

**Question:** Verify that:

$$P(C|A) = P(C|BA)P(B|A) + P(C|B^c A)P(B^c|A)$$

Where,  $AB$  means  $A \cap B$ .

**Answer:**

Take the RHS,

We can expand  $P(C|BA)$  as:

$$P(C \cap B \cap A) / P(B \cap A)$$

Similarly, we can write  $P(C|B^c A)$  as:

$$P(C \cap B^c \cap A) / P(B^c \cap A)$$

We can expand it as:

$$\frac{P(C \cap B \cap A)}{P(B \cap A)} \times \frac{P(B \cap A)}{P(A)} + \frac{P(C \cap B^c \cap A)}{P(B^c \cap A)} \times \frac{P(B^c \cap A)}{P(A)}$$

Which gives:

$$\frac{P(C \cap B \cap A)}{P(A)} + \frac{P(C \cap B^c \cap A)}{P(A)}$$

$P(CBA)$  and  $P(CB^c A)$  are the probabilities of mutually exclusive events.

$$\begin{aligned} &= \frac{1}{P(A)} \times P(CBA \cup CB^c A) \\ &= \frac{P(CA)}{P(A)} \text{ or } \frac{P(C \cap A)}{P(A)} \end{aligned}$$

Thus, proved.

## Quiz-2: Probability and Statistics (30 Marks)

[Instruction: Please state reasons wherever applicable.]

### 5 Marks

**Find the stationary distribution  $\pi$  for Markov Chains with the following transition probability matrix (3 marks). State if  $\pi$  is unique in each case (1 mark). Also which of the two chains are irreducible? Give reasons (1 mark).**

$$P = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} \text{ and } Q = \begin{bmatrix} 0 & 1 \\ 1 & 0 \end{bmatrix}$$

#### Solution

Stationary distribution  $\pi$  for a Discrete Markov Chain, given its transition probability matrix  $P$ , is given as:

$$\pi P = \pi$$

#### Stationary Distribution for P (1.5 marks)

Given  $P = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$

$$\begin{aligned} \pi P &= [\pi_1 \quad \pi_2] \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} \\ &= [\pi_1 \quad \pi_2] \\ &= \pi \end{aligned}$$

This, holds for all  $\pi$  (since  $P$  is an identity matrix).

$\therefore$  Stationary distribution for  $P$

$$= [p \quad 1-p] \text{ where } 0 \leq p \leq 1, p \in \mathbb{R}$$

**Stationary Distribution for Q (1.5 marks)**

Given  $Q = \begin{bmatrix} 0 & 1 \\ 1 & 0 \end{bmatrix}$

$$\begin{aligned}\pi Q &= [\pi_1 \quad \pi_2] \begin{bmatrix} 0 & 1 \\ 1 & 0 \end{bmatrix} \\ &= [\pi_2 \quad \pi_1]\end{aligned}$$

$$\begin{aligned}\pi Q &= \pi \\ \implies [\pi_2 \quad \pi_1] &= [\pi_1 \quad \pi_2] \\ \implies \pi_1 &= \pi_2 = \frac{1}{2} \quad (\because \pi_1 + \pi_2 = 1)\end{aligned}$$

$\therefore$  Stationary distribution for  $Q = \begin{bmatrix} \frac{1}{2} & \frac{1}{2} \end{bmatrix}$ .

**State if  $\pi$  is unique (0.5 + 0.5 mark)**

- Stationary distribution for  $P$  is not unique. (since any  $\pi$  can be its stationary distribution).
- Stationary distribution for  $Q$  is unique.

**Which of the two chains are irreducible? (1 mark)**

We know  $P = I \implies P^n = I^n = I$  where  $I$  is the Identity matrix of order 2.

$$\implies P_{12}^n = 0$$

Thus, state 2 is not accessible from state 1. This is sufficient to show that  $P$  is not an irreducible chain.

$$Q^n = \begin{cases} \begin{bmatrix} 0 & 1 \\ 1 & 0 \end{bmatrix} & \text{when } n \text{ is odd} \\ \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} & \text{when } n \text{ is even} \end{cases}$$

We see that states 1 and 2 communicate with each other *i.e.*,  $P_{12}^n > 0$  and  $P_{21}^n > 0$  (when  $n$  is odd). Thus,  $Q$  is an irreducible markov chain.

