P(X \leq x)]^{j-1} \cdot \binom{n-j+1}{1} P(x \leq X \leq x + \Delta x) \cdot \binom{n-j}{k-j-1} [P(x + \Delta x \leq X \leq y)]^{k-j-1} \\
 & \quad \times \binom{n-k+1}{1} P(y \leq X \leq y + \Delta y) \cdot [1 - F(y + \Delta y)]^{n-k} \\
 = & \underbrace{\binom{n}{j-1} \cdot \binom{n-j+1}{1} \cdot \binom{n-j}{k-j-1} \cdot \binom{n-k+1}{1} \cdot [F(x)]^{j-1} \int_x^{x+\Delta x} f(u) du}_{\int_y^{y+\Delta y} f(u) du} \cdot [F(y) - F(x + \Delta x)]^{k-j-1} \\
 & \rightarrow \frac{n!}{(j-1)! (n-j+1)!} \frac{(n-j+1)!}{1 \cdot (n-j)!} \frac{(n-j)!}{(k-j-1)! (n-k+1)!} \cdot \frac{(n-k+1)!}{1 \cdot (n-k)!} \\
 & = \frac{n!}{(j-1)! (k-j-1)! (n-k)!}
 \end{aligned}$$

let $\Delta x \rightarrow 0, \Delta y \rightarrow 0$

thus, right hand side = $\frac{n!}{(j-1)! (k-j-1)! (n-k)!} [F(x)]^{j-1} f(x) \Delta x [F(y) - F(x)]^{k-j-1}$

$$* f(y) \cdot \Delta y \cdot [1 - F(y)]^{n-k}.$$

left hand side = $\int_x^{x+\Delta x} \int_y^{y+\Delta y} f(u, v) dv du$

$$= f(x, y) \cdot \Delta x \cdot \Delta y$$

Thus, $f(x, y) = \frac{n!}{(j-1)! (k-j-1)! (n-k)!} [F(x)]^{j-1} [F(y) - F(x)]^{k-j-1} [1 - F(y)]^{n-k} f(x) \cdot f(y)$

(7)

Example. Distribution of the range of a random sample.

$X_1, X_2, \dots, X_n \sim f$, $R = X_{(n)} - X_{(1)}$, find $F_R(r) = ?$

Ans. $j=1, k=n$.

$$f_{(X_{(1)}, X_{(n)})}(x, y) = \frac{n!}{(1-1)! (n-1-1)! (n-n)!} [F(x)]^{1-1} \cdot [F(y) - F(x)]^{n-1-1} [1 - F(y)]^{n-n} f(x) f(y)$$

$$= n \cdot (n-1) [F(y) - F(x)]^{n-2} \cdot f(x) f(y), \quad x \leq y.$$

$$R = X_{(n)} - X_{(1)},$$

$$F_R(a) = P(R \leq a) = P(X_{(n)} - X_{(1)} \leq a) \quad \{ \text{only consider } a \geq 0 \}$$

$$= \iint_{\substack{x < y \\ y-x \leq a}} f_{(X_{(1)}, X_{(n)})}(x, y) dx dy \quad \{ \text{here } x \leq y \}$$

$$= \int_{-\infty}^{\infty} \int_x^{x+a} n(n-1) [F(y) - F(x)]^{n-2} f(x) f(y) dy dx$$

$$= \int_{-\infty}^{\infty} n \cdot f(x) \left\{ \int_x^{x+a} (n-1) [F(y) - F(x)]^{n-2} f(y) dy \right\} dx$$

For the above inner integral $\int_x^{x+a} (n-1) [F(y) - F(x)]^{n-2} f(y) dy$,

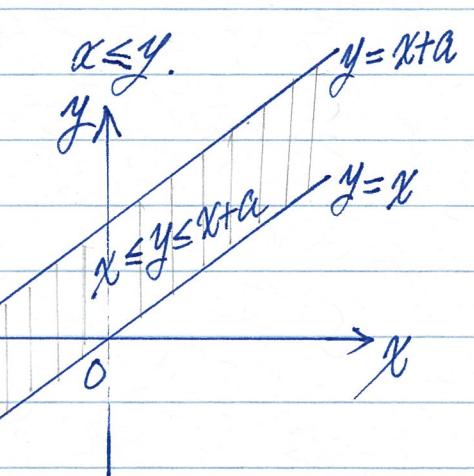
$$\text{Let } u = F(y) - F(x), \quad du = F'(y) dy = f(y) dy.$$

$$\text{thus, above inner integral} = \int_{F(x)-F(x)}^{F(x+a)-F(x)} (n-1) u^{n-2} du = u^{n-1} \Big|_{F(x)-F(x)}^{F(x+a)-F(x)}.$$

$$= [F(x+a) - F(x)]^{n-1}.$$

$$\text{Thus, } F_R(a) = \int_{-\infty}^{\infty} n \cdot f(x) \cdot [F(x+a) - F(x)]^{n-1} dx.$$

Let's consider case $X_1, X_2, \dots, X_n \sim \text{iid } U[0, 1]$. Then $f(x) = \begin{cases} 1, & x \in [0, 1] \\ 0, & \text{else.} \end{cases}$



(8)

$$F(x) = P(X \leq x) = \begin{cases} 0, & x < 0 \\ x, & 0 \leq x \leq 1 \\ 1, & x > 1. \end{cases}$$

$$F_R(a) = P(R \leq a) = P(X_{(n)} - X_{(1)} \leq a) = \begin{cases} 0, & a \leq 0 \\ 1, & a \geq 1 \end{cases}$$

So, we only need to consider $a \in [0, 1]$ below.

$$\begin{aligned} F_R(a) &= \int_{-\infty}^{\infty} n \cdot f(x) [F(x+a) - F(x)]^{n-1} dx \\ &= \int_0^1 n \cdot 1 \cdot [F(x+a) - x]^{n-1} dx \quad \left\{ \begin{array}{l} \text{from } f(x) = \begin{cases} 1, & x \in [0, 1] \\ 0, & \text{else} \end{cases} \\ \text{any } F(x) = x, x \in [0, 1] \end{array} \right\} \\ &= \int_0^{1-a} n [F(x+a) - x]^{n-1} dx + \int_{1-a}^1 n [F(x+a) - x]^{n-1} dx \\ &= \int_0^{1-a} n [x+a - x]^{n-1} dx + \int_{1-a}^1 n [1-x]^{n-1} dx \quad \left\{ \begin{array}{l} \text{if } 0 \leq x \leq 1-a, a \leq x+a \leq 1, \\ \text{then } F(x+a) = x+a. \end{array} \right. \\ &= \int_0^{1-a} n a^{n-1} dx + n \int_{1-a}^1 [1-x]^{n-1} dx \quad \left. \begin{array}{l} \text{if } x \geq 1-a, x+a \geq 1, \\ \text{then } F(x+a) = 1 \end{array} \right\} \\ &= n \cdot a^{n-1} \cdot (1-a) + n \cdot \left[\frac{-(1-x)^n}{n-1+1} \right] \Big|_{1-a}^1 \\ &= n a^{n-1} (1-a) + n \frac{1}{n} [0 + (1-1+a)^n] \\ &= n a^{n-1} (1-a) + a^n, \\ &= n a^{n-1} - n a^n + a^n = n a^{n-1} - (n-1) a^n \end{aligned}$$

$$\text{thus, } f_R'(a) = F_R'(a) = n(n-1) a^{n-2} - n(n-1) a^{n-1} = n(n-1) a^{n-2} (1-a), \quad 0 \leq a \leq 1.$$

$$\text{Put together. } R \sim f_R(a) = \begin{cases} n(n-1) a^{n-2} (1-a), & 0 \leq a \leq 1 \\ 0, & \text{else.} \end{cases}$$

$$\text{i.e. } R \sim \text{Beta}(n-1, 2).$$

(9)

Another solution: $Y_1 = g_1(X_1, X_2)$ $(X_1, X_2) \sim f_{X_1, X_2}(x_1, x_2),$
 $Y_2 = g_2(X_1, X_2)$

$$\begin{cases} X_1 = h_1(Y_1, Y_2) \\ X_2 = h_2(Y_1, Y_2) \end{cases} \quad J(X_1, X_2) = \begin{vmatrix} \frac{\partial g_1}{\partial Y_1} & \frac{\partial g_1}{\partial Y_2} \\ \frac{\partial g_2}{\partial Y_1} & \frac{\partial g_2}{\partial Y_2} \end{vmatrix}$$

Then, $f_{Y_1, Y_2}(y_1, y_2) = f_{X_1, X_2}(x_1, x_2) \cdot |J(x_1, x_2)|^{-1}$

here $f_{X_1, X_2}(x_1, x_2) = n(n-1)(x_2 - x_1)^{n-2}, 0 \leq x_1 \leq x_2 \leq 1.$

Now, let $\begin{cases} Y_1 = X_{(n)} - X_{(1)} \\ Y_2 = X_{(1)} \end{cases}$ i.e. $\begin{cases} Y_1 = X_2 - X_1 \\ Y_2 = X_1 \end{cases}$ $\begin{cases} Y_1 = g_1(X_1, X_2) = X_2 - X_1 \\ Y_2 = g_2(X_1, X_2) = X_1 \end{cases}$

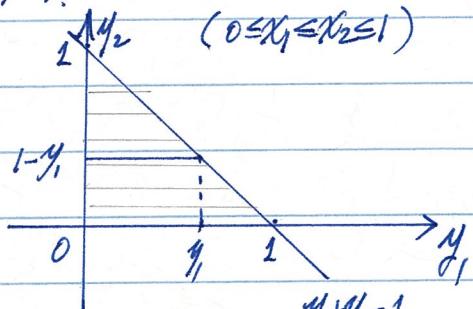
$$J(X_1, X_2) = \begin{vmatrix} -1 & 1 \\ 1 & 0 \end{vmatrix} = -1. \quad |J(x_1, x_2)| = |-1| = 1.$$

From above formula $f_{Y_1, Y_2}(y_1, y_2) = n(n-1)(x_2 - x_1)^{n-2} \cdot 1$

$$= n(n-1)y_1^{n-2}, \quad 0 \leq y_2 \leq y_1 + y_2 \leq 1, \quad 0 \leq y_2 \leq y_1 + y_2 \leq 1$$

i.e. $f_{Y_1, Y_2}(y_1, y_2) = \begin{cases} n(n-1)y_1^{n-2}, & 0 \leq y_1 \leq 1, \\ & 0 \leq y_1 + y_2 \leq 1 \\ 0, & \text{else.} \end{cases}$

$$\begin{aligned} 0 \leq y_1 \leq 1 \\ f_{Y_1}(y_1) &= \int f_{Y_1, Y_2}(y_1, y_2) dy_2 = \int_0^{1-y_1} n(n-1)y_1^{n-2} dy_2 \\ &= n(n-1)y_1^{n-2} \cdot (1-y_1). \end{aligned}$$



Put together. $X_{(n)} - X_{(1)} = Y_1 \sim f_{Y_1}(y) = \begin{cases} n(n-1)y^{n-2}(1-y), & 0 \leq y \leq 1 \\ 0, & \text{else} \end{cases}$

i.e. $R = X_{(n)} - X_{(1)} \sim \text{Beta}(n-1, 2).$