## Quiz: Probability and Statistics

October 28, 2022

### Question 3

Suppose  $X$  is an exponential random variable with parameter  $\lambda$  and CDF denoted by  $F_X(\cdot)$ .  $U$  is a uniform random variable over the interval  $[0, 1]$ . Now consider another random variable  $Y = F_X^{-1}(U)$ . Then derive the expression for the CDF  $F_Y(y)$ .

### Solution

$$\begin{aligned} F_Y(y) &= P(F_X^{-1}(U) \leq y) \\ &= P(U \leq F_X(y)) \\ &= P(U \leq 1 - e^{-\lambda y}) \\ &= F_U(1 - e^{-\lambda y}) \\ &= 1 - e^{-\lambda y} \end{aligned}$$

## Quiz 2: Probability and Statistics

[Instruction: Please state reasons wherever applicable.]

### 1 5 Marks

1. Consider a sequence of random variables  $\{X_n\}$  where  $X_n \sim \text{Exponential}(n)$ . Show that  $X_n$  converges to  $X$  in probability where  $X = 0$  with probability 1. Also show that  $X_n$  converges to  $X$  in distribution (without using the fact that convergence in probability implies convergence in distribution).

**Solution:**

**a.**

$$\begin{aligned}\lim_{n \rightarrow \infty} P(|X_n - 0| \geq \epsilon) &= \lim_{n \rightarrow \infty} P(X_n \geq \epsilon) && [\because X_n \geq 0] \{0.5 \text{ Marks}\} \\ &= \lim_{n \rightarrow \infty} e^{-n\epsilon} && [\because X_n \sim \text{Exponential}(n)] \{1 \text{ Mark}\} \\ &= 0 && \{1 \text{ Mark}\}\end{aligned}$$

*Hence Proved.*

**b.**

$$\begin{aligned}\lim_{n \rightarrow \infty} F_{X_n}(x) &= \lim_{n \rightarrow \infty} 1 - e^{-nx} && \{1 \text{ Mark}\} \\ &= 1 && \{0.5 \text{ Mark}\} \\ &= F_X(x) && [\forall x > 0][\because P_X(0) = 1]\end{aligned}$$

*Note that at  $x = 0$ ,  $F_X(x)$  is discontinuous*

$\implies$  Convergence in distribution doesn't take place at  $x = 0$   $\{0.5 \text{ Mark}\}$

$\implies$  Convergence in distribution takes place  $\forall x > 0$   $\{0.5 \text{ Mark}\}$

*Hence Proved.*

**Marks Division**

- (a) Convergence in probability (2.5M)
  - i. 0.5M for identifying  $P(|X_n - 0| \geq \epsilon) = P(X_n \geq \epsilon)$
  - ii. 1M for getting to  $e^{-n\epsilon}$
  - iii. 1M for getting to final step
- (b) Convergence in distribution (2.5M)
  - i. 1M for writing CDF of  $X_n$
  - ii. 0.5M for getting to 1
  - iii. 0.5M for accounting for discontinuity
  - iv. 0.5M for getting  $F_X(x) = 1 \quad \forall x > 0$

**Note:** Simple stating of final answers without any logical approach will be given 0.

# Probability and Statistics: Quiz 2

October 2022

## Question 1 (10 Marks)

**Given:** Three samples  $u_1 = 0.23$ ,  $u_2 = 0.73$  and  $u_3 = 0.5$  from uniform random variable. We will use the inverse transform method in all the following parts to convert the given sample to the required.

1. Let  $X$  be the random variable denoting the outcome of a fair dice. Now

$$F_x(x) = \begin{cases} 0 & \text{if } x < 1 \\ 1/6 & \text{if } x \geq 1 \text{ and } x < 2 \\ 2/6 & \text{if } x \geq 2 \text{ and } x < 3 \\ 3/6 & \text{if } x \geq 3 \text{ and } x < 4 \\ 4/6 & \text{if } x \geq 4 \text{ and } x < 5 \\ 5/6 & \text{if } x \geq 5 \text{ and } x < 6 \\ 1 & \text{if } x \geq 6 \end{cases}$$

Now we know by the Lemma of inverse transform method that if:

$$X := F^{-1}(U)$$

Then the cdf of  $X$  is  $F$ . Hence, applying the inverse transform method, we get:

$$X = \begin{cases} 1 & \text{if } p < 1/6 \\ 2 & \text{if } p > 1/6 \text{ and } p \leq 2/6 \\ 3 & \text{if } p > 2/6 \text{ and } p \leq 3/6 \\ 4 & \text{if } p > 3/6 \text{ and } p \leq 4/6 \\ 5 & \text{if } p > 4/6 \text{ and } p \leq 5/6 \\ 6 & \text{if } p > 5/6 \text{ and } p \leq 6/6 \end{cases}$$

Where  $p$  is a realization from uniform random variables. Hence, applying this, we get the following samples for  $X$ .

- $u_1 = 0.23$  generates  $x = 2$  as a sample.
- $u_2 = 0.78$  generates  $x = 5$  as a sample.
- $u_3 = 0.5$  generates  $x = 3$  or  $x = 4$  as a sample.

2. Let  $X$  be a random variable with 0.7 as a probability of getting a head. Let head be  $X = 0$  and tail be  $X = 1$

$$F_x(x) = \begin{cases} 0 & \text{if } x < 0 \\ 0.7 & \text{if } x \geq 0 \text{ and } x < 1 \\ 1 & \text{if } x \geq 1 \end{cases}$$

Applying the inverse transform method, we get:

$$X = \begin{cases} 0 & \text{if } p \leq 0.7 \\ 1 & \text{if } p > 0.7 \text{ and } p \leq 1 \end{cases}$$

Where  $p$  is a realization from uniform random variables. Hence, applying this, we get the following samples for  $X$ .

- $u_1 = 0.23$  generates  $x = 0$  i.e Heads as a sample.
- $u_2 = 0.78$  generates  $x = 1$  i.e Tails as a sample.
- $u_3 = 0.5$  generates  $x = 0$  i.e Heads as a sample.

3. Let  $X$  be the exponential random variable with parameter  $\lambda = 1$ . We know

$$f_x(x) = \lambda e^{-\lambda x}$$

Also, cdf of  $f_x(x)$  is written as

$$F_x(x) = 1 - e^{-\lambda x}$$

$$F_x(x) = 1 - e^{-x} \text{ as } \lambda = 1$$

Using the lemma of inverse transform method, we substitute  $F_x(x)$  with  $U$  and thus, we have

$$U = 1 - e^{-x}$$

$$x = -\ln(1 - U)$$

Since,  $U$  and  $1 - U$  are equivalent, since both are uniform random variable over  $(0, 1)$ , we can replace  $U$  and  $1 - U$ .

$$X = -\ln(U)$$

Where  $U$  is a realization from uniform random variables. Hence, applying this, we get the following samples for  $X$ .

- $u_1 = 0.23$  generates  $x = -\ln(0.23)$  or  $x = -\ln(0.77)$  as a sample.
- $u_2 = 0.78$  generates  $x = -\ln(0.78)$  or  $x = -\ln(0.22)$  as a sample.
- $u_3 = 0.5$  generates  $x = -\ln(0.5)$  as a sample.

4. Let  $X$  be the inform random variable in the interval  $[5, 10]$

$$f_x(x) = \frac{1}{10-5} = 1/5$$

$$F_x(x) = \begin{cases} 0 & \text{if } x < 5 \\ \frac{x-5}{5} & \text{if } x \geq 5 \text{ and } x \leq 10 \\ 1 & \text{if } x > 10 \end{cases}$$

Using the lemma of inverse transform method, we substitute  $F_x(x)$  with  $U$  and thus, we have



$$U = \frac{X-5}{5}$$

$$X = 5U + 5$$

- $u_1 = 0.23$  generates  $x = 6.15$  as a sample.
- $u_2 = 0.78$  generates  $x = 8.9$  as a sample.
- $u_3 = 0.5$  generates  $x = 7.5$  as a sample.

## Marking Scheme

- Part 1: 2 Marks
- Part 2: 2 Marks
- Part 3: 3 Marks
- Part 4: 3 Marks

**Note:** Marks will be deducted if the inequalities are incorrect.